

GEOMETRIC DISTRIBUTION :

Consider the geometric series,

$$\sum_{x=0}^{\infty} q^x = \frac{1}{1-q}, \quad \text{if } |q| < 1$$
$$\Rightarrow \sum_{x=0}^{\infty} p q^x = 1, \quad \text{where } p+q=1.$$

For $0 < q < 1$, $p+q=1$,

the terms of the geometric series forms a pmf. Hence, the distribution given by the pmf is known as Geometric Distribution.

Definition: A non-negative integer valued random variable X is said to follow Geometric Distribution if its pmf is given by,

$$f(x) = \begin{cases} p q^x, & x = 0, 1, 2, \dots \\ 0, & \text{otherwise,} \end{cases} \quad \begin{matrix} 0 < p < 1 \\ p+q=1 \end{matrix}$$

Remark:

Consider a sequence of independent Bernoulli trial with probability of success 'p'. Let X be the number of failures before the 1st success,

$$P(X=n) = P(\underbrace{F F \dots F}_{(n-1)} \downarrow S) = q^n p \quad [\text{Due to independence}]$$

A random variable X that has a Geometric Distribution is often referred to a **discrete waiting time random variable** as it represents how long (in terms of failure) one has to wait to get a success.

Mean and Variance :

$$\sum_{x=0}^{\infty} q^x = \frac{1}{1-q}, \quad 0 < q < 1$$

Differentiating with respect to q , $\sum_{x=1}^{\infty} x q^{x-1} = (-1)(1-q)^{-2}(-1) = p^{-2}$

Again, differentiating w.r.t q , $\sum_{x=2}^{\infty} x(x-1) q^{x-2} = (-2)(1-q)^{-3}(-1) = 2p^{-3}$

$$E(X) = \sum_{x=0}^{\infty} x q^x p$$
$$= p q \sum_{x=1}^{\infty} x q^{x-1} = p q p^{-2} = q/p.$$

$$\begin{aligned}
V(X) &= E(X^2) - E^2(X) \\
&= E(X(X-1)) + E(X) - E^2(X) \\
&= \frac{2q^2}{p^2} + \frac{q}{p} - \frac{q^2}{p^2} \\
&= \frac{q^2}{p^2} + \frac{q}{p} = \frac{q}{p} \left(\frac{q}{p} + 1 \right) \\
&= \frac{q}{p} \frac{q+p}{p} = \frac{q}{p^2}
\end{aligned}$$

$$\begin{aligned}
E(X(X-1)) &= \sum x(x-1) q^x p \\
&= pq^2 \sum x(x-1) q^{x-2} \\
&= pq^2 \cdot 2p^{-3} \\
&= 2p^{-2}q^2
\end{aligned}$$

Remark :

Let, Y denote the no of trail required to get the 1st success, In a sequence of independent Bernoulli trials,

$$Y = X + 1.$$

$$P(Y = y) = P(X + 1 = y) = P(X = y - 1) = pq^{y-1}, \quad y = 1, 2, \dots, \infty.$$

$$E(Y) = E(X) + 1 = \frac{q}{p} + 1 = \frac{1}{p}$$

$$V(Y) = V(X) = \frac{q}{p^2}.$$

In a random sampling WR from a population with N distinct members. Let, X denotes no of drawings needed to get n distinct elements in the sample. Find $E(X)$ and $\text{var}(X)$.

Let $Z_i, \{i=1(n)n\}$ denotes the no of drawing needed to get new distinct elements when $(i-1)$ distinct elements have already obtained.

It is to be noted that $Z_1 = 1$

$$X = \sum_{i=1}^n Z_i = 1 + Z_1 + Z_2 + \dots + Z_i + \dots + Z_n$$

Now, Z_i 's are independently distributed with pmf.

$P(Z_i = z) = P[\text{exactly } z \text{ drawings are necessary to get new distinct element when } (i-1) \text{ distinct has already been obtained}]$.

$$\begin{aligned}
&= \underbrace{\left(\frac{i-1}{N}\right) \left(\frac{i-1}{N}\right) \dots \left(\frac{i-1}{N}\right)}_{(z-1)} \frac{N-(i-1)}{N} \\
&= \left(\frac{i-1}{N}\right)^{z-1} \left(1 - \frac{i-1}{N}\right) = q_i^{z-1} p_i
\end{aligned}$$

$$E(z_i) = \frac{q_i^0}{pq} = \frac{N}{(N-i+1)}$$

$$V(z_i) = \frac{q_i^0}{pq^2} = \frac{N(N-1)}{(N-i+1)^2}$$

$$E(X) = \sum_{i=1}^n E(z_i) = \sum_{i=1}^n \frac{N}{(N-i+1)}$$

$$V(X) = \sum_{i=1}^n V(z_i) = \sum_{i=1}^n \frac{N(N-1)}{(N-i+1)^2}$$

CDF: $F_X(x) = P(X \leq x) = \sum_{t=0}^{[x]} pq^t$

$$= p(1 + q + q^2 + \dots + q^{[x]})$$

$$= p \frac{1 - q^{[x]+1}}{1 - q}$$

$$= 1 - q^{[x]+1}$$

Tail function - $P(X > x)$

$$= 1 - P(X \leq x)$$

$$= q^{[x]+1}$$

NOTE : ① If $x_i \stackrel{iid}{\sim} \text{Geo}(p)$, $\sum_{i=1}^n x_i \sim \text{NB}(n, p)$

② If $x_i \stackrel{iid}{\sim} \text{Geo}(p)$, $i=1,2$, Find the conditional distribution of $(x_1 | x_1 + x_2)$

$$\frac{P(x_1 = x, x_2 = y - x)}{P(x_1 + x_2 = y)} = \frac{P(x_1 = x) P(x_2 = y - x)}{P(x_1 + x_2 = y) \sim \text{NB}(2, p)}$$

$$= \frac{\cancel{p} q^x \cancel{p} q^{y-x}}{\binom{y+2-1}{y} \cancel{p^2} q^y}$$

$$= \frac{1}{\binom{y+1}{y}} = \frac{1}{y+1} \sim \text{DU}(0, 1, 2, \dots, y)$$

③ $x_1, x_2, \dots, x_n \stackrel{iid}{\sim} \text{Geo}(p)$, find the distribution of $Y = \min(x_i)$ $i=1(n)$.

$$F_Y(y) = P(Y \leq y) = 1 - P(Y > y)$$

$$= 1 - P(\min(x_1, x_2) > y)$$

$$\begin{aligned}
&= 1 - P(X_1 > y, X_2 > y, \dots, X_n > y) \\
&= 1 - (P(X > y))^n \quad [\because \text{iid}] \\
&= 1 - (q^{[y]+1})^n
\end{aligned}$$

$$\begin{aligned}
P(Y=y) &= P(Y \leq y) - P(Y \leq y-1) \\
&= 1 - (q^{[y]+1})^n - 1 + (q^{[y]})^n \\
&= (q^y)^n - (q^{y+1})^n \\
&= (q^n)^y (1 - q^n) \\
&= Q^y P. \quad \begin{aligned} Q &= q^n \\ P &= 1 - q^n. \end{aligned}
\end{aligned}$$

$$Y = \min \{X_i\} \sim \text{Geo}(P = 1 - q^n).$$

Lack of Memory Property :-

If X has a Geometric Distribution, then $P(X > m+n | X > m) = P(X > n)$. \forall natural no.s m, n .

$$\begin{aligned}
P(X > m+n | X > m) &= \frac{P(X > m+n, X > m)}{P(X > m)} \\
&= \frac{P(X > m+n)}{P(X > m)} \\
&= \frac{q^{m+n+1}}{q^{m+1}} = q^n = q^{n-1+1} = P(X > n-1) \\
&= P(X > n)
\end{aligned}$$

Interpretation:

This theorem states that the probability that more than $(m+n+1)$ trials required for the 1st success given that there have been more than 'm' failures is equal to the unconditional probability of at least n trials are needed before the first success ie, $P(X > m+n | X > m)$ is independent of m ie, the information of no of success in $(m+1)$ trials has been forgotten in the subsequent calculation.

Theorem : Let X be a non-negative integer valued random variable satisfying $P(X > m+n | X > m) = P(X > n) \forall m, n$. Then, X must have a Geometric Distribution.

Let, $p_k = P(X=k)$, $k=0, 1, 2, \dots$ be the pmf of X ,
 Define, $q_m = P(X > m) = \sum_{k=m+1}^{\infty} p_k$

$$P(X > m+n | X > m) = P(X > n)$$

$$\Rightarrow P(X > m+n) = P(X > n-1) P(X > m)$$

$$\Rightarrow q_{m+n} = q_{n-1} q_m.$$

Set $n=1$, $q_{m+1} = q_m q_0$

$$\Rightarrow \frac{q_{m+1}}{q_m} = q_0 \quad \text{--- (i)}$$

Now, $q_m = \underbrace{\frac{q_m}{q_{m-1}} \frac{q_{m-1}}{q_{m-2}} \dots \frac{q_1}{q_0}}_m q_0$

$$= \underbrace{q_0 q_0 \dots q_0}_m q_0 \quad (\text{by (i)})$$

$$= q_0^{m+1}$$

Again, $q_{m-1} = P_m + P_{m+1} + P_{m+2} + \dots$

$$q_m = P_{m+1} + P_{m+2} + \dots$$

$$q_{m-1} - q_m = P_m$$

$$\therefore P_m = q_0^m - q_0^{m+1} = q_0^m (1 - q_0)$$

Now, $q_0 = P(X > 0) = 1 - P(X \leq 0) = 1 - p_0$

So, $P_m = \begin{cases} p_0 (1 - p_0)^m & m = 0, 1, 2, \dots \\ 0 & \text{otherwise,} \end{cases}$

$\therefore X \sim \text{Geo}(p_0).$

Remark:

Among all non-negative integer valued random variables X has the lack of memory property if X has a Geometric Distribution. Hence, the lack of memory property is a characteristic of geometric random variable in the class of all non-negative integer valued random variables.

Show that $E\left(\frac{1}{1+X}\right) = -\frac{p}{q} \log p$

$$E\left(\frac{1}{1+X}\right) = \sum_{x=0}^{\infty} \frac{1}{1+x} p q^x$$

$$= \frac{p}{q} \sum_{x=0}^{\infty} \frac{1}{1+x} q^{x+1}$$

$$\begin{aligned}
 &= \frac{p}{q} \left\{ \frac{q^1}{1} + \frac{q^2}{2} + \dots \right\} \\
 &= -\frac{p}{q} \left[-q - \frac{q^2}{2} - \frac{q^3}{3} - \dots \right] \\
 &= -\frac{p}{q} \log(1-q) \\
 &= -\frac{p}{q} \log p.
 \end{aligned}$$