Project |

$$Q = \sqrt{-2 \log(1-U)}$$
 $P(R \le r) = P(\sqrt{-2 \log(1-U)} \le r) = P(U = 1 - e^{-r/2})$ 
 $= \int_{0}^{1-e^{-r/2}} \sqrt{-2 \log(1-U)} dU$ 
 $Q = 2\pi V$ 
 $P(Q = x) = P(2\pi V = x) = P(V \le x/2\pi)$ 
 $= \int_{0}^{1/2\pi} 2\pi V dV = \frac{x^{2}}{4\pi}$ 

b)  $X = R \cos \theta$  and  $Y = R \sin \theta$ 
 $Y = f(\theta) \longrightarrow \theta = P'(Y) \longrightarrow \theta = 2\pi n - \sin^{-1}(Y/R) + \pi n \in \mathbb{Z}$ 
 $det(\frac{3e}{3x}(f^{-1}(Y))) Px(f^{-1}(Y)) = \longrightarrow Px = PR\theta = P(R)P(\theta) \longrightarrow P(-x)P(2\pi n + \pi)$ 
 $Q = P(-x)P(2\pi n + \pi)$ 

20) Since intra-class is 0, ratio will be 0, so as m increases the ratio will remain the same, implying that the accuracy is good b) In this case o, = oz = oz In but X = X'  $0_{X_1X_1=\frac{1}{8}(X-X_1)}(\sigma^2 I_n)^{-1}(X-X_1)+\frac{1}{2}\ln(1)$ a)  $P(X_1X_1)=\frac{1}{8}(X-X_1)^{-1}(X-X_1)+\frac{1}{2}\ln(1)$ show  $\hat{\lambda}=\frac{1}{6}\sum_{i=1}^{6}X_i$  is MLE of  $\lambda$ L(P(Xi, X))= fi 1xe-x -> log(L)= log(fi 1xe-x)->  $= \sum_{n=1}^{\infty} \ln\left(\frac{\lambda^{x}e^{-x}}{x!}\right) = \sum_{n=1}^{\infty} \left(\ln(e^{-x}) + \ln(\lambda^{x}) - \ln(x!)\right) \longrightarrow$ =-nx+ ZIn(xx1) - ZIn(x!) MLE: set of (log(L))=0 ->  $-n+\frac{1}{\lambda}\sum_{x_i}=0 \longrightarrow \lambda=\frac{1}{\lambda}\sum_{x_i}$ b) P(NO,B)= Bx Na-1e-Bx  $\frac{\beta}{\Gamma(\alpha)} \lambda^{\alpha-1} e^{-\beta \lambda} = \frac{\beta}{\chi_{i,j}} = \frac{\beta}{\Gamma(\alpha)} \lambda^{\alpha-1} e^{-\beta \lambda} = \frac{\beta}{\Pi \chi_{i,j}}$ Likelihood Poisson = TI X:1 [(a) (x a-1+ \( \times \))

Project 1

Second part is Gamma distribution with

&= x+ \(\Six\) and \(\hat{\beta} = \beta + n\)

 $P(\lambda|y) = (\lambda|x + \sum_{i=1}^{n} \chi_i, \beta + N) = F(\hat{\alpha}, \hat{\beta})$ 

c) MAP of a Garma function where the distribution has the highest probability is

 $\Gamma(\alpha, \beta) = \frac{\alpha - 1}{\beta} \rightarrow MAP = \frac{\sum x_i + \alpha - 1}{N + \beta}$ 

(d) A = e-2x

MLE  $\rightarrow X = \lambda(\eta) \rightarrow e^{-2\lambda} = \eta \rightarrow \lambda = -\frac{1}{2}h(\eta) = \lambda(\eta)$ 

N(n) into Poisson P(XIX)=1×e-X where 1=1(1) ->

 $P(X|\lambda(\eta)) = \frac{\lambda(\eta)^{X}e^{-X}}{X!} = \frac{1}{X!}(-\frac{1}{2}\ln(\eta))^{X}e^{\frac{1}{2}\ln(\eta)} \longrightarrow$ 

1n(p(x1λ(n)))=×1n(-½1n(n))+½1n(n)-1n(x1)→

3/n(P(XIX(A))) =0 -> X(-1/n(A))(-1/2)+1/2=0->

A=e-1X

e) Bias of  $\hat{\Lambda} = e^{-2\lambda} - e^{\lambda(1/e^2-1)}$  using  $e^{+} = \sum_{k=0}^{\infty} + \frac{n}{n!}$   $\text{Bias}(\hat{\Lambda}) = \text{E}[\hat{\Lambda}] - \hat{\Lambda} \longrightarrow \text{E}[\hat{\Lambda}] = \sum_{k=0}^{\infty} \hat{\Lambda} P(X) = \sum_{k=0}^{\infty} e^{-2\lambda} \frac{\lambda^{2} e^{-\lambda}}{\lambda!}$  $= e^{-\lambda} \sum_{k=0}^{\infty} (\lambda e^{-2})^{k} = e^{-\lambda} e^{\lambda} e^{-2} = e^{(e^{-2}-1)\lambda}$ 

> Brus(^)=e(e-2-1)\lambda -e-2\lambda

f)  $(-1)^{x}$  is the only unbiased estimator of  $\eta$   $E[\hat{\Lambda}] = \sum_{i=1}^{n} P(x) = \sum_{i=1}^{n} (-1)^{x} \frac{\lambda^{x} e^{-\lambda^{2}}}{x!} = e^{-\lambda} \sum_{i=1}^{n} P(x)^{x} = e^{-\lambda^{2}} \sum_{i=1}^{n} e^{-\lambda^{2}} e^{-\lambda^{2}} = e^{-\lambda^$ 

Brus(A)=E[A]-n=e-2x-e-2x=0

Since Gias = 0 this is indeed inbiased

This Bias is but because it will either be positive or regulive based on X, it will rever approach a true value for n