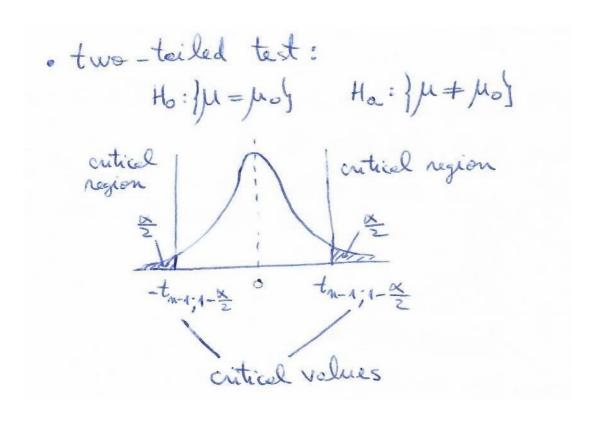
Significance tests (cont....)

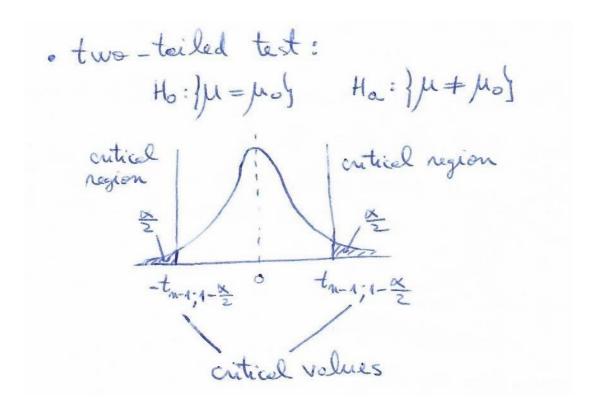
For the acceptance/cutical region approach (also called the critical value approach), we need to compute the test statistic and to find the critical values (i.e. the quantiles) corresponding to a given significance level &.

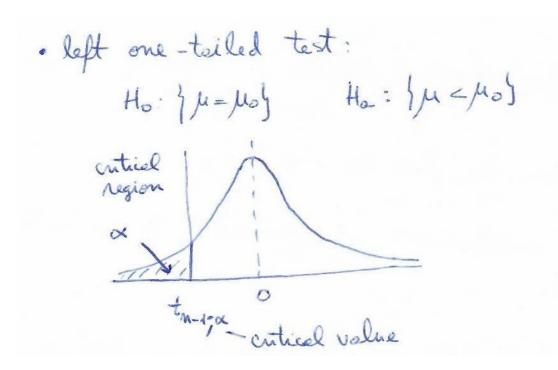
For the one-sample T-tests, depending on the nature of the alternative hypothesis, we use a two-tailed or a one-tailed test:

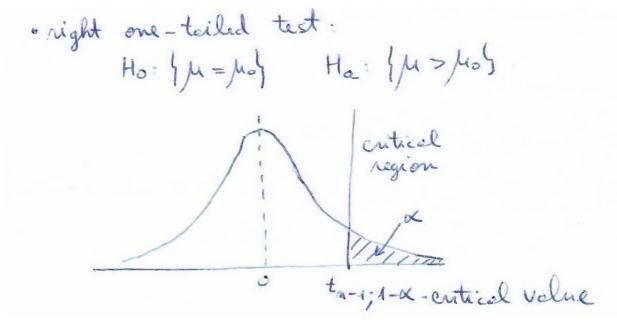
One sample t-test



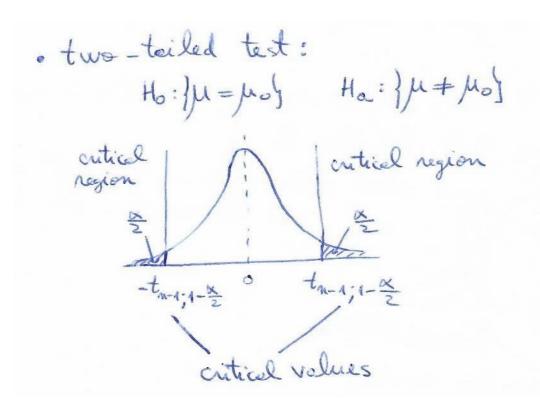
One sample t-test



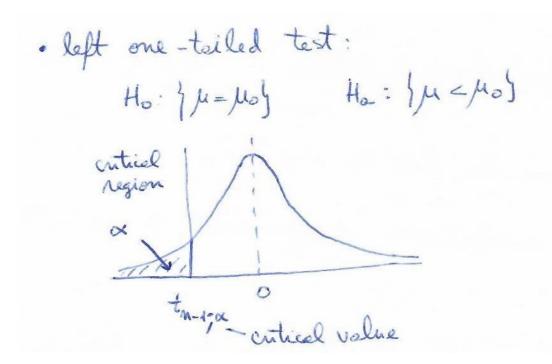


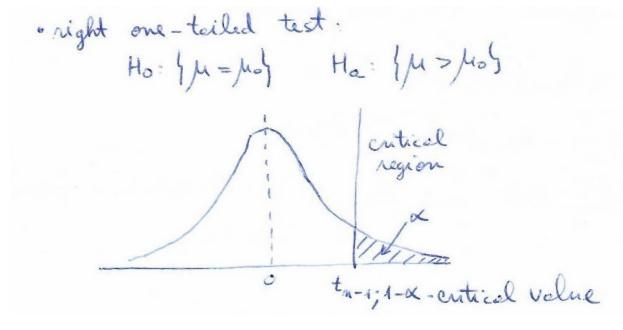


One sample t-test

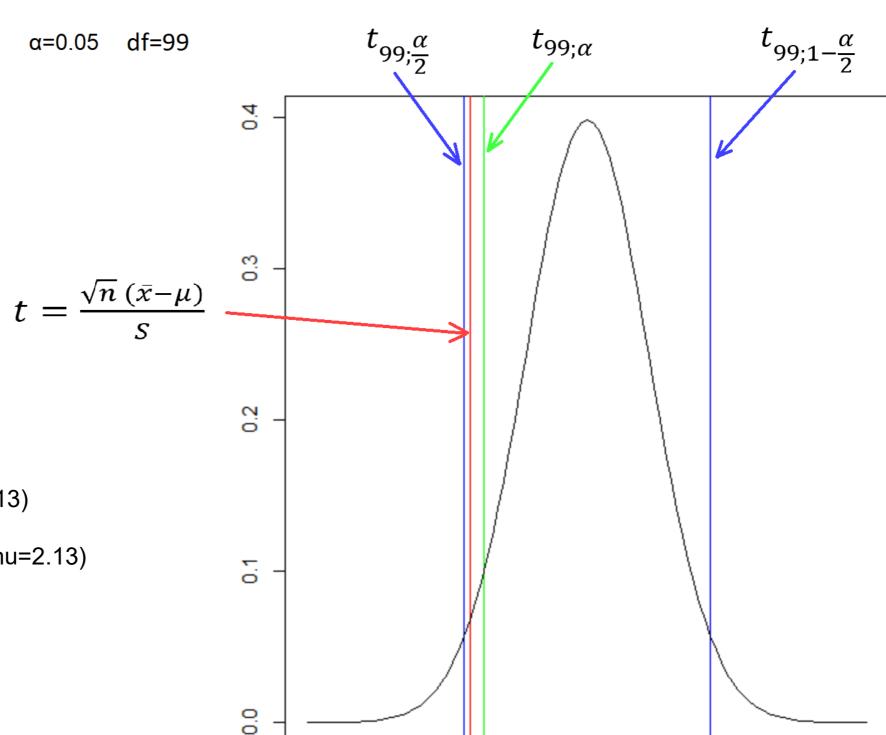


in all cases,
the test statistic
is $t = \sqrt{x} - \mu_0$ is $t = \sqrt{x} - \mu_0$ If $t \in \text{critical}$ region then
Ho is rejected;
otherwise, we
feil to reject Ho





One sample t-test – one-tailed vs. two-tailed



-2

0

x < -rnorm(100,2,2)

a=t.test(x,alternative="less",mu=2.13)

b=t.test(x,alternative="two.sided",mu=2.13)

a\$statistic is -1.877531

b\$statistic is -1.877531

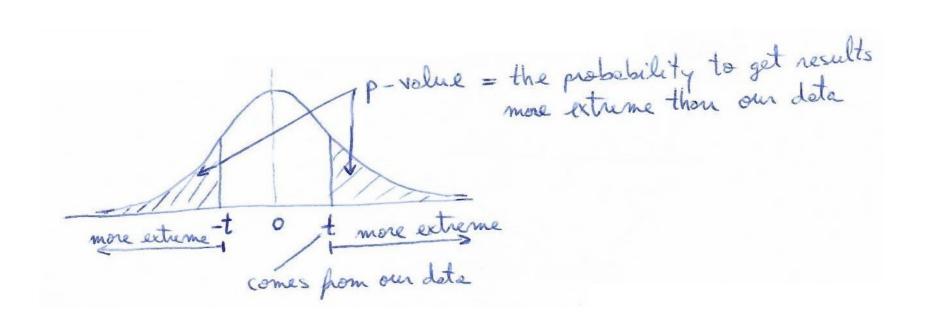
2

It gives the same result as the critical value approach when testing hypothesis (it is used by most statistical software).

The p-value is the probability that, under the assumption that the null hypothesis is true, we obtain a result "more extreme" than the one observed in our date.

P-values one triggers to decide when to reject the mull hypothesis and when to fail to reject it.

For a two-tailed T-test, let's assume that the sample date give a positive test statistic t



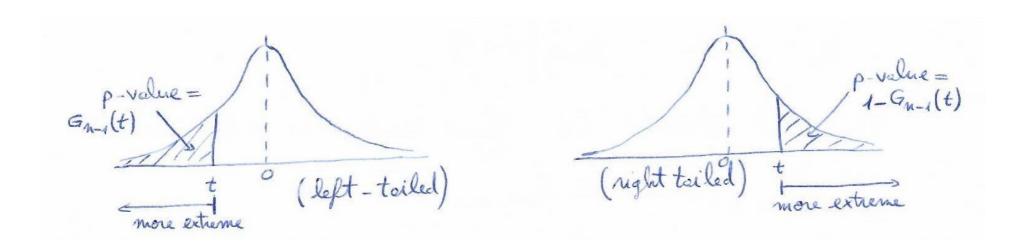
if p-volue < x then the test statistic is "extreme enough" to reject Ho

if p-volue ≥ x then we fail to reject Ho

Obs. The condition p-volue $< \infty$ is equivalent to $t \in \text{critical}$ region of the test.

For the two-tailed T-test, the p-value is as follows: $P = \begin{cases} 2 G_{m-1}(t), & \text{if } t \leq 0 \\ 2 \left(1 - G_{m-1}(t)\right), & \text{if } t > 0 \end{cases}$ where $t = \sqrt{m} \left(\frac{x}{x} - \mu_0\right)$ and G_{m-1} is the distribution function for t_{m-1}

For the left one-tailed T-test, the p-value is $G_{n-1}(t)$ and for the right one-tailed T-test, $1-G_{n-1}(t)$.



Obs. The p-volue approach lets us choose easily the significance level volue.

Ho:
$$\{T_i^2 = T_2^2\}$$
 against Ha: $\{T_i^2 \neq T_2^2\}$ (two-tailed)
(or one-tailed alternative
Ha: $\{T_i^2 > T_2^2\}$ or Ha: $\{T_i^2 < T_2^2\}$)

Let
$$S_1^2 = \frac{1}{n-1} \sum_{i=1}^{m} (X_i - \overline{X})^2$$
, $S_2^2 = \frac{1}{m-1} \sum_{i=1}^{m} (Y_i - \overline{Y})^2$

It can be proven that
$$\frac{(n-1)5^2}{\sqrt{12}} \sim \chi^2(m-1)$$
 => $\frac{(m-1)5^2}{\sqrt{12}} \sim \chi^2(m-1)$ they are independent

$$\frac{1}{m-1} \cdot \frac{(m-1)S_1^2}{\nabla_1^2} = \frac{S_1^2}{S_2^2} \cdot \frac{\nabla_2^2}{\nabla_1^2} \sim F_{m-1, m-1} \quad F \text{ distribution}$$

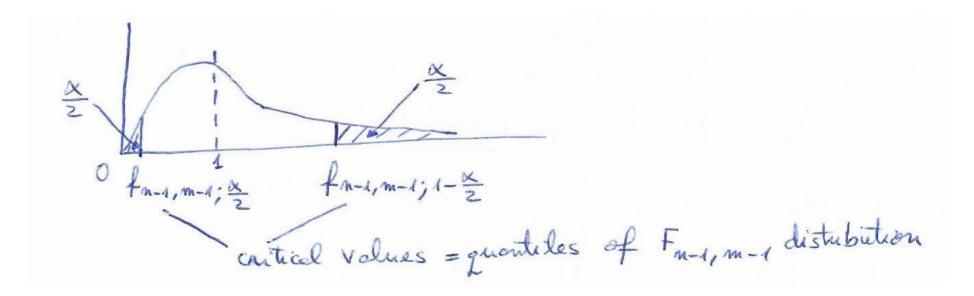
$$\frac{1}{m-1} \cdot \frac{(m-1)S_2^2}{\nabla_2^2} = \frac{S_1^2}{S_2^2} \cdot \frac{\nabla_2^2}{\nabla_1^2} \sim F_{m-1, m-1}$$

if we denote by
$$V = \frac{\overline{V_1^2}}{\overline{V_2^2}}$$
, then our testing problem con be written as:

 $H_0: \{V=1\}$
 $H_a: \{V\neq 1\}$

The test statistic is
$$\frac{S_1^2}{S_2^2}$$
. Under the null hypothesis $(V=1)$, $\frac{S_1^2}{S_2^2}$ has an $F_{m-1,m-1}$ distubution.

For the significance level x, the critical values of the two toiled F-test are:



The p-value =
$$\left(\frac{5^2}{5^2}\right)^{-1}$$
, if $\frac{5^2}{5^2} < 1$ $S_{n-1,m-1}$ is the $F_{m-1,m-1}$ $\left(\frac{5^2}{5^2}\right)^{-1}$, if $\frac{5^2}{5^2} \ge 1$ distribution function

Obs. p-value
$$< \alpha$$
 iff $\frac{5^2}{5^2} \in \overline{W}_{n,m;1-\alpha}(x=1)$

The operating characteristic function
$$OC(V) = P_{Y}(\{x_{1},...,x_{m},Y_{1},...,Y_{m}\}) \in W_{n,m;1-x}(Y=1))$$

$$= P_{Y}(\{x_{1},...,x_{m},Y_{1},...,Y_{m}\}) \in W_{n,m;1-x}(Y=1))$$

$$= P_{Y}(\{x_{1},...,x_{m},Y_{1},...,Y_{m}\}) \in W_{n,m;1-x}(Y=1)$$

$$= P_{Y}(\{x_{1},...,x_{m},Y_{1},...,Y_{m},Y_{1},...,Y_{m}\}) \in W_{n,m;1-x}(Y=1)$$

$$= P_{Y}(\{x_{1},...,x_{m},Y_{1},...,Y_{m},Y_{1},...,Y_{m},Y_{1},...,Y_{m}\}) \in W_{n,m;1-x}(Y=1)$$

$$= P_{Y}(\{x_{1},...,x_{m},Y_{1},...,Y_{m},Y_{1},...,Y_{m},Y_{1},...,Y_{m}\}) \in W_{n,m;1-x}(Y=1)$$

$$= P_{Y}(\{x_{1},...,x_{m},Y_{1},...,Y_{m},Y_{1},...,Y_{m},Y_{1},...,Y_{m},Y_{1},...,Y_{m},Y_{1},...,Y_{m}\}) = W_{n,m;1-x}(Y=1)$$

$$= P_{Y}(\{x_{1},...,x_{m},Y_{1},...,Y_{m},Y_{1},...,Y_{m},Y_{1}$$

1. The case of equal vonionces
$$T_1^2 = T_2^2 = T^2$$

it follows that
$$\overline{X} - \overline{Y} \sim N(\mu_1 - \mu_2, \overline{V}^2(\frac{1}{n} + \frac{1}{m})) \Rightarrow \overline{X} - \overline{Y} - (\mu_1 - \mu_2) \sim N(0, 1)$$

As
$$\frac{1}{T^2}((m-1)S_1^2 + (m-1)S_2^2) \sim \chi^2(m+m-2)$$
, it follows that $\frac{\overline{X} - \overline{Y} - (\mu_1 - \mu_2)}{\sqrt{1 - (1 - \mu_1)(m-1)S_1^2 + (m-1)S_2^2}} \sim t_{m+m-2}$

if we introduce
$$\delta = \mu_1 - \mu_2$$
 then the hypotheses become:
Ho: $1\delta = 01$ against Ha: $1\delta \neq 01$ (for the two-tailed wee)

The test statistic is
$$t = \frac{X - Y}{\sqrt{1 + (m-1)S_1^2 + (m-1)S_2^2}}$$

Under the null hypothesis assumption, t~t(n+m-2)

The test is built now in a similar way to the one-sample T-test

Thus, the acceptance region of Ho is
$$W_{n,m;1-\infty}(\overline{\xi}=0)=\{(x_1,...,x_n,y_1,...,y_m)|-t_{n+m-2;1-\frac{\infty}{2}}\}$$

The p-volue =
$$\begin{cases} 2 - G_{n+m-2}(t), & \text{if } t \leq 0 \\ 2(1 - G_{n+m-2}(t)), & \text{if } t > 0 \end{cases}$$

The operating characteristic function is:
$$OC(\delta) = P_{\delta} \left(-t_{n+m-2}, 1-\frac{\alpha}{2} \leq t \leq t_{n+m-2}, 1-\frac{\alpha}{2} \right)$$

$$= P_{\delta} \left(-t_{n+m-2}, 1-\frac{\alpha}{2} - \Delta \leq t_{n+m-2}, 1-\frac{\alpha}{2} - \Delta \right)$$

$$= G_{n+m-2} \left(t_{n+m-2}, 1-\frac{\alpha}{2} - \Delta \right) - G_{n+m-2} \left(t_{n+m-2}, 1-\frac{\alpha}{2} - \Delta \right),$$
where $\Delta = \frac{\delta}{\sqrt{1+m-2}} \left(\frac{1}{m} + \frac{1}{m} \right) \left((n-1) \cdot S_{1}^{2} + (m-1) \cdot S_{2}^{2} \right)$

The test statistic is
$$t' = \frac{\overline{X} - \overline{Y}}{\sqrt{\frac{S_1^2}{n} + \frac{S_2^2}{m}}}$$

If Ho is true, then
$$t' \sim t_{df}$$
, where
$$df = \frac{\left(\frac{S_{i}^{2}}{n} + \frac{S_{i}^{2}}{m}\right)^{2}}{\frac{1}{n-1}\left(\frac{S_{i}^{2}}{n}\right)^{2} + \frac{1}{m-1}\left(\frac{S_{i}^{2}}{m}\right)^{2}}$$
 (this is the Welch approximation for the degrees of freedom)

Using now the quantile tof; 1-x, the test is similar with the case of equal variances.

$$(X_1, Y_1)_1 \dots, (X_m, Y_m) \sim iid N(\mu, \mathbb{Z}), \text{ where}$$

$$\mu = \begin{pmatrix} \mu_x \\ \mu_y \end{pmatrix}, \quad \mathbb{Z} = \begin{pmatrix} \nabla_x^2 & \nabla_{xy} \\ \nabla_{xy} & \nabla_y^2 \end{pmatrix}$$

Denote
$$\delta = \mu_x - \mu_y$$
.
 $H_0: 1\delta = 0$ against $H_a: 1\delta \neq 0$

Let
$$d_i = X_i - Y_i$$
, $i = I_i \pi$

$$\nabla_{xy} = E((X - \mu_x)(Y - \mu_y)) = E(XY) - \mu_x \mu_y$$

$$Var(X - Y) = E((X - Y)^2) - [E(X - Y)]^2$$

$$= E(X^2 + Y^2 - 2XY) - [\mu_x - \mu_y]^2$$

$$= E(X^2) - \mu_x^2 + E(Y^2) - \mu_y^2 - 2E(XY) + 2\mu_x \mu_y$$

$$\nabla_{x}^2 \qquad \nabla_{y}^2$$

$$= Var(X - Y) = \nabla_x^2 + \nabla_y^2 - 2\nabla_{xy}$$

Thus,
$$d_1,...,d_m \sim iid N(J, T_d^2)$$
, where $J = \mu_x - \mu_y$,
$$T_d^2 = T_x^2 + T_y^2 - 2T_{xy}$$

The estimators for
$$J$$
 and T_d^2 are:
$$\hat{J} = \overline{d} = \frac{1}{m} \sum_{i=1}^{m} d_i$$

$$\hat{T}_d^2 = S_d^2 = \frac{1}{m-1} \sum_{i=1}^{m} (d_i - \overline{d})^2$$

The test is built in a similar way to the one-sample
$$T$$
-test. We have that $\frac{\sqrt{n}(\overline{d}-\overline{\delta})}{s_d} \sim t_{m-1}$. The test statistic is $t=\frac{\sqrt{n}}{s_d}$.

The acceptance region for Ho at the significance level x is:
$$W_{n;1-x} = \left\{ (x_1,y_1), ..., (x_n,y_n) \right| - t_{n-1;1-\frac{x}{2}} \leq \frac{\sqrt{n} d}{s_d} \leq t_{n-1;1-\frac{x}{2}} \right\}$$

$$P(Ha|Ho) = x$$

The p-value =
$$\begin{cases} 2 \cdot G_{n-1}\left(\frac{\sqrt{n} d}{sd}\right), & \text{if } \frac{\sqrt{n} d}{sd} \ge 0 \\ 2\left(1 - G_{n-1}\left(\frac{\sqrt{n} d}{sd}\right)\right), & \text{if } \frac{\sqrt{n} d}{sd} \ge 0 \end{cases}$$

The operating characteristic function is:
$$OC(\delta) = P_{\sigma}\left(-t_{n-1;1-\frac{N}{2}} \leq \frac{\sqrt{\ln d}}{s_d} \leq t_{n-1;1-\frac{N}{2}}\right)$$

$$= P_{\sigma}\left(-t_{n-1;1-\frac{N}{2}} - \frac{\sqrt{n} \delta}{s_d} \leq \frac{\sqrt{\ln (d-\delta)}}{s_d} \leq t_{n-1;1-\frac{N}{2}} - \frac{\sqrt{n} \delta}{s_d}\right)$$

$$= G_{n-1}\left(t_{n-1;1-\frac{N}{2}} - \frac{\sqrt{n} \delta}{s_d}\right) - G_{n-1}\left(-t_{n-1;1-\frac{N}{2}} - \frac{\sqrt{n} \delta}{s_d}\right)$$