Model Inference

by

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Introduction:

Here we provide a general exposition of the maximum likelihood approach and Bayesian method for inference.

A Smoothing Example:

We illustrate the bootstrap in a simple one-dimensional smoothing problem, and show its connection to maximum likelihood. Denote the training data by $Z = z_1, z_2, ..., z_N$, with $z_i = (x_i, y_i)$, i = 1, 2, ..., N. So,

$$\mu(x) = \sum_{j=1}^{7} \beta_j h_j(x).$$

This is a seven-dimensional linear space of functions and $\hat{\beta} = (\mathbf{H}^T \mathbf{H})^{-1} \mathbf{H}^T \mathbf{y}$ obtained by minimizing the squared error over the training set. The corresponding fit is $\hat{\mu}(x) = \sum_{i=1}^{7} \hat{\beta}_i h_i(x)$.

If we simulate new responses by adding Gaussian noise to the predicted values: $\hat{\mu}^*(x) \sim N(\hat{\mu}(x), h(x)^T (\mathbf{H}^T \mathbf{H})^{-1} h(x) \hat{\sigma}^2)$.

Now if we use maximum likelihood approach and Z has a normal distribution with mean μ and variance σ^2 , then $\hat{\beta} = (\mathbf{H}^T \mathbf{H})^{-1} \mathbf{H}^T \mathbf{v}$ which agrees with the least squares estimate.

Bayesian Methods:

In the Bayesian approach to inference we specify a sampling model $Pr(Z|\theta)$ and a prior distribution for the parameters $Pr(\theta)$, the posterior distribution

$$Pr(\theta \mid \mathbf{Z}) = \frac{Pr(\mathbf{Z} \mid \theta) \cdot Pr(\theta)}{\int Pr(\mathbf{Z} \mid \theta) \cdot Pr(\theta) d\theta}$$

which represents our updated knowledge about θ . The function $\mu(x)$ should be smooth, and have guaranteed this by expressing μ in a smooth low-dimensional basis of B splines.

Relationship Between the Bootstrap and Bayesian Inference:

Let $z \sim N(\theta,1)$ and $\theta \mid z \sim N\left(\frac{z}{1+1/\tau},\frac{1}{1+1/\tau}\right)$. Now the larger we take τ , the more concentrated the posterior becomes around the maximum likelihood estimate $\hat{\theta} = z$. his is the same as a parametric bootstrap distribution in which we generate bootstrap values z^* from the maximum likelihood estimate of the sampling density N(z,1)

The EM Algorithm:

The EM algorithm is a popular tool for simplifying difficult maximum likelihood problems.

- 1. Take initial guesses for the parameters $\hat{\mu}_1$, $\hat{\sigma}_1^2$, $\hat{\mu}_2$, $\hat{\sigma}_2^2$, $\hat{\pi}$.
- 2. Expectation Step: compute the responsibilities $\hat{\gamma}_i = \frac{\hat{\pi}\phi_{\hat{\theta}_2}(y_i)}{(1-\hat{\pi})\phi_{\hat{\theta}_1}(y_i) + \hat{\pi}\phi_{\hat{\theta}_2}(y_i)}$, i = 1, 2, ..., N
- 3. Maximization Step: compute the weighted means and variances:

$$\hat{\mu}_1 = \frac{\sum_{i=1}^N (1-\widehat{\gamma}_i) y_i}{\sum_{i=1}^N (1-\widehat{\gamma}_i)}, \quad \hat{\sigma}_1^2 = \frac{\sum_{i=1}^N (1-\widehat{\gamma}_i) (y_i - \widehat{\mu}_1)^2}{\sum_{i=1}^N (1-\widehat{\gamma}_i)}, \quad \hat{\sigma}_2^2 = \frac{\sum_{i=1}^N \widehat{\gamma}_i (y_i - \widehat{\mu}_2)^2}{\sum_{i=1}^N \widehat{\gamma}_i}, \quad \hat{\sigma}_2^2 = \frac{\sum_{i=1}^N \widehat{\gamma}_i (y_i - \widehat{\mu}_2)^2}{\sum_{i=1}^N \widehat{\gamma}_i}, \quad \hat{\sigma}_2^2 = \frac{\sum_{i=1}^N \widehat{\gamma}_i (y_i - \widehat{\mu}_2)^2}{\sum_{i=1}^N \widehat{\gamma}_i}, \quad \hat{\sigma}_2^2 = \frac{\sum_{i=1}^N \widehat{\gamma}_i (y_i - \widehat{\mu}_2)^2}{\sum_{i=1}^N \widehat{\gamma}_i}, \quad \hat{\sigma}_2^2 = \frac{\sum_{i=1}^N \widehat{\gamma}_i (y_i - \widehat{\mu}_2)^2}{\sum_{i=1}^N \widehat{\gamma}_i}, \quad \hat{\sigma}_2^2 = \frac{\sum_{i=1}^N \widehat{\gamma}_i (y_i - \widehat{\mu}_2)^2}{\sum_{i=1}^N \widehat{\gamma}_i}, \quad \hat{\sigma}_2^2 = \frac{\sum_{i=1}^N \widehat{\gamma}_i (y_i - \widehat{\mu}_2)^2}{\sum_{i=1}^N \widehat{\gamma}_i}, \quad \hat{\sigma}_2^2 = \frac{\sum_{i=1}^N \widehat{\gamma}_i (y_i - \widehat{\mu}_2)^2}{\sum_{i=1}^N \widehat{\gamma}_i}, \quad \hat{\sigma}_2^2 = \frac{\sum_{i=1}^N \widehat{\gamma}_i (y_i - \widehat{\mu}_2)^2}{\sum_{i=1}^N \widehat{\gamma}_i}, \quad \hat{\sigma}_2^2 = \frac{\sum_{i=1}^N \widehat{\gamma}_i (y_i - \widehat{\mu}_2)^2}{\sum_{i=1}^N \widehat{\gamma}_i}, \quad \hat{\sigma}_2^2 = \frac{\sum_{i=1}^N \widehat{\gamma}_i (y_i - \widehat{\mu}_2)^2}{\sum_{i=1}^N \widehat{\gamma}_i}, \quad \hat{\sigma}_2^2 = \frac{\sum_{i=1}^N \widehat{\gamma}_i (y_i - \widehat{\mu}_2)^2}{\sum_{i=1}^N \widehat{\gamma}_i}, \quad \hat{\sigma}_2^2 = \frac{\sum_{i=1}^N \widehat{\gamma}_i (y_i - \widehat{\mu}_2)^2}{\sum_{i=1}^N \widehat{\gamma}_i}, \quad \hat{\sigma}_2^2 = \frac{\sum_{i=1}^N \widehat{\gamma}_i (y_i - \widehat{\mu}_2)^2}{\sum_{i=1}^N \widehat{\gamma}_i}, \quad \hat{\sigma}_2^2 = \frac{\sum_{i=1}^N \widehat{\gamma}_i (y_i - \widehat{\mu}_2)^2}{\sum_{i=1}^N \widehat{\gamma}_i}, \quad \hat{\sigma}_2^2 = \frac{\sum_{i=1}^N \widehat{\gamma}_i (y_i - \widehat{\mu}_2)^2}{\sum_{i=1}^N \widehat{\gamma}_i}, \quad \hat{\sigma}_2^2 = \frac{\sum_{i=1}^N \widehat{\gamma}_i (y_i - \widehat{\mu}_2)^2}{\sum_{i=1}^N \widehat{\gamma}_i}, \quad \hat{\sigma}_2^2 = \frac{\sum_{i=1}^N \widehat{\gamma}_i (y_i - \widehat{\mu}_2)^2}{\sum_{i=1}^N \widehat{\gamma}_i}, \quad \hat{\sigma}_2^2 = \frac{\sum_{i=1}^N \widehat{\gamma}_i (y_i - \widehat{\mu}_2)^2}{\sum_{i=1}^N \widehat{\gamma}_i}, \quad \hat{\sigma}_2^2 = \frac{\sum_{i=1}^N \widehat{\gamma}_i (y_i - \widehat{\mu}_2)^2}{\sum_{i=1}^N \widehat{\gamma}_i}, \quad \hat{\sigma}_2^2 = \frac{\sum_{i=1}^N \widehat{\gamma}_i (y_i - \widehat{\mu}_2)^2}{\sum_{i=1}^N \widehat{\gamma}_i}, \quad \hat{\sigma}_2^2 = \frac{\sum_{i=1}^N \widehat{\gamma}_i (y_i - \widehat{\mu}_2)^2}{\sum_{i=1}^N \widehat{\gamma}_i}, \quad \hat{\sigma}_2^2 = \frac{\sum_{i=1}^N \widehat{\gamma}_i (y_i - \widehat{\mu}_2)^2}{\sum_{i=1}^N \widehat{\gamma}_i}, \quad \hat{\sigma}_2^2 = \frac{\sum_{i=1}^N \widehat{\gamma}_i (y_i - \widehat{\mu}_2)^2}{\sum_{i=1}^N \widehat{\gamma}_i}, \quad \hat{\sigma}_2^2 = \frac{\sum_{i=1}^N \widehat{\gamma}_i (y_i - \widehat{\mu}_2)^2}{\sum_{i=1}^N \widehat{\gamma}_i}, \quad \hat{\sigma}_2^2 = \frac{\sum_{i=1}^N \widehat{\gamma}_i (y_i - \widehat{\mu}_2)^2}{\sum_{i=1}^N \widehat{\gamma}_i}, \quad \hat{\sigma}_2^2 = \frac{\sum_{i=1}^N \widehat{\gamma}_i (y_i - \widehat{\mu}_2)^2}{\sum_{i=1}^N \widehat{\gamma}_i}, \quad \hat{\sigma}_2^2 = \frac{\sum_{i=1}^N \widehat{\gamma}_i (y_i - \widehat{\mu}_$$

4. Iterate steps 2 and 3 until convergence.

MCMC for Sampling from the Posterior:

One would like to draw samples from the resulting posterior distribution, in order to make inferences about the parameter. **Gibbs sampling** is an MCMC procedure.

- 1. Take some initial values $\theta^{(0)} = (\mu_1^{(0)}, \mu_2^{(0)})$.
- 2. Repeat for t = 1, 2, ..., n

(a) For i = 1, 2,..., N generate
$$\Delta_i^{(t)} \in \{0,1\}$$
 with $\Pr(\Delta_i^{(t)} = 1) = \hat{\gamma}_i(\theta^{(t)})$

(b) Set
$$\hat{\mu}_1 = \frac{\sum_{i=1}^{N} \left(1 - \Delta_i^{(t)}\right) \cdot y_i}{\sum_{i=1}^{N} \left(1 - \Delta_i^{(t)}\right)}$$
 and $\hat{\mu}_2 = \frac{\sum_{i=1}^{N} \Delta_i^{(t)} \cdot y_i}{\sum_{i=1}^{N} \Delta_i^{(t)}}$ and generate $\mu_1^{(t)} \sim N(\hat{\mu}_1, \hat{\sigma}_1^2)$ and $\mu_2^{(t)} \sim N(\hat{\mu}_2, \hat{\sigma}_2^2)$

3. Continue step 2 until the joint distribution of $(\Delta^{(t)}, \mu_1^{(t)}, \mu_2^{(t)})$ doesn't change.