

Exam 1: Solution  
Stat 428  
March 3, 2015

1. Consider the probability density function  $f(x) = \lambda e^{-\lambda(x-\eta)}$ , for  $x \geq \eta$ , where  $\lambda$  and  $\eta$  are positive constants.
  - a. Find the cumulative distribution function  $F(x)$ .
  - b. Given two draws from a uniform distribution on  $(0,1)$ ,  $u_1 = .343$  and  $u_2 = .879$ , use the inverse cdf method to obtain two variates  $x_1$  and  $x_2$  from  $f(x)$ .
  - c. Let  $\lambda = 2$  and  $\eta = 1$ , and the probability density function  $g(x) = e^{-x}$  for  $x > 0$ . Find a constant  $c$  that satisfies  $f(x)/g(x) \leq c$  for  $x > 1$ , and write *R* code for generating a sample of size  $m = 10000$  from  $f(x)$  using acceptance rejection sampling with proposals drawn from  $g(x)$  (you may use `rexp()`).
  - d. Using `rexp()`, discuss how you could use a simple transformation method by drawing observations from a well chosen exponential distribution and transforming them into a sample from  $f(x) = 2e^{-2(x-1)}$  for  $x > 1$ .

**Solution:**

a. Find out cdf. For  $x \geq \eta$ ,

$$F(x) = \int_{\eta}^x f(t)dt = \int_{\eta}^x \lambda e^{-\lambda(t-\eta)} dt = \lambda \times \frac{1}{(-\lambda)} e^{-\lambda(t-\eta)} \Big|_{\eta}^x = -e^{-\lambda(t-\eta)} \Big|_{\eta}^x = 1 - e^{-\lambda(x-\eta)} \quad (1)$$

b. Find out inverse cdf.

$$\begin{aligned} u = F(x) = 1 - e^{-\lambda(x-\eta)} &\Rightarrow e^{-\lambda(x-\eta)} = 1 - u \\ &\Rightarrow -\lambda(x - \eta) = \ln(1 - u) \\ &\Rightarrow x = -\frac{1}{\lambda} \ln(1 - u) + \eta \end{aligned} \quad (2)$$

Plug in  $u_1$  and  $u_2$ , then we get:

$$x_1 = -\frac{1}{\lambda} \ln(1 - u_1) + \eta = -\frac{1}{\lambda} \ln(0.657) + \eta$$

$$x_2 = -\frac{1}{\lambda} \ln(1 - u_2) = -\frac{1}{\lambda} \ln(0.121) + \eta$$

It is ok if you plug in  $\lambda$ ,  $\eta$  given in part c. Then the answer would be  $x_1 = 1.210$  and  $x_2 = 2.056$ .

c. Since  $f(x) = 2e^{-2(x-1)}$  and  $g(x) = e^{-x}$ , for  $x \geq 1$ ,

$$\frac{f(x)}{g(x)} = \frac{2e^{-2(x-1)}}{e^{-x}} = 2e^{-x+2} \leq 2e \quad (3)$$

Therefore, we can set  $c = 2e \approx 5.437$ . Then,  $\frac{f(x)}{cg(x)} = e^{-x+1}$

```
m=0 #our sample size
n=0 #how many times we draw from g(x)
x=NULL
while (m<=10000)
{
  u=runif(1)
  xx=rexp(1)
  if (xx>=1) #notice that the supports of g(x) and f(x) are not the same
  {
    if (u<exp(1-xx)) #simplify f(x)/cg(x)
    {
      x=c(x, xx)
      m=m+1
    }
  }
  n=n+1
}
```

**Note:** You can also use the method on classnotes. Noticing that  $c = 2e \approx 5.437$ , you may want to draw roughly 54370 samples.

d.  $f(x) = 2e^{-2(x-1)}$  for  $x > 1$  is a shifted exponential distribution for  $\exp(2)$ . That is, it just shifts the density curve 1 unit to the left. Therefore, what we can do is, simply draw samples  $X_i$  from  $\exp(2)$ , or  $\text{rexp}(2)$ . Transform samples to be  $Y_i = X_i + 1$ . Then  $Y_i$  is what we want from target distribution.

2. Recall that if  $X$  has a chi-squared distribution with  $w$  degrees of freedom and  $Y$  is independent of  $X$  and has a chi-squared distribution with  $v$  degrees of freedom, then  $F = \frac{X/w}{Y/v}$  has an F-distribution with  $w$  and  $v$  degrees of freedom. Write an R function for drawing from an F distribution with user supplied  $w$  and  $v$  that makes use of the `rnorm()` function for drawing from a normal distribution.

**Solution:**

Notice that the chi-squared distribution with freedom  $w$  is the sum of  $w$ 's square of independent standard normal variables. That is, if  $X \sim \chi^2(w)$ , then  $X = S_1^2 + S_2^2 + \dots + S_w^2$ , and  $S_i \sim N(0, 1), i = 1, 2, \dots, w$ . It's the same thing for  $Y \sim \chi^2(v)$ .

Then transform  $X, Y$ , two independent Chi-square samples, to get a sample from F distribution, using  $F = \frac{X/w}{Y/v}$ .

```
rfdist=function(w,v){
  xchisq=0
  ychisq=0
  for (i in 1:w)
    xchisq=xchisq+(rnorm(1))^2
  for (i in 1:v)
    ychisq=ychisq+(rnorm(1))^2
  sample=xchisq/w/(ychisq/v)
  return(sample)
}
```

3. Let  $\theta = \int_0^{\pi/2} e^{\sin x} dx$ .

- Write R code to construct an estimate  $\hat{\theta}$  of  $\theta$  using Monte Carlo integration with respect to a well chosen uniform distribution.
- Derive an upper bound for the variance  $\hat{\theta}$ . It doesn't have to be the precise variance, but a reasonably close upper bound.

**Solution:**

a.

- Draw  $m$  (say,  $m=10000$ ) samples  $X_1, X_2, \dots, X_m$  from  $f(x) \sim \text{Unif}[0, \frac{\pi}{2}]$
- Transform them into:  $g(X_i) = e^{\sin X_i}, i = 1, 2, \dots, m$ .
- Estimate the expectation value of  $g(X_i)$  using:

$$\hat{\theta} = \int_0^{\frac{\pi}{2}} e^{\sin t} dt = E_f(g(X)) \approx \frac{1}{m} \sum_{i=1}^m \frac{g(X_i)}{f(X_i)} = \frac{\frac{\pi}{2} - 0}{m} \sum_{i=1}^m g(X_i) \quad (4)$$

R code:

```
xx=runif(10000,0,pi/2)
theta_est=(pi/2)*mean(exp(sin(xx)))
theta_est
```

- b. Since  $X_i$ 's are i.i.d from  $\text{Unif}[0, \frac{\pi}{2}]$ . After some derivation, we try to bound  $\text{Var}(e^{\sin X})$ . Notice that since  $X \in [0, \frac{\pi}{2}]$ ,  $e^{\sin X} \in [0, 1]$ .

$$\text{Var}(\hat{\theta}) = \text{Var}\left(\frac{\frac{\pi}{2} - 0}{m} \sum_{i=1}^m e^{\sin X_i}\right) = \frac{\pi^2}{4m^2} \text{Var}\left(\sum_{i=1}^m e^{\sin X_i}\right) = \frac{\pi^2}{4m} \text{Var}(e^{\sin X}) \quad (5)$$

$$\begin{aligned} \text{Var}(e^{\sin X}) &= E(e^{2\sin X}) - E^2(e^{\sin X}) \\ &\leq E(e^2) - E^2(1) = e^2 - 1 \end{aligned} \quad (6)$$

Note: Other reasonable answers are  $e^2, (e-1)^2$  etc.

4. Let  $\phi(x) = ce^x$  for  $0 < x < \pi/2$ , and  $\theta = \int_0^{\pi/2} e^{\sin x} dx$ .
- Find  $c$  so that  $\phi(x)$  is a probability density function.
  - Using the inverse cdf sampling technique to draw observations from  $\phi(x)$ , write R code to construct an estimator  $\hat{\theta}$  of  $\theta$  using importance sampling with importance function  $\phi(x)$ .
  - Assuming the estimator of (4b) uses the same number of draws  $m$  as the estimator of (3a), discuss which you would expect to be more efficient, and why.

**Solution:**

a.

$$\begin{aligned} \int_0^{\pi/2} \phi(t) dt &= \int_0^{\pi/2} ce^t dt = ce^t \Big|_0^{\pi/2} = c(e^{\pi/2} - 1) = 1 \\ \Rightarrow c &= \frac{1}{e^{\pi/2} - 1} \approx 0.2624 \end{aligned} \quad (7)$$

b. First of all, find out cdf of  $\phi(x)$ . For  $x \in (0, \frac{\pi}{2})$ ,

$$\begin{aligned} F(x) &= \int_0^x \phi(t) dt = \int_0^x ce^t dt = ce^t \Big|_0^x = c(e^x - 1) \\ &= \frac{e^x - 1}{e^{\pi/2} - 1} \end{aligned} \quad (8)$$

Then, find out inverse cdf,

$$\begin{aligned} u = F(X) &= \frac{e^X - 1}{e^{\pi/2} - 1} \Rightarrow e^X - 1 = u(e^{\pi/2} - 1) \\ &\Rightarrow e^X = u(e^{\pi/2} - 1) + 1 \\ &\Rightarrow X = \ln(u(e^{\pi/2} - 1) + 1) \end{aligned} \quad (9)$$

Let  $g(x) = e^{\sin x}$ . We first got samples from  $\phi(x)$  using inverse cdf sampling technique. Then, we can estimate  $\theta$  using importance sampling with importance function  $\phi(x)$ .

$$\theta = \int_0^{\pi/2} e^{\sin x} dx = \int_0^{\pi/2} \frac{g(x)}{\phi(x)} \times \phi(x) dx \quad (10)$$

R code:

```
#sample size=10000
#first get 10000 samples from phi(x)
c=1/(exp(pi/2)-1)
```

```

u=runif(10000)
x=log(u/c+1)
xx=exp(sin(x))/(c*exp(x))  #importance sampling samples
theta_hat=mean(xx)          #estimate

```

c. The method in (4b) is more efficient. Compared with uniform distribution in (3a),  $\phi(x)$  and  $g(x)$  are more similar in shape (they are both exponential-like density functions). Therefore, we can expect that the importance sampling in (4b) is more efficient.