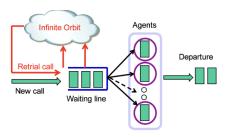
#### Retrial Queues

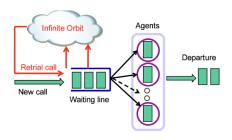
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PMCS Project , Spring '24

### Analogy



### Analogy



- ► Busy Call Centres
- ► TCP Packet Transmission
- LAN

#### **Notation**

- $ightharpoonup \lambda$  : arrival rate of primary calls
- $\blacktriangleright \mu$  : rate of repeated calls
- $\triangleright$  B(x): service distribution
- ightharpoonup C(t): no of busy servers at time t
- $\triangleright$  N(t): no of sources of repeated calls
- $\blacktriangleright \xi(t)$ : age of current process
- ho  $\beta(t) = \int_0^\infty e^{-sx} dB(x)$ : Laplace transform of service time
- $b(x) = \frac{B'(x)}{1-B(x)}$ : Hazard rate
- $k(z) = \sum_{0}^{\infty} k_{n} z^{n} = \beta(\lambda \lambda z)$

$$k_n = \int_0^\infty \frac{\lambda x^n}{n!} e^{-x} dB(x)$$

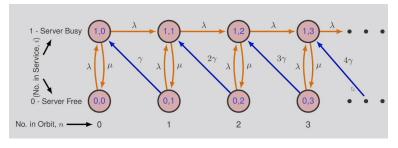
is distribution of number of primary calls that arrive during service time of a call

# M/M/1

Service Time distribution

$$B(x) = 1 - e^{-\nu x}$$

State transitions



# M/M/1 (State Transitions)

From a state (0, n), only transitions into the following states are possible:

- 1. (1, n) with rate  $\lambda$ ;
- 2. (1, n-1) with rate  $\nu$ .

Reaching state (0, n) is possible only from state (1, n) with rate  $\nu$ . From a state (1, n), only transitions into the following states are possible:

- 1. (1, n+1) with rate  $\lambda$ ;
- 2. (0, n) with rate  $\nu$ .

Reaching state (1, n) is possible only from the states:

- 1. (0, n) with rate  $\lambda$ ;
- 2. (0, n + 1) with rate  $(n + 1)\mu$ ;
- 3. (1, n-1) with rate  $\lambda$ .

## M/M/1 (Limiting Distribution)

The statistical probability equations are given by

$$(\lambda + n\mu)p_{0n} = \nu p_{1n},$$
  
 $(\lambda + \nu)p_{1n} = \lambda p_{0n} + (n+1)\mu p_{0,n+1} + \lambda p_{1,n-1}$ 

The partial generating functions are

$$p_0(z) = \frac{(1-\rho)^{\frac{\lambda}{\mu}+1}}{(1-\rho z)^{\frac{\lambda}{\mu}}}.$$

$$p_1(z) = \frac{\rho}{(1-\rho z)}p_0(z)$$

Mean number of jobs in queue

$$E[N(t)] = \frac{\rho(\lambda + \rho\mu)}{(1 - \rho)\mu}$$

The stationary distribution of the number of sources of repeated calls  $q_n = PN(t) = n$  has the generating function

$$p(z) = p_0(z) + p_1(z) = (1 + \rho - \rho z)(\frac{1 - \rho}{1 - \rho z})^{\frac{\lambda}{\mu} + 1}.$$

$$E[N(t)] = \sum np_n = p'(1)$$

Mean number of jobs in system

$$E[K(t)] = \frac{\rho(\lambda + \mu)}{(1 - \rho)\mu}$$

$$Q(z) = p_0(z) + zp_1(z) = (\frac{1-\rho}{1-\rho z})^{\frac{\lambda}{\mu}+1}$$
  $E[K(t)] = Q'(1)$ 

Blocking probability

$$p_b = \rho = \frac{\lambda}{\nu}$$

$$P(\mathsf{Server\ busy}) = \sum_n p_{1n} = p_(1)$$

Mean sojourn time

$$W = \frac{\rho(\lambda + \mu)}{(1 - \rho)\mu\lambda}$$

Found using Little's Law

Recurrence Conditions

```
ho < 1 for positive recurrence 
ho = 1 for null recurrence 
ho > 1 for transience
```

Proved by examining mean sojourn time

Mean no of retrials per job

$$E[R] = \frac{\rho(\lambda + \rho\mu)}{(1 - \rho)\lambda}$$

If a job spends time T inside the system, it retries after  $X_i$  which are  $Exp(\mu)$ . The no of retries is a stopping time for X.  $T = \sum_{i=0}^{R} X_i$ 

$$\implies E[T] = E[X]E[R]$$
$$\implies E[R] = \mu E[T]$$

#### M/M/1 (Embedded DTMC)

- General service times do not have the memoryless property.
- Convert CTMC to an Embedded DTMC by taking  $N_i = N(\eta_i)$  i.e no of calls in orbit at the time  $\eta_i$  of  $i^{th}$  departure.

$$N_i = N_{i-1} - B_i + \nu_i$$

- $\triangleright$   $B_i$  is indicator for repeated calls
- $\triangleright \nu_i$  is no of jobs that arrive during service

$$P\{\nu_i = n\} = k_n = \int_0^\infty \frac{(\lambda x)^n}{n!} e^{-\lambda x} dB(x)$$

#### Embedded DTMC

One-step transition probabilities  $r_{mn} = P\{N_i = n \mid N_{i-1} = m\}$  are given by the formula

$$r_{mn} = \frac{\lambda}{\lambda + m\mu} k_{n-m} + \frac{m\mu}{\lambda + m\mu} k_{n-m+1}$$

Mean Length of queue

$$E[N] = \rho + \frac{\lambda \rho}{1 - \rho} \left( \frac{1}{\mu} + \frac{1}{\nu} \right)$$

#### Special Case: M/M/2

lacktriangle Consider the basic case of multi server model (Taking u=1)

$$p_{0j} = P\{C(t) = 0, N(t) = j\}$$

$$p_{1j} = P\{C(t) = 1, N(t) = j\}$$

$$p_{2j} = P\{C(t) = 2, N(t) = j\}$$

be the limiting distributions. These probabilities satisfy the following statistic probability equations

$$(\lambda + j\mu)p_{0j} = p_{1j}$$
  
 $(\lambda + 1 + j\mu)p_{1j} = \lambda p_{0j} + (j+1)\mu p_{0,j+1} + 2p_{2j}$   
 $(\lambda + 2)p_{2j} = \lambda p_{1j} + (j+1)\mu p_{1,j+1} + \lambda p_{2,j-1}$ 

and normalizing condition

$$\sum_{j=0}^{\infty} (p_{0j} + p_{1j} + p_{2j}) = 1$$

# Simulating M/M/1

- ► The arrival of jobs , the retrial of jobs in orbit and the processing of the jobs is simulated by generating random numbers.
- ➤ This gives us the pre-computed values of various parameters like arrival time , departure time, number of retrials etc . These will be used to compute various performence metrics.

# Simulating M/M/1

Parameters:  $\lambda=5, \nu=7, \mu=1$ 

Parameter	Expected Values	Simulated Values
Mean Sojourn Time	3	3.16
Blocking Probability	0.71	0.705
Mean No of jobs	15	15.85
Mean No of pings	3	3.02

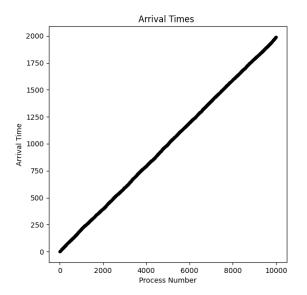


Figure: Arrival Times of the Jobs

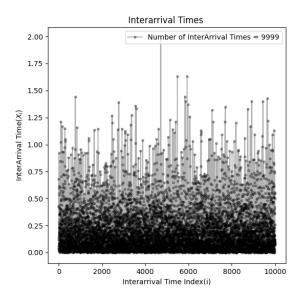


Figure: Interarrival Times

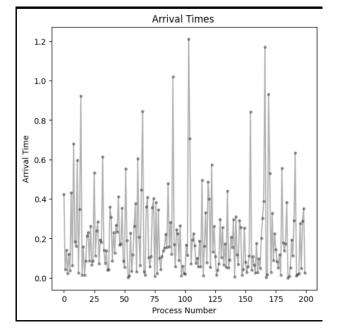


Figure: Interarrival Times

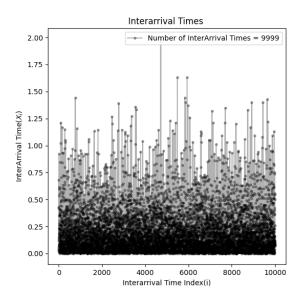


Figure: Interarrival Times

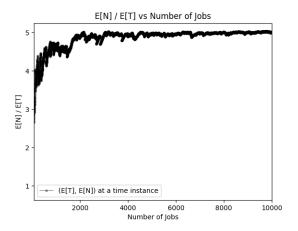


Figure: Verification of Little's Law

Little's Law is verified by:

E[N]/E[T] converges to  $\lambda$ 

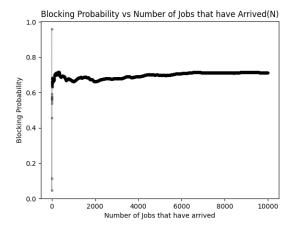


Figure: Blocking Probability vs Number of Jobs arrived(N)

The Blocking Probability stagnates to 0.70.