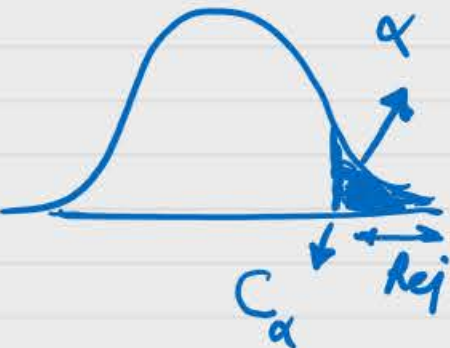


①  $H_0: \beta_j \leq 0$   
 $H_1: \beta_j > 0 \rightarrow$

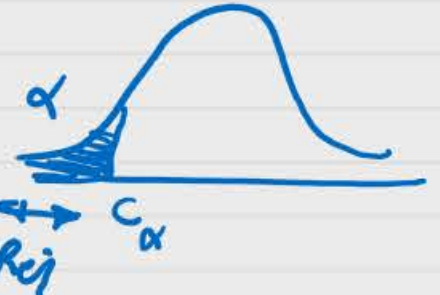


$$t = \frac{\hat{\beta}_j - \beta_{jH_0}}{Se(\hat{\beta}_j)} = \frac{\hat{\beta}_j - 0}{Se(\hat{\beta}_j)} \quad \checkmark$$

$$|t| > |c_\alpha| \Rightarrow \text{Rej } H_0 \Leftarrow P\text{-value} < \alpha$$

::

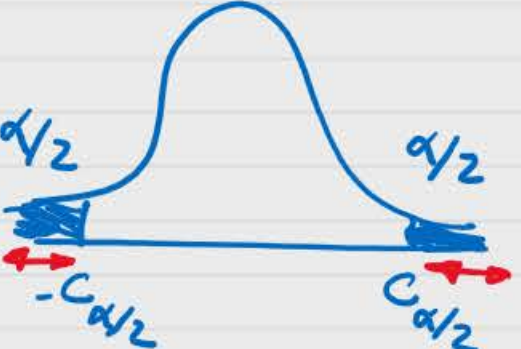
②  $H_0: \beta_j \geq 0$   
 $H_1: \beta_j < 0$



$$t = \frac{\hat{\beta}_j - 0}{Se(\hat{\beta}_j)} \quad \checkmark$$

$$|t| > |c_\alpha| \rightarrow \text{Rej } H_0 \Leftarrow P\text{-value} < \alpha$$


③  $H_0: \beta_j = 0$   
 $H_1: \beta_j \neq 0$



$$t = \frac{\hat{\beta}_j - 0}{Se(\hat{\beta}_j)} \quad \text{use R or manually}$$

$$|t| > |c_{\alpha/2}| \rightarrow \text{Rej } H_0 \Leftarrow P\text{-value} < \alpha$$

④  $H_0: \beta_j = a_j$   
 $H_1: \beta_j \neq a_j$



$$|t| > |c_{\alpha/2}| \rightarrow \text{Rej } H_0 \Leftarrow P\text{-value} < \alpha$$

$$t = \frac{\hat{\beta}_j - a_j}{Se(\hat{\beta}_j)}$$

⑤ linear combination of  $\beta_j$   
 $H_0: \beta_1 - \beta_2 = 0$   
 $H_1: \begin{cases} \beta_1 - \beta_2 \neq 0 \\ < 0 \\ > 0 \end{cases}$

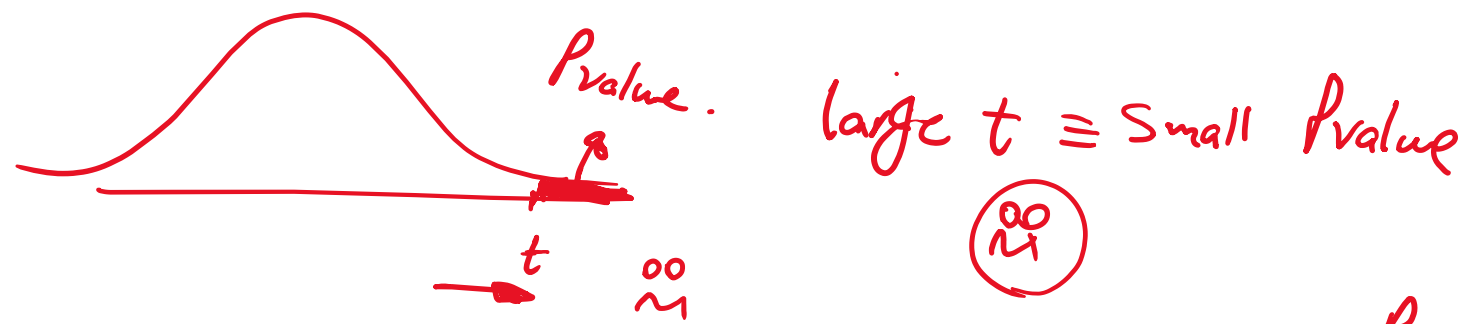
$\rightarrow$  use  $\theta$  method  
 $\theta = \beta_1 - \beta_2$  or  $\beta_2 = \beta_1 + \theta$   
 Rewrite the model with  $\theta$   
 $\rightarrow$  find the  $\theta$   $H_0: \theta = 0$   
 $H_1: \theta \neq 0$

## Class 16 – Multiple Regression Model Inference (Part II)

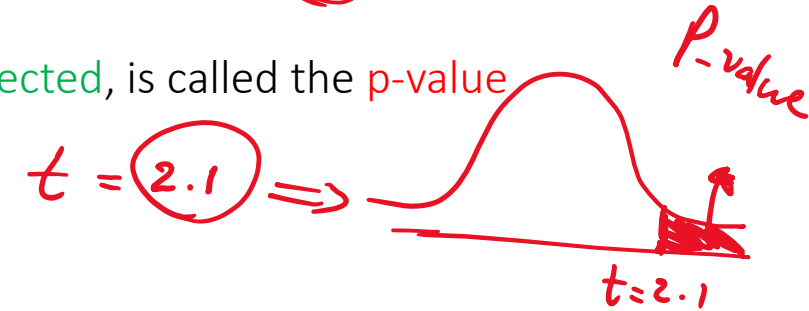
Pedram Jahangiry



## Computing p-values for t-tests



- ✓ ☐ Recall 1: The smallest significance level at which the null hypothesis is still rejected, is called the p-value of the hypothesis test
- ✓ ☐ Recall 2: p-value is the corresponding significance level of the test statistic.
- ☐ A small p-value is evidence against the null hypothesis (a good thing!) because one would reject the null hypothesis even at small significance levels
- ✓ ☐ A large p-value is evidence in favor of the null hypothesis (a bad thing!)
- ✓ ☐ P-values are more informative than tests at fixed significance levels
- ✓ ☐ The p-value is the significance level at which one is indifferent between rejecting and not rejecting the null hypothesis.



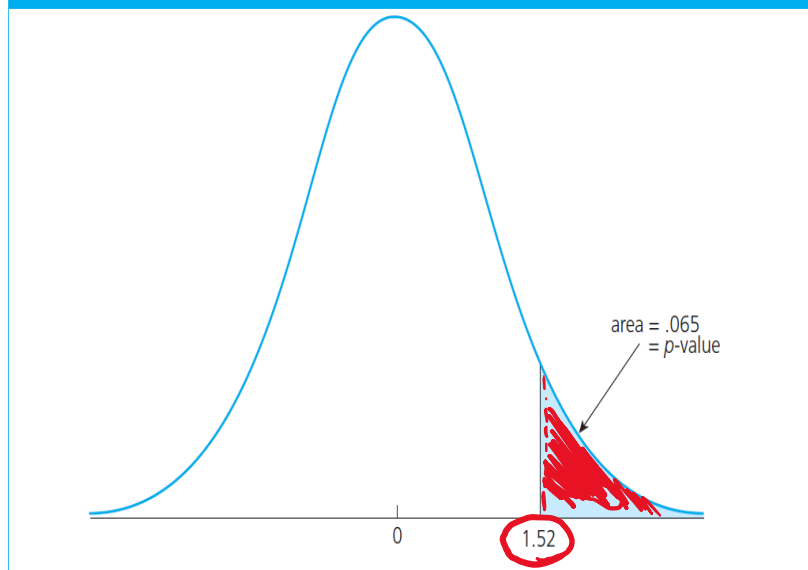
$$\frac{\beta_j}{se} \rightarrow t \rightarrow P\text{-value}$$

## Computing and Using $p$ -values (cont'd)

We said that **p-value** is the corresponding significance level of the test statistic.

P-values for one-tailed tests:

FIGURE C.7 The  $p$ -value when  $t = 1.52$  for the one-sided alternative



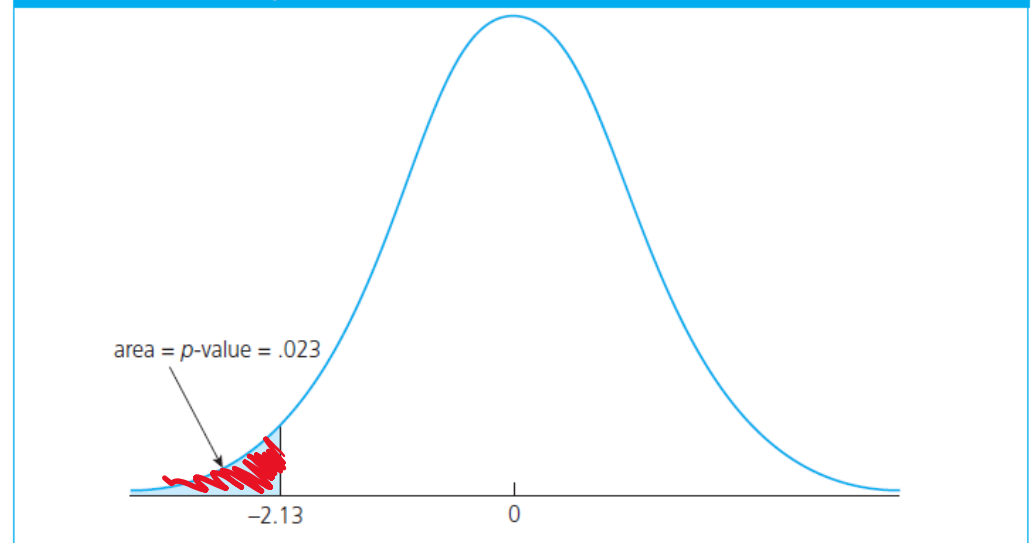
$$p_{value} = P(T > 1.52) = 1 - CDF(1.52) = 0.065$$

$n=200$

$$1 - P_t(1.52, df) =$$

$\downarrow$   
100

FIGURE C.8 The  $p$ -value when  $t = -2.13$  with 19 degrees of freedom for the one-sided alternative  $\mu < 0$ .



$$p_{value} = P(T < -2.13) = CDF(-2.13) = 0.023$$

$n=20$

$$P_t(-2.13, df) = 0.023$$

$\downarrow$   
6

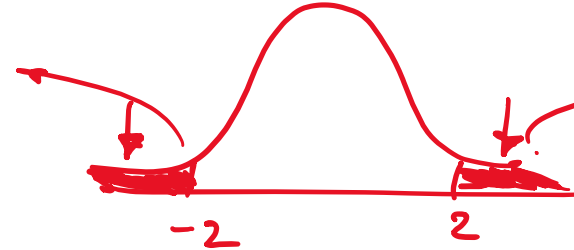
# Computing and Using p-values (cont'd)

P-values for Two-tailed tests:

$$2 \times \text{COF}(-2)$$

$t = 2 \Rightarrow$  P-value for 2-tailed test

$$(1 - \text{CDF}(2)) \times 2$$



$$p_{\text{value}} = P(|T| > |2|) =$$

$$2(1 - \text{CDF}(2)) = 0.06,$$

$$df = 20$$

A null hypothesis is **rejected** if and only if the corresponding p-value is **smaller** than the significance level.

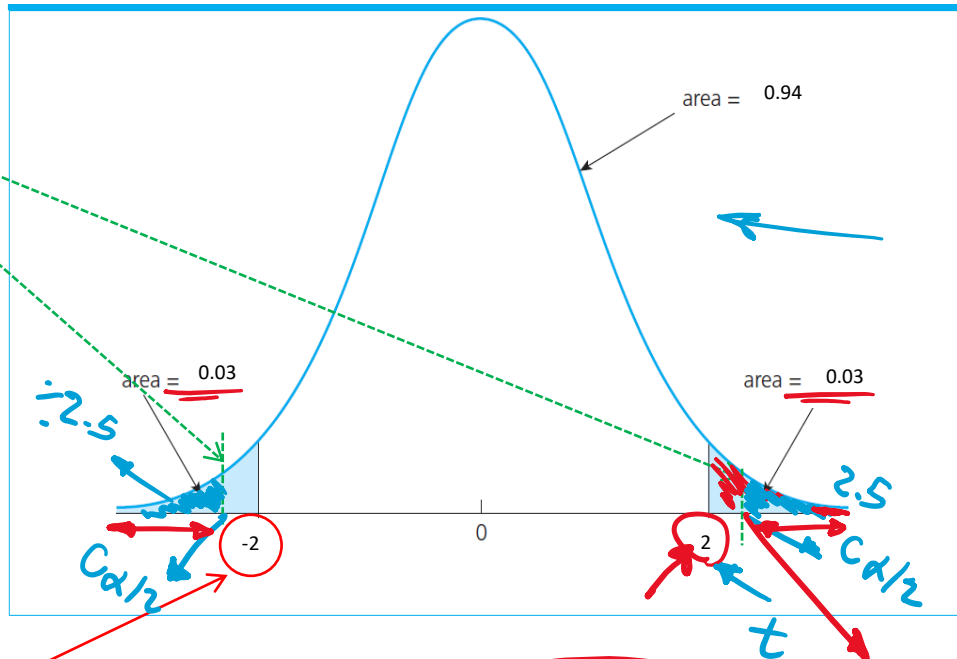
Do you reject the null here?

$$p_{\text{value}} = 6\%$$

$$\alpha = 5\%$$

$$\alpha = 5\%$$

These would be the critical values for a 5% significance level



Value of test statistic

$$t < c_{\alpha/2}$$

$$|t| > |c| \Rightarrow \text{rej}$$

$\equiv$

$$p_{\text{value}} < \alpha \Rightarrow \text{rej}$$

↓  
rej



## Economic / Practical significance VS. Statistical significance

$$\text{int} = \beta_0 + \beta_1 \text{ height} + \dots + u$$

$\Rightarrow R^2 = 90\%$

$\hat{\beta}_1$  was significant

$$\hat{\beta}_1 = 0.0000001$$

- ✓ ☐ If a variable is statistically significant, discuss the **magnitude** of the coefficient to get an idea of its economic or practical importance
- ✓ ☐ The fact that a coefficient is statistically significant does not necessarily mean it is economically or practically significant!
- ✓ ☐ If a variable is statistically and economically important but has the **"wrong" sign**, the regression model might be misspecified
- ✓ ☐ If a variable is **NOT** statistically significant at the usual levels (10%, 5%, or 1%), one may think of dropping it from the regression
- ✓ ☐ If the **sample size is small**, effects might be imprecisely estimated so that the case for dropping insignificant variables is less strong

$$\text{wage} \sim \text{educ}$$

$$\text{wage} \sim \text{educ} + \text{gender}$$





## Confidence intervals

Recall: CI is two-sided by nature

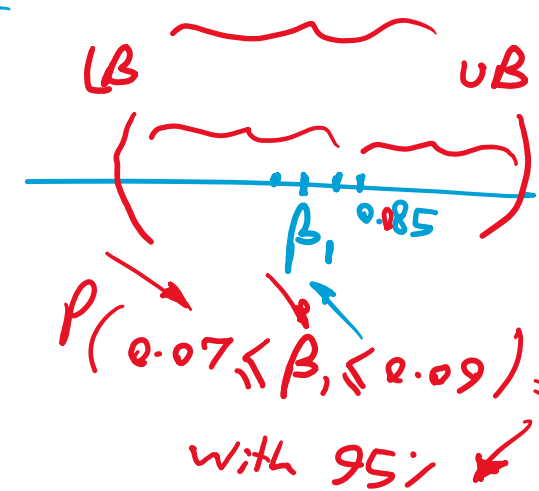
$$\hat{\beta}_j \pm c * se(\hat{\beta}_j)$$

$$\log(\text{wage}) = \beta_0 + \beta_1 \text{educ} + u$$

$$\log(\hat{\text{wage}}) = 0.1 + 0.085 \text{educ}$$

$$\rightarrow \beta_1$$

Prob



$$P\left(\underbrace{\hat{\beta}_j - c_{0.05} \cdot se(\hat{\beta}_j)}_{LB} \leq \beta_j \leq \underbrace{\hat{\beta}_j + c_{0.05} \cdot se(\hat{\beta}_j)}_{UB}\right) = 0.95$$

Lower bound of the Confidence interval

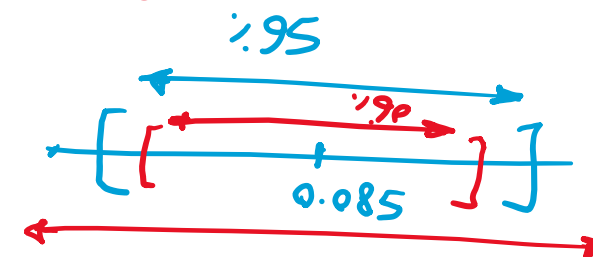
Upper bound of the Confidence interval

Confidence level

CI will contain  $\beta_j$  with 90% Prob ✓  
 $\beta_j$  will lie in CI = 90% Prob ✓

$$c_{0.05} \approx 1.96$$

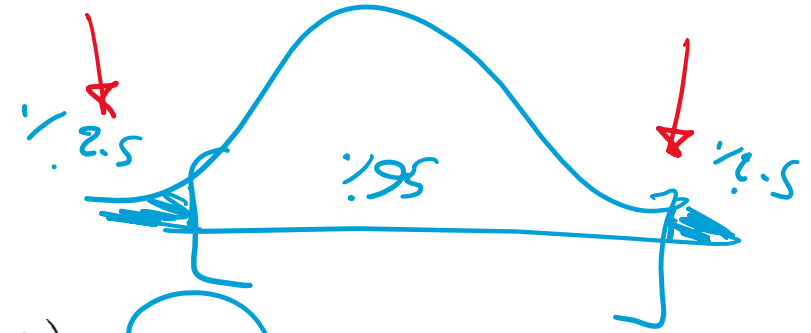
Interpretation of the confidence interval:



- ✓ The bounds of the interval are random
- ✓ In repeated samples, the interval will contain the population regression coefficient ( $\beta$ ) in  $(1 - \alpha)\%$  of the cases

two-tailed

## Confidence intervals for typical confidence levels



$$P(\hat{\beta}_j - c_{0.01} \cdot se(\hat{\beta}_j) \leq \beta_j \leq \hat{\beta}_j + c_{0.01} \cdot se(\hat{\beta}_j)) = 0.99$$

$$P(\hat{\beta}_j - c_{0.05} \cdot se(\hat{\beta}_j) \leq \beta_j \leq \hat{\beta}_j + c_{0.05} \cdot se(\hat{\beta}_j)) = 0.95$$

$$P(\hat{\beta}_j - c_{0.10} \cdot se(\hat{\beta}_j) \leq \beta_j \leq \hat{\beta}_j + c_{0.10} \cdot se(\hat{\beta}_j)) = 0.90$$

Use rules of thumb  $c_{0.01} = 2.576, c_{0.05} = 1.96, c_{0.10} = 1.645$

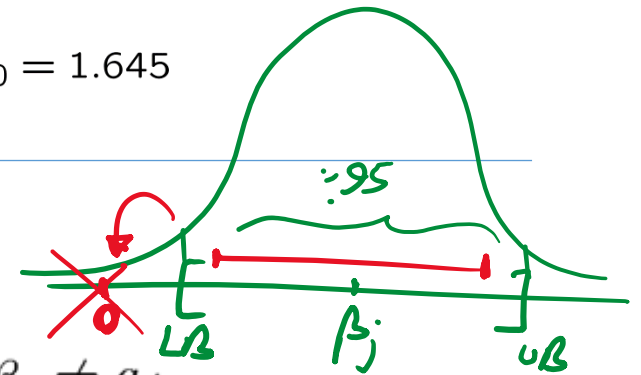
## Relationship between confidence intervals and hypotheses tests

$$H_0: \beta_j = a_j$$

$a_j \notin \text{interval} \Rightarrow$  reject  $H_0: \beta_j = a_j$  in favor of  $H_1: \beta_j \neq a_j$

$$\beta_j = 0$$

$$0 \notin [ ]$$





## Example: Model of firms' R&D expenditures

$$\textcircled{1} \rightarrow t_{\beta_1} = \frac{1.084 - 0}{0.06} \approx \textcircled{18}$$

1% 5%  $\rightarrow C_{\alpha/2} = \textcircled{2.9}$

$t > 2.9 \rightarrow$   
 $\textcircled{\text{Ref}}$

Spending on R&D

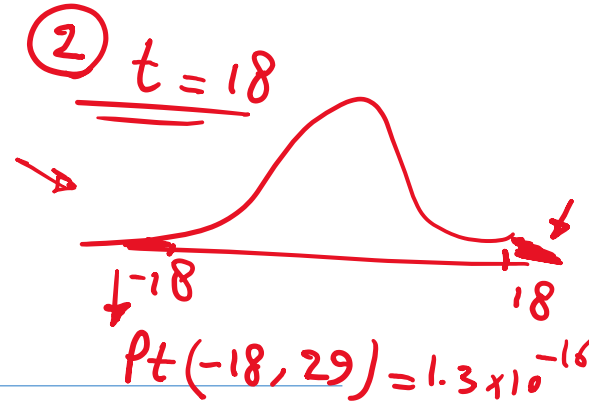
Annual sales

Profits as percentage of sales

$$\widehat{\log(rd)} = -4.38 + 1.084 \log(sales) + .0217 \text{ profmarg}$$

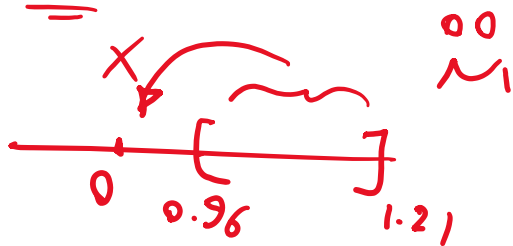
(.47) (.060) (.0128)

$$n = 32, R^2 = .918. \quad df = 32 - 2 - 1 = 29 \Rightarrow c_{0.05} = 2.045$$



What are the CI for  $\beta_1$  and  $\beta_2$ ?

$\textcircled{3}$



$$1.084 \pm 2.045(.060) = (.961, 1.21)$$

$$.0217 \pm 2.045(.0218) = (-.0045, .0479)$$

The effect of sales on R&D is relatively **precisely** estimated as the interval is narrow. Moreover, the effect is **significantly different from zero** because zero is outside the interval.

This effect is **imprecisely** estimated as the interval is very wide. It is **not even statistically significant** because zero lies in the interval.

R

# chapter 4: MRM, Inference

```
library(wooldridge)
library(stargazer)
```

# Example 4-8

```
MRM <- lm(log(rd) ~ log(sales) + profmarg, rdchem)
summary(MRM)
```

# finding critical values

```
df <- nobs(MRM) - 2 - 1
alpha <- 0.05
qt(1 - alpha/2, df)
```

# Look at t\_stat

```
summary(MRM)$coefficients[, "t value"]
```

```
# Confidence Interval
confint(MRM, level = 1 - alpha)
```

&gt; summary(MRM)

Call:

lm(formula = log(rd) ~ log(sales) + profmarg, data = rdchem)

Residuals:

Min	1Q	Median	3Q	Max
-0.97681	-0.31502	-0.05828	0.39020	1.21783

Coefficients:

	Estimate	Std. Error	t value	Pr(> t )
(Intercept)	-4.37827	0.46802	-9.355	2.93e-10 ***
log(sales)	1.08422	0.06020	18.012	< 2e-16 ***
profmarg	0.02166	0.01278	1.694	0.101

Signif. codes: 0 '\*\*\*' 0.001 '\*\*' 0.01 '\*' 0.05 '.' 0.1 ' ' 1

Residual standard error: 0.5136 on 29 degrees of freedom

Multiple R-squared: 0.918, Adjusted R-squared: 0.9123

F-statistic: 162.2 on 2 and 29 DF, p-value: &lt; 2.2e-16

&gt; qt(1 - alpha/2, df)

[1] 2.04523

# Look at t\_stat

```
> summary(MRM)$coefficients[, "t value"]
```

(Intercept)	log(sales)	profmarg
-9.354916	18.011791	1.694150

# Confidence Interval

&gt; confint(MRM, level = 1 - alpha)

	2.5 %	97.5 %
(Intercept)	-5.335478450	-3.4210681
log(sales)	0.961107256	1.2073325
profmarg	-0.004487722	0.0477991

 $\alpha_{rej} = 5\%$ 

?

 $\alpha = 5\%$   
Fail to Rej $\alpha/2$   
 $2.045$ Rej  
Rej  
Fail to Rej

## Testing hypotheses about a linear combination of the parameters

Example: Return to education at two-year vs. at four-year colleges

Number of years attending a **2-year** college      Number of years at a **4-year** college      Months in the workforce

$$\log(\text{wage}) = \beta_0 + \beta_1 jc + \beta_2 \text{univ} + \beta_3 \text{exper} + u$$

Data  
Suggest  
 $\hat{\beta}_{\text{univ}} > \hat{\beta}_{jc}$

$$\widehat{\log(\text{wage})} = 1.472 + .0667 jc + .0769 \text{univ} + .0049 \text{exper}$$

(.021) (.0068) (.0023) (.0002)

$n = 6,763, R^2 = .222.$

- ✓ The **hypothesis** of interest is whether one year at a junior college is worth one year at a university
- ✓ the **alternative** of interest is one-sided: a year at a junior college is worth less than a year at a university

$$\beta_{\text{univ}} > \beta_{jc}$$

## Testing hypotheses about a linear combination of the parameters

- ✓ The **hypothesis** of interest is whether one year at a junior college is worth one year at a university
- ✓ the **alternative** of interest is one-sided: a year at a junior college is worth less than a year at a university

Test  $H_0: \beta_1 - \beta_2 = 0$  against  $H_1: \beta_1 - \beta_2 < 0$   $\rightarrow \beta_{jc} - \beta_{univ} < 0 \Rightarrow \beta_{univ} > \beta_{jc}$

A possible test statistic would be:

$$\frac{\hat{\beta}_j - \beta_{H_0}}{se(\hat{\beta}_j)}$$

$$t = \frac{\hat{\beta}_1 - \hat{\beta}_2}{se(\hat{\beta}_1 - \hat{\beta}_2)} \\ = \frac{(\hat{\beta}_1 - \hat{\beta}_2) - 0}{se(\hat{\beta}_1 - \hat{\beta}_2)}$$

The difference between the estimates is normalized by the estimated standard deviation of the difference. The null hypothesis would have to be rejected if the statistic is “too negative” to believe that the true difference between the parameters is equal to zero.

## Testing hypotheses about a linear combination of the parameters (cont'd)

Impossible to compute with standard regression output because

$$se(\hat{\beta}_1 - \hat{\beta}_2) = \sqrt{Var(\hat{\beta}_1 - \hat{\beta}_2)} = \sqrt{Var(\hat{\beta}_1) + Var(\hat{\beta}_2) - 2Cov(\hat{\beta}_1, \hat{\beta}_2)}$$

$$Var(X+Y) = ?$$

Usually not available in regression output

$$t = \frac{0}{0}$$

Alternative method:

Define  $\theta_1 = \beta_1 - \beta_2$  and test  $H_0 : \theta_1 = 0$  against  $H_1 : \theta_1 < 0$

w	jc	univ	exper	Total
?	?	?	?	?
?	?	?	?	?
?	?	?	?	?

$$\log(wage) = \beta_0 + (\theta_1 + \beta_2)jc + \beta_2univ + \beta_3exper + u$$

$$\log(w) = \beta_0 + \theta_1jc + \beta_2(jc + univ) + \beta_3exper + u$$

Insert into original regression

a new regressor (= total years of college)

$$\beta_1 - \beta_2 < 0$$

$\hat{\theta}$  is sig or not

## Testing hypotheses about a linear combination of the parameters (cont'd)

Estimation results

$$\widehat{\log(wage)} = 1.472 - \underbrace{.0102}_{\beta_1} \underbrace{jc}_{\theta} + \underbrace{.0769}_{\beta_2} \underbrace{totcoll}_{\beta_3} + .0049 \text{ exper}$$

(.021)
(.0069)
(.0023)
(.0002)

Total years of college

$$n = 6.763, R^2 = .222$$

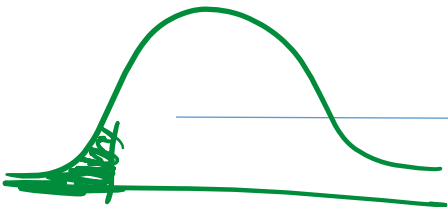
$$t = -.0102 / .0069 = -1.48$$

$$p\text{-value} = P(t\text{-ratio} < -1.48) = .070$$

Hypothesis is rejected at 10% level but not at 5% level

Confidence Interval for  $\theta_1 = \beta_1 - \beta_2 \longrightarrow -.0102 \pm 1.96(.0069) = (-.0237, .0003)$

This method works always for single linear hypotheses



-1.48

$P_t(-1.48, 6759)$

