

MAP 531: Homework

Paul-Antoine GIRARD & Adrien TOULOUSE

Problem 1: Estimating parameters of a Poisson distribution

We recall that the Poisson distribution with parameter $\theta > 0$ has a pdf given by $(p(\theta, k), k \in \mathbb{N})$ w.r.t the counting measure on \mathbb{N} :

$$p(\theta, k) = e^{-\theta} \frac{\theta^k}{k!}$$

Question 1

The poisson distribution is a discrete distribution since it has a countable number of possible values (\mathbb{N}).

In statistics, we use this distribution to compute the probability of a given number of (rare) events in a time period.

For example a poisson distribution can model:

- The number of patients arriving in an emergency room between 9 and 10am.
- The number of network failures per day.
- In quality control, the number of manufacturing defect.

Question 2

We assume that \mathbb{X} follows a Poisson distribution with parameter $\theta > 0$.

We will use the fact that $e^\theta = \sum_{i=0}^{\infty} (\frac{\theta^i}{i!}), \forall \theta \in \mathbb{R}$

$$\mathbb{E}[\mathbb{X}] = \sum_{i=0}^{\infty} (i * p(\theta, i)) = \sum_{i=0}^{\infty} (i * e^{-\theta} \frac{\theta^i}{i!}) = \theta * e^{-\theta} \sum_{i=1}^{\infty} (\frac{\theta^{i-1}}{(i-1)!}) = \theta * e^{-\theta} \sum_{i=0}^{\infty} (\frac{\theta^i}{i!}) = \theta * e^{-\theta} * e^\theta = \theta$$

$$\begin{aligned} \mathbb{E}[\mathbb{X}^2] &= \sum_{i=0}^{\infty} (i^2 * p(\theta, i)) = \sum_{i=0}^{\infty} (i^2 * e^{-\theta} \frac{\theta^i}{i!}) = \theta * e^{-\theta} \sum_{i=1}^{\infty} (i \frac{\theta^{i-1}}{(i-1)!}) = \theta * e^{-\theta} \sum_{i=0}^{\infty} ((i+1) \frac{\theta^i}{i!}) \\ &= \theta * e^{-\theta} [\sum_{i=0}^{\infty} (i \frac{\theta^i}{i!}) + \sum_{i=0}^{\infty} (\frac{\theta^i}{i!})] = \theta * e^{-\theta} [\theta \sum_{i=0}^{\infty} (\frac{\theta^i}{i!}) + e^\theta] = \theta * e^{-\theta} [\theta * e^\theta + e^\theta] = \theta(\theta + 1) \end{aligned}$$

$$\mathbb{V}(\mathbb{X}) = \mathbb{E}[\mathbb{X}^2] - \mathbb{E}[\mathbb{X}]^2 = \theta(\theta + 1) - \theta^2 = \theta$$

Question 3

We are provided with n independent observations of a Poisson random variable of parameter $\theta \in \Theta = \mathbb{R}_+^*$. Our observations are $X_k \sim Pois(\theta), \forall k \in 1, \dots, n$.

The corresponding statistical model is:

$$\mathcal{M}^n = (\mathbb{N}^n, \mathcal{P}(\mathbb{N}^n), \{\mathbb{P}_\theta^n, \theta \in \Theta\})$$

with $\mathbb{P}_\theta^n = \mathbb{P}_\theta \otimes \dots \otimes \mathbb{P}_\theta$ (n times)

We are trying to estimate the parameter θ .

Question 4

The likelihood function is the function on θ that makes our n observations most likely.

Using the independance of the X_k :

$$l(\theta) = \prod_{k=1}^n e^{-\theta} \frac{\theta^{X_k}}{X_k!}$$

$$L(\theta) = \log(l(\theta)) = \sum_{k=1}^n (-\theta + X_k \log(\theta) - \log(X_k!)) = -n\theta + \log(\theta) \sum_{k=1}^n X_k - \sum_{k=1}^n \log(X_k!)$$

By derivating with respect to θ , we have:

$$L'(\theta) = -n + \frac{\sum_{k=1}^n X_k}{\theta}$$

$$L''(\theta) = -\frac{\sum_{k=1}^n X_k}{\theta^2} < 0$$

Since, the second derivative of the log-likelihood function is negative, the function is concave and admits a global maximum given by:

$$L'(\theta) = 0 \Leftrightarrow -n + \frac{\sum_{k=1}^n X_k}{\theta} = 0 \Leftrightarrow \hat{\theta}_{MLE} = \bar{X}$$

So, the maximum likelihood estimator is:

$$\hat{\theta}_{MLE} = \bar{X}$$

Question 5

Since the X_k are iid, we have that:

$$\mathbb{E}[\bar{X}] = \frac{1}{n} \sum_{k=1}^n \mathbb{E}[X_k] = \mathbb{E}[X_1] = \theta$$

$$\mathbb{V}(\bar{X}) = \frac{1}{n^2} \sum_{k=1}^n \mathbb{V}(X_k) = \frac{1}{n} \mathbb{V}[X_1] = \frac{\theta}{n}$$

Applying the central limit theorem, we have that $\sqrt{n}(\hat{\theta}_{MLE} - \theta)$ converges towards a Gaussian $\mathcal{N}(0, \theta)$.

Question 6

The weak law of large numbers gives us that $\hat{\theta}_{MLE}$ converges in probability towards θ .

By continuous mapping, $\sqrt{\hat{\theta}_{MLE}}$ converges in probability towards $\sqrt{\theta}$. Then, by Slutsky's theorem, we have that $\sqrt{n} \frac{(\hat{\theta}_{MLE} - \theta)}{\sqrt{\hat{\theta}_{MLE}}}$ converges in law towards a gaussian $\mathcal{N}(0, 1)$.

Let's check this result in R by simulating 1000 times our random variable $\sqrt{n} \frac{(\hat{\theta}_{MLE} - \theta)}{\sqrt{\hat{\theta}_{MLE}}}$ with a sample size of 100:

```
estim <- function(x, theta){  
  n <- length(x)  
  est <- sqrt(n) * (mean(x) - theta) / sqrt(mean(x))  
  return(est)}  

```

```

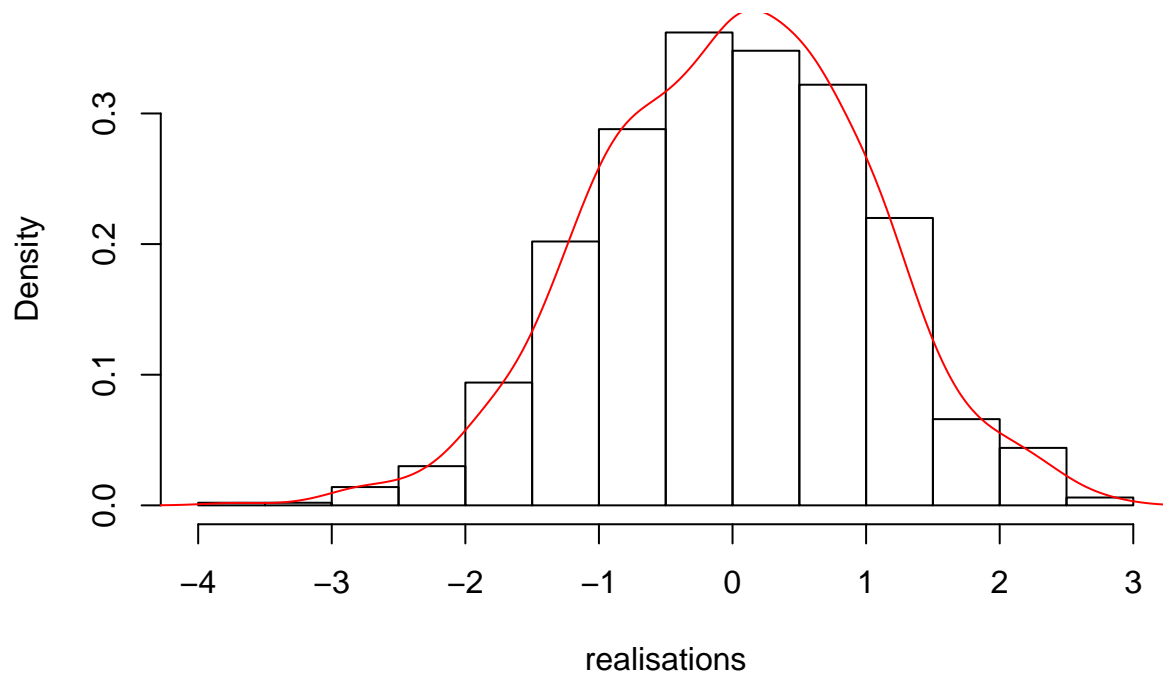
set.seed(42)
Nattempts = 1e3
nsample = 100
theta = 3

samples <- lapply(1:Nattempts, function(i) rpois(nsample, theta))
realisations <- sapply(samples, function(x) estim(x, theta))

hist(realisations, probability = TRUE)
d = density(realisations, kernel='gaussian')
lines(d, col = 'red')

```

Histogram of realisations

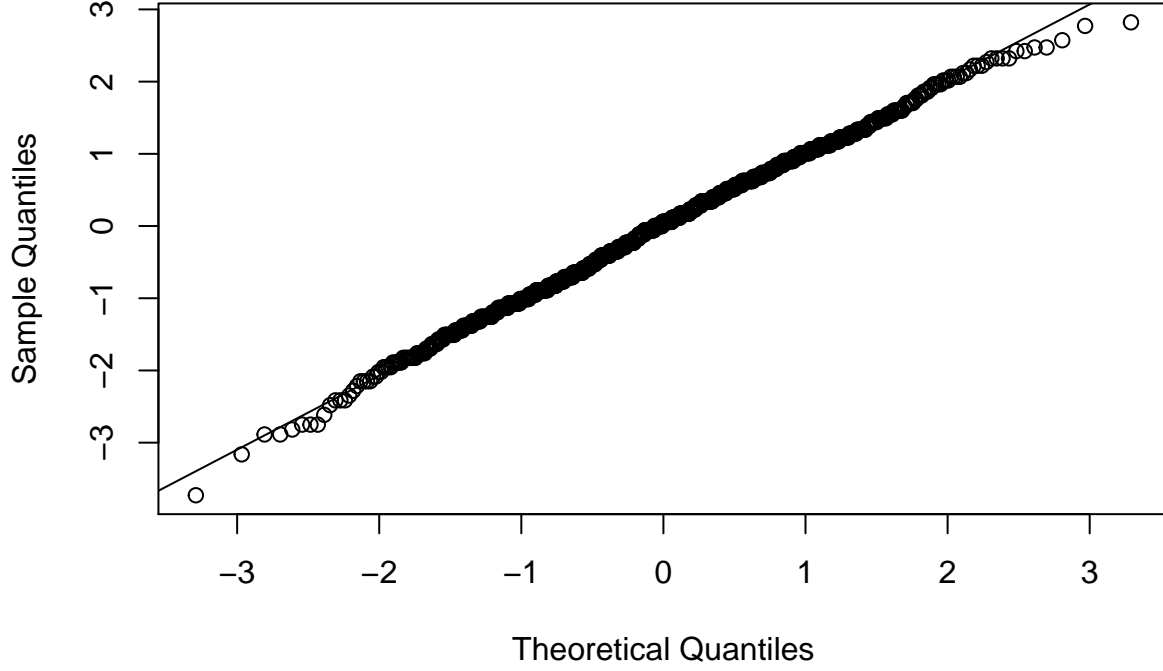


```

qqnorm(realisations)
qqline(realisations)

```

Normal Q-Q Plot



This confirms what we found theoretically: the random variable $\sqrt{n} \frac{(\hat{\theta}_{MLE} - \theta)}{\sqrt{\hat{\theta}_{MLE}}}$ follows a standard gaussian distribution.

Question 7

Let $Z_n = \sqrt{n} \frac{(\hat{\theta}_{MLE} - \theta)}{\sqrt{\hat{\theta}_{MLE}}}$ be our random variable.

Denote z_{α} the α -quantile for the standard Normal distribution for $\alpha \in (0, 1)$.

$$\lim_{n \rightarrow +\infty} \mathbb{P}(-z_{1-\alpha/2} \leq Z_n \leq z_{1-\alpha/2}) \geq 1-\alpha \Leftrightarrow \lim_{n \rightarrow +\infty} \mathbb{P}\left(-z_{1-\alpha/2} \sqrt{\frac{\hat{\theta}_{MLE}}{n}} \leq \hat{\theta}_{MLE} - \theta \leq z_{1-\alpha/2} \sqrt{\frac{\hat{\theta}_{MLE}}{n}}\right) \geq 1-\alpha$$

For $\alpha \in (0, 1)$, an asymptotic confidence interval for θ of level α is therefore :

$$\left[\hat{\theta}_{MLE} - z_{1-\alpha/2} \frac{\sqrt{\hat{\theta}_{MLE}}}{\sqrt{n}}; \hat{\theta}_{MLE} + z_{1-\alpha/2} \frac{\sqrt{\hat{\theta}_{MLE}}}{\sqrt{n}} \right]$$

Question 8

We apply the δ -method with $g(x) = 2\sqrt{x}$

We have: $g'(x) = \frac{1}{\sqrt{x}}$

So,

$$\sqrt{n}(g(\hat{\theta}_{MLE}) - g(\theta)) \xrightarrow{d} \mathcal{N}(0, g'(\theta)^2 \times \theta) \Leftrightarrow \sqrt{n}(g(\hat{\theta}_{MLE}) - g(\theta)) \xrightarrow{d} \mathcal{N}(0, 1)$$

Question 9

Let $W_n = \sqrt{n}(2\sqrt{\hat{\theta}_{MLE}} - 2\sqrt{\theta})$ be our random variable.

We know by the last question that $W_n \xrightarrow{d} \mathcal{N}(0, 1)$.

$$\begin{aligned} \lim_{n \rightarrow +\infty} \mathbb{P}(-z_{1-\alpha/2} \leq W_n \leq z_{1-\alpha/2}) &\geq 1 - \alpha \Leftrightarrow \lim_{n \rightarrow +\infty} \mathbb{P}\left(-\frac{z_{1-\alpha/2}}{2\sqrt{n}} \leq \sqrt{\hat{\theta}_{MLE}} - \sqrt{\theta} \leq \frac{z_{1-\alpha/2}}{2\sqrt{n}}\right) \geq 1 - \alpha \\ &\Leftrightarrow \lim_{n \rightarrow +\infty} \mathbb{P}\left(\sqrt{\hat{\theta}_{MLE}} - \frac{z_{1-\alpha/2}}{2\sqrt{n}} \leq \sqrt{\theta} \leq \sqrt{\hat{\theta}_{MLE}} + \frac{z_{1-\alpha/2}}{2\sqrt{n}}\right) \geq 1 - \alpha \end{aligned}$$

When n goes towards infinity, $\frac{z_{1-\alpha/2}}{2\sqrt{n}}$ goes to 0. Since $\sqrt{\hat{\theta}_{MLE}}$ is positive, there exists a n_0 such that $\sqrt{\hat{\theta}_{MLE}} - \frac{z_{1-\alpha/2}}{2\sqrt{n}}$ is positive and we can take the squares in the inequality without changing the order of the inequalities:

$$\Leftrightarrow \lim_{n \rightarrow +\infty} \mathbb{P}\left(\left(\sqrt{\hat{\theta}_{MLE}} - \frac{z_{1-\alpha/2}}{2\sqrt{n}}\right)^2 \leq \theta \leq \left(\sqrt{\hat{\theta}_{MLE}} + \frac{z_{1-\alpha/2}}{2\sqrt{n}}\right)^2\right) \geq 1 - \alpha$$

For $\alpha \in (0, 1)$, an asymptotic confidence interval for θ of level α is therefore:

$$\left[\left(\sqrt{\hat{\theta}_{MLE}} - \frac{z_{1-\alpha/2}}{2\sqrt{n}}\right)^2; \left(\sqrt{\hat{\theta}_{MLE}} + \frac{z_{1-\alpha/2}}{2\sqrt{n}}\right)^2\right]$$

Question 10

Based on the first moment of a poisson distribution, we easily have that:

$$\hat{\theta}_{MME} = \bar{X}$$

We can remark that $\hat{\theta}_{MME} = \hat{\theta}_{MLE}$

Based on the second moment of a poisson distribution, we have:

$$n^{-1} \sum_{k=1}^n X_k^2 = \hat{\theta}_2(\hat{\theta}_2 + 1)$$

Let's define the function $h(x) = x(x+1)$

Its inverse on \mathbb{R}_+^* is $h^{-1}(x) = \frac{1}{2}[-1 + \sqrt{4x+1}]$ and this gives us:

$$\hat{\theta}_2 = \frac{1}{2}[-1 + \sqrt{(4n^{-1} \sum_{k=1}^n X_k^2) + 1}]$$

Question 11

$\mathbb{E}[\hat{\theta}_{MLE}] = \frac{1}{n} \sum_{k=1}^n \mathbb{E}[X_k]$ by linearity of the expectation. So,

$$\mathbb{E}[\hat{\theta}_{MLE}] = \frac{1}{n} * n * \theta = \theta$$

Therefore, $\hat{\theta}_{MLE}$ is an unbiased estimator of θ , ie. $b_{\theta}^*(\hat{\theta}_{MLE}) = 0$

$\mathbb{V}(\hat{\theta}_{MLE}) = \frac{1}{n^2} \sum_{i=1}^n \mathbb{V}(X_i)$ by independance of the X_k .

$$\mathbb{V}(\hat{\theta}_{MLE}) = \frac{1}{n^2} * n * \theta = \frac{\theta}{n}$$

The quadratic risk Q is:

$$Q = b_{\theta}^*(\hat{\theta}_{MLE})^2 + \mathbb{V}^*(\hat{\theta}_{MLE}) = 0 + \frac{\theta}{n} = \frac{\theta}{n}$$

Question 12

$\hat{\theta}_{MLE}$ is an unbiased estimator. So the Cramer-Rao bound is given by:

$$\frac{1}{I_n(\theta^*)} = \frac{1}{\mathbb{E}[-L''(\theta^*)]}$$

By derivating the log-likelihood function with respect to θ , we have:

$$\begin{aligned} L'(\theta^*) &= -n + \frac{\sum_{i=1}^n X_k}{\theta} \\ -L''(\theta^*) &= \frac{\sum_{i=1}^n X_k}{\theta^2} \end{aligned}$$

Therefore,

$$\mathbb{E}[-L''(\theta^*)] = \frac{\sum_{i=1}^n \mathbb{E}[X_k]}{\theta^2} = \frac{n}{\theta}$$

Finally,

$$\frac{1}{I_n(\theta^*)} = \frac{\theta}{n} = \mathbb{V}(\hat{\theta}_{MLE})$$

We can conclude that our estimator $\hat{\theta}_{MLE}$ is efficient.

Question 13

$$\begin{aligned} \hat{\theta}_2 &= \frac{1}{n} \sum_{i=1}^n (X_i - \bar{X}_n)^2 = \frac{1}{n} \sum_{i=1}^n (X_i - \theta + \theta - \bar{X}_n)^2 = \frac{1}{n} \sum_{i=1}^n [(X_i - \theta)^2 + (\theta - \bar{X}_n)^2 + 2(X_i - \theta)(\theta - \bar{X}_n)] \\ &= \frac{1}{n} \sum_{i=1}^n (X_i - \theta)^2 + (\theta - \bar{X}_n)^2 + \frac{2}{n} (\theta - \bar{X}_n) \sum_{i=1}^n (X_i - \theta) = \frac{1}{n} \sum_{i=1}^n (X_i - \theta)^2 + (\theta - \bar{X}_n)^2 + 2(\theta - \bar{X}_n)(\bar{X}_n - \theta) \\ &= \frac{1}{n} \sum_{i=1}^n (X_i - \theta)^2 - (\theta - \bar{X}_n)^2 \end{aligned}$$

Question 14

$$\begin{aligned} \mathbb{E}[(\theta - \bar{X}_n)^2] &= \mathbb{E}[\theta^2 - 2\theta\bar{X}_n + \bar{X}_n^2] = \theta^2 - 2\theta\mathbb{E}[\bar{X}_n] + \mathbb{E}[\bar{X}_n^2] \\ &= -\theta^2 + \mathbb{V}(\bar{X}_n) + \mathbb{E}[\bar{X}_n^2] = -\theta^2 + \frac{\theta}{n} + \theta^2 = \frac{\theta}{n} \end{aligned}$$

$$\begin{aligned} \mathbb{E}[\hat{\theta}_2] &= \mathbb{E}\left[\frac{1}{n} \sum_{i=1}^n (X_i - \theta)^2 - (\theta - \bar{X}_n)^2\right] = \frac{1}{n} \sum_{i=1}^n \mathbb{E}[(X_i - \theta)^2] - \mathbb{E}[(\theta - \bar{X}_n)^2] \\ &= \frac{1}{n} \sum_{i=1}^n \mathbb{V}(X_i) - \frac{\theta}{n} = \theta\left(1 - \frac{1}{n}\right) \end{aligned}$$

Therefore the bias is:

$$b_{\hat{\theta}_2} = -\frac{\theta}{n}$$

We can get an unbiased estimator $\hat{\theta}_3$ by defining $\hat{\theta}_3 = (1 - \frac{1}{n})^{-1}\hat{\theta}_2$

Question 15

Using the previous questions, we know that:

$$\hat{\theta}_2 = \frac{1}{n} \sum_{i=1}^n (X_i - \theta)^2 - (\theta - \bar{X}_n)^2$$

therefore, we have:

$$\sqrt{n}(\hat{\theta}_2 - \theta) = \frac{1}{\sqrt{n}} \sum_{i=1}^n (X_i - \theta)^2 - \sqrt{n}(\theta - \bar{X}_n)^2 - \sqrt{n}\theta = \sqrt{n}(\bar{Y}_n - \theta) - \sqrt{n}(\theta - \bar{X}_n)^2$$

where:

$$\begin{aligned} \forall i \in \llbracket 1, n \rrbracket, Y_i &= (X_i - \theta)^2 \\ \bar{Y}_n &= \frac{1}{n} \sum_{i=1}^n Y_i \end{aligned}$$

Since:

$$\mathbb{E}[Y_i] = \mathbb{V}(X_i) = \theta$$

and

$$\mathbb{V}(Y_i) = 2\theta^2 + \theta$$

We can apply the central limit theorem, and we have that $\sqrt{n}(\bar{Y}_n - \theta)$ converges towards a Gaussian $\mathcal{N}(0, 2\theta^2 + \theta)$.

We also have that:

$$\sqrt{n}(\theta - \bar{X}_n)^2 = \sqrt{n}(\bar{X}_n - \theta)^2 = \sqrt{n}(\bar{X}_n - \theta)(\bar{X}_n - \theta)$$

Applying the central limit theorem, we have that $\sqrt{n}(\bar{X}_n - \theta)$ converges towards a Gaussian $\mathcal{N}(0, \theta)$.

On the other hand, applying the law of large numbers: $(\theta - \bar{X}_n)$ converges in probability towards 0.

Applying Slutsky's theorem, $\sqrt{n}(\theta - \bar{X}_n)^2$ converges in distribution towards the constant 0. Therefore, it converges in probability towards 0.

Now, we can apply Slutsky's theorem to $\sqrt{n}(\bar{Y}_n - \theta) - \sqrt{n}(\theta - \bar{X}_n)^2$ which gives us finally that $\sqrt{n}(\hat{\theta}_2 - \theta)$ converges in distribution towards a Gaussian $\mathcal{N}(0, 2\theta^2 + \theta)$.

We can now compute an other asymptotic confidence interval centered in $\hat{\theta}_2$.

Let $V_n = \sqrt{n}(\hat{\theta}_2 - \theta)$ be our random variable.

We know by the last question that $\frac{V_n}{\sqrt{2\theta^2 + \theta}} \xrightarrow{d} \mathcal{N}(0, 1)$.

$$\lim_{n \rightarrow +\infty} \mathbb{P}(-z_{1-\alpha/2} \leq \frac{V_n}{\sqrt{2\theta^2 + \theta}} \leq z_{1-\alpha/2}) \geq 1 - \alpha \Leftrightarrow \lim_{n \rightarrow +\infty} \mathbb{P}(** \leq ** \leq **) \geq 1 - \alpha$$

$$\Leftrightarrow \lim_{n \rightarrow +\infty} \mathbb{P}(** \leq ** \leq **) \geq 1 - \alpha$$

For $\alpha \in (0, 1)$, an asymptotic confidence interval for θ of level α is therefore:

$$[**; **]$$

Question 16

Let $s \in \mathbb{R}$. The probability generating function of the Poisson distribution is given by:

$$G_{\mathbb{X}}(s) = \mathbb{E}[\exp(s\mathbb{X})] = \sum_{k=0}^{\infty} e^{ks} e^{-\theta} \frac{\theta^k}{k!} = e^{-\theta} \sum_{k=0}^{\infty} \frac{(\theta e^s)^k}{k!} = e^{-\theta} e^{\theta e^s} = e^{\theta(e^s - 1)}$$

In order to compute the first and second moment of the Poisson distribution, we can now use the moment generating function. Let's compute its first and second order derivatives.

$$\begin{aligned} G'_{\mathbb{X}}(s) &= \theta e^s e^{\theta(e^s - 1)} \\ G''_{\mathbb{X}}(s) &= \theta[e^s e^{\theta(e^s - 1)} + \theta e^{2s} e^{\theta(e^s - 1)}] = \theta e^s [e^{\theta(e^s - 1)} + \theta e^{\theta(e^s - 1)}] \end{aligned}$$

Then, we have:

$$\begin{aligned} \mathbb{E}[\mathbb{X}] &= G'_{\mathbb{X}}(0) = \theta \\ \mathbb{E}[\mathbb{X}^2] &= G''_{\mathbb{X}}(0) = \theta(1 + \theta) \\ \mathbb{V}(\mathbb{X}) &= \mathbb{E}[\mathbb{X}^2] - \mathbb{E}[\mathbb{X}]^2 = \theta(1 + \theta) - \theta^2 = \theta \end{aligned}$$

We will now show that: $\mathbb{V}[(\mathbb{X}_i - \theta)^2] = 2\theta^2 + \theta$

$$\begin{aligned} G_{\mathbb{X}}^{(3)}(s) &= (1 + 3\theta e^s + \theta^2 e^{2s}) \theta e^{s + \theta(e^s - 1)} \\ G_{\mathbb{X}}^{(4)}(s) &= (1 + \theta^3 e^{3s} + 6\theta^2 e^{2s} + 7\theta e^s) \theta e^{s + \theta(e^s - 1)} \\ \mathbb{E}[\mathbb{X}^3] &= G_{\mathbb{X}}^{(3)}(0) = \theta + 3\theta^2 + \theta^3 \\ \mathbb{E}[\mathbb{X}^4] &= G_{\mathbb{X}}^{(4)}(0) = \theta^4 + 6\theta^3 + 7\theta^2 + \theta \\ \mathbb{V}[(\mathbb{X}_i - \theta)^2] &= \mathbb{E}[(\mathbb{X} - \theta)^4] - \mathbb{E}[(\mathbb{X} - \theta)^2]^2 = \mathbb{E}[\mathbb{X}^4] - 4\theta \mathbb{E}[\mathbb{X}^3] + 6\theta^2 \mathbb{E}[\mathbb{X}^2] - 4\theta^3 \mathbb{E}[\mathbb{X}] + \theta^4 - \text{Var}(\mathbb{X})^2 \\ &= \theta^4 + 6\theta^3 + 7\theta^2 + \theta - 4\theta(\theta + 3\theta^2 + \theta^3) + 6\theta^2(\theta + \theta^2) - 4\theta^4 + \theta^4 - \theta^2 = 2\theta^2 + \theta \end{aligned}$$

Problem 2: Analysis of the USJudgeRatings dataset

This exercise is open. You are asked to use the tools we have seen together to analyze the USJudgeRatings data set. This data set is provided in the package datasets. Your analysis should be reported here and include:

- an introduction
- a general description of the data
- the use of descriptive statistics
- the use of all techniques we have seen together that might be relevant
- a conclusion

Overall, your analysis, including the graphs and the codes should not exceed 15 pages in pdf.

Introduction

We are given to analyse a dataset, named USJudgeratings, containing various ratings of state judges in the US Superior Court made by lawyers. The different variables given help us to determine if a judge is worthy staying in the US Superior Court or not. We will start by doing a general description of the data and applying descriptive statistics to better apprehend the data.

General description

We start by uploading our data.

```
data(USJudgeRatings)
```

First, let's see how the dataset is organized.

```
str(USJudgeRatings)
```

```
## 'data.frame':   43 obs. of  12 variables:
## $ CONT: num   5.7 6.8 7.2 6.8 7.3 6.2 10.6 7 7.3 8.2 ...
## $ INTG: num   7.9 8.9 8.1 8.8 6.4 8.8 9 5.9 8.9 7.9 ...
## $ DMNR: num   7.7 8.8 7.8 8.5 4.3 8.7 8.9 4.9 8.9 6.7 ...
## $ DILG: num   7.3 8.5 7.8 8.8 6.5 8.5 8.7 5.1 8.7 8.1 ...
## $ CFMG: num   7.1 7.8 7.5 8.3 6 7.9 8.5 5.4 8.6 7.9 ...
## $ DECI: num   7.4 8.1 7.6 8.5 6.2 8 8.5 5.9 8.5 8 ...
## $ PREP: num   7.1 8 7.5 8.7 5.7 8.1 8.5 4.8 8.4 7.9 ...
## $ FAMI: num   7.1 8 7.5 8.7 5.7 8 8.5 5.1 8.4 8.1 ...
## $ ORAL: num   7.1 7.8 7.3 8.4 5.1 8 8.6 4.7 8.4 7.7 ...
## $ WRIT: num   7 7.9 7.4 8.5 5.3 8 8.4 4.9 8.5 7.8 ...
## $ PHYS: num   8.3 8.5 7.9 8.8 5.5 8.6 9.1 6.8 8.8 8.5 ...
## $ RTEN: num   7.8 8.7 7.8 8.7 4.8 8.6 9 5 8.8 7.9 ...
```

The data is stored in a dataframe.

```
dim(USJudgeRatings)
```

```
## [1] 43 12
```

We are provided with $n = 43$ observations and $p = 12$ quantitative variables.

We can have a full view of the dataset by using the kable function:

```
library(knitr)
library(kableExtra)
kable(USJudgeRatings, 'latex', caption = "Ratings of US judges", booktabs = T) %>%
  kable_styling(latex_options = "striped", font_size = 6)
```

The variables are all numeric. An observation in this dataset represents the different ratings received by a judge (given by his name) in the US Superior Court. Let's define properly what each variable means.

```
colnames(USJudgeRatings)
```

```
## [1] "CONT" "INTG" "DMNR" "DILG" "CFMG" "DECI" "PREP" "FAMI" "ORAL" "WRIT"
## [11] "PHYS" "RTEN"
```

The variables are:

- *CONT* : The number of contacts of the lawyer with judge.
- *INTG* : The judicial integrity of the judge.
- *DMNR* : Demeanor of the judge.
- *DILG* : Diligence of the judge.
- *CFMG* : Case flow managed by the judge.
- *DECI* : Prompt decisions taken by the judge.
- *PREP* : How the judge is prepared trial.
- *FAMI* : The judge's familiarity with law.
- *ORAL* : The judge's sound oral rulings.
- *WRIT* : The judge's sound written rulings.
- *PHYS* : The judge's physical ability.
- *RTEN* : Scaling if the judge is worthy to retain in the US Superior court.

Table 1: Ratings of US judges

	CONT	INTG	DMNR	DILG	CFMG	DECI	PREP	FAMI	ORAL	WRIT	PHYS	RTEN
AARONSON,L.H.	5.7	7.9	7.7	7.3	7.1	7.4	7.1	7.1	7.1	7.0	8.3	7.8
ALEXANDER,J.M.	6.8	8.9	8.8	8.5	7.8	8.1	8.0	8.0	7.8	7.9	8.5	8.7
ARMENTANO,A.J.	7.2	8.1	7.8	7.8	7.5	7.6	7.5	7.5	7.3	7.4	7.9	7.8
BERDON,R.I.	6.8	8.8	8.5	8.8	8.3	8.5	8.7	8.7	8.4	8.5	8.8	8.7
BRACKEN,J.J.	7.3	6.4	4.3	6.5	6.0	6.2	5.7	5.7	5.1	5.3	5.5	4.8
BURNS,E.B.	6.2	8.8	8.7	8.5	7.9	8.0	8.1	8.0	8.0	8.0	8.6	8.6
CALLAHAN,R.J.	10.6	9.0	8.9	8.7	8.5	8.5	8.5	8.5	8.6	8.4	9.1	9.0
COHEN,S.S.	7.0	5.9	4.9	5.1	5.4	5.9	4.8	5.1	4.7	4.9	6.8	5.0
DALY,J.J.	7.3	8.9	8.9	8.7	8.6	8.5	8.4	8.4	8.4	8.5	8.8	8.8
DANNEHY,J.F.	8.2	7.9	6.7	8.1	7.9	8.0	7.9	8.1	7.7	7.8	8.5	7.9
DEAN,H.H.	7.0	8.0	7.6	7.4	7.3	7.5	7.1	7.2	7.1	7.2	8.4	7.7
DEVITA,H.J.	6.5	8.0	7.6	7.2	7.0	7.1	6.9	7.0	7.0	7.1	6.9	7.2
DRISCOLL,P.J.	6.7	8.6	8.2	6.8	6.9	6.6	7.1	7.3	7.2	7.2	8.1	7.7
GRILLO,A.E.	7.0	7.5	6.4	6.8	6.5	7.0	6.6	6.8	6.3	6.6	6.2	6.5
HADDEN,W.L.JR.	6.5	8.1	8.0	8.0	7.9	8.0	7.9	7.8	7.8	7.8	8.4	8.0
HAMILL,E.C.	7.3	8.0	7.4	7.7	7.3	7.3	7.3	7.2	7.1	7.2	8.0	7.6
HEALEY,A.H.	8.0	7.6	6.6	7.2	6.5	6.5	6.8	6.7	6.4	6.5	6.9	6.7
HULL,T.C.	7.7	7.7	6.7	7.5	7.4	7.5	7.1	7.3	7.1	7.3	8.1	7.4
LEVINE,I.	8.3	8.2	7.4	7.8	7.7	7.7	7.7	7.8	7.5	7.6	8.0	8.0
LEVISTER,R.L.	9.6	6.9	5.7	6.6	6.9	6.6	6.2	6.0	5.8	5.8	7.2	6.0
MARTIN,L.F.	7.1	8.2	7.7	7.1	6.6	6.6	6.7	6.7	6.8	6.8	7.5	7.3
MCGRATH,J.F.	7.6	7.3	6.9	6.8	6.7	6.8	6.4	6.3	6.3	6.3	7.4	6.6
MIGNONE,A.F.	6.6	7.4	6.2	6.2	5.4	5.7	5.8	5.9	5.2	5.8	4.7	5.2
MISSAL,H.M.	6.2	8.3	8.1	7.7	7.4	7.3	7.3	7.3	7.2	7.3	7.8	7.6
MULVEY,H.M.	7.5	8.7	8.5	8.6	8.5	8.4	8.5	8.5	8.4	8.4	8.7	8.7
NARUK,H.J.	7.8	8.9	8.7	8.9	8.7	8.8	8.9	9.0	8.8	8.9	9.0	9.0
O'BRIEN,F.J.	7.1	8.5	8.3	8.0	7.9	7.9	7.8	7.8	7.8	7.7	8.3	8.2
O'SULLIVAN,T.J.	7.5	9.0	8.9	8.7	8.4	8.5	8.4	8.3	8.3	8.3	8.8	8.7
PASKEY,L.	7.5	8.1	7.7	8.2	8.0	8.1	8.2	8.4	8.0	8.1	8.4	8.1
RUBINOW,J.E.	7.1	9.2	9.0	9.0	8.4	8.6	9.1	9.1	8.9	9.0	8.9	9.2
SADEN,G.A.	6.6	7.4	6.9	8.4	8.0	7.9	8.2	8.4	7.7	7.9	8.4	7.5
SATANIELLO,A.G.	8.4	8.0	7.9	7.9	7.8	7.8	7.6	7.4	7.4	7.4	8.1	7.9
SHEA,D.M.	6.9	8.5	7.8	8.5	8.1	8.2	8.4	8.5	8.1	8.3	8.7	8.3
SHEA,J.F.JR.	7.3	8.9	8.8	8.7	8.4	8.5	8.5	8.5	8.4	8.4	8.8	8.8
SIDOR,W.J.	7.7	6.2	5.1	5.6	5.6	5.9	5.6	5.6	5.3	5.5	6.3	5.3
SPEZIALE,J.A.	8.5	8.3	8.1	8.3	8.4	8.2	8.2	8.1	7.9	8.0	8.0	8.2
SPONZO,M.J.	6.9	8.3	8.0	8.1	7.9	7.9	7.9	7.7	7.6	7.7	8.1	8.0
STAPLETON,J.F.	6.5	8.2	7.7	7.8	7.6	7.7	7.7	7.7	7.5	7.6	8.5	7.7
TESTO,R.J.	8.3	7.3	7.0	6.8	7.0	7.1	6.7	6.7	6.7	6.7	8.0	7.0
TIERNEY,W.L.JR.	8.3	8.2	7.8	8.3	8.4	8.3	7.7	7.6	7.5	7.7	8.1	7.9
WALL,R.A.	9.0	7.0	5.9	7.0	7.0	7.2	6.9	6.9	6.5	6.6	7.6	6.6
WRIGHT,D.B.	7.1	8.4	8.4	7.7	7.5	7.7	7.8	8.2	8.0	8.1	8.3	8.1
ZARRILLI,K.J.	8.6	7.4	7.0	7.5	7.5	7.7	7.4	7.2	6.9	7.0	7.8	7.1

Descriptive dataset analysis

```
summary(USJudgeRatings)
```

```
##           CONT           INTG           DMNR           DILG
## Min.      : 5.700   Min.      :5.900   Min.      :4.300   Min.      :5.100
## 1st Qu.: 6.850   1st Qu.:7.550   1st Qu.:6.900   1st Qu.:7.150
## Median : 7.300   Median :8.100   Median :7.700   Median :7.800
## Mean     : 7.437   Mean     :8.021   Mean     :7.516   Mean     :7.693
## 3rd Qu.: 7.900   3rd Qu.:8.550   3rd Qu.:8.350   3rd Qu.:8.450
## Max.     :10.600   Max.     :9.200   Max.     :9.000   Max.     :9.000
##           CFMG           DECI           PREP           FAMI
## Min.      :5.400   Min.      :5.700   Min.      :4.800   Min.      :5.100
## 1st Qu.:7.000   1st Qu.:7.100   1st Qu.:6.900   1st Qu.:6.950
## Median :7.600   Median :7.700   Median :7.700   Median :7.600
## Mean     :7.479   Mean     :7.565   Mean     :7.467   Mean     :7.488
## 3rd Qu.:8.050   3rd Qu.:8.150   3rd Qu.:8.200   3rd Qu.:8.250
## Max.     :8.700   Max.     :8.800   Max.     :9.100   Max.     :9.100
##           ORAL           WRIT           PHYS           RTEN
## Min.      :4.700   Min.      :4.900   Min.      :4.700   Min.      :4.800
## 1st Qu.:6.850   1st Qu.:6.900   1st Qu.:7.700   1st Qu.:7.150
## Median :7.500   Median :7.600   Median :8.100   Median :7.800
## Mean     :7.293   Mean     :7.384   Mean     :7.935   Mean     :7.602
## 3rd Qu.:8.000   3rd Qu.:8.050   3rd Qu.:8.500   3rd Qu.:8.250
## Max.     :8.900   Max.     :9.000   Max.     :9.100   Max.     :9.200
```

All the variables (except the variable CONT) seem to be ranged between 0 and 10.

The last variable, RTEN, is our response variable. It says if the lawyers think that the judge is worthy staying in the US Superior Court or not.

We can observe that each variable seems to follow a symmetric distribution, since median and mean are always close.

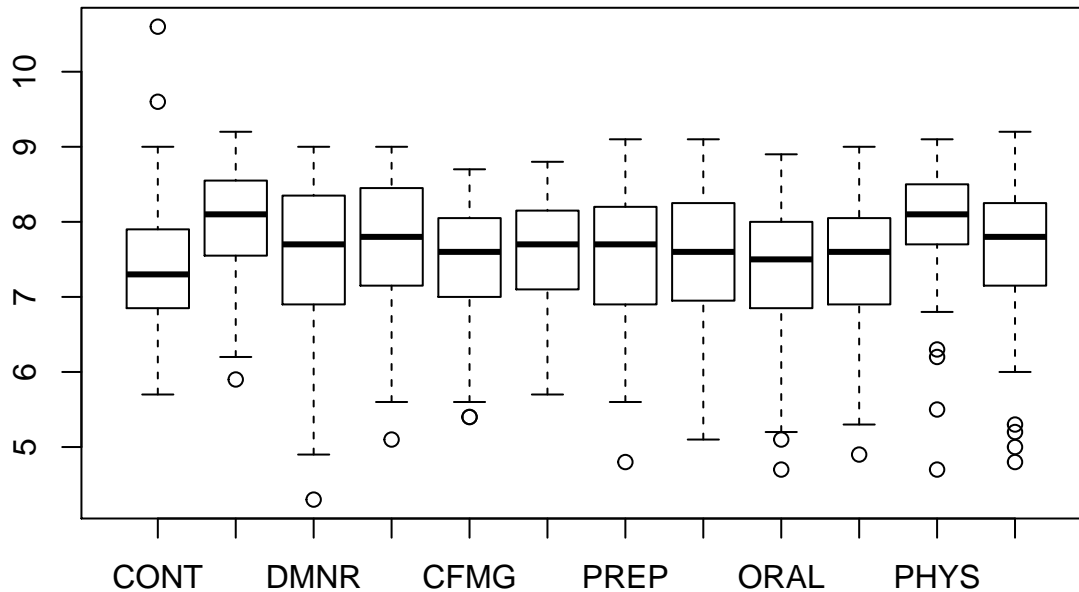
Let's inspect the dataframe for missing values, outliers and errors:

```
sum(is.na(USJudgeRatings))
```

```
## [1] 0
```

There are no missing values in the dataframe.

```
Outvals = boxplot(USJudgeRatings)
```



We observe the presence of outliers for 10 of the 12 variables (with larger values for CONT and with lower values for the other variables).

Let's take a closer look at these outliers.

```
max(USJudgeRatings$CONT)
```

```
## [1] 10.6
```

```
rownames(USJudgeRatings)[which.max(USJudgeRatings$CONT)]
```

```
## [1] "CALLAHAN,R.J."
```

The judge with the biggest number of contacts of lawyer is judge Callahan with a a number of 10.6 contacts.

```
min(USJudgeRatings$RTEN)
```

```
## [1] 4.8
```

```
rownames(USJudgeRatings)[which.min(USJudgeRatings$RTEN)]
```

```
## [1] "BRACKEN,J.J."
```

The judge with the lowest rating for worthiness of retention is judge Bracken with a rating of 4.8.

```
max(USJudgeRatings$RTEN)
```

```
## [1] 9.2
```

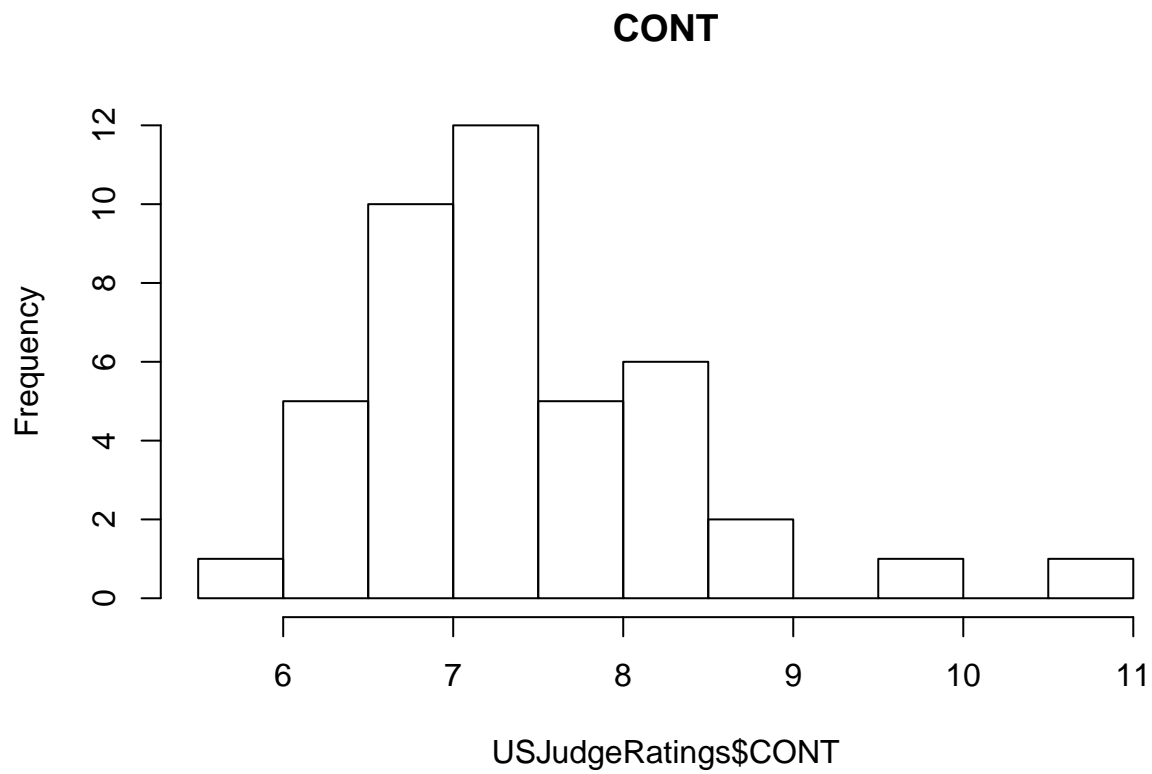
```
rownames(USJudgeRatings)[which.max(USJudgeRatings$RTEN)]
```

```
## [1] "RUBINOW,J.E."
```

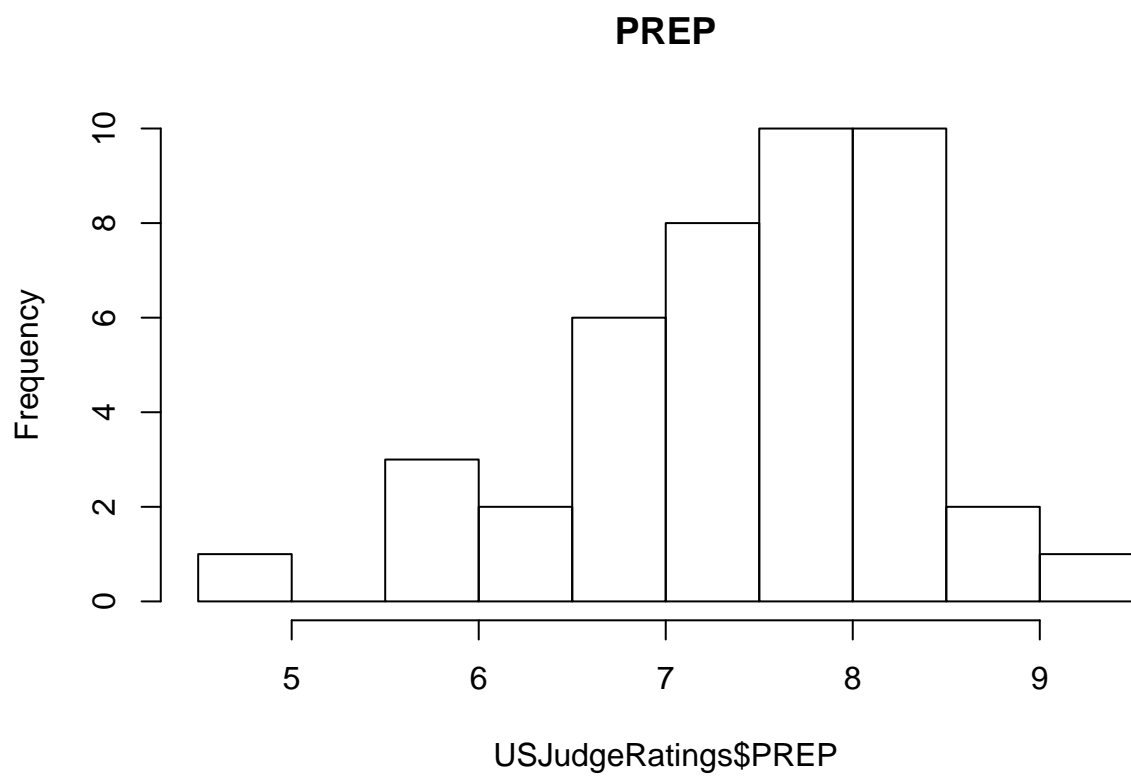
The judge with the highest rating for worthiness of retention is judge Rubinow with a rating of 9.2.

We are not provided with extra information and nothing indicated that these outliers correspond to mistakes. Thus, we will assume that they aren't mistakes and keep them in our analysis.

```
hist(USJudgeRatings$CONT, main="CONT")
```

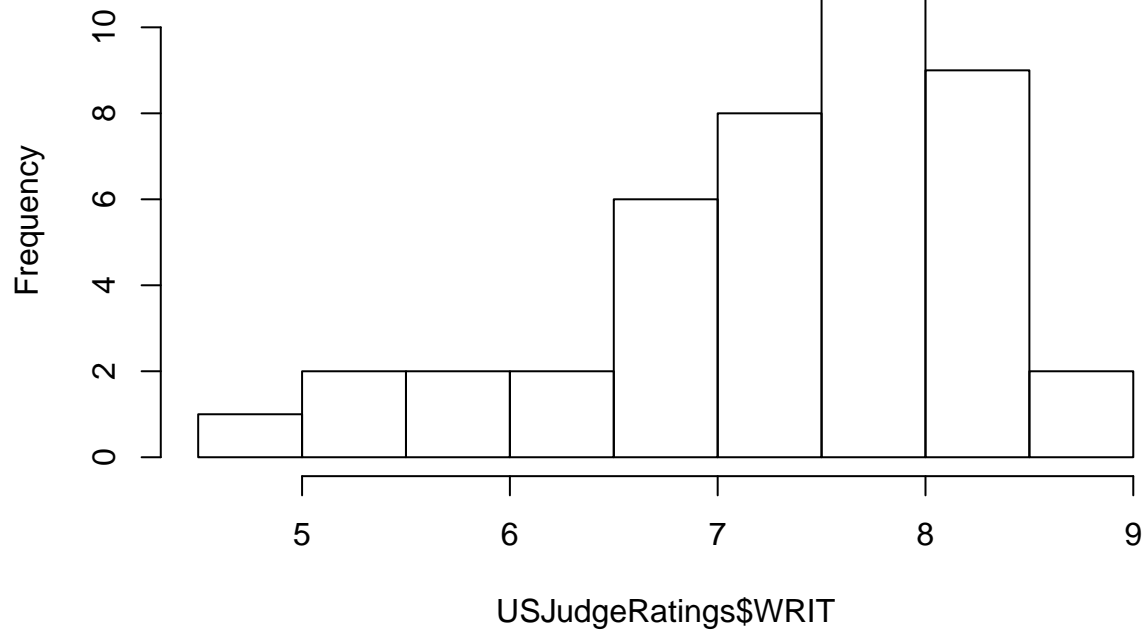


```
hist(USJudgeRatings$PREP, main="PREP" )
```



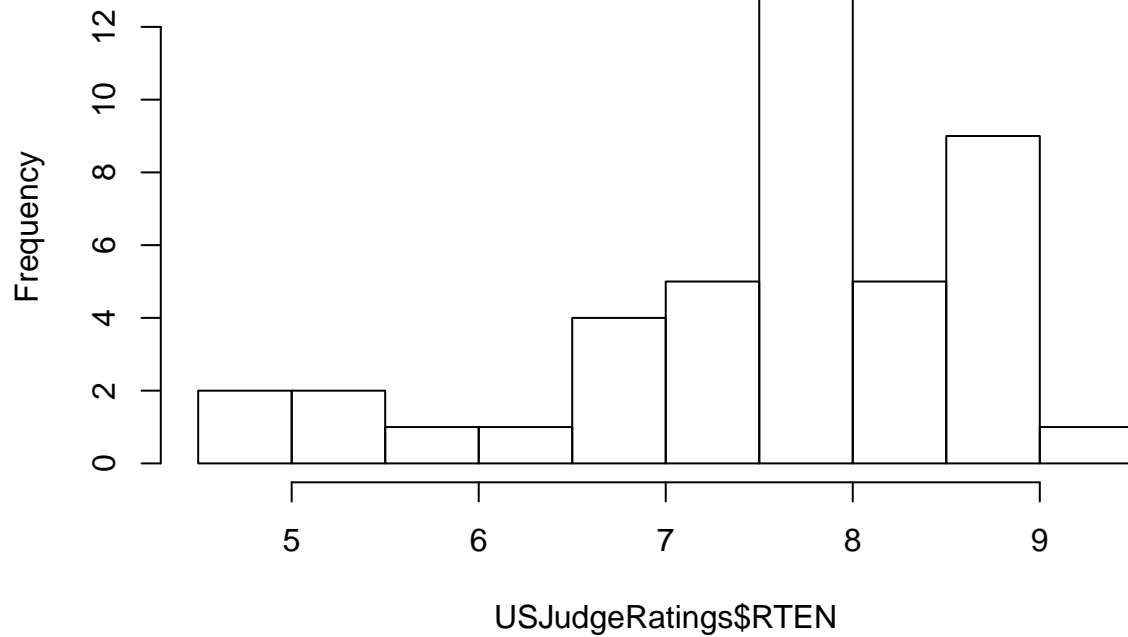
```
hist(USJudgeRatings$WRIT, main="WRIT")
```

WRIT



```
hist(USJudgeRatings$RTEN, main="RTEN")
```

RTEN



```
round(var(USJudgeRatings), 2)
```

```
##      CONT  INTG  DMNR  DILG  CFMG  DECI  PREP  FAMI  ORAL  WRIT  PHYS  RTEN
## CONT  0.89 -0.10 -0.17  0.01  0.11  0.07  0.01 -0.02 -0.01 -0.04  0.05 -0.03
```

```
## INTG -0.10 0.59 0.85 0.60 0.54 0.50 0.64 0.64 0.71 0.67 0.54 0.79
## DMNR -0.17 0.85 1.31 0.86 0.80 0.74 0.93 0.91 1.05 0.98 0.85 1.19
## DILG 0.01 0.60 0.86 0.81 0.74 0.69 0.84 0.82 0.87 0.83 0.69 0.92
## CFMG 0.11 0.54 0.80 0.74 0.74 0.68 0.79 0.76 0.83 0.78 0.71 0.88
## DECI 0.07 0.50 0.74 0.69 0.68 0.64 0.73 0.72 0.77 0.73 0.66 0.82
## PREP 0.01 0.64 0.93 0.84 0.79 0.73 0.91 0.90 0.95 0.90 0.76 1.00
## FAMI -0.02 0.64 0.91 0.82 0.76 0.72 0.90 0.90 0.94 0.90 0.75 0.98
## ORAL -0.01 0.71 1.05 0.87 0.83 0.77 0.95 0.94 1.02 0.96 0.85 1.09
## WRIT -0.04 0.67 0.98 0.83 0.78 0.73 0.90 0.90 0.96 0.92 0.77 1.02
## PHYS 0.05 0.54 0.85 0.69 0.71 0.66 0.76 0.75 0.85 0.77 0.88 0.94
## RTEN -0.03 0.79 1.19 0.92 0.88 0.82 1.00 0.98 1.09 1.02 0.94 1.21
```

```
round(sqrt(diag(var(USJudgeRatings))),2)
```

```
## CONT INTG DMNR DILG CFMG DECI PREP FAMI ORAL WRIT PHYS RTEN
## 0.94 0.77 1.14 0.90 0.86 0.80 0.95 0.95 1.01 0.96 0.94 1.10
```

```
print('The smallest standard deviation is: ')
```

```
## [1] "The smallest standard deviation is: "
```

```
min(round(sqrt(diag(var(USJudgeRatings))),2))
```

```
## [1] 0.77
```

```
print('The largest standard deviation is: ')
```

```
## [1] "The largest standard deviation is: "
```

```
max(round(sqrt(diag(var(USJudgeRatings))),2))
```

```
## [1] 1.14
```

Regarding the dispersion, we look at the interquartile range (given by the boxplots) and the empirical standard deviation. Overall, the dispersions are not very high (around 1). We find that the variables DMNR and RTEN have the largest standard deviation, while the DECI variable has the smallest.

Let's measure the correlations between the 11 first variables and the variable RTEN. For this we use the correlations function and the pairs function to visualize the scatter plots of the variables two by two.

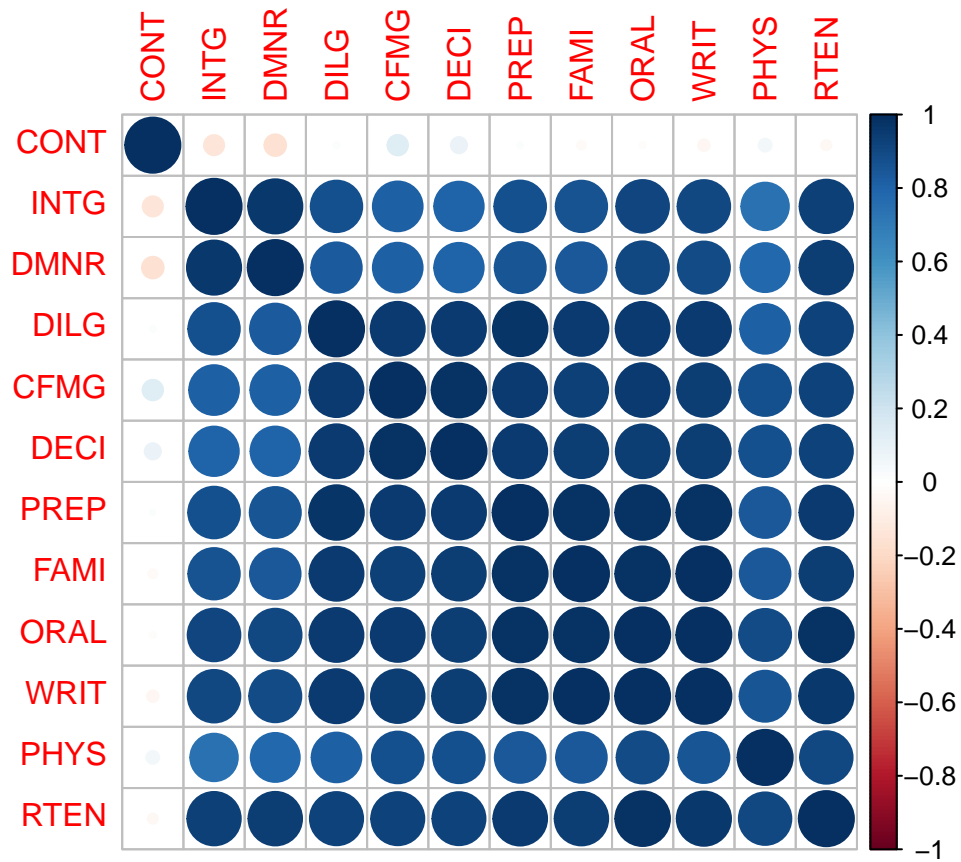
```
round(cor(USJudgeRatings),2)
```

```
##      CONT  INTG  DMNR DILG CFMG DECI PREP  FAMI  ORAL  WRIT PHYS  RTEN
## CONT  1.00 -0.13 -0.15 0.01 0.14 0.09 0.01 -0.03 -0.01 -0.04 0.05 -0.03
## INTG -0.13  1.00  0.96 0.87 0.81 0.80 0.88  0.87  0.91  0.91 0.74  0.94
## DMNR -0.15  0.96  1.00 0.84 0.81 0.80 0.86  0.84  0.91  0.89 0.79  0.94
## DILG  0.01  0.87  0.84 1.00 0.96 0.96 0.98  0.96  0.95  0.96 0.81  0.93
## CFMG  0.14  0.81  0.81 0.96 1.00 0.98 0.96  0.94  0.95  0.94 0.88  0.93
## DECI  0.09  0.80  0.80 0.96 0.98 1.00 0.96  0.94  0.95  0.95 0.87  0.92
## PREP  0.01  0.88  0.86 0.98 0.96 0.96 1.00  0.99  0.98  0.99 0.85  0.95
## FAMI -0.03  0.87  0.84 0.96 0.94 0.94 0.99  1.00  0.98  0.99 0.84  0.94
## ORAL -0.01  0.91  0.91 0.95 0.95 0.95 0.98  0.98  1.00  0.99 0.89  0.98
## WRIT -0.04  0.91  0.89 0.96 0.94 0.95 0.99  0.99  0.99  1.00 0.86  0.97
## PHYS  0.05  0.74  0.79 0.81 0.88 0.87 0.85  0.84  0.89  0.86 1.00  0.91
## RTEN -0.03  0.94  0.94 0.93 0.93 0.92 0.95  0.94  0.98  0.97 0.91  1.00
```

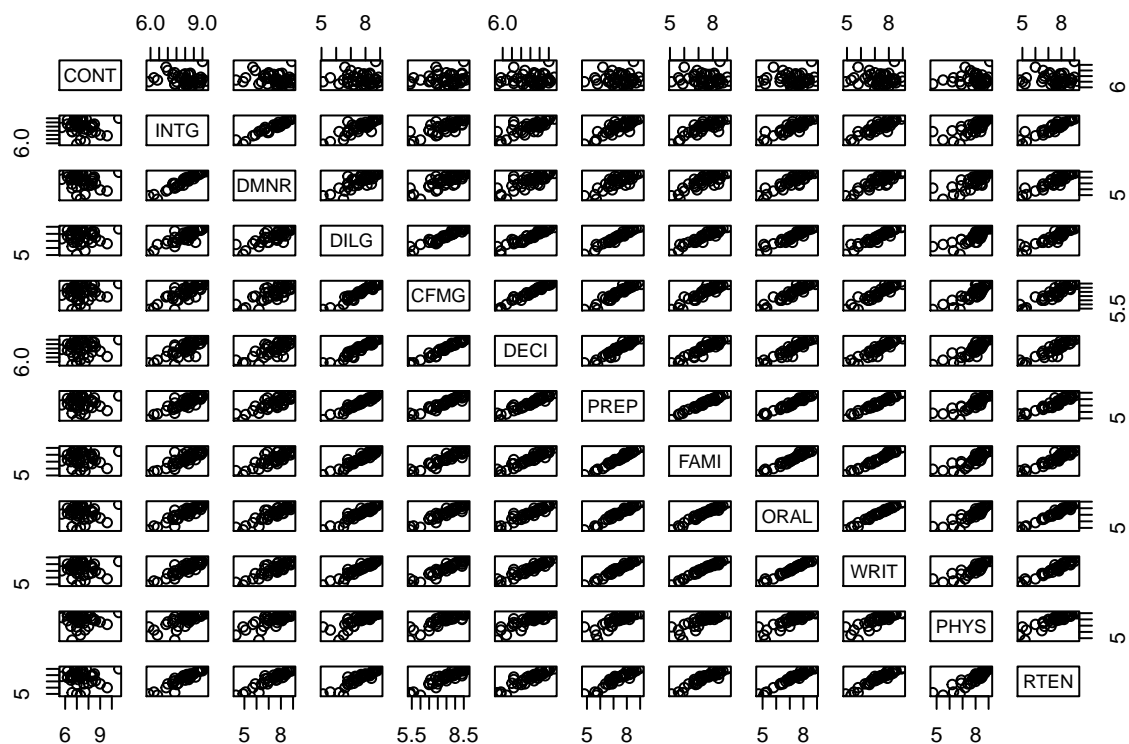
```
library(corrplot)
```

```
## corrplot 0.84 loaded
```

```
corrplot(cor(USJudgeRatings))
```



```
pairs(USJudgeRatings)
```

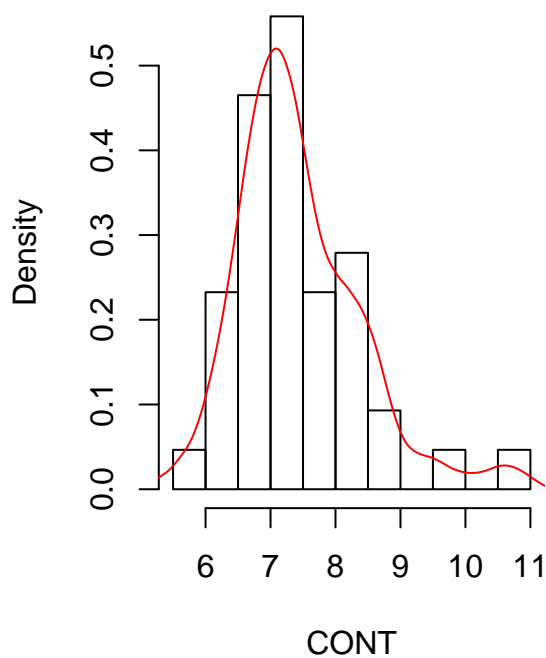



All the variables have a strong positive correlation two by two except the variable CONT which is not correlated to all the other variables. The number of contacts of a lawyer with the judge doesn't seem to explain the ratings received by the judge.

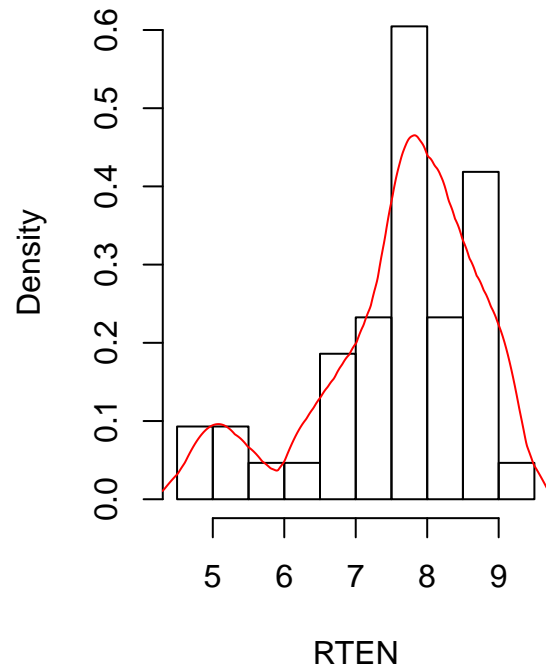
```
par(mfrow=c(1,2))
hist(USJudgeRatings$CONT, probability= TRUE, main="Histogram of CONT", xlab="CONT")
d = density(USJudgeRatings$CONT, kernel = 'c', bw = 0.3)
lines(d, col="red")

hist(USJudgeRatings$RTEN, probability= TRUE, main="Histogram of RTEN", xlab="RTEN")
d = density(USJudgeRatings$RTEN, kernel = 'o', bw = 0.3)
lines(d, col="red")
```

Histogram of CONT

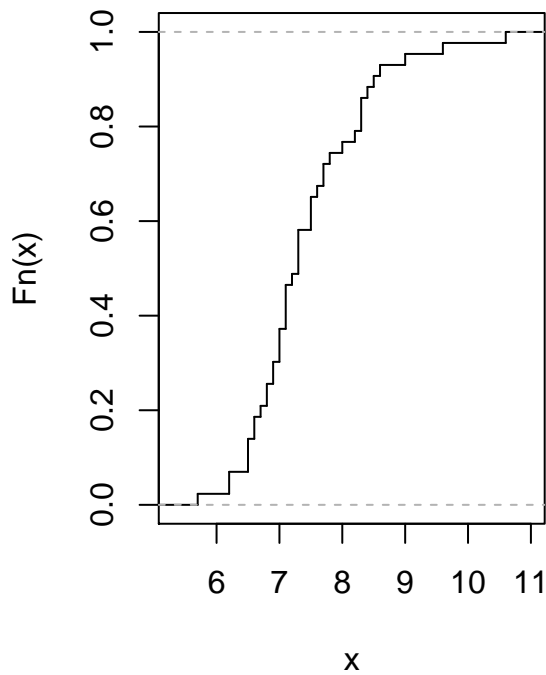


Histogram of RTEN

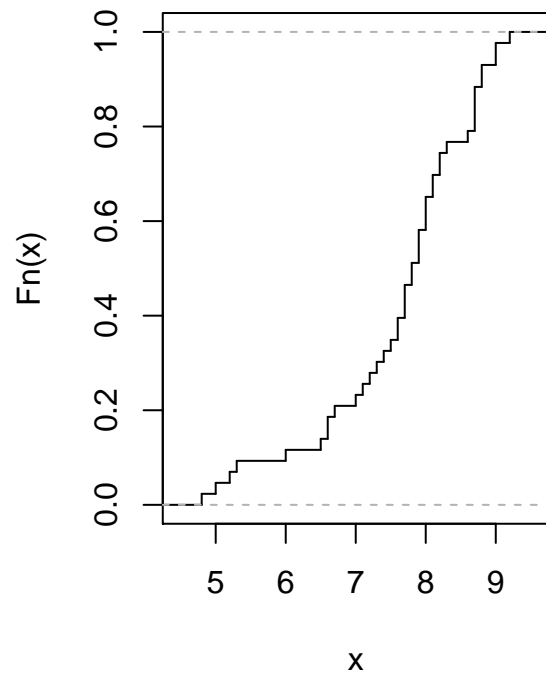


```
par(mfrow=c(1,2))
plot(ecdf(USJudgeRatings$CONT), verticals = TRUE, do.points = FALSE, main = "ECDF CONT")
plot(ecdf(USJudgeRatings$RTEN), verticals = TRUE, do.points = FALSE, main = "ECDF RTEN")
```

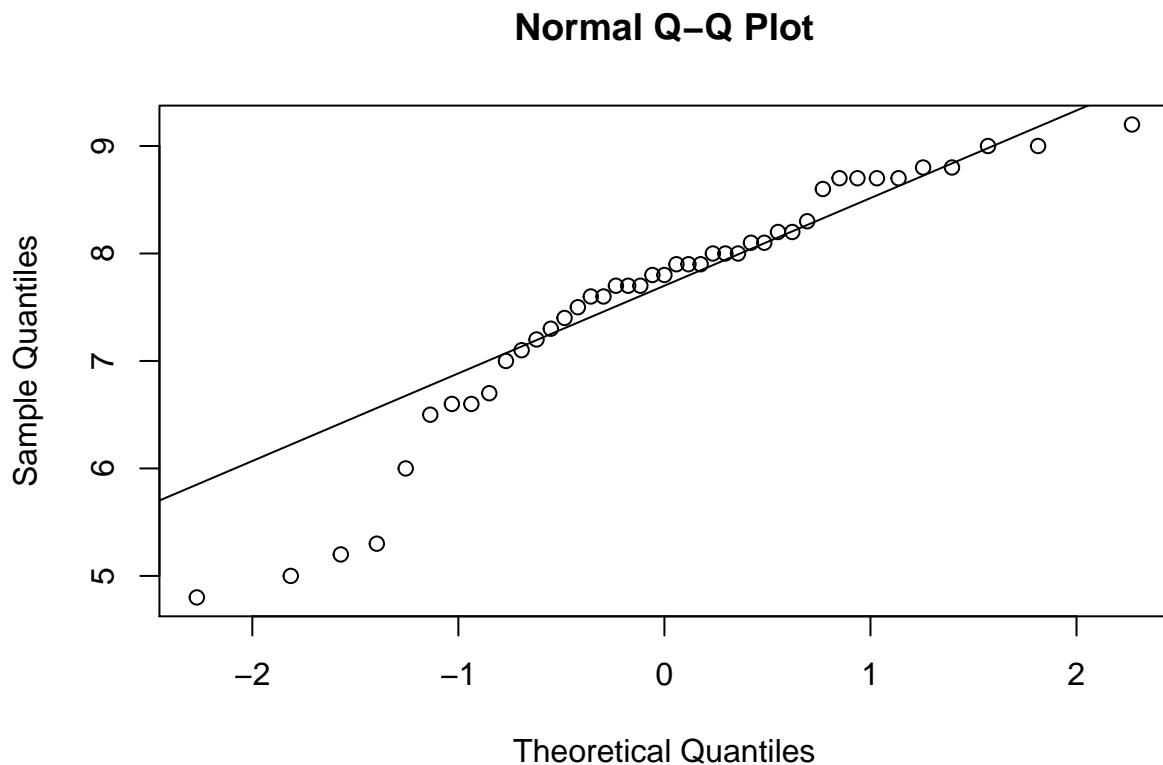
ECDF CONT



ECDF RTEN



```
qqnorm(USJudgeRatings$RTEN)
qqline(USJudgeRatings$RTEN)
```



The QQ plots suggests that the RTEN variable seems to follow a Gaussian distribution except for lower values.

```
library(e1071)
kurtosis
```

```
## function (x, na.rm = FALSE, type = 3)
## {
##   if (any(is.na(x))) {
##     if (na.rm)
##       x <- x[!is.na(x)]
##     else return(NA)
##   }
##   if (!(type %in% (1:3)))
##     stop("Invalid 'type' argument.")
##   n <- length(x)
##   x <- x - mean(x)
##   r <- n * sum(x^4)/(sum(x^2)^2)
##   y <- if (type == 1)
##     r - 3
##   else if (type == 2) {
##     if (n < 4)
##       stop("Need at least 4 complete observations.")
##     ((n + 1) * (r - 3) + 6) * (n - 1)/((n - 2) * (n - 3))
##   }
##   else r * (1 - 1/n)^2 - 3
##   y
## }
## <bytecode: 0x7f9ddb8d02a8>
## <environment: namespace:e1071>
```

skewness

```
## function (x, na.rm = FALSE, type = 3)
## {
##   if (any(ina <- is.na(x))) {
##     if (na.rm)
##       x <- x[!ina]
##     else return(NA)
##   }
##   if (!(type %in% (1:3)))
##     stop("Invalid 'type' argument.")
##   n <- length(x)
##   x <- x - mean(x)
##   y <- sqrt(n) * sum(x^3)/(sum(x^2)^(3/2))
##   if (type == 2) {
##     if (n < 3)
##       stop("Need at least 3 complete observations.")
##     y <- y * sqrt(n * (n - 1))/(n - 2)
##   }
##   else if (type == 3)
##     y <- y * ((1 - 1/n))^(3/2)
##   y
## }
## <bytecode: 0x7f9ddb934f90>
## <environment: namespace:e1071>
```

Conclusion

Our analysis has