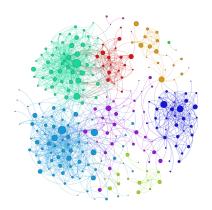
Foundations of Statistical Learning

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名人名言

"All models are wrong but some are useful." – George Box

"The existence of a problem in knowledge depends on the future being different from the past, while the possibility of a solution of the problem depends on the future being like the past." – Frank Knight

- Given variables x and y, suppose we are interested in predicting the value of y based on the value of x.
 - x: feature; input; predictor; independent variable
 - ▶ y: target; output; response; dependent variable
- For simplicity, assume there exists a function f such that $y = f(x)^1$. f is the **target function** that we want to learn: to predict the value of y is to learn f^2 .

¹i.e., given x, y is completely determined.

²In the statistics and econometrics literature, **learning** is called **estimation**. In this lecture, we use the two terms interchangeably.

- Let the observed data be $\mathcal{D} = \{(x_1, y_1), \dots, (x_N, y_N)\}.$
- Start with a set of candidate hypotheses which you think are likely to represent f:

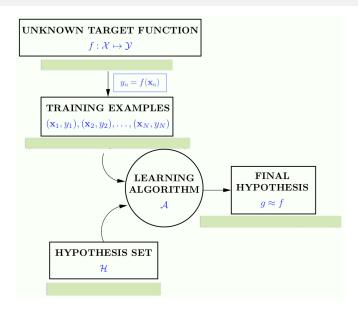
$$\mathcal{H} = \{h_1, h_2, \ldots\}$$

is called a **hypothesis set** or a **model**³.

• Based on \mathcal{D} , use an algorithm to select a hypothesis g from \mathcal{H} . Goal: $g \approx f$.

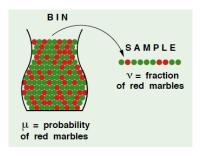
³Let $\theta \in \Theta$ be a set of parameters. If h_1, h_2, \ldots are functions of θ , such that $h_1 = h(\theta_1), h_2 = h(\theta_2), \ldots$, then we say $\mathcal{H} = \{h_1, h_2, \ldots\}$ is **parametrized** by θ and can be written as $\mathcal{H} = \{h(\theta) : \theta \in \Theta\}$.





Is Learning Feasible?

- bin with red and green marbles.
- pick a sample of N marbles independently.
- μ : probability of picking a red marble
- $m{\circ}$ ν : fraction of red marbles in the sample



- Can we say anything about μ after observing ν ?
 - No, sample can be mostly green while bin is mostly red.
 - Yes, sample frequency ν is likely close to bin frequency μ .
 - possible vs. probable

Probability to the Rescue

Hoeffding's Inequality

Let z_1, \ldots, z_N be N independent Bernoulli random variables with $\Pr(z_i=1)=\mu$ and $\Pr(z_i=0)=1-\mu$. Let $\nu=\frac{1}{N}\sum_{i=1}^N z_i$. Then for any $\epsilon>0^{a,b}$,

$$\Pr\left(|\nu - \mu| > \epsilon\right) \le 2e^{-2\epsilon^2 N}$$

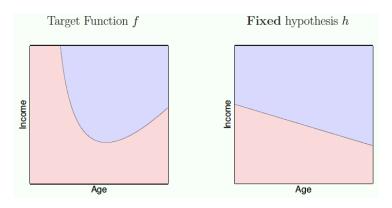
 ${}^{\mathrm{a}}\mathrm{Note}$ that ν is random, but observed. μ is fixed, but unobserved.

 b e.g., draw a sample of N=1000 and observe $\nu.$ Then,

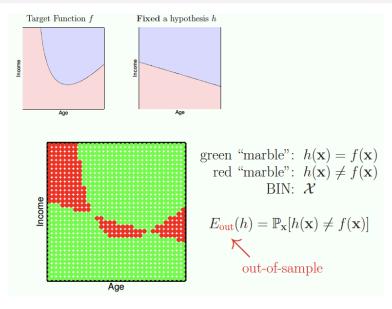
$$\Pr\left(|\nu - \mu| > 0.05\right) \le 0.014$$

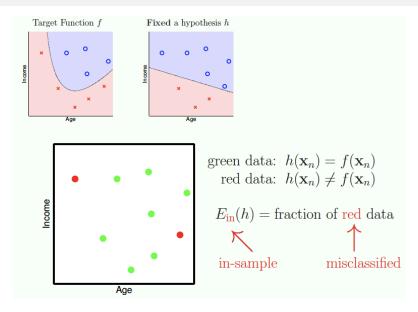
 $\Pr\left(|\nu - \mu| > 0.10\right) \le 0.000000004$

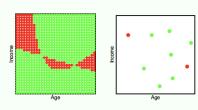
As we will see, learning is feasible in a probabilistic sense.



In learning, the unknown is an entire function f







Unknown f and P(x), fixed h

Learning

input space \mathcal{X}

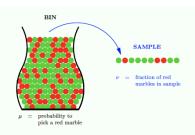
 \mathbf{x} for which $h(\mathbf{x}) = f(\mathbf{x})$

x for which $h(\mathbf{x}) \neq f(\mathbf{x})$

 $P(\mathbf{x})$

data set \mathcal{D}

Out-of-sample Error: $E_{\text{out}}(h) = \mathbb{P}_{\mathbf{x}}[h(\mathbf{x}) \neq f(\mathbf{x})]$ In-sample Error: $E_{\text{in}}(h) = \frac{1}{N} \sum_{n=1}^{N} \llbracket h(\mathbf{x}) \neq f(\mathbf{x}) \rrbracket$



Bin Model

Bin

- green marble
- red marble

randomly picking a marble sample of N marbles

 $\mu = \text{probability of picking a red marble}$ ν = fraction of red marbles in the sample



According to Hoeffding's inequality,

$$\Pr\left(\left|E_{in}\left(h\right) - E_{out}\left(h\right)\right| > \epsilon\right) \le 2e^{-2\epsilon^2 N} \tag{1}$$

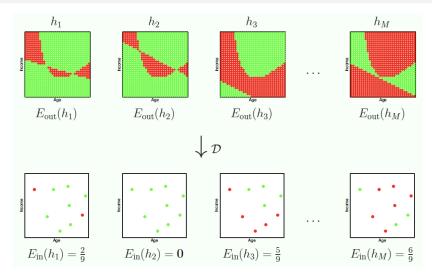
- $-E_{in}$: in-sample error; training error; empirical error; empirical risk
- $-E_{out}$: out-of-sample error; expected error; prediction error; risk
- (1) says that for a given h, given large enough N,
 - $E_{in} \approx E_{out}$.
 - If $E_{in} \approx 0$, then $E_{out} \approx 0$. In this case we have learned something about $f: f \approx h$ over \mathcal{X} .

The key assumptions that are needed for (1) to hold are:

- The data set \mathcal{D} is a random sample, i.e. data points are drawn *i.i.d.* from the underlying distribution.
 - ▶ In order to say something about unobserved data, we need to assume that they resemble the observed data.
 - Assumptions on the data generating process (here: drawn *i.i.d.*) is always necessary if we want to say anything beyond our observed data.
- **2** h is fixed (before \mathcal{D} is generated).

If the assumptions are satisfied, then (1) says that a sample \mathcal{D} can be used to assess whether or not h is close to f.

• However, this is verification, not learning.



If we pick the hypothesis with minimum E_{in} , will E_{out} be small?

- If you toss a fair coin 10 times, what is the probability that you will get 10 heads?
 - ► ≈ 0.1%
- If you toss 1000 fair coins 10 times each, what is the probability that *some* coin will get 10 heads?
 - ► ≈ 62%

Let $g \in \mathcal{H} = \{h_1, \ldots, h_M\}.$

$$\Pr(|E_{in}(g) - E_{out}(g)| > \epsilon) \leq \Pr\{|E_{in}(h_1) - E_{out}(h_1)| > \epsilon$$

$$\text{or } |E_{in}(h_2) - E_{out}(h_2)| > \epsilon$$

$$\cdots$$

$$\text{or } |E_{in}(h_M) - E_{out}(h_M)| > \epsilon\}$$

$$\leq \sum_{m=1}^{M} \Pr(|E_{in}(h_m) - E_{out}(h_m)| > \epsilon)$$

$$\leq 2|\mathcal{H}|e^{-2\epsilon^2N}$$
(2)

- , where $|\mathcal{H}| = M$ is the size of \mathcal{H} .
 - (2) is valid for any $g \in \mathcal{H}$, regardless how g is selected.
 - Note g is *not* fixed before the data is generated: the selection of g depends on \mathcal{D} .

Let $\delta \equiv 2 |\mathcal{H}| e^{-2\epsilon^2 N}$. (2) \Rightarrow with probability at least $1 - \delta$,

$$E_{out}(g) \le E_{in}(g) + \sqrt{\frac{1}{2N} \ln \frac{2|\mathcal{H}|}{\delta}}$$
 (3)

• (3) is referred to as a **generalization bound**.



The feasibility of learning is split into two questions:

- **1** Can we make sure that $E_{out}(g)$ is close enough to $E_{in}(g)$?
- ② Can we make $E_{in}(g)$ small enough?
 - ullet $|\mathcal{H}|$ can be thought of as a measure of the **complexity** of \mathcal{H} .
 - Tradeoff:
 - Small $|\mathcal{H}| \Rightarrow E_{in}(g) \approx E_{out}(g)$
 - ▶ Large $|\mathcal{H}|$ \Rightarrow more likely to find g such that $E_{in}\left(g\right) \approx 0$

Effective Number of Hypotheses

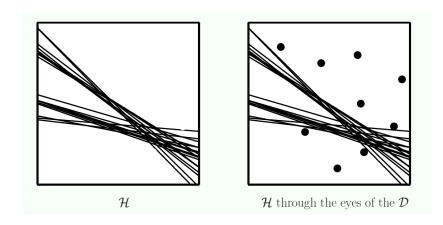
- In practice, hypothesis sets are typically infinite in size.
- ullet How to derive the generalization bound when ${\cal H}$ is infinite?

Idea:

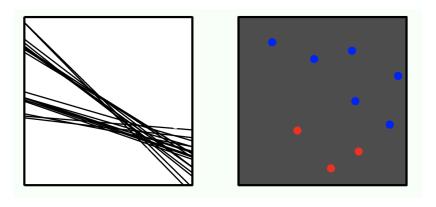
- On a given data set \mathcal{D} , many $h \in \mathcal{H}$ will look the same, i.e., they map \mathcal{D} into the same set of values.
- ullet These hypotheses are identical from the data's perspective. Therefore there are "effectively" fewer than $|\mathcal{H}|$ hypotheses⁴.

⁴Since the data is all we have.

Effective Number of Hypotheses



Effective Number of Hypotheses



From the point of view of \mathcal{D} , the entire \mathcal{H} is just one dichotomy.

Growth Function

- Consider binary target functions and hypothesis sets that contain $h: \mathcal{X} \to \{-1, +1\}.$
- If $h \in \mathcal{H}$ is applied to a finite sample $x_1, \ldots, x_N \in \mathcal{X}$, we get an N-tuple $h(x_1), \ldots, h(x_N)$ of ± 1 's.
- Such an N-tuple is called a dichotomy since it splits x_1, \ldots, x_N into two groups: those points for which h is -1 and those for which h is +1.
- Each $h \in \mathcal{H}$ generates a dichotomy on x_1, \dots, x_N , but two different h's may generate the same dichotomy if they happen to give the same pattern of ± 1 's on this particular sample.

VC Dimension

- The **growth function** for a hypothesis set \mathcal{H} , denoted $m_{\mathcal{H}}(N)$, is the maximum possible number of dichotomies \mathcal{H} can generate on a data set of N points⁵.
- If \mathcal{H} is capable of generating all possible dichotomies on x_1, \ldots, x_N , then \mathcal{H} shatters x_1, \ldots, x_N , in which case $m_{\mathcal{H}}(N) = 2^N$.
- The Vapnik-Chervonenkis (VC) dimension of \mathcal{H} , denoted $d_{vc}(\mathcal{H})$, is the size of the largest data set that \mathcal{H} can shatter⁶.
 - ▶ $d_{vc}(\mathcal{H})$ is the largest value of N for which $m_{\mathcal{H}}(N) = 2^{N}$.

⁶See (Appendix I) for a detailed introduction to growth function and VC dimension.



⁵rather than over the entire input space \mathcal{X} .

VC Inequality

The Vapnik-Chervonenkis Inequality

Let \mathcal{H} be a set of binary-valued hypotheses. For any $g \in \mathcal{H}$,

$$\Pr\left(\left|E_{in}\left(g\right) - E_{out}\left(g\right)\right| > \epsilon\right) \le 4m_{\mathcal{H}}\left(2N\right)e^{-\frac{1}{8}\epsilon^{2}N} \tag{4}$$

VC Bound

VC Generalization Bound

(4) \Rightarrow for any tolerance $\delta > 0$,

$$E_{out}(g) \leq E_{in}(g) + \sqrt{\frac{8}{N} \ln \frac{4m_{\mathcal{H}}(2N)}{\delta}}$$
 (5)

with probability $\geq 1 - \delta$.

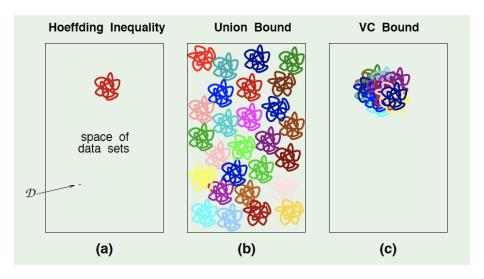
VC Bound

• We can prove that:

$$m_{\mathcal{H}}\left(N\right) \leq \begin{cases} N^{d_{vc}(\mathcal{H})} + 1\\ \left(\frac{eN}{d_{vc}(\mathcal{H})}\right)^{d_{vc}(\mathcal{H})} & N \geq d_{vc}\left(\mathcal{H}\right) \end{cases} \tag{6}$$

• The VC inequality and VC generalization bound establish the feasibility of learning with infinite hypothesis sets: with enough data, each and every hypothesis in an infinite \mathcal{H} with a finite VC dimension will generate well from E_{in} to E_{out} .

VC Bound



Training versus Testing

If we have an independent test set not used for selecting g from \mathcal{H} , and on which we can evaluate the performance of g, then

Training:
$$\Pr(|E_{in}(g) - E_{out}(g)| > \epsilon) \le 4m_{\mathcal{H}}(2N) e^{-\frac{1}{8}\epsilon^2 N}$$

Testing:
$$\Pr(|E_{test}(g) - E_{out}(g)| > \epsilon) \le 2e^{-2\epsilon^2 N}$$

- The generalization bound for test error is much tighter.
- The test set is not biased, whereas the training set has an *optimistic* bias, since it is used to choose a hypothesis that looks good *on it*.
- The price for a test set is fewer data for training.

Approximation-Generalization Tradeoff

(5) and (6)
$$\Rightarrow$$

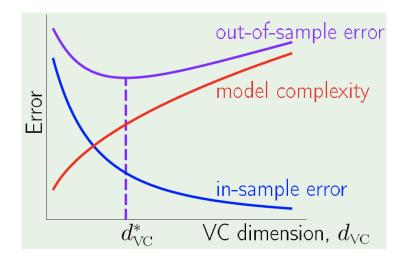
$$E_{out}\left(g
ight) \leq E_{in}\left(g
ight) + \sqrt{rac{8}{N}\lnrac{4m_{\mathcal{H}}\left(2N
ight)}{\delta}} \ \leq E_{in}\left(g
ight) + \underbrace{\sqrt{rac{8}{N}\lnrac{4\left(\left(2N
ight)^{d_{vc}}+1
ight)}{\delta}}}_{\Omega\left(d_{vc}
ight)}$$

 $\Omega(d_{vc})$ can be viewed as a penalty for model complexity.

- More complex \mathcal{H} $(d_{vc} \uparrow) \Rightarrow$ better chance of approximating f in sample $(E_{in} \approx 0)$
- Less complex \mathcal{H} $(d_{vc}\downarrow)$ \Rightarrow better chance of generalizing out of sample $(E_{in}\approx E_{out})$



Approximation-Generalization Tradeoff



Approximation-Generalization Tradeoff

- VC analysis shows the choice of \mathcal{H} needs to strike a balance between approximating f on the training data and generalizing to new data.
- If \mathcal{H} is too simple, we may fail to approximate f well on the training data and end up with a large in-sample error term.
- \bullet If ${\cal H}$ is too complex, we may fail to generalize well because of the large model complexity term.

Learning as Optimization

The choice of error measure that quantifies how well a hypothesis h approximates the target function f matters for the learning process and can affect the final hypothesis g that is chosen.

Formally,

$$E_{out}(h) = \mathbb{E}\left[\ell\left(h\left(x\right), f\left(x\right)\right)\right]$$
$$E_{in}(h) = \frac{1}{N} \sum_{i=1}^{N} \ell\left(h\left(x_{i}\right), f\left(x_{i}\right)\right)$$

, where $\ell(h(x), f(x))$ is a **loss function** that measures the difference between h(x) and $f(x)^7$.

⁷Hence E_{out} = expected loss. E_{in} =average loss on the training data.



Learning as Optimization

• Some commonly used loss functions are:⁸

$$\ell(x,y) = (x-y)^2$$
 squared-error loss $\ell(x,y) = |x-y|$ absolute-error loss $\ell(x,y) = \mathcal{I}(x \neq y)$ zero-one loss

- We have used zero-one loss in VC analysis when dealing with binary target functions. For real-valued functions, a common choice is to use the squared-error loss.
- The process of learning is also a process of optimization: we choose g
 by minimizing an objective function, which is the error measure based
 on our chosen loss function.

⁸The squared-error loss is also called **quadratic loss** or **L2 loss**. The absolute-error loss is also called **linear loss** or **L1 loss**.



Bias-Variance Decomposition

Bias-variance decomposition provides another way of looking at the approximation-generalization tradeoff.

Consider a real-valued target function f. Let $g \in \mathcal{H}$ be the hypothesis chosen to approximate f. Then

$$E_{out}(g) = \mathbb{E}\left[\left(g(x) - f(x)\right)^{2}\right]$$

$$= \mathbb{V}\left(g(x)\right) + \mathbb{E}\left[\left(g(x) - f(x)\right)\right]^{2}$$

$$= \mathbb{V}(g) + \left[\operatorname{bias}(g)\right]^{2}$$
(7)

, where bias $(g) \doteq \mathbb{E} [(g(x) - f(x))].$

Bias-Variance Decomposition

9,10

⁹Note: the expectation is with respect to both x and \mathcal{D} , since g depends on \mathcal{D} . I.e.,

bias
$$(g) = \mathbb{E}_{x} \left[\mathbb{E}_{\mathcal{D}} \left[\left(g(x) - f(x) \right) \right] \right] = \mathbb{E}_{x} \left[\mathbb{E}_{\mathcal{D}} \left[g(x) \right] - f(x) \right]$$

= $\mathbb{E}_{x} \left[\bar{g}(x) - f(x) \right]$

, where $\bar{g}\left(x\right) \doteq \mathbb{E}_{\mathcal{D}}\left[g\left(x\right)\right]$. Similarly,

$$\mathbb{V}\left(g\right) = \mathbb{E}_{\mathsf{x}}\left[\mathbb{E}_{\mathcal{D}}\left[\left(g\left(\mathsf{x}\right) - \bar{g}\left(\mathsf{x}\right)\right)^{2}\right]\right]$$

10 If

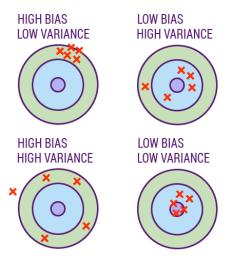
$$y = f(x) + e$$

, where $\mathbb{E}\left[e\right]=0$, then

$$\begin{aligned} E_{out}\left(g\right) &= \mathbb{E}\left[\left(y - g\left(x\right)\right)^{2}\right] \\ &= \mathbb{V}(g) + \left[\mathsf{bias}(g)\right]^{2} + \mathbb{V}\left(e\right) \end{aligned}$$

- Intuitively, bias arises if the model \mathcal{H} does not contain f^{11} . Thus there will be error even if we fit the model on the entire population.
- $\mathbb{V}(g)$ refers to the amount by which g would change if we estimate it using a different data set. The variance term arises because we have limited data. The g that we select based on a limited sample is almost never the same as the g that we would select if we have access to the entire population.
 - ▶ In general, the variance term decreases as sample size increases.

 $^{^{11}}$ which is almost always the case: our model hardly ever contains the true f.

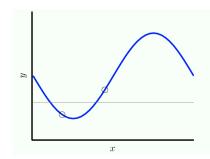


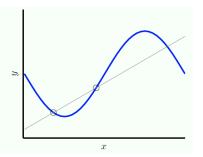
$$y = f(x) = \sin(\pi x)$$

• Two models:

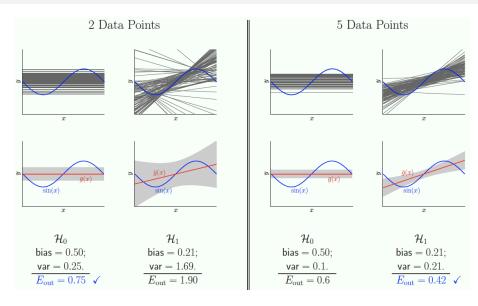
$$\mathcal{H}_0: h(x) = b$$

$$\mathcal{H}_1: h(x) = ax + b$$

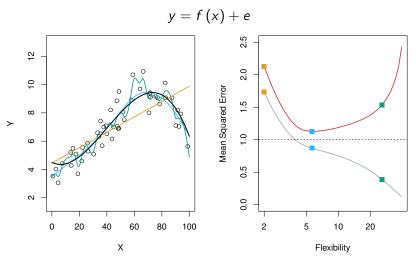




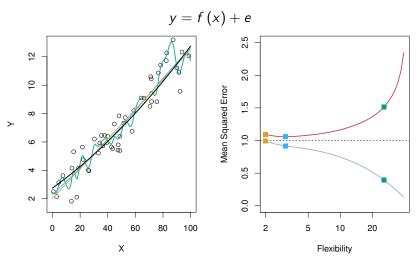
2 data points



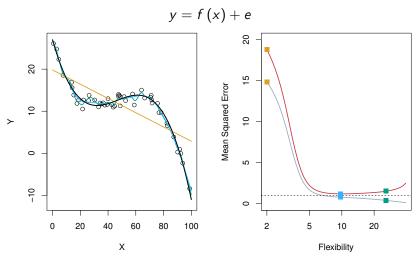
- In general, as model complexity increases, the bias will decrease and the variance will increase, leading to the bias-variance trade-off.
 - More complex models tend to have higher variance because they have the capacity to follow the data more closely. Thus using a different set of data points may cause g to change considerably.
 - The challenge lies in finding a model for which both the bias and the are low.



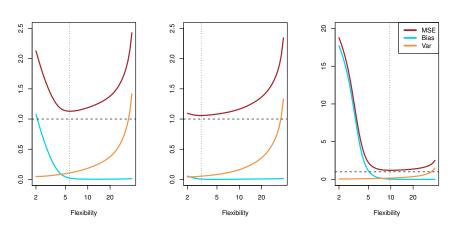
Left: f (black), linear fit (orange), smoothing spline fits (blue & green) Right: E_{in} (grey), E_{out} (red), Var(e) (dashed)



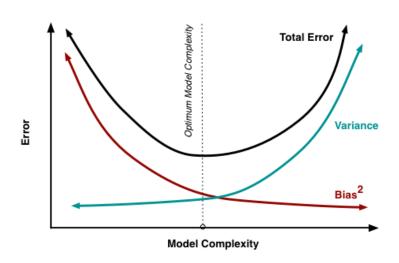
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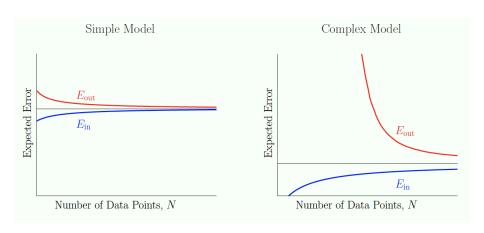
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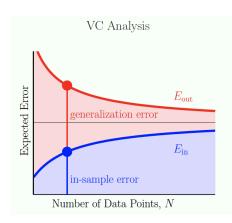
Bias-variance decomposition for the three examples

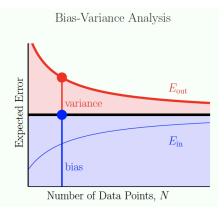


Learning Curve



Learning Curve





Noisy Targets

If y is not uniquely determined by x, i.e. if there does not exist a deterministic function f such that y = f(x), then the relation between x and y needs to be described by a joint-distribution p(x, y) = p(y|x) p(x).

Three approaches to learning and prediction when y is "noisy" 12:

- Learn p(y|x): in this case we have a **target distribution** rather than a target function¹³.
- ② Find a deterministic function f such that y = f(x) + e, where e is an error term, and let f be our target function.
- **1** Let p(x, y) be our target distribution, from which we can calculate $p(y|x) = \frac{p(x,y)}{p(x)}$.



¹²We say y is a noisy target when conditional on x, y is not completely determined.

¹³In Bayesian terms, p(y|x) is the **posterior distribution** of y.

- To learn p(y|x), let the hypothesis set \mathcal{H} be a set of conditional probability distributions: $\mathcal{H} = \{q_1(y|x), q_2(y|x), \ldots\}.$
 - H is said to be a probabilistic model¹⁴.
- Goal: select a $q(y|x) \in \mathcal{H}$ that approximates p(y|x) well.
- What is a suitable loss function for quantifying how well q(y|x) approximates p(y|x)?
 - ▶ <u>Need</u>: a measure of (dis)similarity between probability distributions.

 $^{^{14}}$ In general, hypothesis sets consisting of (conditional) probability distributions are called probabilistic models.



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KL Divergence

Let p and q be two distributions of x. The **Kullback-Leibler (KL)** divergence of q from p^{16} , is defined as

$$D_{KL}(p||q) = \sum_{x} \log \left(\frac{p(x)}{q(x)}\right) p(x)$$
 (8)

, or in the case of continuous random variables,

$$D_{KL}(p||q) = \int \log\left(\frac{p(x)}{q(x)}\right) p(x) dx$$
 (9)

¹⁶Also called **relative entropy**. See Appendix II for an introduction to information theory, entropy, and KL divergence.

KL Divergence as Loss Function

KL divergence can be interpreted as a measure of dissimilarity between two distributions: $D_{KL}(p||q) \ge 0$ if and only if $p = q^{17}$.

$$D_{KL}(p||q) = \mathbb{E}_{x \sim p(x)} [\log p(x) - \log q(x)]$$

$$= -\mathbb{E}_{x \sim p(x)} [\log q(x) - \log p(x)]$$

$$\geq -\log \mathbb{E}_{x \sim p(x)} \left[\frac{q(x)}{p(x)} \right]$$

$$= -\log \int \frac{q(x)}{p(x)} p(x) dx = 0$$

Therefore, KL divergence can be used as a loss function to quantify the difference between probability distributions.

¹⁷Note that KL divergence is not symmetric: $D_{KL}\left(p||q\right) \neq D_{KL}\left(q||p\right)$. Therefore it is not a proper distance measure.

KL Divergence as Loss Function

Now suppose we are given data $\mathcal{D} = \{x_1, \dots, x_N\} \sim^{i.i.d.} p(x)$ and want to learn p(x) based on \mathcal{D} .

For any hypothesis distribution q(x), using KL divergence as a loss function, we have:

$$E_{out}(q) = \mathbb{E}\left[\log p(x) - \log q(x)\right] \tag{10}$$

$$E_{in}(q) = \frac{1}{N} \sum_{i=1}^{N} (\log p(x_i) - \log q(x_i))$$
 (11)

KL Divergence as Loss Function

Since p is fixed, choosing a q to minimize (10) and (11) is equivalent to minimizing¹⁸:

$$E_{out}(q) = -\mathbb{E}\left[\log q(x)\right] \tag{12}$$

$$E_{in}(q) = -\frac{1}{N} \sum_{i=1}^{N} \log q(x_i)$$
 (13)

• cross-entropy loss: $\ell(q(x), p(x)) = -\log q(x)$

 $^{^{18}(12)}$ and (13) are the out-of-sample and in-sample expressions for **cross entropy**. Given a fixed true distribution p, minimizing the KL divergence of any distribution q from p is the same as minimizing their cross entropy. See Appendix II.



Maximum Likelihood

Given observed data \mathcal{D} and a probability distribution q, the **likelihood** function is defined as the probability of observing \mathcal{D} according to q:

$$\mathcal{L}(q) = \Pr_{q}(\mathcal{D}) = \prod_{i=1}^{N} q(x_i)$$
(14)

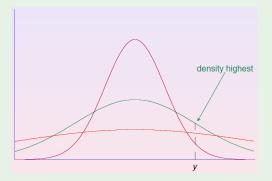
⇒ the log likelihood function

$$\log \mathcal{L}(q) = \sum_{i=1}^{N} \log q(x_i)$$
 (15)

Let the hypothesis set \mathcal{H} be a set of probability distributions. The **maximum likelihood estimation (MLE)** method chooses a distribution from \mathcal{H} that maximizes the (log) likelihood function.

Maximum Likelihood

Suppose we only observe a single data point, y, drawn from an underlying distribution. We want to learn the underlying distribution based on this one data point. Our hypothesis set consists of the following three distributions:

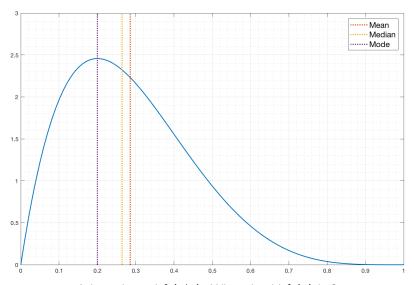


Then according to MLE, we would choose the green distribution.

Maximum Likelihood as Minimum KL Divergence

Minimizing the empirical KL divergence (cross entropy) is equivalent to maximizing the (log) likelihood function.

- Now suppose from a hypothesis set $\mathcal{H} = \{q_1(y|x), q_2(y|x), \ldots\}$, we have selected a q(y|x) to approximate p(y|x) by minimizing the KL divergence. Let's write q(y|x) as $\widehat{p}(y|x)$.
- Armed with $\hat{p}(y|x)$ our estimate of p(y|x) how should we make a prediction of y given a value of x?
- For continuous y, let $\hat{y}(x)$ denote our prediction of y given x. There are many choices: $\hat{y}(x)$ can be
 - mean of $\widehat{p}(y|x)$
 - median of $\widehat{p}(y|x)$
 - ▶ mode of $\widehat{p}(y|x)$
 - **.** . . .
- It depends on the loss function that we use.



A hypothetical $\widehat{p}(y|x)$. What should $\widehat{y}(x)$ be?

Given $\hat{p}(y|x)$, let $\hat{y}(x)$ be the solution to

$$\widehat{y}(x) = \arg\min_{c} \mathbb{E}_{\widehat{p}(y|x)} [\ell(y,c)|x]$$

Then

$$\ell(y,c) = (y-c)^2 \Rightarrow \widehat{y} = \mathbb{E}[y|x]$$

$$\ell(y,c) = |y-c| \Rightarrow \widehat{y} = \text{Median}(y|x)$$

$$\ell(y,c) = \mathcal{I}(y \neq c) \Rightarrow \widehat{y} = \text{Mode}(y|x)$$

, where the mean, median, and mode are with respect to $\hat{p}(y|x)$.

For discrete or categorical y, a common choice is to use the 0-1 loss¹⁹:

$$\ell(y, \hat{y}) = \mathcal{I}(y \neq \hat{y}) \text{ for } y \in \{1, \dots, K\}$$

 $^{^{19}}$ In the classification setting, the 0-1 loss is also called the **misclassification loss**.



Given $\widehat{p}(y|x)$, using the 0-1 loss for prediction, we have:

$$\widehat{y}(x) = \underset{c \in \{1, \dots, K\}}{\min} \mathbb{E}_{\widehat{p}(y|x)} \left[\mathcal{I}(y \neq c) | x \right]$$

$$= \underset{c \in \{1, \dots, K\}}{\arg\min} \widehat{p}(y \neq c|x)$$

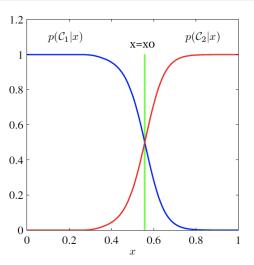
$$= \underset{c \in \{1, \dots, K\}}{\arg\max} \widehat{p}(y = c|x)$$

$$(16)$$

, i.e., we predict y to be the value (class, category) that has the highest posterior probability²⁰. This is called the **Bayes classifier**.



²⁰according to the estimated $\widehat{p}(y|x)$.



 $y \in \{C_1, C_2\}$. The Bayes classifier classifies y to be C_1 for $x < x_0$ and C_2 for $x > x_0$. The green line $x = x_0$ is called a **decision boundary**.

The loss function we use here is separate and can be different from the loss function that we use for learning p(y|x). This is because for predicting noisy targets, we essentially have two stages:

- Learning p(y|x)
- ② Making prediction of y based on the estimated p(y|x)

These two stages are called learning and prediction²¹.

²¹Also called inference and decision.

Decision Theory

How to make a prediction of y based on its probability distribution is a subject of decision theory, which is concerned with how to make optimal decisions given the appropriate probabilities.

Fingerprint Verification

Consider the problem of fingerprint verification. Let $y \in \{-1,1\}$ denote whether the fingerprint belongs to the person of interest or not. Let \widehat{y} be our prediction. There are two types of error we can make here:

		y	
		+1	-1
\widehat{y}	+1	no error	false positive
	-1	false negative	no error

Decision Theory

Fingerprint Verification

Loss functions can be used to control which type of error we want to minimize: the overall error rate, the false positive rate (FPR), or the false negative rate (FNR).

The choice of $\ell(y, \hat{y})$ depends on our needs.

Decision Theory

Fingerprint Verification

If $\ell(y, \hat{y}) = \mathcal{I}(y \neq \hat{y})$, then the decision rule is the Bayes classifier^a

predict
$$\hat{y} = \begin{cases} 1 & \text{if } p(y=1) \ge p(y=0) \\ 0 & \text{if } p(y=1) < p(y=0) \end{cases}$$

, which minimizes the overall error rate.

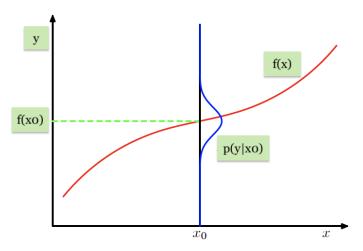
^aAssuming we know p(y).

- The second approach to learning and prediction when y is "noisy" is to find a deterministic function f such that y = f(x) + e, and let f be our target function (see Page 49).
- Then let \hat{f} be our estimated f. Our prediction of y for any value of x will just be $\hat{y} = \hat{f}(x)$.
- What should f be? Ideally, f should be the function that best predicts y in the *underlying population*. Then we try to learn this f using our observed sample \mathcal{D} . Finally, we use \widehat{f} to make predictions of y given x.
- This approach combines the two stages learning and prediction –
 into one problem: directly learning a function f that maps each x into
 a prediction of y.

- What is the function that produces the best prediction of *y* given *x* in the underlying population?
- The answer, again, depends on the loss function, i.e. on what we mean by "best."
- A common choice for continuous y is to use the squared-error loss, which $\Rightarrow f(x) = \mathbb{E}[y|x]^{22}$.
- The conditional expectation function $\mathbb{E}[y|x]$ is known as the **regression function**.

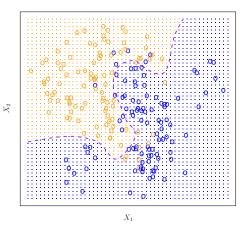
Thus in this approach, instead of learning p(y|x), we only learn a *moment* of p(y|x), which is $\mathbb{E}[y|x]$.





The regression function f(x), which minimizes the expected squared error loss, is given by the mean of the conditional distribution p(y|x).

When y is discrete or categorical, this approach tries to learn the decision boundaries directly.



Rather than estimating p(y|x) and using it to derive a decision rule (e.g., the Bayes classifier), this approach focuses on learning directly the f (here the purple boundary) that best separates y = orange and y = blue.

 $x = (x_1, x_2), y \in \{\text{orange, blue}\}$

Learning p(x,y)

- The third approach is to learn the entire joint distribution p(x, y) (see page 49).
- Let $\hat{p}(x,y)$ be an estimate of p(x,y). Once we have $\hat{p}(x,y)$, we can use it to calculate $\hat{p}(y|x)$, which in turn, would allow us to make a prediction of y given x.

Generative vs. Discriminative Models

- Models of the joint distribution p(x, y) are called **generative** models²³, while models of p(y|x) or f(x) are called **discriminative** models²⁴.
- While discriminative models are mainly used for prediction tasks, generative models allows us to do more than just making predictions of y given x. We can, for example, generate new data points $\{(x_i, y_i)\}$ by drawing from $\hat{p}(x, y)$. These new data points are called **synthetic data**, since they are not real, observed data. The process of generating synthetic data is called **simulation**.

²³Approach 3 on page 49

²⁴Approach 1 and 2 on page 49

- **Scientific models**²⁵ are an important type of generative models that describe the **causal mechanisms** that generate p(x, y).
- While scientific models can be used for prediction, the goal of learning causal mechanisms is distinct from the goal of prediction.

²⁵Also called **causal models**.

Scientific vs. Statistical Model

If you want to predict where Mars will be in the night sky^a, you may do very well with a model in which Mars revolves around the Earth. You can estimate, from data, how fast Mars goes around the Earth and where it should be tonight. But the estimated model does not describe the actual causal mechanisms. Nor does it need to: if our only goal is prediction, then we often do not need a scientific model.

^aThis example is taken from Shalizi (2019).

- Because scientific models describe causal mechanisms, what we learn from one set of data $\mathcal{D} \sim p(x,y)$ can be potentially used to explain and predict data drawn from another distribution, say p(u,v), if $\{x,y\}$ and $\{u,v\}$ share similar underlying causal mechanisms.
 - In other words, what we learn from one observed phenomenon can be used to explain and predict other related phenomena.
 - For example, we can learn individuals' risk aversion from their investment behavior, which in turn, can help explain and predict their career choices.

- Good scientific models²⁶ can potentially deliver better predictive performance than statistical models trained on single data sets, because they can be learned from a combination of data from various sources that share the same underlying causal mechanisms.
 - ▶ Apples falling down trees and the earth orbiting around the sun both inform us of the gravitational constant.

²⁶Think of quantum mechanics!

- Consider binary target functions and hypothesis sets that contain $h: \mathcal{X} \to \{-1, +1\}^{27}$.
- Let $\mathcal{H}(x_1,\ldots,x_N) = \{(h(x_1),\ldots,h(x_N))|h\in\mathcal{H}\}$ denote the dichotomies generated by \mathcal{H} on $x_1,\ldots,x_N\in\mathcal{X}$.

²⁷The following analysis is all based on binary target functions.

Definition

The **growth function** for a hypothesis set ${\mathcal H}$ is defined by

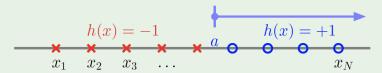
$$m_{\mathcal{H}}(N) = \max_{x_1,...,x_N \in \mathcal{X}} |\mathcal{H}(x_1,...,x_N)|$$

, i.e., $m_{\mathcal{H}}(N)$ is the maximum possible number of dichotomies \mathcal{H} can generate on a data set of N points^a.

Note: $m_{\mathcal{H}}(N) \leq 2^N$. If \mathcal{H} is capable of generating all possible dichotomies on x_1, \ldots, x_N , then \mathcal{H} shatters x_1, \ldots, x_N , in which case $m_{\mathcal{H}}(N) = 2^N$.

^arather than over the entire input space \mathcal{X} .

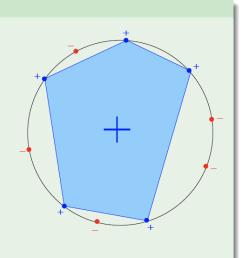
Positive Rays



- $\mathcal{H} = \{h(x) = sign(x a)\}$
- ullet There are N+1 dichotomies depending on where you put a.
- $m_{\mathcal{H}}(N) = N + 1$

Convex Sets

- \mathcal{H} consists of all $h: \mathbb{R}^2 \to \{-1, +1\}$ that are are positive inside some convex set and negative elsewhere.
- If N points lie on a circle, then any dichotomy on these points can be generated by an h that is positive inside the polygon that connects the +1 points. Hence the N points are shattered by H.
- $m_{\mathcal{H}}(N) = 2^N$



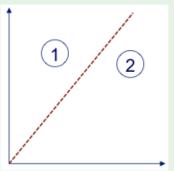
Definition

The **Vapnik-Chervonenkis (VC) dimension** of \mathcal{H} , denoted $d_{vc}(\mathcal{H})$, is the size of the largest data set that \mathcal{H} can shatter.

- $d_{vc}(\mathcal{H})$ is the largest value of N for which $m_{\mathcal{H}}(N) = 2^{N}$.
- If arbitrarily large finite sets can be shattered by \mathcal{H} , then $d_{vc}(\mathcal{H}) = \infty$.
- \exists some shattered set of size $d \Rightarrow d_{vc}(\mathcal{H}) \geq d$.
- No set of size d+1 is shattered $\Rightarrow d_{vc}(\mathcal{H}) \leq d$.

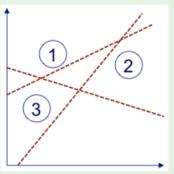
Hyperplanes in \mathbb{R}^2

- ullet ${\cal H}$ is set of lines (linear separators) in ${\mathbb R}^2$
- can find can an *h* consistent with 2 data points no matter how they are labeled:



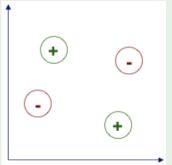
Hyperplanes in \mathbb{R}^2

• can find can an *h* consistent with 3 non-collinear data points no matter labeling:



Hyperplanes in \mathbb{R}^2

• cannot find can an h consistent with 4 data points for some labeling:



• Hence $d_{vc}(\mathcal{H}) = 3^a$.

^aIn general, d_{vc} (hyperplanes in \mathbb{R}^d) = d+1

Appendix I: Sauer's Lemma

Sauer's lemma

If $d_{vc}(\mathcal{H}) < \infty$, then

$$m_{\mathcal{H}}(N) \leq \sum_{i=0}^{d_{vc}(\mathcal{H})} \binom{N}{i}$$
 (17)

• If the VC dimension is finite, then $m_{\mathcal{H}}(N)$ can be bounded by a polynomial in N and the order of the polynomial is $d_{vc}(\mathcal{H})$.

Appendix I: Sauer's Lemma

We can prove that 28

$$\sum_{i=0}^{d} \left(\begin{array}{c} N \\ i \end{array}\right) \leq \begin{cases} N^d + 1 \\ \left(\frac{eN}{d}\right)^d \end{cases}$$

Therefore, $m_{\mathcal{H}}(N)$ can be further bounded by:

$$m_{\mathcal{H}}\left(\mathcal{N}\right) \leq \mathcal{N}^{d_{vc}(\mathcal{H})} + 1$$
 (18)

, or

$$m_{\mathcal{H}}(N) \le \left(\frac{eN}{d_{VC}(\mathcal{H})}\right)^{d_{VC}(\mathcal{H})}$$
 (19)

²⁸The second inequality requires $N \ge d$.

Appendix II: Information Theory

Consider a random variable x. How much information is received when we observe a specific value of x?

- Depends on 'degree of surprise': a highly improbable value conveys more information than a very likely one.
- If we know an event is certain to happen, we would receive no information when we observe it happens.

Appendix II: Information Theory

Let h(.) denote the information content of an event. h(.) should satisfy

- **1** h(a) should be inversely correlated with p(a).
- ② For two unrelated events a and b, such that p(ab) = p(a)p(b), we should have h(ab) = h(a) + h(b).

 \Rightarrow we can let:

$$h(a) = \log \frac{1}{p(a)}$$

Appendix II: Entropy

For a discrete random variable x with probability distribution p(x), the average amount of information transmitted by x is:

$$\mathbb{H}(p) = E_p[h(x)] = \sum_{x} p(x) \log \frac{1}{p(x)}$$

 $\mathbb{H}\left(p\right)$ is called the **entropy**²⁹ of probability distribution p .

 Distributions that are sharply peaked around a few values will have a relatively low entropy, while those that are spread more evenly across many values will have higher entropy.



²⁹More precisely, **information entropy**, or **Shannon entropy**.

Appendix II: Entropy

- Historically, information entropy is developed to describe the average amount of information needed to specify the state of a random variable.
- Specifically, if we use base—2 logarithm in the definition of $\mathbb{H}(p)$, then $\mathbb{H}(p)$ is a lower bound on the average number of bits needed to encode a random variable with probability distribution p.
- Achieving this bound would require using an optimal coding scheme designed for p, which assigns shorter codes to higher probability events and longer codes to less probable events.

Appendix II: Entropy

- Suppose a random variable has 8 states, each being equally likely. Then we can code these 8 states as 000, 001, 010, 011, 100, 101, 110, 111. In this case, the average length of the code needed to encode the variable is 3, which is equivalent to its entropy $\mathbb{H}=8\times\frac{1}{8}\log_28=3$.
- If the probabilities of the 8 states are given by $\left(\frac{1}{2},\frac{1}{4},\frac{1}{8},\frac{1}{16},\frac{1}{64},\frac{1}{64},\frac{1}{64},\frac{1}{64}\right), \text{ then the optimal coding scheme is 0, 10, }$ 110, 1110, 111100, 1111101, 1111110, 1111111. Under this coding scheme, the average length of the code needed to encode the variable is $\frac{1}{2}\times 1+\frac{1}{4}\times 2+\frac{1}{8}\times 3+\cdots=2, \text{ which is equivalent to its entropy }$ $\mathbb{H}=\frac{1}{2}\log_22+\frac{1}{4}\log_24+\frac{1}{8}\log_28+\cdots=2.$

Appendix II: Cross Entropy

If p is the distribution of x, but we use distribution q to describe x instead, then the average amount of information needed to specify x as a result of using q instead of p is

$$\mathbb{H}\left(p,q\right) = \mathbb{E}_{p}\left[\log\frac{1}{q\left(x\right)}\right] = \sum_{x} p\left(x\right)\log\frac{1}{q\left(x\right)}$$

 $\mathbb{H}(p,q)$ is called the **cross entropy** of p and q. In information theory³⁰, it can be interpreted as the average number of bits needed to encode a random variable using a coding scheme designed for probability distribution q rather than the true distribution p.

³⁰Using base-2 logarithm.

Appendix II: KL Divergence

The **relative entropy** of q with respect to p, or the **Kullback-Leibler** (KL) divergence of q from p, is defined as

$$D_{KL}(p||q) = \mathbb{H}(p,q) - \mathbb{H}(p)$$
$$= \sum_{x} \log \left(\frac{p(x)}{q(x)}\right) p(x)$$

, or in the case of continuous random variables,

$$D_{KL}(p||q) = \int \log\left(\frac{p(x)}{q(x)}\right) p(x) dx$$

KL divergence represents the average additional information required to specify x as a result of using q instead of the true distribution p.

As an example to illustrate the concepts of entropy, cross entropy, and relative entropy (KL divergence), let's play the following games³¹:

Game 1

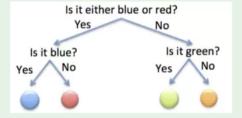
I will draw a coin from a bag of 4 coins: a blue, a red, a green, and an orange coin. Your goal is to guess which color it is with the fewest questions.



³¹Source of this example.

Game 1

One of the best strategies is this:

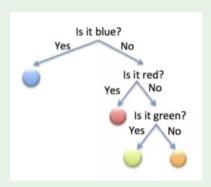


Using this strategy, the expected number of questions needed to guess the coin is 2. This is the entropy 30 of the probability distribution

$$p_1 = \left(\frac{1}{4}, \frac{1}{4}, \frac{1}{4}, \frac{1}{4}\right).$$

Game 2

Now suppose the coins in the bag have the following distribution: 1/2 of them are blue, 1/4 are red, 1/8 are green, and 1/8 are orange. The optimal strategy now looks like this:



Under this strategy, the expected number of questions to guess a coin is $\frac{1}{2} \times 1 + \frac{1}{4} \times 2 + 2 \times \frac{1}{8} \times 3 = 1.75$. This is the entropy³⁰ of the probability distribution $p_2 = \left(\frac{1}{2}, \frac{1}{4}, \frac{1}{8}, \frac{1}{8}\right)$.

Using Game 1 Strategy on Game 2

What if we still use the strategy for game 1 to play game 2?

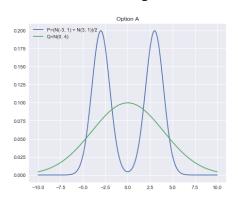
Then the expected number of questions needed to guess the coin is $\frac{1}{2} \times 2 + \frac{1}{4} \times 2 + 2 \times \frac{1}{8} \times 2 = 2$. This is the cross entropy for using game 1 strategy (optimized for p_1) on game 2 (with probability distribution p_2).

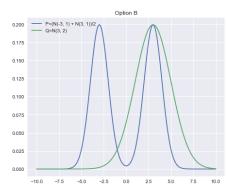
Obviously, using Game 1 strategy on Game 2 is not optimal. The additional expected number of questions we need to ask as a result of not using the optimal strategy is 2-1.75=0.25. This is the KL divergence of p_1 from p_2 .

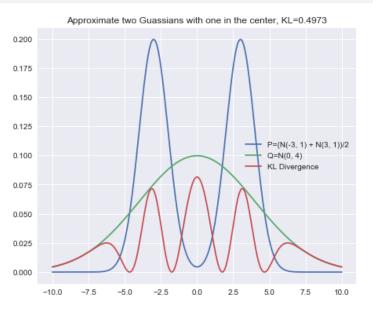
KL divergence can be interpreted as a measure of dissimilarity between two distributions. It satisfies $D_{KL}(p||q) \ge 0$ if and only if p = q.

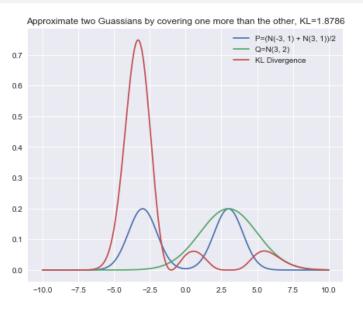
Therefore, KL divergence can be used as a loss function to quantify the difference between probability distributions.

Which of the following Q distributions better approximates P?









Acknowledgement

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