
MATH 2625: Biostatistical Methods

Homework 1, due Thursday, January 23

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Theory

1. Define the negative sign counting function to be

$$S^-(x) = \sum_{i=1}^n 1(x_i < 0).$$

Show that a sign test constructed using $S^-(x)$ is equivalent to one we discussed in class.

We know that when we treat $S^-(x)$ as a random variable, it takes on a binomial distribution as such:

$$S^-(x) \sim \text{Bin}(n, p = 1/2).$$

This is because the null hypothesis is defined As

$$H_0 = P(x_i < 0) = 1/2.$$

Let s_0 be a realization of $S^-(x)$. To calculate the p-value for the sign test, we use the following probability calculation:

$$p = P(S^-(x) \leq s_0 | H_0) + P(S^-(x) \geq n - s_0 | H_0).$$

Suppose, like the good mathematicians we are, we rigirously constructed this proof for the positive sign test too. Let s_1 be a realization of the random variable created by the positive sign counting function.

The p-value for the positive sign test would be

$$p = P(S^+(x) \leq s_1 | H_0) + P(S^+(x) \geq n - s_1 | H_0).$$

Since the binomial distribution is symmetric when $p = 1/2$, this means that $p = 1 - p$. This also means that

$$P(X = k) = P(X = n - k).$$

This means that

$$P(S^+(x) \leq s_1 | H_0) = P(S^-(x) \geq n - s_0 | H_0),$$

and also that

$$P(S^+(x) \geq n - s_1 | H_0) = P(S^+(x) \leq s_1 | H_0).$$

This means that the p-value generated by the positive and negative sign functions are the same.

2. Consider the p -value constructions for both the Sign Test and Wilcoxon Signed Rank Tests:

$$p_S = \begin{cases} P(\Sigma^+ \leq S^+(d)|H_0) + P(\Sigma^+ \geq n - S^+(d)|H_0) & S^+(d) < \frac{n}{2} \\ P(\Sigma^+ \geq S^+(d)|H_0) + P(\Sigma^+ \leq n - S^+(d)|H_0) & S^+(d) > \frac{n}{2} \end{cases}$$

$$p_W = \begin{cases} P(\Omega^+ \leq W^+(d)|H_0) + P(\Omega^+ \geq \frac{n(n+1)}{2} - W^+(d)|H_0) & W^+(d) < \frac{n(n+1)}{4} \\ P(\Omega^+ \geq W^+(d)|H_0) + P(\Omega^+ \leq \frac{n(n+1)}{2} - W^+(d)|H_0) & W^+(d) > \frac{n(n+1)}{4} \end{cases}$$

- (a) What value does p_S or p_W take on when $S^+(d) = \frac{n}{2}$ or when $W^+(d) = \frac{n(n+1)}{4}$?
- (b) Can you just double the one-sided probability, i.e. the upper (or lower) tail calculation in p_S or p_W , to attain the p -value? Explain why or why not.
- (a) When $S^+(d) = \frac{n}{2}$, then $p_S = 1$. When we substitute $\frac{n}{2}$ in for $S^+(d)$ (into either equation), we get

$$p_S = P(\Sigma^+ \leq \frac{n}{2}|H_0) + P(\Sigma^+ \geq \frac{n}{2}|H_0).$$

Since Σ^+ is binomial, these two statements sum to 1 when $S^+(d) = \frac{n}{2}$, no matter what the size of n .

We observe something similar in p_W , namely when $W^+(d) = \frac{n(n+1)}{4}$, $p_W = 1$. When we substitute in $W^+(d) = \frac{n(n+1)}{4}$ into the p -value constructions (either the positive or negative end), we observe

$$p_W = P(\Omega^+ \leq \frac{n(n+1)}{4}|H_0) + P(\Omega^+ \geq \frac{n(n+1)}{2} - \frac{n(n+1)}{4}|H_0).$$

This simplifies into

$$p_W = P(\Omega^+ \leq \frac{n(n+1)}{4}|H_0) + P(\Omega^+ \geq \frac{n(n+1)}{4}|H_0).$$

We can observe that this sum equals 1.

- (b) Yes, we can double the one-sided probability in either test. This is because the null hypothesis H_0 states that the chance of having a positive or negative sign, in both the ranked and unranked sign test, is expected to be equal ($p = \frac{1}{2}$). In both of these nonparametric tests, we are calculating the probability of having this value or greater (absolutely, since it would be less if the values were negative). Since both Σ^+ and Ω^+ (or their respective negatives) are symmetric, we are able to double the probability from one tail.

Case Studies

For each of the following, create a structured abstract no longer than 2 pages in length (including figures, tables, and references). The Background section is provided for each and should be included in your write-up. You must write the Methods, Results, and Conclusion sections.

1. In this case study, you will examine data from a crossover study examining the impact of exposure to altitude on heart rate on older and susceptible passengers. The data is in file `hrPaired.txt` where the variable `ID` denotes the subject, the variable `Control` is the average heart rate during the control day, and the variable `Flight` is the average heart rate during the flight day. Additional information of this case study is in the Background section below.

Background

Older and susceptible passengers and those with preexisting disease are flying with increasing frequency and in-flight cardiac emergencies are a more frequent occurrence. While commercial airplanes fly at altitudes of around 34,000 feet, Federal Aviation Administration (FAA) regulations limit cabin pressurization to an equivalent of between 7,000 and 9,000 feet, with the typical pressurization implemented by most aircraft equal to 8,000 feet. Pressurization to this equivalent level is selected to balance preventing acute altitude-related health symptoms among flyers with operational demands on the aircraft. However, comprehensive acute and longer-term health effects of cabin pressurization have not been well characterized. In particular, the impacts of short term exposure to altitude on cardiovascular health is of particular interest for study among older and vulnerable passengers. To examine possible effects, we conducted a block-randomized crossover design study of the physiological effects under simulated cabin altitudes in a hypobaric pressure chamber among such passengers. The goal of this study is to assess the changes in heart rate between simulated cabin conditions on a flight day versus control conditions. Under flight day conditions, the chamber was pressurized to the equivalent of 7,000 feet altitude. On the control days, the chamber remained at sea level.

Methods

The trial involved 33 randomly selected participants who were older and more susceptible to in-flight cardiac emergencies. Data was analyzed at the nominal level using a paired T-test, with a supporting sensitivity analysis through the Wilcoxon sign-rank test. A mean baseline heart rate of 78.75 beats per minute (BPM) was observed (SD: 9.12) along with a median of 77.82 BPM and median absolute difference of 10.35 BPM. In-flight heart rate was simulated in a hypobaric pressure chamber. We observed a mean in-flight heart rate of 81.21 BPM (SD: 10.56 BPM) along with a median of 82.05 BPM. The MAD for in-flight heart rate was 14.21 BPM. We observed a mean difference between heart rates of 2.46 BPM (SD: 6.00 BPM). The median difference was 1.75 BPM, while the MAD was 5.22 BPM.

Results

A summary of the results as stated above are provided in Table 1, including both baseline and in flight heart rates and summaries of the observed differences. Figure 1 (below) shows that these differences are fairly normal in distribution, despite the small number of participants.

Table 1: Heart Rate Summary Statistics (in BPM)

Measure	Baseline Heart Rate	In-Flight Heart Rate	Difference
Median	77.82	82.05	1.750
MAD	10.35	14.21	5.22
Mean	78.75	81.21	2.462
Standard Deviation	9.123	10.560	5.998

As shown in Table 2, we found evidence to support the claim that for people who are older and more susceptible to cardiac emergencies, their in-flight heart rate is different to their baseline heart rate ($p = 0.0246$, $df = 32$, 95% CI = (0.335, 4.589)). This result is consistent with the sensitivity analysis (see Table 3), which also found marginally significance in the difference between heart rates ($p = 0.0253$, $W = 405$, 95% CI = (0.344, 4.315)).

Table 2: Paired T-Test Results Indicate Marginally Significant Difference in Heart Rate

T-Value	df	Mean Difference (BPM)	95% Confidence Interval	p-value
2.3582	32	2.462	(0.335, 4.589)	0.0246

Table 3: Wilcoxon Test Confirms Findings of T-Test of Significant Difference in Heart Rate

Test Statistic	Hodges-Lehmann Estimator	95% Confidence Interval	p-value
$W = 405$	2.157	(0.344, 4.315)	0.0253

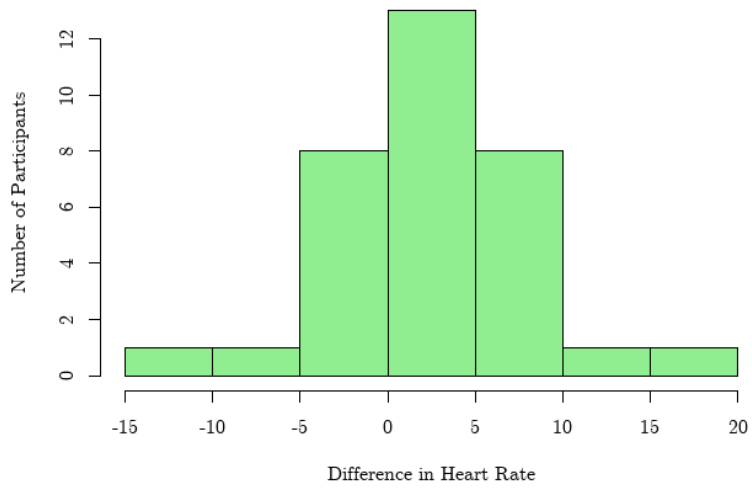


Figure 1: Histogram of Heart Rate Differences Between Baseline and In-Flight Conditions

2. This case study focuses on the effects of different surgical procedures on infant development as measured by the Bayley Scales. The data is in the file `heart.txt` where the variable `treatment` contains the grouping variable with the labels `DCHA` and `Low-flow` and the variable `pdi` and `mdi` contain measurements for the scales themselves. Additional information of this case study is in the Background section below.

Background

The Bayley Scales of Infant Development yield scores on two indices—the Psychomotor Development Index (PDI) and the Mental Development Index (MDI)—that can be used to assess a child’s level of functioning at approximately one year of age. As part of a study investigating the development and neurologic status of children who had undergone reparative heart congenital heart disease. Specifically, the study was on infants with D-transposition of the great arteries who underwent an arterial-switch operation. D-transposition of the great arteries is a birth defect where the child’s arteries formed incorrectly and are transposed, i.e. connected to the wrong ventricles. The children in the study were randomized to one of two different treatment groups, known as “DCHA” and “low-flow bypass.” The groups differed in the specific way in which the reparative surgery was performed. Deep hypothermic circulatory arrest (DHCA) is a surgical technique that involves cooling the body to temperatures below 20°C (68°F), and stopping blood circulation and brain function for up to one hour during which time the reparative heart surgery is performed—in this case switching the ventricles the arteries are connected to. In low-flow cardiopulmonary bypass, circulation to the brain is continuously maintained, though at a reduced rate, while the reparative heart surgery is performed. While some physicians feel low-flow bypass is preferable, it has its own risks associated with brain injury. Thus, this study aims to compare PDI in the DCHA group to that in the low-flow group as well as comparing MDI between the two groups.