

NEGATIVE BINOMIAL DISTRIBUTION :-

Note that, $(1-q)^{-r} = \sum_{x=0}^{\infty} \binom{x+r-1}{x} q^x$ for $|q| < 1$ (Taylor's Expansion)
and $r > 0$ (real no)

This is called Binomial Theorem for negative index.

$\sum_{x=0}^{\infty} \binom{x+r-1}{x} p^r q^x = 1$ and the terms of the series forms a pmf.

The probability distribution given by this pmf is known as Negative Binomial Distribution.

Definition - A counting random variable is said to follow a negative Binomial Distribution if its pmf is given by,

$$f(x) = \begin{cases} \binom{x+r-1}{x} p^r q^x, & x=0, 1, 2, \dots \quad 0 < p < 1 \\ 0 & \text{otherwise} \end{cases} \quad q = 1-p, \quad r > 0 \text{ (real no)}$$

Moments -

Note that, $(1-q)^{-r} = \sum_{x=0}^{\infty} \binom{x+r-1}{x} q^x$

Differentiating w.r.t q , $(-r)(1-q)^{-r-1}(-1) = \sum_{x=0}^{\infty} \binom{x+r-1}{x} x q^{x-1}$
 $\Rightarrow r p^{-r-1} = \sum_{x=1}^{\infty} \binom{x+r-1}{x} x q^{x-1}$

Again, differentiating w.r.t q , $r(r+1)(1-q)^{-r-2} = \sum_{x=2}^{\infty} \binom{x+r-1}{x} x(x-1) q^{x-2}$

$$\begin{aligned} E(x) &= \sum_{x=0}^{\infty} x \binom{x+r-1}{x} p^r q^x \\ &= p^r q \sum_{x=1}^{\infty} x \binom{x+r-1}{x} q^{x-1} \\ &= p^r q r p^{-r-1} = \frac{rq}{p} \end{aligned}$$

(Derivative of sum = Sum of derivatives but it is not always true)

$$\begin{aligned} V(x) &= E(x^2) - E^2(x) = E(x(x-1)) + E(x) - E^2(x) \\ &= p^r q^2 \sum_{x=2}^{\infty} x(x-1) \binom{x+r-1}{x} q^{x-2} + \frac{rq}{p} - \frac{r^2 q^2}{p^2} \\ &= p^r q^2 r(r+1) p^{-r-2} + \frac{rq}{p} - \frac{r^2 q^2}{p^2} \end{aligned}$$

$$= r(r+1) \frac{q^2}{p^2} + \frac{rq}{p} - \frac{r^2q^2}{p^2}$$

$$= \frac{rq}{p} \left[\frac{rq}{p} + \frac{q}{p} + 1 - \frac{rq}{p} \right]$$

$$= \frac{rq}{p^2}$$

$$E(x) = \frac{rq}{p} < \frac{rq}{p^2} = var(x)$$

Recursion Relation of Negative Binomial Distribution :-

$$u_m = E[x - E(x)]^m = E\left[x - \frac{rq}{p}\right]^m$$

$$= \sum_{x=0}^{\infty} \left(x - \frac{rq}{p}\right)^m \binom{x+r-1}{x} p^r q^x$$

Differentiating wrt p, we get,

$$\begin{aligned} \frac{\partial u_m}{\partial p} &= \sum_{x=0}^{\infty} \left(x - \frac{rq}{p}\right)^m \binom{x+r-1}{x} r p^{r-1} q^x + \sum_{x=0}^{\infty} \left(x - \frac{rq}{p}\right)^m \binom{x+r-1}{x} p^m x q^{x-1} (-1) \\ &\quad + \sum_{x=0}^{\infty} m \left(x - \frac{rq}{p}\right)^{m-1} \binom{x+r-1}{x} p^r q^x (-r) \frac{p(-1) - (1-p)}{p^2} = \frac{-1}{p^2} \\ &= \sum_{x=0}^{\infty} \left(x - \frac{rq}{p}\right)^{m-1} \binom{x+r-1}{x} p^r q^x \left(\frac{mr}{p^2}\right) + \\ &\quad \sum_{x=0}^{\infty} \left(x - \frac{rq}{p}\right)^m \binom{x+r-1}{x} p^m q^x \left[\frac{r}{p} - \frac{x}{mr}\right] \\ &= \frac{mr}{p^2} u_{m-1} - \frac{1}{mr} u_{m+1} \\ &= - \left(x - \frac{rq}{p}\right) \frac{1}{mr} \end{aligned}$$

$$\Rightarrow \frac{1}{mr} u_{m+1} = \frac{mr}{p^2} u_{m-1} - \frac{\partial u_m}{\partial p}$$

$$\Rightarrow u_{m+1} = mr \left(\frac{mr}{p^2} u_{m-1} - \frac{\partial u_m}{\partial p} \right)$$

This is the recursion relation of negative binomial distribution.

Now, $u_0 = 1, u_1 = 0$

$$\text{If } m=1, u_2 = q \left(\frac{mr}{p^2} u_0 - 0 \right) = \frac{qr}{p^2}$$

$$m_2 = 2, u_3 = q \left(\frac{2r}{m_2} u_1 - \frac{\partial u_2}{\partial p} \right)$$

$$\frac{\partial}{\partial p} \frac{qr}{p^2} = r - \frac{p(-1) - (1-p)2p}{p^4}$$

$$\begin{aligned}
&= q \left(\frac{2n}{p^2} \cdot 0 - n \frac{(p-2)}{p^3} \right) \\
&= q n \frac{(2-p)}{p^3} \\
\text{skewness} &= \frac{\mu_3}{\mu_2^{3/2}} = \frac{q n \frac{(2-p)}{p^3}}{\left(\frac{q n}{p^2} \right)^{3/2}} = \frac{n}{\sqrt{qn}} > 0
\end{aligned}$$

$$\begin{aligned}
&= n \frac{-p^2 - 2p + 2p^2}{p^4} \\
&= n \frac{p^2 - 2p}{p^4} \\
&= n \frac{(p-2)}{p^3}
\end{aligned}$$

\Rightarrow the distribution is positively skewed.

$$\begin{aligned}
\text{Kurtosis} &= \frac{\mu_4}{\mu_2^2} - 3 = \frac{nq(1+4q+q^2) + 3n^2q^2}{\frac{q^2 n^2}{p^4}} - 3 \\
&= \frac{1+4q+q^2}{q^2} + 3 - 3 > 0
\end{aligned}$$

$$\begin{aligned}
\mu_4 &= q \left\{ \frac{3n}{p^2} \mu_2 - \frac{\partial \mu_3}{\partial p} \right\} \\
&= q \left\{ 3q \left(\frac{n}{p^2} \right)^2 - \left\{ -\frac{n}{p^2} (6p^{-2} - 6p^{-1} + 1) \right\} \right\} \\
&= 3 \left(\frac{q n}{p^2} \right)^2 + \frac{nq}{p^2} \left(\frac{6}{p^2} - \frac{6}{p} + 1 \right) \\
&= \frac{nq(1+4q+q^2) + 3n^2q^2}{p^4}
\end{aligned}$$

① If $X \sim NB(n, p)$, then show that $E(X^k) = \frac{nq}{p} E(Y+1)^{k-1}$ where $Y \sim NB(n+1, p)$. Hence, find $E(X)$, $\text{Var}(X)$.

$$\begin{aligned}
E(X^k) &= \sum_{x=0}^{\infty} x^k \binom{n+r-1}{x} p^n q^r x \\
&= \sum_{x=1}^{\infty} x^{k-1} \frac{(x+r-1) \dots (n+1)r}{(x-1)!} p^n q^r x \\
&= \frac{nq}{p} \sum_{x=1}^{\infty} x^{k-1} \frac{(x-1+r+1-1) \dots (n+1)}{(x-1)!} p^{n+1} q^{x-1} \\
&= \frac{nq}{p} \sum_{x-1=y=0}^{\infty} (y+1)^{k-1} \frac{(y+r+1-1) \dots (n+1)}{y!} p^{n+1} q^y \\
&= \frac{nq}{p} \sum_{y=0}^{\infty} (y+1)^{k-1} \underbrace{\left(\frac{y+r+1-1}{y} \right)}_{\text{pmf of } NB(n+1, p)} p^{n+1} q^y
\end{aligned}$$

then show that the mean and variance can be written as the corresponding formula for binomial distribution.

$$E(X) = \frac{r}{p} E(Y+1) = \frac{rq}{p}$$

$$\begin{aligned} E(X^2) &= \frac{rq}{p} (E(Y)+1) = \frac{rq}{p} \left(\frac{(r+1)q}{p} + 1 \right) \\ &= \frac{rq}{p} \frac{rq+q+p}{p} = \frac{rq(rq+1)}{p^2} \end{aligned}$$

② Derive the PMF of NB distribution in the following form,

$$f(x) = \binom{r}{x} p^x q^{r-x}, \quad x=0,1,2,\dots$$

$$P+Q=1$$

$$\begin{aligned} f(x) &= \binom{x+r-1}{x} p^r q^x \\ &= \frac{(x+r-1) \dots (r+1)r}{x!} \left(\frac{1}{p}\right)^{-r} q^x \\ &= \frac{(-r)(-r-1) \dots (-x-r-1)}{x!} \left(\frac{1}{p}\right)^{-r} (-q)^x \\ &= \binom{-r}{x} \left(-\frac{q}{p}\right)^x \left(\frac{1}{p}\right)^{-r-x} \\ &= \binom{-r}{x} (p)^x Q^{-r-x} \quad \text{where } P = -\frac{q}{p}, \quad Q = \frac{1}{p} \end{aligned}$$

which is in the form of Binomial Distribution.

$$E(X) = np = (-r)\left(-\frac{q}{p}\right) = \frac{rq}{p}$$

$$V(X) = npq = (-r)\left(-\frac{q}{p}\right)\left(\frac{1}{p}\right) = \frac{rq}{p^2}$$

Probability Model:

Consider a sequence of independent Bernoulli trials with probability of success p . Let X be the no of failure before the 5th success.

So, then

$$P(X=x)$$

= P [that there are $(r-1)$ success in the first $(x-r+1)$ trials and the $(x+r)$ th trial results in a success]

$$= P[(r-1) \text{ success in } (x+r-1) \text{ trials}] P(\text{success})$$

$$= \binom{x+r-1}{x} p^{r-1} q^x p$$

$$= \binom{x+r-1}{x} p^r q^x, \quad x=0,1,2,\dots$$

This is the pmf of $NB(r, p)$ distribution.

Let Y be the no of trials needed to get the 5th success then

Let, t be the no of trials needed to get the total success. Then
 $Y = X+t$ with probability 1.

$$P(Y=y) = P(X=y-t) = \binom{y-r}{y-n} p^r q^{y-n}, \quad y=n, n+1, \dots$$

$$E(Y) = E(X)+t = \frac{rq}{p} + t = \frac{t}{p}$$

$$V(Y) = V(X) = \frac{rq}{p^2} \quad [\because \text{no change due to linear shift}]$$

Cumulative Distribution Function (CDF) :

$$P(X \leq K) = \sum_{x=0}^K \binom{x+r-1}{x} p^r q^x = g(p) \text{ (say)}$$

Differentiating w.r.t p ,

$$\begin{aligned} g'(p) &= \sum_{n=0}^K \binom{n+r-1}{n} np^{n-1} (1-p)^n - \sum_{n=0}^K \binom{n+r-1}{n} p^n n (1-p)^{n-1} \\ &= \sum_{n=0}^K \binom{n+r-1}{n} np^{n-1} (1-p)^n - \sum_{n=0}^K \frac{(n+r-1)!}{(n-1)! (r-1)!} p^n (1-p)^{n-1} \\ &= \sum_{n=0}^K \binom{n+r-1}{n} np^{n-1} (1-p)^n - r \sum_{n=1}^K \binom{n+r-1}{n-1} p^n (1-p)^{n-1} \\ &\quad \downarrow p^{n-1} \{1-(1-p)\} \\ &\quad - r \sum_{n=1}^K \binom{n+r-1}{n-1} (1-p)^{n-1} p^{n-1} \\ &\quad + r \sum_{n=1}^K \binom{n+r-1}{n-1} p^{n-1} (1-p)^n \\ &= \sum_{n=0}^K \left\{ \binom{n+r-1}{n} + \binom{n+r-1}{n-1} \right\} np^{n-1} (1-p)^n - r \sum_{n=1}^K \binom{n+r-1}{n-1} (1-p)^{n-1} p^{n-1} \\ &= \sum_{n=0}^K \binom{n+r}{n} np^{n-1} (1-p)^n - r \sum_{n=1}^K \binom{n+r-1}{n-1} (1-p)^{n-1} p^{n-1} \\ &= \sum_{n=0}^K \binom{n+r}{n} np^{n-1} (1-p)^n - r \sum_{n=1}^{K-1} \binom{n+r}{n} (1-p)^{n-1} p^{n-1} \\ &= \binom{K+r}{K} np^{K-1} (1-p)^K \quad [= g(p)] \end{aligned}$$

$$g'(p) = \binom{r+k}{k} np^{k-1} (1-p)^k$$

$$g'(t) = \binom{r+k}{k} nt^{k-1} (1-t)^k$$

$$\int_0^p g'(t) dt = \int_0^p r \binom{r+k}{k} t^{k-1} (1-t)^k dt$$

$$\Rightarrow g(p) - g(0) = \frac{\int_0^p t^{k-1} (1-t)^{k+1-1} dt}{B(r, k+1)} = I_p(r, k+1)$$

$$P\left(\frac{n+r}{k}\right) = \frac{(n+r)!}{k! n! (r-1)!} = \frac{1}{B(n, k+1)}$$

$$\Rightarrow g(p) = \sum_{x=0}^k \binom{n+r-1}{x} p^n q^x = g(p) = I_p(n, k+1)$$

It is an increasing function of p and decreasing function of q .

Mode :-

$$f(x) = \begin{cases} \binom{x+r-1}{x} p^n q^x & , x=0, 1, \dots, p+q=1 \\ 0 & \text{otherwise} \end{cases}$$

$$\frac{f(x)}{f(x-1)} = \frac{\binom{n+r-1}{x} p^n q^x}{\binom{n+r-2}{x-1} p^n q^{x-1}} = q \cdot \frac{x+r-1}{x}$$

This is according as $\frac{x+r-1}{x} q \geq 1$

$$\Rightarrow (x+r-1)q \geq x$$

$$\Rightarrow x(q-1) \geq (r-1)q$$

$$\Rightarrow x \geq \frac{(r-1)q}{q-1}$$

Case I - $\frac{(r-1)q}{q-1}$ is not an integer., let $k = \left[\frac{(r-1)q}{q-1} \right]$ (say)

$$\frac{f(1)}{f(0)} > 1, \frac{f(2)}{f(1)} > 1, \frac{f(3)}{f(2)} > 1, \dots, \frac{f(k)}{f(k-1)} > 1, \frac{f(k+1)}{f(k)} < 1,$$

$$\therefore f(0) < f(1) < \dots < f(k-1) < f(k) > f(k+1) > f(k+2) > \dots$$

$\Rightarrow f(k)$ is the maximum.

$k = \left[\frac{(r-1)q}{q-1} \right]$ is the mode.

Case II: $\frac{(r-1)q}{q-1}$ is an integer. ($= k$)

$$f(0) < f(1) < f(2) < \dots < f(k-1) = f(k) > f(k+1) > f(k+2) > \dots$$

$\therefore f(k)$ or $f(k-1)$ be the maximum.

$k = \frac{(r-1)q}{q-1}$ is the mode.

Tail Function :-

$$P(X > k) = 1 - P(X \leq k) = 1 - I_p(n, k+1)$$

$$= \int_0^p t^{n-1} (1-t)^{k+1-1} dt$$

$$\begin{aligned}
&= 1 - \frac{\int_0^P t^{n-1} (1-t)^{k+1-1} dt}{\int_0^1 t^{n-1} (1-t)^{k+1-1} dt} \\
&= \frac{\int_0^P (1-t)^{k+1-1} + t^{n-1} dt}{B(n, k+1)} \\
&= \frac{\int_0^P -z^{k+1-1} (1-z)^{n-1} dz}{B(n, k+1)} \\
&= \frac{\int_0^q z^{k+1-1} (1-z)^{n-1} dz}{B(n, k+1)} \\
&= I_q(k+1, n)
\end{aligned}$$

$1-t = z$
 $-dt = dz$
 $\begin{array}{c|cc} t & 1 & P \\ \hline z & 0 & q \end{array}$

③ Let X be the intensity of an accident proneness of an individual and let Y be the no of accident as the individual. Let,

$$f_X(x) = \begin{cases} \frac{\theta^\alpha}{\Gamma(\alpha)} e^{-\theta x} x^{\alpha-1}, & x \geq 0 \\ 0, & \text{otherwise} \end{cases}, \quad \theta > 0, \alpha > 0 \quad (\text{real no})$$

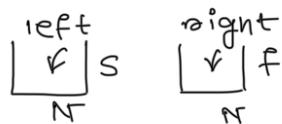
$$f_{Y|X}(y|x) = \begin{cases} e^{-x} \frac{x^y}{y!}, & y=0,1,2,\dots \\ 0 & \text{otherwise} \end{cases}$$

$$\begin{aligned}
f_Y(y) &= \int_{-\infty}^{\infty} f(y|x) f(x) dx, \quad y \in \mathbb{N} \\
&= \int_0^{\infty} e^{-x} \frac{x^y}{y!} \frac{\theta^\alpha}{(\alpha-1)!} e^{-\theta x} x^{\alpha-1} dx \\
&= \frac{\theta^\alpha}{y! (\alpha-1)!} \int_0^{\infty} e^{-x(1+\theta)} x^{y+\alpha-1} dx \\
&= \frac{\theta^\alpha}{y! (\alpha-1)!} \frac{(\alpha+y-1)!}{(1+\theta)^{y+\alpha}} \\
&= \left(\frac{y \alpha^{-1}}{y} \right) \left(\frac{\theta}{1+\theta} \right)^\alpha \left(1 - \frac{\theta}{1+\theta} \right)^y, \quad y=0,1,2,\dots
\end{aligned}$$

$$Y \sim NB(\alpha, \frac{\theta}{1+\theta})$$

A certain mathematician always carries one match box in his right pocket and one in his left. When he wants a match he selects a pocket at random and choose a match. Suppose initially each box contained exactly N matches and consider the moment when for the first time our mathematician discovers that one box is empty. At that point in time, find the probability that the other box contain exactly r matches.

matches. Find $E(X)$.



Let us identify the success with the choice of the left pocket box. The left pocket box will be empty when the right pocket box contains r matches iff exactly $(n-r)$ failure preceding $(N+r)$ of success. A similar argument applies to the right pocket box.

$$P_{nr} = 2 \binom{N-r + \frac{N+1}{2} - 1}{N-r} \left(\frac{1}{2}\right)^{N+1} \left(\frac{1}{2}\right)^{N-r} \quad [\text{left} + \text{right}]$$

$$= \binom{2N-r}{N-r} \left(\frac{1}{2}\right)^{2N-r}, \quad r=0, 1, \dots, N$$

Clearly, P_{nr} forms a pmf of a distribution.

$$\begin{aligned} M_r &= \sum_{r=0}^N r P_{nr} = \sum_{r=0}^N r \binom{2N-r}{N-r} \left(\frac{1}{2}\right)^{2N-r} \\ (N-M_r) &= \sum_{r=0}^N (N-r) P_{nr} \\ &= \sum_{r=0}^N (N-r) \frac{(2N-r)!}{(N-r)! N!} \left(\frac{1}{2}\right)^{2N-r} \\ &= \sum_{r=0}^{N-1} \frac{(2N-r)}{2} \frac{(2N-r-1)!}{(N-r-1)! N!} \left(\frac{1}{2}\right)^{2N-r-1} \\ &= \sum_{r=0}^{N-1} \frac{2N-r}{2} \binom{2N-r-1}{N-r-1} \frac{1}{2^{2N-r-1}} \\ &= \sum_{r=0}^{N-1} \frac{(2N+1)-(r+1)}{2} P_{nr+1} \\ &= \frac{2N+1}{2} \sum_{r=0}^{N-1} P_{nr+1} - \frac{1}{2} \sum_{r=0}^{N-1} (r+1) P_{nr+1} \\ &= \frac{2N+1}{2} \sum_{r+1=k=1}^N P_k - \frac{1}{2} \sum_{r=0}^{N-1} (r+1) P_{nr+1} \\ \Rightarrow (N-M_r) &= \frac{2N+1}{2} (1-P_0) - \frac{M_r}{2} \\ \Rightarrow -\frac{M_r}{2} &= \frac{2N+1}{2} \left\{ 1 - \binom{2N}{N} \left(\frac{1}{2}\right)^{2N} \right\} - N \\ \Rightarrow \frac{M_r}{2} &= N - \frac{2N+1}{2} \left\{ 1 - \binom{2N}{N} \left(\frac{1}{2}\right)^{2N} \right\} \\ &= N - \frac{2N+1}{2} + \frac{2N+1}{2} \binom{2N}{N} \left(\frac{1}{2}\right)^{2N} \end{aligned}$$

$$= \frac{2^{N+2} - 2^{N+1}}{2} + \frac{2^{N+1}}{2} \binom{2N}{N} \left(\frac{1}{2}\right)^{2N}$$

$$\Rightarrow M = (2N+1) \binom{2N}{N} \left(\frac{1}{2}\right)^{2N} - 1$$

Relationship with Binomial Distribution :-

$$\text{If } X \sim NB(n, p), \quad P(X \leq k) = P(U \geq n) \quad . \quad U \sim \text{Bin}(k+n, p)$$

$$= P(Y \leq k) \quad , \quad Y \sim \text{Bin}(k+1, q)$$

X denotes the number of failure preceding the n th success.
 $(n+x)$ denotes the no of trials required to get the n th success.

$$P(X \leq k) = P(X+n \leq k+n)$$

$$= P(\text{no of trials required to get } n\text{th success} \leq n+k)$$

$$\text{If we define } U \text{ as the no of success in } (n+k) \text{ trials,}$$

$$= P(U \geq n) \quad \text{So, } U \sim \text{Bin}(k+n, p)$$

$$= P(-U \leq -n) \quad (k+n-U) \sim \text{Bin}(k+n, q)$$

$$= P(k+n-U \leq k) \quad \Rightarrow V \sim \text{Bin}(k+n, q)$$

$$= P(V \leq k)$$

Let X be a random variable with pmf,

$$f(x) = \begin{cases} c \binom{2n-x}{n} 2^x, & x=0,1,\dots,n \\ 0, & \text{otherwise} \end{cases}$$

Determine c such that it is a pmf.

$$\sum_{x=0}^n f(x) = 1$$

$$\Rightarrow c \sum_{x=0}^n \binom{2n-x}{n} 2^x = 1$$

$$\Rightarrow c \sum_{y=0}^n \binom{n+y}{n} 2^{n-y} = 1$$

$$\Rightarrow c 2^{2n+1} \sum_{y=0}^n \underbrace{\binom{y+(n+1)-1}{y}}_{= P(Y \leq n), Y \sim NB(n+1, \frac{1}{2})} \left(\frac{1}{2}\right)^{n+1} \left(\frac{1}{2}\right)^y = 1 \quad \text{--- (1)}$$

$$= P(Y \leq n), Y \sim NB(n+1, \frac{1}{2})$$

$$= P(Y \leq n), Y \sim \text{Bin}(2n+1, \frac{1}{2})$$

$$= \sum_{y=0}^n \binom{2n+1}{y} \left(\frac{1}{2}\right)^{2n+1}$$

$y = 0, 1, \dots, 2n+1$

$$= \frac{1}{2}$$

$$\Rightarrow c 2^{2n+1} \frac{1}{2} = 1$$

$$\Rightarrow c = 2^{-2n}$$

$$f(x) = 2^{-2n} \binom{2n-x}{n} 2^x = \binom{2n-x}{n} \left(\frac{1}{2}\right)^{2n-x}$$

Alternative, by ①, $c 2^{2n+1} \underbrace{P(Y \leq n)}_{I_{Y_2}(n+1, n+1)} = 1$, $Y \sim NB(n+1, \frac{1}{2})$

$$\Rightarrow c 2^{2n+1} I_{Y_2}(n+1, n+1) = 1$$

$$\Rightarrow c 2^{2n+1} \frac{\int_0^{Y_2} t^n (1-t)^n dt}{B(n+1, n+1)} = 1$$

$$\Rightarrow c 2^{2n+1} \frac{\int_0^{Y_2} t^n (1-t)^n dt}{2 \int_0^{Y_2} t^n (1-t)^n dt} = 1$$

$$\Rightarrow c 2^{2n+1} \frac{1}{2} = 1$$

$$\Rightarrow c = \left(\frac{1}{2}\right)^{2n}$$

Putting the value of c , we proof it is a pmf.

$$\begin{aligned} & \sum_x f(x) \\ &= 2^{-2n} \sum_{n=0}^{\infty} \binom{2n-x}{n} 2^x \quad n-x=y \\ &= 2^{-2n} \sum_{y=0}^n \binom{n+y}{y} 2^{n-y} \\ &= 2^{-2n} 2^{2n+1} \sum_{y=0}^n \binom{y+(n+1)-1}{y} \left(\frac{1}{2}\right)^{n+1} \left(\frac{1}{2}\right)^y \\ &= 2 P(Y \leq n) \\ &= 2 I_{Y_2}(n+1, n+1) \\ &= 2 \frac{1}{2} \\ &= 1. \end{aligned}$$

Example:

In a series of indefinite Bernoulli trials with success probability p . Find the probability that 'a' successes will appear before 'b' failures. (Pascal format).

Let λ denote the no of failures preceding a th success.

$$X \sim NB(a, p)$$

$$P(X \leq b-1) = \sum_{x=0}^{b-1} \binom{a+b-1}{x} p^a q^x$$

Y denotes the no of failures in $(a+b-1)$ trials.

$$Y \sim \text{Bin}(a+b-1, q)$$

$$P(Y \leq b-1) = \sum_{y=0}^{b-1} \binom{a+b-1}{y} q^y p^{a+b-1-y}$$

$$\text{So, } P(X \leq b-1) = P(Y \leq b-1)$$

$$\Rightarrow \sum_{i=0}^{b-1} \binom{a+i-1}{a-i} p^a q^i = \sum_{i=0}^{b-1} \binom{a+b-1}{i} q^i p^{a+b-1-i}$$

$$\Rightarrow \sum_{i=0}^{b-1} \binom{a+i-1}{a-i} = \sum_{i=0}^{b-1} \binom{a+b-1}{i} p^{b-1-i}$$

Prove that $X_1 \sim NB(r_1, p)$ > independent
 $X_2 \sim NB(r_2, p)$

$$X_1 + X_2 \sim NB(r_1 + r_2, p)$$

$$P(X_1 + X_2 = y) = \sum_x P(X_1 = x, X_2 = y - x)$$

$$= \sum_x P(X_1 = x) P(X_2 = y - x)$$

$$= \sum_x \binom{x+r_1-1}{r_1-1} p^{r_1} q^x \binom{y-x+r_2-1}{r_2-1} p^{r_2} q^{y-x}$$

$$= \sum_x \binom{x+r_1-1+y-x+r_2-1}{r_1+r_2-2} p^{r_1+r_2} q^y$$

$$= \sum_y \binom{y+r_1+r_2-2+1}{r_1+r_2-2} p^{r_1+r_2} q^y$$

$$= \sum_y \binom{y+r_1+r_2-1}{y} p^{r_1+r_2} q^y$$

$$Y \sim NB(r_1 + r_2, p)$$

$X_i \sim NB(r_i, p)$ $i=1, 2$. Find the conditional distribution of $(X_1 | X_1 + X_2 = y)$

$$P(X_1 = x | X_1 + X_2 = y) = \frac{P(X_1 = x) P(X_2 = y - x)}{P(X_1 + X_2 = y)}$$

$$\begin{aligned}
&= \frac{\binom{x+r_1-1}{x} p^{r_1} q^x \binom{y-x+r_2-1}{y-x} p^{r_2} q^{y-x}}{\binom{y+r_1+r_2-1}{y} p^{r_1+r_2} q^y} \\
&= \frac{\binom{x+r_1-1}{r_1-1} \binom{y-x+r_2-1}{r_2-1}}{\binom{y+r_1+r_2-1}{r_1+r_2-1}}
\end{aligned}$$

It is independent of parameters.

Remark:

In $\text{Bin}(n,p)$ distribution we take a fixed number of trials, then we count no of success but in $\text{NB}(r,p)$ distribution we fix up the no of success and then we count the no of trials required to get the Roth success. Hence, the NB distribution (Pascal's) is also known as Inverse Binomial Sampling.

$$X \sim \text{Bin}(n,p), \text{ show that } E\left(\frac{1}{1+x}\right) = \frac{p}{q(n-1)} \{1 - p^{n-1}\}$$

$$\begin{aligned}
E\left(\frac{1}{1+x}\right) &= \sum_{x=0}^{\infty} \frac{1}{1+x} \binom{x+r-1}{x} p^x q^x \\
&= \sum_{x=0}^{\infty} \frac{(x+r-1)!}{(x+1)! (r-1)!} p^x q^x \\
&= \sum_{x=0}^{\infty} \frac{(\overline{x+1} + \overline{r-1} - 1)!}{(x+1)! (r-1)!} p^x q^x \\
&= \frac{p}{q(r-1)} \sum_{x=0}^{\infty} \frac{(\overline{x+1} + \overline{r-1} - 1)!}{(x+1)! (r-2)!} p^{x-1} q^{r-1} \\
&= \frac{p}{q(r-1)} \sum_{x=1}^{\infty} \frac{(\overline{y} + \overline{r-1} - 1)!}{y! (r-2)!} p^{y-1} q^y \\
&= \frac{p}{q(r-1)} [1 - P(X=0)] \\
&= \frac{p}{q(r-1)} (1 - p^{r-1})
\end{aligned}$$

Mean Deviation about mean :-

$$\begin{aligned}
MD_M(x) &= E[|X - M|] \\
&= \sum_{x=0}^{\infty} |x - M| f(x)
\end{aligned}$$

$$\begin{aligned}
&= \sum_{x=0}^{[m]} (m-x) f(x) + \sum_{x=[m+1]}^{\infty} (n-m) f(x) \\
&= 2 \sum_{x=0}^{[m]} (m-x) f(x) \\
&= 2 \sum_{x=0}^{[m]} \left(\frac{m}{p} - x \right) f(x) \\
&= \frac{2}{p} \sum_{x=0}^{[m]} (m - px) f(x) \\
&= \frac{2}{p} \sum_{x=0}^{[m]} \left[(n+x) q f(x) - x f(x) \right]
\end{aligned}$$

Alternative,

$$\begin{aligned}
g(n) &= n f(n) \\
g(n+1) &= (n+1) f(n) \\
&= (n+1) \binom{n+1+n-1}{n+1} p^n q^{n+1} \\
&= (n+1) \frac{(2n)!}{(n+1)! (n-1)!} p^n q^{n+1} \\
&= (n+n) q \frac{(2n-1)!}{n! (n-1)!} p^n q^n \\
&= (n+n) q f(n)
\end{aligned}$$

$$\begin{aligned}
MD_{m,n}(x) &= \frac{2}{p} \sum_{x=0}^{[m]} \left\{ (n+x) q f(x) - x f(n) \right\} \\
&= \frac{2}{p} \sum_{x=0}^{[m]} \left\{ g(n+1) - g(n) \right\} \\
&= \frac{2}{p} \sum_{x=[m]+1}^{\infty} \left\{ g(n+1) - g(n) \right\} \\
&= \frac{2}{p} \left\{ -g[m] \right\}
\end{aligned}$$