

CSEN1022: Machine Learning

Probabilistic Classifiers

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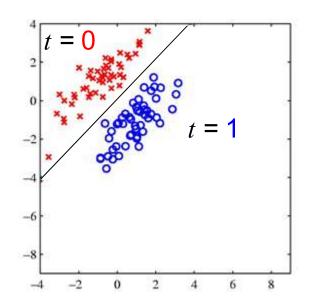
Linear Classification

Classification Problem

Input vector: x

Target variable (vector): t

- Discriminant Functions
 - Least Squares
 - Fisher's Linear Discriminant
 - Perceptron



- Probabilistic Generative Models
- Probabilistic Discriminative Models

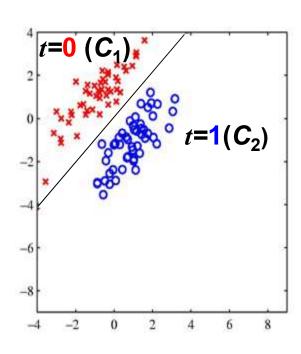
Bayes' Theorem

$$p(C_k|\mathbf{x}) = \frac{p(\mathbf{x}|C_k)p(C_k)}{p(\mathbf{x})}$$

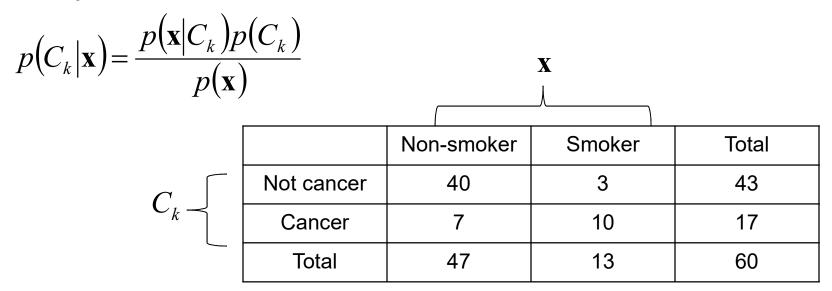
 $p(C_k)$: Prior probability

 $p(\mathbf{x}|C_k)$: Likelihood

 $p(C_k|\mathbf{x})$: Posterior probability (Knowing \mathbf{x} , what is the probability of C_k ?)



Example



If a patient is a smoker, would he have cancer?

$$p(Smoker|Cancer)=10/17, p(Cancer)=17/60, p(Smoker)=13/60$$

 $p(Cancer|Smoker)=(10/17).(17/60)/(13/60)=10/13$

p(Not cancer|Smoker)=3/13

C = Cancer

$$p(C_k|\mathbf{x}) = \frac{p(\mathbf{x}|C_k)p(C_k)}{p(\mathbf{x})}$$

- Inference Stage: Find $p(C_k|\mathbf{x})$
- Decision Stage: $k^* = \arg \max_k p(C_k | \mathbf{x})$
- Probabilistic Generative Model: Learns $p(\mathbf{x}|C_k)$ and $p(C_k)$
- Probabilistic Discriminative Model: Learns $p(C_k|\mathbf{x})$ directly

- Sometimes it is difficult to estimate $p(\mathbf{x}|C_k)$ for high dimensional data
- A simple solution is to assume independence between the features
- Naïve Bayes approximation

$$p(\mathbf{x}|C_k) \approx \prod_{j=1}^D p(x_j|C_k)$$

where D is the dimensionality of the input data

Example:

Consider the data about car theft given in the table below

Example No.	Color	Туре	Origin	Stolen?
1	Red	Sports	Domestic	Yes
2	Red	Sports	Domestic	No
3	Red	Sports	Domestic	Yes
4	Black	Sports	Domestic	No
5	Black	Sports	Imported	Yes
6	Black	SUV	Imported	No
7	Black	SUV	Imported	Yes
8	Black	SUV	Domestic	No
9	Red	SUV	Imported	No
10	Red	Sports	Imported	Yes

Using Naïve Bayes classifier, predict whether a Red Domestic SUV is stolen or not. Note that since the data is discrete, you can use the frequentist statistics to compute the needed probabilities.

Solution:

Since the goal is to classify a Red Domestic SUV as stolen or not, we first define two classes C_1 and C_2 , corresponding to Stolen = Yes and Stolen = No, respectively.

To classify the given car with attributes \mathbf{x} , we need to compute $p(C_1|\mathbf{x})$: $p(\text{Stolen} = \text{Yes} \mid \text{Color} = \text{Red}, \text{Type} = \text{SUV}, \text{Origin} = \text{Domestic})$

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and p(C_2|\mathbf{x}):

p(\text{Stolen = No | Color = Red, Type = SUV, Origin = Domestic})
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and find which conditional probability is larger. If the first one is larger, then our prediction is Stolen = Yes. If the second one is larger, then our prediction is Stolen = No. Note that \mathbf{x} here is 3 dimensional corresponding to Color, Type and Origin.

Since
$$p(C_1|\mathbf{x}) = \frac{p(\mathbf{x}|C_1)p(C_1)}{p(\mathbf{x})}$$

We need to compute $p(\mathbf{x}|C_1) = p(\text{Color} = \text{Red}, \text{Type} = \text{SUV}, \text{Origin} = \text{Domestic} \mid \text{Stolen} = \text{Yes}).$

Using the Naïve Bayes assumption which assumes that the dimensions of the input data (the attributes of the car) are independent, we can re-write $p(\mathbf{x}|C_1)$ as

$$p(\mathbf{x}|C_1) = \prod_{i=1}^{D} p(x_i|C_1)$$

= $p(\text{Color} = \text{Red} \mid \text{Stolen} = \text{Yes}) p(\text{Type} = \text{SUV} \mid \text{Stolen} = \text{Yes}) p(\text{Origin} = \text{Domestic} \mid \text{Stolen} = \text{Yes})$

Similarly, $p(\mathbf{x}|C_2)$ can be re-written as $p(\mathbf{x}|C_2) = \prod_{i=1}^{D} p(x_i|C_2)$

= $p(\text{Color} = \text{Red} \mid \text{Stolen} = \text{No}) p(\text{Type} = \text{SUV} \mid \text{Stolen} = \text{No}) p(\text{Origin} = \text{Domestic} \mid \text{Stolen} = \text{No})$

• From the available data in the table and using frequentist statistics:

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p(\text{Color} = \text{Red} \mid \text{Stolen} = \text{Yes}) = 3/5 (out of the 5 stolen cars, 3 were red) p(\text{Color} = \text{Red} \mid \text{Stolen} = \text{No}) = 2/5 (out of the 5 non-stolen cars, 2 were red) p(\text{Type} = \text{SUV} \mid \text{Stolen} = \text{Yes}) = 1/5 (out of the 5 stolen cars, 1 was SUV) p(\text{Type} = \text{SUV} \mid \text{Stolen} = \text{No}) = 3/5 p(\text{Origin} = \text{Domestic} \mid \text{Stolen} = \text{Yes}) = 2/5 p(\text{Origin} = \text{Domestic} \mid \text{Stolen} = \text{No}) = 3/5
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Therefore,
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p(\text{Color} = \text{Red} \mid \text{Stolen} = \text{Yes}) p(\text{Type} = \text{SUV} \mid \text{Stolen} = \text{Yes}) p(\text{Origin} = \text{Domestic} \mid \text{Stolen} = \text{Yes}) = (3/5) \times (1/5) \times (2/5) = 0.048
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And

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p(\text{Color} = \text{Red} \mid \text{Stolen} = \text{No}) \ p(\text{Type} = \text{SUV} \mid \text{Stolen} = \text{No}) \ p(\text{Origin} = \text{Domestic} \mid \text{Stolen} = \text{No}) = (2/5) \times (3/5) \times (3/5) = 0.144
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Also
$$p(Stolen = Yes) = 5/10$$
 and $p(Stolen = No) = 5/10$

To classify the given car, we need to compare $p(C_1|\mathbf{x})$ to $p(C_2|\mathbf{x})$ such that

If
$$\frac{p(C_1|\mathbf{x})}{p(C_2|\mathbf{x})} > 1$$
 then $\mathbf{x} \in C_1$, otherwise $\mathbf{x} \in C_2$

$$\therefore \frac{p(C_1|\mathbf{x})}{p(C_2|\mathbf{x})} = \frac{p(\mathbf{x}|C_1)p(C_1)}{p(\mathbf{x}|C_2)p(C_2)}$$

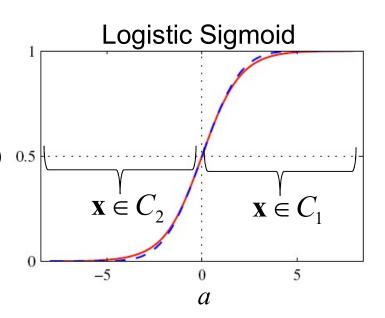
Therefore, for the this problem

$$\frac{p(C_1|\mathbf{x})}{p(C_2|\mathbf{x})} = \frac{p(\mathbf{x}|C_1)p(C_1)}{p(\mathbf{x}|C_2)p(C_2)} = \frac{0.048 \times 0.5}{0.144 \times 0.5} = 0.333$$

Therefore, our prediction is C_2 which is that the car is not stolen.

 We first represent the posterior probability as

$$p(C_1|\mathbf{x}) = \frac{p(\mathbf{x}|C_1)p(C_1)}{p(\mathbf{x}|C_1)p(C_1) + p(\mathbf{x}|C_2)p(C_2)} \stackrel{\sigma(a)}{=} \frac{0.5}{1 + \exp(-a)}$$



where

$$a = \ln \frac{p(\mathbf{x}|\mathcal{C}_1)p(\mathcal{C}_1)}{p(\mathbf{x}|\mathcal{C}_2)p(\mathcal{C}_2)}$$

If
$$a > 0 \rightarrow \frac{p(\mathbf{x}|\mathcal{C}_1)p(\mathcal{C}_1)}{p(\mathbf{x}|\mathcal{C}_2)p(\mathcal{C}_2)} > 1 \rightarrow p(C_1|\mathbf{x}) > p(C_2|\mathbf{x}) \rightarrow \mathbf{x} \in C_1$$

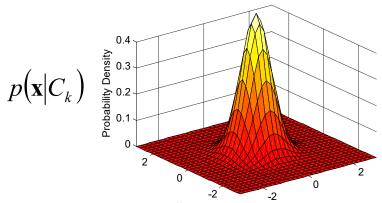
If
$$a < 0 \rightarrow \frac{p(\mathbf{x}|\mathcal{C}_1)p(\mathcal{C}_1)}{p(\mathbf{x}|\mathcal{C}_2)p(\mathcal{C}_2)} < 1 \rightarrow p(C_1|\mathbf{x}) < p(C_2|\mathbf{x}) \rightarrow \mathbf{x} \in C_2$$

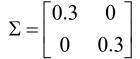
Assume Gaussian distribution for class-conditional densities $p(\mathbf{x}|\mathcal{C}_k)$

$$p(\mathbf{x}|\mathcal{C}_k) = \frac{1}{(2\pi)^{D/2}} \frac{1}{|\mathbf{\Sigma}|^{1/2}} \exp\left\{-\frac{1}{2}(\mathbf{x} - \boldsymbol{\mu}_k)^{\mathrm{T}} \mathbf{\Sigma}^{-1} (\mathbf{x} - \boldsymbol{\mu}_k)\right\} \qquad \begin{array}{c} \boldsymbol{\mu}_k : \mathsf{Mean} \\ \boldsymbol{\Sigma} : \mathsf{Covar} \end{array}$$

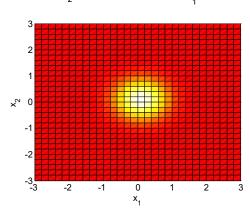
: Covariance Matrix (Common for both classes)

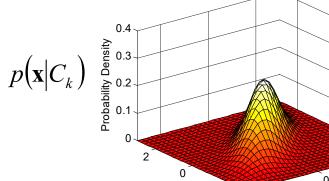
Example





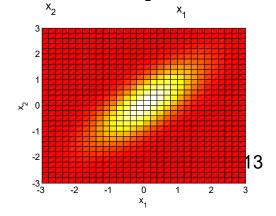
Diagonal Covariance





$$\Sigma = \begin{bmatrix} 1 & 0.75 \\ 0.75 & 1 \end{bmatrix}$$

Non-diagonal Covariance



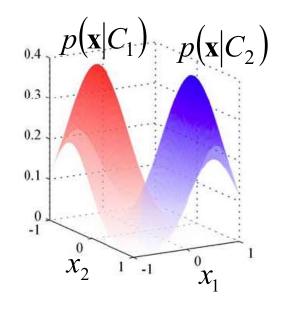
 Using Gaussian assumption and logistic sigmoid representation, we can show that

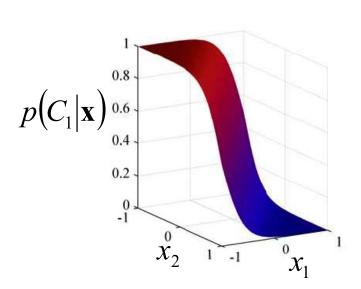
$$p(\mathcal{C}_1|\mathbf{x}) = \sigma(\mathbf{w}^{\mathrm{T}}\mathbf{x} + w_0)$$

where

$$\mathbf{w} = \mathbf{\Sigma}^{-1}(\mu_1 - \mu_2)$$

$$w_0 = -\frac{1}{2}\mu_1^{\mathrm{T}}\mathbf{\Sigma}^{-1}\mu_1 + \frac{1}{2}\mu_2^{\mathrm{T}}\mathbf{\Sigma}^{-1}\mu_2 + \ln\frac{p(C_1)}{p(C_2)}$$





• Using the previous assumptions and a training dataset $\{x_n, t_n\}$, we can estimate the values of the parameters

$$\mu_k \quad \Sigma \quad p(\mathcal{C}_k)$$

- Maximum Likelihood Estimation (MLE)
 - Let t_n = 1 denote class C_1 and t_n = 0 denote class C_2
 - Let $p(\mathcal{C}_1) = \pi \rightarrow p(\mathcal{C}_2) = 1 \pi$
 - For $\mathbf{x}_n \in C_1$ $p(\mathbf{x}_n, \mathcal{C}_1) = p(\mathcal{C}_1)p(\mathbf{x}_n|\mathcal{C}_1) = \pi \mathcal{N}(\mathbf{x}_n|\boldsymbol{\mu}_1, \boldsymbol{\Sigma})$
 - For $\mathbf{X}_n \in C_2$

$$p(\mathbf{x}_n, \mathcal{C}_2) = p(\mathcal{C}_2)p(\mathbf{x}_n|\mathcal{C}_2) = (1 - \pi)\mathcal{N}(\mathbf{x}_n|\boldsymbol{\mu}_2, \boldsymbol{\Sigma})$$

- Maximum Likelihood Estimation (MLE)
 - We define the likelihood of the data as the probability of observing the available data
 - Let **D** denote the available data, where **D** = $\{\mathbf{x}_1, \mathbf{x}_2, ..., \mathbf{x}_N\}$
 - The probability of observing the data $p(\mathbf{D})$ (Given the parameters)

$$p(\mathbf{D}) = p(\mathbf{x}_1, \mathbf{x}_2, ..., \mathbf{x}_N)$$

Assuming independence between the input vectors

$$p(\mathbf{D}) = p(\mathbf{x}_1)p(\mathbf{x}_2) \dots p(\mathbf{x}_N) = \prod_{n=1}^{N} p(\mathbf{x}_n) = \prod_{i=1}^{N_1} p(\mathbf{x}_i, C_1) \prod_{j=1}^{N_2} p(\mathbf{x}_j, C_2)$$

 Given the assumption of the previous slide, the likelihood can be modeled as

$$p(\mathbf{t}|\pi, \boldsymbol{\mu}_1, \boldsymbol{\mu}_2, \boldsymbol{\Sigma}) = \prod_{n=1}^{N} \left[\pi \mathcal{N}(\mathbf{x}_n | \boldsymbol{\mu}_1, \boldsymbol{\Sigma})\right]^{t_n} \left[(1-\pi) \mathcal{N}(\mathbf{x}_n | \boldsymbol{\mu}_2, \boldsymbol{\Sigma})\right]^{1-t_n}$$

Maximum Likelihood Estimation

MLE Solution:

- Find the parameters that maximize the likelihood of the data
- Given that the maximum of this function is achieved at the same value that maximizes the log of the function, we use the Loglikelihood of the function for convenience
- Remember that ln(AB) = ln(A) + ln(B), $ln(A^m) = m ln(A)$

$$\ln p(\mathbf{t}|\pi, \mathbf{\mu}_{1}, \mathbf{\mu}_{2}, \Sigma) = \ln \prod_{n=1}^{N} \left[\pi \mathscr{N}(\mathbf{x}_{n}|\mathbf{\mu}_{1}, \Sigma)\right]^{t_{n}} \left[(1-\pi)\mathscr{N}(\mathbf{x}_{n}|\mathbf{\mu}_{2}, \Sigma)\right]^{t_{n}} \\
= \sum_{n=1}^{N} \ln \left(\left[\pi \mathscr{N}(\mathbf{x}_{n}|\mathbf{\mu}_{1}, \Sigma)\right]^{t_{n}} \left[(1-\pi)\mathscr{N}(\mathbf{x}_{n}|\mathbf{\mu}_{2}, \Sigma)\right]^{t_{n}}\right) \\
= \sum_{n=1}^{N} t_{n} \ln \left(\pi \mathscr{N}(\mathbf{x}_{n}|\mathbf{\mu}_{1}, \Sigma)\right) + (1-t_{n}) \ln \left((1-\pi)\mathscr{N}(\mathbf{x}_{n}|\mathbf{\mu}_{2}, \Sigma)\right) \\
= \sum_{n=1}^{N} t_{n} \ln (\pi) + t_{n} \ln \left(\mathscr{N}(\mathbf{x}_{n}|\mathbf{\mu}_{1}, \Sigma)\right) + (1-t_{n}) \ln \left((1-\pi)\right) + (1-t_{n}) \ln \left(\mathscr{N}(\mathbf{x}_{n}|\mathbf{\mu}_{2}, \Sigma)\right)$$

Maximum Likelihood Estimation

$$\ln p(\mathbf{t}|\pi, \mathbf{\mu}_1, \mathbf{\mu}_2, \Sigma)$$

$$= \sum_{n=1}^{N} t_n \ln(\pi) + t_n \ln(\mathcal{N}(\mathbf{x}_n|\mathbf{\mu}_1, \Sigma)) + (1 - t_n) \ln((1 - \pi)) + (1 - t_n) \ln(\mathcal{N}(\mathbf{x}_n|\mathbf{\mu}_2, \Sigma))$$

Derivative w.r.t. parameters and equate with zero

$$\pi \qquad \qquad \pi = \frac{1}{N} \sum_{n=1}^{N} t_n = \frac{N_1}{N} = \frac{N_1}{N_1 + N_2}$$

$$\mu_1 \qquad \qquad \mu_1 = \frac{1}{N_1} \sum_{n=1}^{N} t_n \mathbf{x}_n$$

$$\mu_2 \qquad \qquad \mu_2 = \frac{1}{N_2} \sum_{n=1}^{N} (1 - t_n) \mathbf{x}_n \qquad \mathbf{S}_1 = \frac{1}{N_1} \sum_{n \in \mathcal{C}_1} (\mathbf{x}_n - \mu_1) (\mathbf{x}_n - \mu_1)^{\mathrm{T}}$$

$$\Sigma \qquad \qquad \Sigma = \frac{N_1}{N} \mathbf{S}_1 + \frac{N_2}{N} \mathbf{S}_2 \qquad \mathbf{S}_2 = \frac{1}{N_2} \sum_{n \in \mathcal{C}_2} (\mathbf{x}_n - \mu_2) (\mathbf{x}_n - \mu_2)^{\mathrm{T}}$$

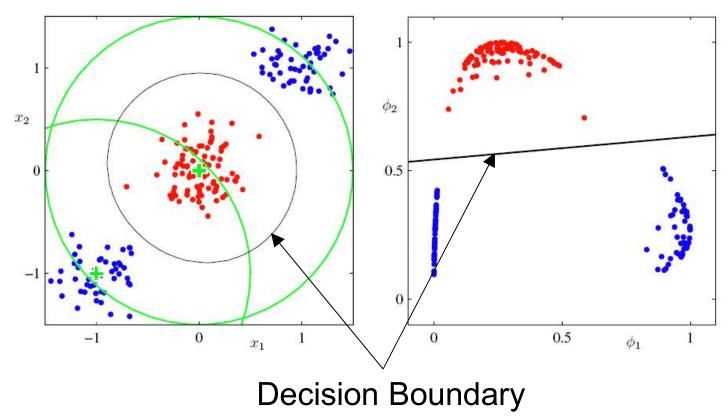
• Using the expression in slide 8, $p(C_1|\mathbf{x}) = \sigma(\mathbf{w}^T\mathbf{x} + w_0)$ can be obtained

$$p(C_k|\mathbf{x}) = \frac{p(\mathbf{x}|C_k)p(C_k)}{p(\mathbf{x})}$$

- Inference Stage: Find $p(C_k|\mathbf{x})$
- Decision Stage: $k^* = \max_k p(C_k|\mathbf{x})$
- Probabilistic Generative Model: Learns $p(\mathbf{x}|C_k)$ and $p(C_k)$
- Probabilistic Discriminative Model: Learns $p(C_k|\mathbf{x})$ directly

Probabilistic Discriminative Models

- Learn $p(C_k|\mathbf{x})$ directly
- Using a nonlinear transformation of the data $\phi(\mathbf{x})$ (basis function)

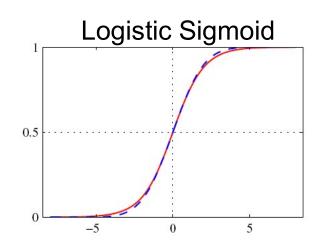


Logistic Regression

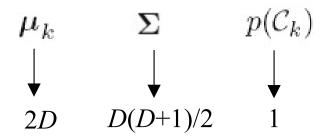
Assumption

$$p(C_1|\phi) = y(\phi) = \sigma\left(\mathbf{w}^{\mathrm{T}}\phi\right)$$

where $\sigma(a) = \frac{1}{1 + \exp(-a)}$



- Logistic Regression vs. MLE for Gaussian Generative Model (For 2 classes)
 - No. of Parameters for Logistic Regression: D parameters
 - No. of Parameters for MLE: D(D+5)/2 + 1 parameters



Logistic Regression

$$p(C_1|\phi) = y(\phi) = \sigma(\mathbf{w}^T\phi)$$

• Using maximum likelihood to estimate w from the training dataset $\{\varphi_n, t_n\}$. Likelihood function

$$p(\mathbf{t}|\mathbf{w}) = \prod_{n=1}^{N} y_n^{t_n} \{1 - y_n\}^{1 - t_n}$$

 Minimizing the negative of the log-likelihood is equivalent to maximizing the log-likelihood

$$E(\mathbf{w}) = -\ln p(\mathbf{t}|\mathbf{w}) = -\sum_{n=1}^{N} \{t_n \ln y_n + (1 - t_n) \ln(1 - y_n)\}$$

Logistic Regression

We make use of the property

$$\frac{d\sigma}{da} = \sigma(1 - \sigma)$$

• Take the derivative of $E(\mathbf{w})$ w.r.t. \mathbf{w}

$$\nabla E(\mathbf{w}) = -\sum_{n=1}^{N} \left\{ t_n \frac{1}{\sigma_n} \sigma_n (1 - \sigma_n) \phi_n + (1 - t_n) \frac{1}{1 - \sigma_n} (-\sigma_n (1 - \sigma_n)) \phi_n \right\}$$

$$= -\sum_{n=1}^{N} \left\{ t_n (1 - \sigma_n) \phi_n - (1 - t_n) \sigma_n \phi_n \right\}$$

$$= \sum_{n=1}^{N} (y_n - t_n) \phi_n$$

Optimal w can be found using gradient descent

$$\mathbf{w}^{(\tau+1)} = \mathbf{w}^{(\tau)} - \eta \frac{\partial E}{\partial \mathbf{w}^{(\tau)}}$$