

Simulating Random Variables

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Monte Carlo (MC) Methods play a central role in statistical analysis, including power and type-I error rate analysis and in both Likelihood and Bayesian inference. The basic idea underlying MC methods is very simple: **we use samples** from the distribution of random variables (or functions thereof) to evaluate features of that distribution (e.g., expected value, variance, quantiles) are very difficult or cannot be derived analytically.

The first step in a MC study is then to sample random variables. In this note I review a few basic algorithms for simulating random variables and provide examples using R.

1. Random numbers & seeds

Computers do not generate random numbers. Everything in an algorithm is deterministic. However, there are well established algorithms that will generate sequences of numbers that resemble IID sequences from uniform random variables. A random number generator implements an algorithm that, starting from a seed (an integer) generates a deterministic sequence of numbers that resemble IID samples from uniform distributions. Given the seed the sequence is deterministic but we regard it as random. Try in R the following code

```
set.seed(1923)
runif(5)
runif(4)
set.seed(1923)
runif(5)
runif(4)
```

2. Sampling Bernoulli Random Variables

A Bernoulli random variable (X) with success probability θ can be sampled as follows: sample $U \sim \text{uniform}(n=1, \min=0, \max=1)$, set X to be 0 if $U < \theta$, 1 otherwise, that is $X = \text{ifelse}(U < \theta, 0, 1)$. Why does it work? $p(X = 0) = p(U < \theta) = \int_0^\theta 1 d\theta = \theta$.

3. Transformations of random variables

Suppose that $Y = g(X)$ where X is a random variable with pdf $f_X(x)$ and $g(X)$ is a one-to-one monotonic map, then

$$f_Y(y) = f_X(X(y)) \left| \frac{dX(y)}{dy} \right| \quad [1]$$

where $X(y) = g^{-1}(y)$.

Here $Y = g(X)$ is a transformation and $X(Y) = g^{-1}(Y)$ is the inverse transformation.

Example: sampling an exponential random variable.

The pdf of an exponential random variable is given by

$$f_Y(y) = \lambda e^{-\lambda y} \quad \lambda, y > 0$$

Result: if $X \sim U(0,1)$ and $Y = -\frac{\log(X)}{\lambda}$ then $Y \sim \text{Exponential}(\lambda)$.

Proof:

(1) here $Y = g(X) = -\frac{\log(X)}{\lambda}$; therefore $X = g^{-1}(Y) = e^{-\lambda Y}$.

(2) If $X \sim U(0,1) \Rightarrow f_X(x) = 1(0 \leq x \leq 1)$, then using [1]

(3) $f_Y(y) = 1(g^{-1}(Y) \in [0,1]) \left| \frac{dX(y)}{dy} \right| = 1(y > 0) e^{-\lambda y} |-\lambda| = \lambda e^{-\lambda y} \quad y > 0$

Generalization: The above results generalizes to random vectors, in which case $\left| \frac{dX(y)}{dy} \right|$ needs to be replaced with the absolute value of the Jacobian of the transformation (i.e., the absolute value of the determinant of the matrix of first derivatives of the inverse-transformation) evaluated at y . Note: this result holds for one-to-one maps. It would not hold, for instance for $g(X) = X^2$.

Example:

```
lambda=5;n=1e5
x=runif(n)
y=-log(x)/lambda
hist(y,n)
mean(y); 1/lambda
var(y); 1/(lambda^2)
y2=rexp(n=n,rate=lambda)
par(mfrow=c(1,2))
hist(y,50); hist(y2,50)
```

[See map of distributions from Casella and Berger in our website]