# **Machine Learning II**

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Week 3 Lecture - 20th October 2023

Non-linear regression methods 2

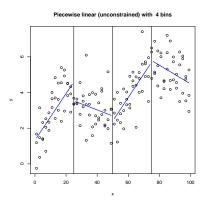


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### **Local Regression**

One of the non-linear function prediction methods we considered last week used piecewise linear regression, where a regression line is fitted to each interval



$$y_i = f(x_i) = f_0(x_i)C_0(x_i) + f_0(x_i)C_0(x_i) + \cdots + f_K(x_i)C_K(x_i) + \epsilon_i$$
 where each  $f_k(x_i) = \beta_{0k} + \beta_{1k}x$  is a linear function

We quickly moved on to adding constraints such as the piecewise function should be continuous and smooth.

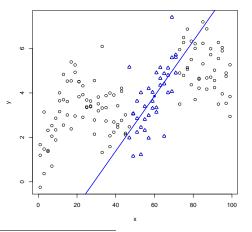
This week we will take this idea in another direction: Instead of four disjoint intervals each with a quarter of the points, we will use many overlapping intervals. We consider a *window* containing a quarter of the data points passing through the range of x.

To estimate  $f(x_0)$ , instead of finding the disjoint interval which contains  $x_0$ , we define a window containing take the  $\frac{1}{4}n$  points nearest to  $x_0$ .

Because we are using  $\frac{1}{4}n$  points, the *span* of the window is equal to  $\frac{1}{4}$ .

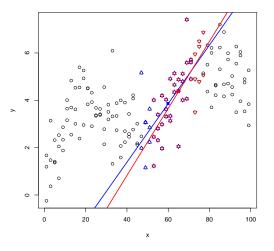
Example:  $x_0 = 60$  and n = 150, so  $\lceil \frac{n}{4} \rceil = 38^1$ .

The regression line using the 38 points nearest to  $x_0$  is f(x) = -3.6 + 0.12x, giving  $f(x_0) = f(60) = 3.6$ .



<sup>&</sup>lt;sup>1</sup>non-integers are rounded up.

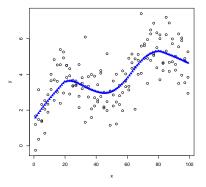
To predict f(65) we use the 38 nearest points to  $x_0 = 65$  which includes 31 of the blue points on the previous slide.



Blue:  $x_0 = 60$ 

Red:  $x_0 = 65$ 

We then repeat for all the points on a grid and plot them to give the predictor function f(x).



Note that the method does not require that  $x_0$  is a value in our data, we just need to identify which points are nearest to  $x_0$ .

The local regression algorithm, called *loess* (locally estimated least squares) or *lowess*, has a couple of important difference to the above explanation.

In the above example we computed the regression line giving equal weight to all points in the window.

Loess uses weighted regression, so that the points close to  $x_0$  have a large influence on f.

Points further away, but still within the window, have less influence.

Points outside the window have zero influence.

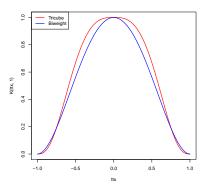
The weight function is defined using the distance of each  $x_i$  from  $x_0$ . The function used to calculate the weights is called a Kernel  $K(\delta)$ .

In loess  $\delta \propto x_i - x_0$ , so the further away a data point is from  $x_0$ , the smaller  $K(\delta)$  is.

There are lots of possibilities for the Kernel. They are all symmetric about the y axis, have a maximum at  $\delta$ =0 and are monotonic decreasing in  $|\delta|$ . There is a computational advantage when  $K(\delta) = 0$  when  $\delta$  is outside the window.

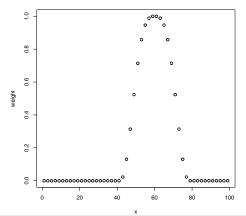
The two most common are Kernel functions for loess fitting are Biweight  $K(\delta)=(1-\delta^2)^2$  for  $|\delta|\leqslant 1$  and Tricube  $K(\delta)=(1-|\delta|^3)^3$  for  $|\delta|\leqslant 1$  (the default in R)

The unweighted example corresponds to using a *uniform* or *boxcar* kernel.  $K(\delta) = 1$  for  $|\delta| \le 1$ 



For a given  $x_0$  the weight of the *i*-th point is  $w_i = K\left(\frac{x_i - x_0}{s}\right)$  where s is the distance  $s = \max_i |x_i - x_0|$  using the furthest  $x_i$  still inside the window.

Here are the weights for each data point with a span of 0.4 and  $x_0 = 60$ .As the weight is more localised than when using the uniform kernel it makes sense to use a wider span.



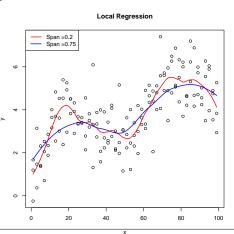
**Weighted least squares**: In principle the regression is carried out using standard least squares estimation, except that each residual is multiplied by the weight. Minimise  $RSS = \sum_{i=1}^{n} w_i (y_i - f_{\beta}(x_i))^2$ 

The other difference in the loess implementation is that usually the regression is a locally quadratic regression instead of locally linear regression. The loess function in R takes an argument degree= to specify locally constant, linear or quadratic.

N.B. If a uniform kernel is used and a constant estimated (degree=0) rather than estimating a linear/quadratic regression line, then the method is equivalent to K-Nearest Neighbours regression, which was covered in ML1.

### **Comments on local regression**

As with spline smoothing we need to specify the level of smoothing. This is done by defining the span. As the span is usually within the range [0,1], we have a reasonable starting point for the parameter unlike specifying  $\lambda$  in spline smoothing. The default in R is 0.75.



Loess is a good all-round method which can be used to gain a first impression of the relationship between *x* and *y*.

In R some plot methods add a loess curve by default, or make it easy to specify as an option. Example: The diagnostic plots for an R object of class lm add a red loess curve as a visual guide

#### **Generalised additive Models**

All the non-linear regression methods considered so far have used just one predictor variable X. We will now consider how we can adapt these methods to multiple predictor variables  $X_1, X_2, \ldots, X_p$ .

A multiple linear regression model has the form.

$$Y_i = \beta_0 + \beta_1 x_{i1} + \beta_2 x_{i2} + \ldots + \beta_p x_{ip} + \epsilon_i.$$

Non-linear equivalent is the generalised additive Model (GAM)

$$Y_i = \beta_0 + f_1(X_{i1}) + f_2(X_{i2}) + \ldots + f_p(X_{ip}) + \epsilon_i.$$

Note: there is another class of statistical models called generalised linear models (GLMs), which include logistic regression. These are not the same!

There is no constraint that each  $f_j$  has to be the same type of function. So  $f_1$  could be a quadratic function,  $f_2$  a smoothing spline function and  $f_3$  a loess function.

As with linear models the predictor variables can be either numeric or factor variables.

For convenience, we will assume for the rest of this lecture that each component function is a smoothing spline. The following methods generalise easily to any univariate linear or non-linear regression method.

How do we estimate each of the functions in the GAM?

We can easily estimate  $f_1(x_{i1})$  fitting univariate spline smoothing on  $(x_{i1}, y_i)$ .

To estimate  $f_2(x_{i2})$  we need to remove the fitted effect that  $f_1(x_{i1})$  has on  $y_i$  before using spline smoothing on  $X_2$ .

This means we calculate  $\eta_{i2} = y_i - f_1(x_{i1})$ , and fit univariate spline smoothing on  $(x_{i2}, \eta_{i2})$ .

To estimate  $f_3(x_{i3})$  we need to remove the fitted effect that  $f_1(x_{i1})$  and  $f_2(x_{i2})$  before using spline smoothing has on  $y_i$  before using spline smoothing on  $X_2$ . This means we calculate  $\eta_{i3} = y_i - f_1(x_{i1}) - f_2(x_{i2})$ , and fit univariate spline smoothing on  $(x_{i3}, \eta_{i3})$ .

The  $\eta$  values obtained by subtracting out the other components are called partial residuals.

The k-th partial residuals for a GAM model is defined as

$$\eta_{ik} = y_i - \beta_0 - f_1(x_{i1}) - \dots - f_{(k-1)}(x_{i(k-1)}) - f_{(k+1)}(x_{i(k+1)}) - \dots f_p(x_{ip})$$
for  $i = 1, \dots, n$ .

Notice all the other functions are removed **except**  $f_k(x_{ik})$ 

The full residuals are the error terms when the full model is fitted, the k-th partial residuals is the error term when everything except for  $f_k$  is fitted.

The method for obtaining all p estimating functions is called *backfitting*. An insight to the method is below, but the technical details are omitted.

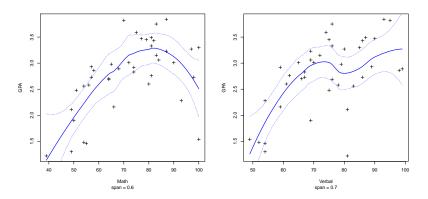
## **Backfitting algorithm for GAMs**

- 1) Set the constant parameter  $\widehat{\beta}_0 = \overline{y}$  this remains fixed throughout the algorithm.
- 2) Initialise all the component functions to  $f_k(x) = 0$
- 3) Iterate over  $k \in \{1, \dots, p\}$ 
  - a) Obtain the k-th partial residuals  $\eta_{ik}$ .
  - b) Obtain the spline smoothing function  $f_k(x)$ , which fits the  $\eta_{ik}$  partial residuals.
- 4) Repeat Step 3 until convergence, i.e. the change in the fitted values for each the component  $f_k(x)$  is less than a specified tolerance.

## **Example**

The GPA (grade point average) for matriculating freshmen in US universities is the grade attained when completing the 1st year at university. GPA is a continuous number on a scale from 1 (bad) to 4 (very good). The data set Freshmen contains the GPA score along with the percentage scores for a mathematics test and a verbal test, as part of their entrance exam.

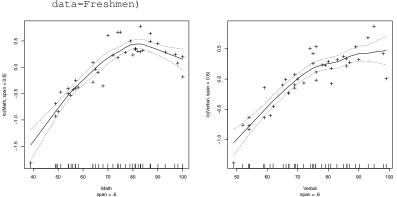
#### Loess curves for one variable models



The local regression *GPA dependent on math score* ignoring the verbal score seems to be non-linear, as does *GPA dependent on verbal score* ignoring the math score.

#### GAM with loess curve for Math and Verbal

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\label{eq:gam.mod=gam(GPA~lo(Math,span=0.6)+lo(Verbal,span=0.6),} $$ data=Freshmen)
```

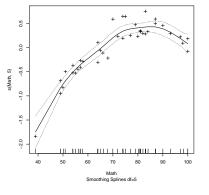


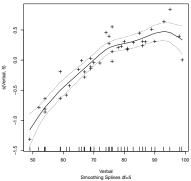
The dip in predicted GPA for verbal scores around 80 has disappeared when Math and Verbal are simultaneously fitted.

NB: the points plotted on the *y*-axis are the partial residuals and not GPA scores.

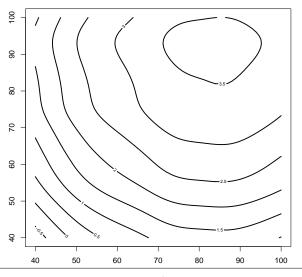
### GAM with spline smoothing curve for Math and Verbal

gam.mod=gam(GPA~s(Math, 5)+s(Verbal, 5), data=Freshmen)





## Contour plot of the predicted values using spline smoothing.



You can specify different combinations such as spline smoothing for Math and a quadratic polynomial for Verbal

gam.mod=gam(GPA~s(Math, 5)+poly(Verbal, 2), data=Freshmen)

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#### **Comments**

- GAMs allow us to fit a non-linear predictor to each variable. We will find non-linear relationships that standard multiple linear regression will miss. We do not need to manually try out many different transformations on each variable individually.
- ► The non-linear fits can potentially make more accurate predictions for the response Y.
- ▶ Because the model is additive, we can examine the effect of each X<sub>j</sub> on Y individually while holding all of the other variables fixed.
- ▶ The smoothness of the function  $f_j$  for the variable  $X_j$  can be controlled independently for each variable.

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#### Comments ctd.

- ▶ Up to now we have assumed all of the variables are continuous, or at least numeric. Factor variables are usually modelled using one parameter for each factor level. These can be directly added to the GAM with no need for fitting a function. Eg. In the GPA data, if there were another variable for type of high school: public vs private.
- ▶ A GAM is restricted to be additive. Important interactions might be missed. We can manually add an interaction term to the GAM model by adding a predictor for  $X_j X_k$ . This is normally only done if an interaction term between two variables is expected, because checking for all combinations of interactions adds a large computational overhead.

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