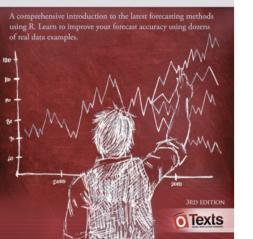
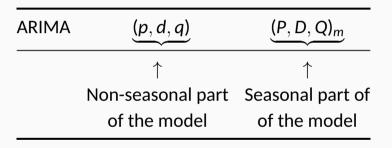
Rob J Hyndman George Athanasopoulos

# FORECASTING PRINCIPLES AND PRACTICE



## 9. ARIMA models

9.9 Seasonal ARIMA models
OTexts.org/fpp3/



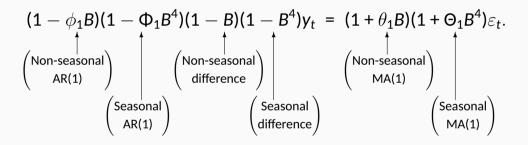
where m = number of observations per year.

E.g., ARIMA(1, 1, 1)(1, 1, 1)<sub>4</sub> model (without constant)

E.g., ARIMA $(1, 1, 1)(1, 1, 1)_4$  model (without constant)

$$(1 - \phi_1 B)(1 - \Phi_1 B^4)(1 - B)(1 - B^4)y_t = (1 + \theta_1 B)(1 + \Theta_1 B^4)\varepsilon_t.$$

E.g., ARIMA $(1, 1, 1)(1, 1, 1)_4$  model (without constant)



E.g., ARIMA $(1, 1, 1)(1, 1, 1)_4$  model (without constant)

$$(1-\phi_1 B)(1-\Phi_1 B^4)(1-B)(1-B^4) y_t \ = \ (1+\theta_1 B)(1+\Theta_1 B^4) \varepsilon_t.$$

All the factors can be multiplied out and the general model written as follows:

$$\begin{aligned} y_{t} &= (1 + \phi_{1})y_{t-1} - \phi_{1}y_{t-2} + (1 + \Phi_{1})y_{t-4} \\ &- (1 + \phi_{1} + \Phi_{1} + \phi_{1}\Phi_{1})y_{t-5} + (\phi_{1} + \phi_{1}\Phi_{1})y_{t-6} \\ &- \Phi_{1}y_{t-8} + (\Phi_{1} + \phi_{1}\Phi_{1})y_{t-9} - \phi_{1}\Phi_{1}y_{t-10} \\ &+ \varepsilon_{t} + \theta_{1}\varepsilon_{t-1} + \Theta_{1}\varepsilon_{t-4} + \theta_{1}\Theta_{1}\varepsilon_{t-5}. \end{aligned}$$

The seasonal part of an AR or MA model will be seen in the seasonal lags of the PACF and ACF.

## ARIMA $(0,0,0)(0,0,1)_{12}$ will show:

- a spike at lag 12 in the ACF but no other significant spikes.
- The PACF will show exponential decay in the seasonal lags; that is, at lags 12, 24, 36, ....

## ARIMA $(0,0,0)(1,0,0)_{12}$ will show:

- exponential decay in the seasonal lags of the ACF
- a single significant spike at lag 12 in the PACF.

# How does ARIMA() work for seasonal models?

#### A seasonal ARIMA process

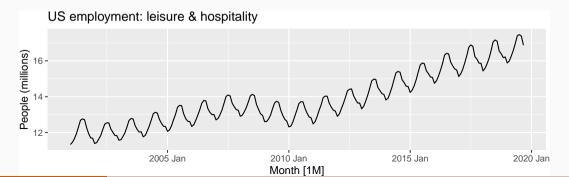
$$\Phi(B)\phi(B)(1-B)^d(1-B)^Dy_t = c + \Theta(B)\theta(B)\varepsilon_t$$

Need to select appropriate orders: d, D, p, q, P, Q, and whether to include the intercept c.

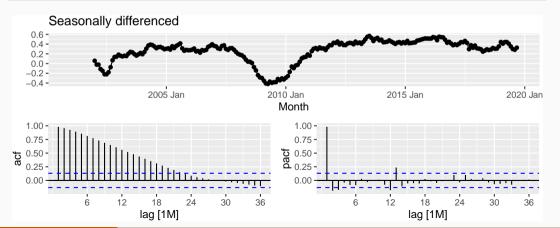
## Hyndman and Khandakar (JSS, 2008) algorithm:

- Select no. differences d via KPSS test and D using seasonal strength.
- Select p, q, P, Q and c by minimising AICc.
- Use stepwise search to traverse model space.

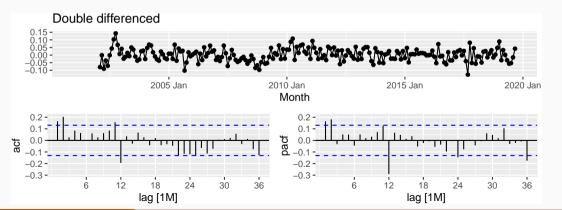
```
leisure <- us_employment |>
  filter(Title == "Leisure and Hospitality", year(Month) > 2000) |>
  mutate(Employed = Employed / 1000) |>
  select(Month, Employed)
autoplot(leisure, Employed) +
  labs(title = "US employment: leisure & hospitality", y = "People (millions)")
```



```
leisure |>
    gg_tsdisplay(difference(Employed, 12), plot_type = "partial", lag = 36) +
    labs(title = "Seasonally differenced", y = "")
```

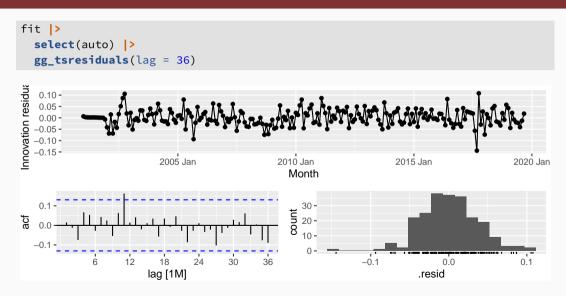


```
leisure |>
   gg_tsdisplay(difference(Employed, 12) |> difference(),
   plot_type = "partial", lag = 36) +
   labs(title = "Double differenced", y = "")
```



```
fit <- leisure |>
  model(arima012011 = ARIMA(Employed ~ pdq(0, 1, 2) + PDQ(0, 1, 1)),
        arima210011 = ARIMA(Employed ~ pdq(2, 1, 0) + PDQ(0, 1, 1)),
        auto = ARIMA(Employed, stepwise = FALSE, approx = FALSE))
fit |>
  pivot_longer(everything(),
    names to = "Model name".
    values to = "Orders")
## # A mable: 3 x 2
```

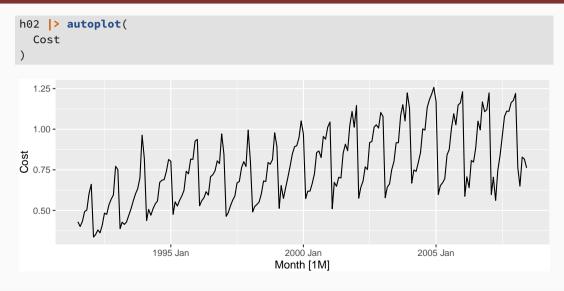
```
glance(fit) |>
  arrange(AICc) |>
  select(.model:BIC)
```



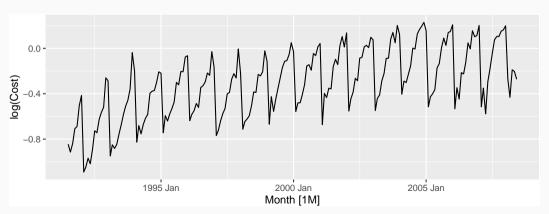
```
augment(fit) |>
  filter(.model == "auto") |>
  features(.innov, ljung_box, lag = 24, dof = 4)
```

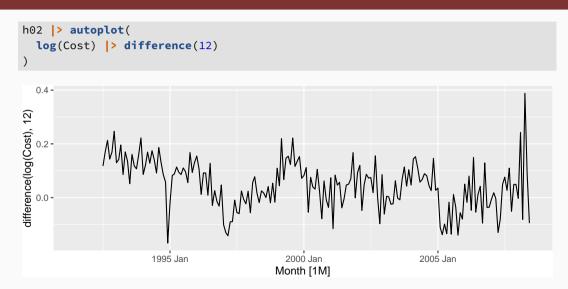
```
forecast(fit, h = 36) |>
  filter(.model == "auto") |>
  autoplot(leisure) +
  labs(title = "US employment: leisure & hospitality", y = "People (millions)")
     US employment: leisure & hospitality
      17.5 -
People (millions)
                                                                     level
                                                                        80
  15.0 -
                                                                        95
  12.5 -
  2000 Jan
                                 Month
```

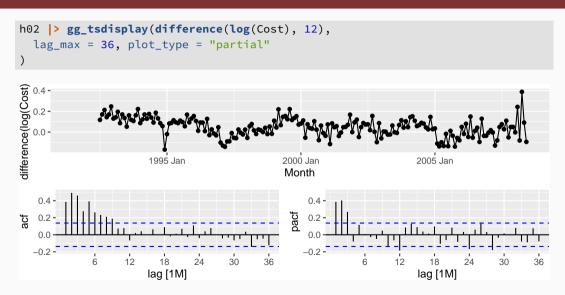
```
h02 <- PBS |>
filter(ATC2 == "H02") |>
summarise(Cost = sum(Cost) / 1e6)
```



```
h02 |> autoplot(
  log(Cost)
)
```



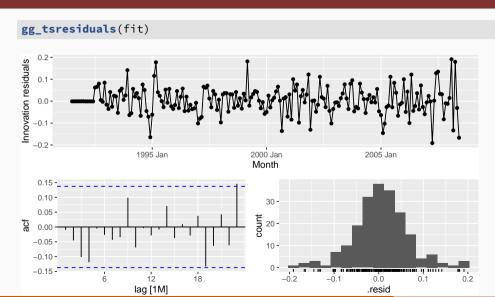




- Choose D = 1 and d = 0.
- Spikes in PACF at lags 12 and 24 suggest seasonal AR(2) term.
- Spikes in PACF suggests possible non-seasonal AR(3) term.
- Initial candidate model: ARIMA(3,0,0)(2,1,0)<sub>12</sub>.

.model	AICc
ARIMA(3,0,1)(0,1,2)[12]	-485
ARIMA(3,0,1)(1,1,1)[12]	-484
ARIMA(3,0,1)(0,1,1)[12]	-484
ARIMA(3,0,1)(2,1,0)[12]	-476
ARIMA(3,0,0)(2,1,0)[12]	-475
ARIMA(3,0,2)(2,1,0)[12]	-475
ARIMA(3,0,1)(1,1,0)[12]	-463

```
fit <- h02 |>
 model(best = ARIMA(log(Cost) \sim 0 + pdq(3, 0, 1) + PDQ(0, 1, 2)))
report(fit)
## Series: Cost
## Model: ARIMA(3,0,1)(0,1,2)[12]
## Transformation: log(Cost)
##
## Coefficients:
##
     ar1 ar2 ar3 ma1 sma1
                                               sma2
  -0.160 0.5481 0.5678 0.383 -0.5222 -0.1768
##
## s.e. 0.164 0.0878 0.0942 0.190 0.0861 0.0872
##
## sigma^2 estimated as 0.004278: log likelihood=250
## ATC=-486 ATCc=-485 BTC=-463
```



.model lb\_stat lb\_pvalue

## <chr> <dbl> <dbl> \*dbl> \*dbl| \*db

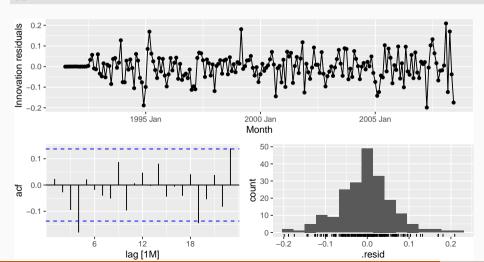
##

```
augment(fit) |>
  features(.innov, ljung_box, lag = 36, dof = 6)

## # A tibble: 1 x 3
```

```
fit <- h02 |> model(auto = ARIMA(log(Cost)))
report(fit)
## Series: Cost
## Model: ARIMA(2,1,0)(0,1,1)[12]
## Transformation: log(Cost)
##
## Coefficients:
##
         ar1 ar2 sma1
   -0.8491 -0.4207 -0.6401
##
## s.e. 0.0712 0.0714 0.0694
##
## sigma^2 estimated as 0.004387: log likelihood=245
## ATC=-483 ATCc=-483 BTC=-470
```





.model lb\_stat lb\_pvalue

## <chr> <dbl> <dbl> \*dbl> \*dbl> ## 1 auto 59.3 0.00332

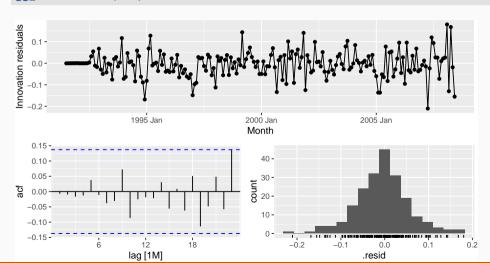
##

```
augment(fit) |>
  features(.innov, ljung_box, lag = 36, dof = 3)
## # A tibble: 1 x 3
```

```
fit <- h02 |>
 model(best = ARIMA(log(Cost),
   stepwise = FALSE, approximation = FALSE,
   order constraint = p + q + P + 0 \le 9)
report(fit)
## Series: Cost
## Model: ARIMA(4,1,1)(2,1,2)[12]
## Transformation: log(Cost)
##
## Coefficients:
##
           arl ar2 ar3 ar4 mal sar1 sar2 sma1 sma2
## -0.0425 0.210 0.202 -0.227 -0.742 0.621 -0.383 -1.202 0.496
## s.e. 0.2167 0.181 0.114 0.081 0.207 0.242 0.118 0.249 0.213
##
## sigma^2 estimated as 0.004049: log likelihood=254
## ATC=-489 ATCc=-487 BTC=-456
```

28

#### gg\_tsresiduals(fit)



```
augment(fit) |>
features(.innov, ljung_box, lag = 36, dof = 9)
```

```
## # A tibble: 1 x 3
## .model lb_stat lb_pvalue
## <chr> <dbl> <dbl>
## 1 best 36.5 0.106
```

Training data: July 1991 to June 2006

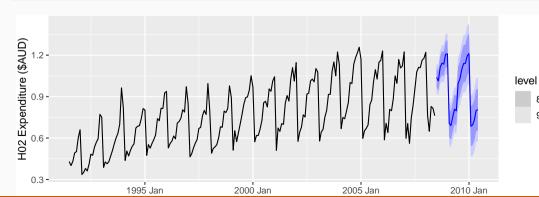
Test data: July 2006-June 2008

```
fit <- h02 |>
  filter index(~ "2006 Jun") |>
  model(
    ARIMA(log(Cost) \sim 0 + pdq(3, 0, 0) + PDQ(2, 1, 0)),
    ARIMA(log(Cost) \sim 0 + pdg(3, 0, 1) + PDO(2, 1, 0)),
    ARIMA(log(Cost) \sim 0 + pdg(3, 0, 2) + PDO(2, 1, 0)),
    ARIMA(log(Cost) \sim 0 + pdq(3, 0, 1) + PDQ(1, 1, 0))
    # ... #
fit |>
  forecast(h = "2 years") |>
  accuracy(h02)
```

.model	RMSE
ARIMA(3,0,1)(1,1,1)[12]	0.0619
ARIMA(3,0,1)(0,1,2)[12]	0.0621
ARIMA(3,0,1)(0,1,1)[12]	0.0630
ARIMA(2,1,0)(0,1,1)[12]	0.0630
ARIMA(4,1,1)(2,1,2)[12]	0.0631
ARIMA(3,0,2)(2,1,0)[12]	0.0651
ARIMA(3,0,1)(2,1,0)[12]	0.0653
ARIMA(3,0,1)(1,1,0)[12]	0.0666
ARIMA(3,0,0)(2,1,0)[12]	0.0668

- Models with lowest AICc values tend to give slightly better results than the other models.
- AICc comparisons must have the same orders of differencing.
   But RMSE test set comparisons can involve any models.
- Use the best model available, even if it does not pass all tests.

```
fit <- h02 |>
  model(ARIMA(Cost ~ 0 + pdq(3, 0, 1) + PDQ(0, 1, 2)))
fit |>
  forecast() |>
  autoplot(h02) + labs(y = "H02 Expenditure ($AUD)")
```



80 95