

## Uniform / Rectangular Distribution :

A continuous random variable  $X$  is said to have a uniform (or rectangular) distribution over  $(a, b)$  if its pdf is given by,

$$f(x) = \text{constant, say } k, \quad x \in (a, b)$$

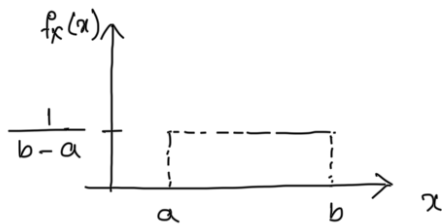
$$1 = \int_a^b f(x) dx \Rightarrow k \int_a^b dx = 1$$

$$\Rightarrow k = \frac{1}{b-a}$$

$$f(x) = \frac{1}{b-a} = \frac{1}{\text{length of interval } I}$$

**Definition -** A continuous random variable  $X$  is said to follow a uniform distribution over  $(a, b)$  if its PDF is given by

$$f_X(x) = \begin{cases} \frac{1}{b-a}, & a < x < b \\ 0, & \text{otherwise} \end{cases}$$



The graph of the pdf  $f(x)$  looks like a rectangle, that's why the distribution is known as Rectangular Distribution.

**Notation -**  $X \sim U(a, b)$  or  $X \sim \text{Rect}(a, b)$

Density  $f_X(x)$  can be greater than 1.

**CDF :**  $F(x) = \int_{-\infty}^x f(t) dt.$

$$= \begin{cases} 0 & x < a \\ \int_a^x \frac{1}{b-a} dt & a \leq x < b \\ \int_a^b \frac{1}{b-a} dt & x \geq b \end{cases} = \begin{cases} 0 & x < a \\ \frac{x-a}{b-a} & a \leq x < b \\ 1 & x \geq b \end{cases}$$

1) Let  $X \sim U(0, 1)$ . does  $E(\frac{1}{x})$  exist?

$$\begin{aligned}
E\left(\frac{1}{x}\right) &= \int_0^1 \frac{1}{x} dx = \lim_{a \rightarrow 0^+} \int_a^1 \frac{1}{x} dx \\
&= \lim_{a \rightarrow 0^+} [\log |x|]_a^1 \\
&= \lim_{a \rightarrow 0^+} (\ln 1 - \ln a) \\
&= \lim_{a \rightarrow 0^+} \ln \frac{1}{a} \\
&= \infty
\end{aligned}$$

$E\left(\frac{1}{x}\right)$  does not exist.

2) Let  $X \sim U(0, n)$   $\forall n \in \mathbb{N}$ , let  $Y = X - [X]$ , show that  $Y \sim U(0, 1)$

$$\begin{aligned}
F_Y(y) &= P(Y \leq y) = P[X - [X] \leq y] \\
&= \sum_{k=0}^{n-1} P[X \leq [x] + y, k \leq X < k+1], [x] = k, \text{ say} \\
&= \sum_{k=0}^{n-1} P[X \leq k+y, k \leq X < k+1] \\
&= \sum_{k=0}^{n-1} P[k \leq X \leq k+y] \\
&= \sum_{k=0}^{n-1} \int_k^{k+y} \frac{1}{n} dx, 0 < y < 1 \\
&= \sum_{k=0}^{n-1} \frac{y}{n} \\
&= y, 0 < y < 1
\end{aligned}$$

$$F_Y(y) = \begin{cases} 0, & y < 0 \\ y, & 0 \leq y < 1 \\ 1, & y \geq 1 \end{cases} \quad f_Y(y) = \begin{cases} 1, & 0 < y < 1 \\ 0, & \text{ow.} \end{cases}$$

$$\Rightarrow Y = X - [X] \sim U(0, 1)$$

3) If  $X \sim U\{0, 1, \dots, n\}$  and  $Y \sim U(0, 1)$ . Find the CDF of  $Z = X + Y$  and find  $f_{XZ}$ .

$Z$  can take any value in interval  $(0, n+1)$   
 $F_Z(z) = P(Z \leq z) = P(X + Y \leq z)$

$$\begin{pmatrix} 0 \\ P(X=0, Y \leq z) \end{pmatrix}$$

$$\begin{cases} z \leq 0 \\ 0 \leq z < 1 \end{cases}$$

$$= \begin{cases} P(X=0, Y \leq 1) + P(X=1, Y \leq Z-1) & , 1 \leq Z < 2 \\ P(X=0, Y \leq 1) + P(X=1, Y \leq 1) + P(X=2, Y \leq Z-2) & , 2 \leq Z < 3 \\ P(X=0, Y \leq 1) + P(X=1, Y \leq 1) + \dots + P(X=K-1, Y \leq 1) \\ \quad + P(X=K, Y \leq Z-K) & , K \leq Z < K+1 \end{cases}$$

$$= \begin{cases} 0 & , Z < 0 \\ \frac{K}{n+1} + \frac{1}{n+1} (Z-K) & , K \leq Z < K+1, K=0,1,\dots,n \\ 1 & , Z \geq n+1 \end{cases}$$

$$f_Z(z) = \begin{cases} \frac{1}{n+1} & , 0 \leq z < n+1 \\ 0 & , \text{otherwise} \end{cases}$$

$$Z = X+Y \sim U(0, n+1)$$

4) Praguel and his girlfriend decide to meet at a certain location. If each of them arrives at a time uniformly distributed between 12 noon to 1 pm. Find the probability that the first one to arrive has to wait longer than 10 minutes.

Let,  $X$  and  $Y$  be the arrival time in minutes of Praguel and his gf after 12 noon.

$$X \sim U(0, 60)$$

$$Y \sim U(0, 60)$$

$$P(|X-Y| > 10) = \iint_{|x-y| > 10} f_{X,Y}(x,y) dx dy$$

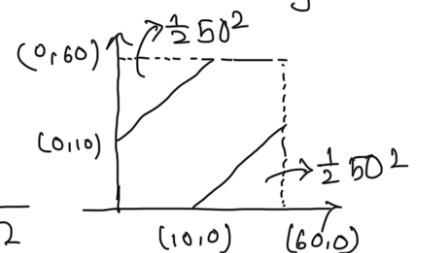
$$= \iint_A \frac{1}{60^2} dx dy$$

$$= \frac{\iint_A dx dy}{60^2}$$

$$= \frac{\text{Bounded by the region A}}{\text{The area bounded by } \Omega}$$

$$= \frac{50^2}{60^2} = \frac{5}{6}$$

$$A = \{(x,y) : |x-y| > 10, 0 < x,y < 60\}$$



$$\int_a^b \int_c^d f(x,y) dx dy = \text{base area} \times h$$

$$\Rightarrow h \int_a^b \int_c^d dx dy = \text{base area} \times h$$

$$\Rightarrow \int_a^b \int_c^d dx dy = \text{Area}$$

### Median :

let  $M_e$  be the median of the distribution.

$$\text{By definition, } \int_a^{M_e} f(x) dx = \int_{M_e}^b f(x) dx = \frac{1}{2}$$

$$\Rightarrow \int_{M_e}^b f(x) dx = \frac{1}{2}$$

$$\Rightarrow \frac{\beta - M_e}{\beta - \alpha} = \frac{1}{2}$$

$$\Rightarrow \beta - M_e = \frac{\beta - \alpha}{2}$$

$$\Rightarrow M_e = \beta - \frac{\beta - \alpha}{2} = \frac{\beta + \alpha}{2}$$

$$\text{Mean} = \text{median} = \frac{\alpha + \beta}{2}$$

Mode does not exist for uniform distribution.

### Quantile Deviation :

let  $Q_1, Q_3$  be the 1st and 3rd quantile of the distribution,

By definition,

$$\int_a^{Q_1} f(x) dx = \frac{1}{4}$$

$$\Rightarrow \frac{Q_1 - \alpha}{\beta - \alpha} = \frac{1}{4}$$

$$\begin{aligned} \Rightarrow Q_1 &= \frac{\beta - \alpha}{4} + \alpha \\ &= \frac{3\alpha + \beta}{4} \end{aligned}$$

$$\int_a^{Q_3} f(x) dx = \frac{3}{4}$$

$$\Rightarrow \frac{Q_3 - \alpha}{\beta - \alpha} = \frac{3}{4}$$

$$\begin{aligned} \Rightarrow Q_3 &= \frac{3}{4}(\beta - \alpha) + \alpha \\ &= \frac{3\beta + \alpha}{4} \end{aligned}$$

$$\begin{aligned} QD &= \frac{Q_3 - Q_1}{2} = \frac{3\beta + \alpha - 3\alpha - \beta}{8} \\ &= \frac{2\beta - 2\alpha}{8} = \frac{\beta - \alpha}{4} \end{aligned}$$

$$\frac{\quad}{8} - \frac{\quad}{4}$$

Remark :

Consider a uniform distribution over the region  $\Omega \in \mathbb{R}^n$ , then the PDF of  $X$  is ,

$$f(x) = \begin{cases} c & x \in \Omega \\ 0 & x \in \mathbb{R}^n - \Omega \end{cases}$$

$$\text{Then, } 1 = \int_{\Omega} f(x) dx \Rightarrow c \int_{\Omega} f(x) dx = 1$$

$$\Rightarrow c = \frac{1}{\text{Measure of } \Omega}$$

The probability that one randomly selected point  $\tilde{x}$  falls in a region  $A \subset \mathbb{R}^n = P(\tilde{x} \in A) = \int_A f(x) dx$

$$= c \int_A dx$$

$$= \frac{\text{The measure of } A}{\text{The measure of } \Omega}$$

**Moments :**

$$\begin{aligned} \mu'_n &= E(x^n) = \int_a^b x^n \frac{1}{b-a} dx = \left[ \frac{x^{n+1}}{(n+1)(b-a)} \right]_a^b \\ &= \frac{b^{n+1} - a^{n+1}}{(b-a)(n+1)} \end{aligned}$$

$$E(x) = \mu'_1 = \frac{b+a}{2}$$

$$\begin{aligned} V(x) &= \mu'_2 - \mu'_1{}^2 = \frac{b^3 - a^3}{3(b-a)} - \frac{(b+a)^2}{4} \\ &= \frac{(b-a)^2}{12} \end{aligned}$$

5) let  $X$  be a RV defined on  $[0,1]$  if  $P(x < X < y) \propto (y-x) \forall 0 < x < y < 1$ , show that  $X \sim U(0,1)$

let  $F(x)$  be the DF of  $X$ , ,  $0 \leq x \leq y \leq 1$

$$P(x \leq X \leq y) = K(y-x)$$

$$\Rightarrow F(y) - F(x) = K(y-x)$$

$$\Rightarrow \frac{F(y) - F(x)}{y-x} = K \quad \forall x < y$$

$$y - x$$

$$\Rightarrow \lim_{y \rightarrow x} \frac{F(y) - F(x)}{y - x} = k$$

$$\Rightarrow F'(x) = k$$

$$\Rightarrow F(x) = kx + C, \quad 0 \leq x \leq 1.$$

$$F(0) = 0 \Rightarrow C = 0$$

$$F(1) = 1 \Rightarrow k = 1.$$

The CDF of  $X$  is  $F_X(x) =$