

Name: Imisi Raphael Aiyetan
Course: Econometrics 512

Problem Set 5

Solution

Problem 1

Question (a):

```
. * The regular robust standard errors
.
. xtreg lrpdi lsales, fe vce(robust)
```

```
Fixed-effects (within) regression      Number of obs   =      28
Group variable: id                    Number of groups =       7

R-sq:  within = 0.0070                  Obs per group: min =       3
      between = 0.4037                      avg =      4.0
      overall  = 0.0449                      max =       5

                                         F(1,6)          =      0.40
corr(u_i, Xb)  = 0.1999                  Prob > F         =     0.5510
```

(Std. Err. adjusted for 7 clusters in id)

lrpdi	Coef.	Robust Std. Err.	t	P> t	[95% Conf. Interval]	
lsales	.0042463	.0067238	0.63	0.551	-.0122061	.0206988
_cons	4.512929	.0306475	147.25	0.000	4.437938	4.587921
sigma_u	.06705105					
sigma_e	.0135427					
rho	.96080464	(fraction of variance due to u_i)				

```

. * Clustering by i, the of cross-sectional observation
.
. xtreg lrpdi lsales, fe vce(cluster id)

Fixed-effects (within) regression              Number of obs   =        28
Group variable: id                          Number of groups =         7

R-sq:  within = 0.0070                      Obs per group: min =         3
        between = 0.4037                      avg =        4.0
        overall = 0.0449                      max =         5

                                         F(1,6)          =        0.40
corr(u_i, Xb) = 0.1999                     Prob > F         =    0.5510

```

(Std. Err. adjusted for 7 clusters in id)

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sigma_u	.06705105					
sigma_e	.0135427					
rho	.96080464	(fraction of variance due to u_i)				

Interpretation:

Ignoring in fixed effect, we realized that regression for robust standard and when we grouped individuals in cluster the result of the two analyses are the same. By implication, when we account for unknown structure of variation and heteroskedasticity/differences across clusters of observation the results are the same However, the estimated parameter for *lsales* is insignificant for both methods. We present the STATA output in Appendix 1 and 2. Also, the programming codes are shown in Appendix 13.

Question (b):

```
. egen lsale_bar = mean(lsales)
```

```
.  
. egen lrpdi_bar =mean(lrpdi)
```

```
.  
. gen z = lsales - lsale_bar
```

```
.  
. gen q = lrpdi - lrpdi_bar
```

```
.  
. reg z q
```

Source	SS	df	MS	Number of obs =	28
Model	.071949325	1	.071949325	F(1, 26) =	1.22
Residual	1.52932909	26	.05882035	Prob > F =	0.2789
				R-squared =	0.0449
				Adj R-squared =	0.0082
Total	1.60127842	27	.059306608	Root MSE =	.24253

z	Coef.	Std. Err.	t	P> t	[95% Conf. Interval]	
q	.7995552	.7229344	1.11	0.279	-.6864578	2.285568
_cons	6.82e-08	.0458337	0.00	1.000	-.0942124	.0942126

```
. egen l_sale_bar = mean(lsales), by(id)
```

```
. egen l_rpdidi_bar =mean(lrpdi), by(id)
```

```
. gen x = lsales - l_sale_bar
```

```
. gen y = lrpdi - l_rpdidi_bar
```

```
. reg y x
```

Source	SS	df	MS	Number of obs = 28		
Model	.000026032	1	.000026032	F(1, 26) = 0.18		
Residual	.003668096	26	.000141081	Prob > F = 0.6711		
				R-squared = 0.0070		
				Adj R-squared = -0.0311		
Total	.003694129	27	.00013682	Root MSE = .01188		

y	Coef.	Std. Err.	t	P> t	[95% Conf. Interval]	
x	.0042463	.0098853	0.43	0.671	-.0160732	.0245659
_cons	-1.54e-07	.0022447	-0.00	1.000	-.0046142	.0046139

Interpretation:

Estimating the fixed effect model using demean approach, we find that the estimated parameter of x and q are different. Although the results for both approaches are insignificant, but it suggests that when the fixed effect model is estimated and we account for difference/heteroskedasticity across group, the results are different. We present the STATA output in Appendix 2 and 3. Also, the programming codes are shown in Appendix 13

Question (C):

```
. reg lrpdi lsales d2 d3 d4 d5 d6 d7
```

Source	SS	df	MS	Number of obs = 28		
Model	.108981639	7	.015568806	F(7, 20) = 87.36		
Residual	.003564294	20	.000178215	Prob > F = 0.0000		
				R-squared = 0.9683		
				Adj R-squared = 0.9572		
Total	.112545933	27	.004168368	Root MSE = .01335		

lrpdi	Coef.	Std. Err.	t	P> t	[95% Conf. Interval]	
lsales	.0050718	.0109675	0.46	0.649	-.017806	.0279495
d2	.0241732	.0095558	2.53	0.020	.0042403	.0441062
d3	.0675644	.0095742	7.06	0.000	.047593	.0875359
d4	.1227823	.0097271	12.62	0.000	.1024918	.1430727
d5	.1309612	.0097047	13.49	0.000	.1107177	.1512048
d6	.1532011	.0096012	15.96	0.000	.1331732	.1732289
d7	.177454	.0095976	18.49	0.000	.1574336	.1974743
_cons	4.412576	.0488529	90.32	0.000	4.310671	4.514481

```
. reg lrpdi lsales i.id
```

Source	SS	df	MS	Number of obs = 28		
Model	.108877836	7	.015553977	F(7, 20) = 84.81		
Residual	.003668096	20	.000183405	Prob > F = 0.0000		
Total	.112545933	27	.004168368	R-squared = 0.9674		
				Adj R-squared = 0.9560		
				Root MSE = .01354		

lrpdi	Coef.	Std. Err.	t	P> t	[95% Conf. Interval]	
lsales	.0042463	.011271	0.38	0.710	-.0192646	.0277572
id						
2	.024285	.009697	2.50	0.021	.0040575	.0445125
3	.0676848	.0097162	6.97	0.000	.0474172	.0879524
4	.1245132	.0104325	11.94	0.000	.1027514	.146275
5	.1285638	.0095564	13.45	0.000	.1086295	.1484982
6	.1533331	.0097443	15.74	0.000	.1330068	.1736594
7	.1775845	.0097406	18.23	0.000	.157266	.1979029
_cons	4.416218	.0501929	87.98	0.000	4.311518	4.520919

Interpretation:

The most interesting part of these analyses is that when we account for individual and group effects, lsales has a significant impact on lrpdi. Comparing the two methods, we realize that the estimated parameters of the indicator variables are similar, and they are all significant at 5% level of significance. Therefore, estimating the fixed effect model with individual dummy variables and in clusters, the results are similar and significant. We present the STATA output in Appendix 3 and 4. Also, the programming codes are shown in Appendix 13

How do all of these results compare?

Using different approaches to estimate the fixed effect model, we realize that they generate the same results based on the estimated parameter. However, the most efficient method in this case is the fixed effect regression when we account for individual dummy variables or differences across groups. Note that these results could be as a result of the data we considered in this analysis. Both demean and robust standard errors approach give insignificant results.

Problem 2

Question a(i):

```
egen E = max(grant), by(fcode)
```



```
. mean (hrsemp) if year==1987 & E==0
```

Mean estimation Number of obs = **98**

	Mean	Std. Err.	[95% Conf. Interval]	
hrsemp	9.296744	1.663396	5.995363	12.59813

```
. mean (hrsemp) if year==1988 & E==0
```

Mean estimation Number of obs = **96**

	Mean	Std. Err.	[95% Conf. Interval]	
hrsemp	9.671083	1.855213	5.98802	13.35415

```
. mean (hrsemp) if year==1987 & E==1
```

Mean estimation Number of obs = **31**

	Mean	Std. Err.	[95% Conf. Interval]	
hrsemp	7.591087	3.650933	.1348872	15.04729

```
. mean (hrsemp) if year==1988 & E==1
```

Mean estimation Number of obs = **31**

	Mean	Std. Err.	[95% Conf. Interval]	
hrsemp	35.97834	6.637497	22.42276	49.53392

. ttest hrsemp, by(E)

Two-sample t test with equal variances

Group	Obs	Mean	Std. Err.	Std. Dev.	[95% Conf. Interval]	
0	194	9.481984	1.241371	17.2903	7.033589	11.93038
1	62	21.78471	4.172989	32.85815	13.44031	30.12912
combined	256	12.46155	1.414802	22.63684	9.675366	15.24774
diff		-12.30273	3.217664		-18.63943	-5.966032

diff = mean(0) - mean(1) t = -3.8235
Ho: diff = 0 degrees of freedom = 254

Ha: diff < 0 Ha: diff != 0 Ha: diff > 0
Pr(T < t) = 0.0001 Pr(|T| > |t|) = 0.0002 Pr(T > t) = 0.9999

. ttest hrsemp if year==1987, by(E)

Two-sample t test with equal variances

Group	Obs	Mean	Std. Err.	Std. Dev.	[95% Conf. Interval]	
0	98	9.296744	1.663396	16.46678	5.995363	12.59813
1	31	7.591087	3.650933	20.32753	.1348872	15.04729
combined	129	8.886858	1.532252	17.40304	5.855035	11.91868
diff		1.705658	3.597038		-5.412231	8.823546

diff = mean(0) - mean(1) t = 0.4742
Ho: diff = 0 degrees of freedom = 127

Ha: diff < 0 Ha: diff != 0 Ha: diff > 0
Pr(T < t) = 0.6819 Pr(|T| > |t|) = 0.6362 Pr(T > t) = 0.3181

```
. ttest hrsemp if year==1988, by(E)
```

Two-sample t test with equal variances

Group	Obs	Mean	Std. Err.	Std. Dev.	[95% Conf. Interval]	
0	96	9.671083	1.855213	18.1773	5.98802	13.35415
1	31	35.97834	6.637497	36.95602	22.42276	49.53392
combined	127	16.09254	2.352765	26.51432	11.43649	20.74859
diff		-26.30726	4.970328		-36.14415	-16.47036

diff = mean(0) - mean(1) t = -5.2929
Ho: diff = 0 degrees of freedom = 125

Ha: diff < 0 Ha: diff != 0 Ha: diff > 0
Pr(T < t) = 0.0000 Pr(|T| > |t|) = 0.0000 Pr(T > t) = 1.0000

Interpretation:

It is evident from the result that when considering E=0 for 1987 and 1988 the mean values are quite close. However, for E=1 the mean values are far from each other. This effect could be as a result different number of observations. Similarly, for the diff-in-diff estimation, we find that when considering 1987 the result is insignificant based on the t-test. However, for other estimations, the results are significant at 5% level of significance.

Question a(ii):

```
. reg hrsemp grant d88 E
```

Source	SS	df	MS	Number of obs =	256
Model	19608.6794	3	6536.22646	F(3, 252) =	14.83
Residual	111060.074	252	440.714578	Prob > F =	0.0000
				R-squared =	0.1501
				Adj R-squared =	0.1399
Total	130668.753	255	512.426483	Root MSE =	20.993

hrsemp	Coef.	Std. Err.	t	P> t	[95% Conf. Interval]	
grant	28.01292	6.125444	4.57	0.000	15.94933	40.0765
d88	.374339	3.014608	0.12	0.901	-5.562698	6.311376
E	-1.705658	4.325932	-0.39	0.694	-10.22524	6.813929
_cons	9.296744	2.120634	4.38	0.000	5.120321	13.47317

Interpretation:

The result shows that grant has significant positive impact on hrsemp. Other variables in this model are insignificant to hrsemp. We present the STATA output in Appendix 5, 6, and 7. Also, the programming codes are shown in Appendix 14

Question a(iii):

```
. xtreg hrsemp grant d88, fe
```

```
Fixed-effects (within) regression      Number of obs   =    256
Group variable: fcode                 Number of groups =    131

R-sq:  within = 0.4711                Obs per group:  min =     1
      between = 0.0692                      avg =    2.0
      overall  = 0.1495                      max =     2

corr(u_i, Xb) = -0.0235                F(2,123)        =   54.77
                                         Prob > F         =   0.0000
```

hrsemp	Coef.	Std. Err.	t	P> t	[95% Conf. Interval]	
grant	27.87793	3.129216	8.91	0.000	21.68384	34.07202
d88	.5093233	1.558337	0.33	0.744	-2.57531	3.593956
_cons	8.833036	.9462108	9.34	0.000	6.96007	10.706
sigma_u	19.405599					
sigma_e	10.683421					
rho	.76740875	(fraction of variance due to u_i)				

```
F test that all u_i=0:      F(130, 123) =    6.55      Prob > F = 0.0000
```

Interpretation:

The result shows that when estimating fixed effect model grant has significant positive impact on hrsemp. Other variables in the model are insignificant to hrsemp. We present the STATA output in Appendix 7. Also, the programming codes are shown in Appendix 14

Do you get the answer in each of these cases? Why or why not?

The results show that when we account for dummy variable for being a firm that receives treatment the estimated parameters of grant are close to the one in the fixed effect regression and they both significant. Generally, the results from two approaches are the same. The reason for this result is that the fixed effect model is exactly the same with the regression model since q_i and E_i are considered to be dummy variables in the two model.

Question 2b(i):

```
. xtreg hrsemp t, fe
```

```
Fixed-effects (within) regression      Number of obs   =      390
Group variable: fcode                  Number of groups =      135

R-sq:  within = 0.0872                  Obs per group:  min =      1
        between = 0.0117                  avg =      2.9
        overall = 0.0301                  max =      3

corr(u_i, Xb) = -0.0164                  F(1,254)         =      24.27
                                          Prob > F          =      0.0000
```

hrsemp	Coef.	Std. Err.	t	P> t	[95% Conf. Interval]	
t	5.825537	1.182386	4.93	0.000	3.497008	8.154066
_cons	3.241778	2.566541	1.26	0.208	-1.812633	8.296189
sigma_u	20.071123					
sigma_e	18.973517					
rho	.52808937	(fraction of variance due to u_i)				

```
F test that all u_i=0:      F(134, 254) =      3.27      Prob > F = 0.0000
```

```
.
. predict hrsemp_res
(option xb assumed; fitted values)
```

```
. xtreg grant t, fe
```

```
Fixed-effects (within) regression      Number of obs   =      390
Group variable: fcode                  Number of groups =      135

R-sq:  within = 0.0774                  Obs per group:  min =      1
        between = 0.0253                  avg =      2.9
        overall = 0.0564                  max =      3

corr(u_i, Xb) = -0.0251                  F(1,254)         =      21.31
                                          Prob > F          =      0.0000
```

grant	Coef.	Std. Err.	t	P> t	[95% Conf. Interval]	
t	.1087379	.0235548	4.62	0.000	.0623502	.1551255
_cons	-.0675878	.0511292	-1.32	0.187	-.1682789	.0331034
sigma_u	.17032656					
sigma_e	.37797988					
rho	.16878723	(fraction of variance due to u_i)				

```
F test that all u_i=0:      F(134, 254) =      0.57      Prob > F = 0.9998
```

```
.
. predict grant_res
(option xb assumed; fitted values)
```

. xtreg d88 t, fe

```

Fixed-effects (within) regression              Number of obs   =       390
Group variable: fcode                        Number of groups =       135

R-sq:  within = 0.0000                      Obs per group:  min =        1
          between = 0.3217                      avg =       2.9
          overall = 0.0001                      max =        3

                                          F(1,254)        =       0.00
corr(u_i, Xb) = 0.0568                      Prob > F         =     0.9569

```

d88	Coef.	Std. Err.	t	P> t	[95% Conf. Interval]	
t	-.0019417	.0358725	-0.05	0.957	-.0725872	.0687037
_cons	.3295494	.0778666	4.23	0.000	.1762031	.4828957
sigma_u	.08343789					
sigma_e	.57563965					
rho	.02057762	(fraction of variance due to u_i)				

F test that all u_i=0: F(134, 254) = 0.03 Prob > F = 1.0000

```

.
. predict d88_res
(option xb assumed; fitted values)
. xtreg E t, fe
note: t omitted because of collinearity

```

```

Fixed-effects (within) regression              Number of obs   =       390
Group variable: fcode                        Number of groups =       135

R-sq:  within = .                          Obs per group:  min =        1
          between = .                      avg =       2.9
          overall = .                      max =        3

                                          F(1,254)        =       .
corr(u_i, Xb) = .                          Prob > F         =       .

```

E	Coef.	Std. Err.	t	P> t	[95% Conf. Interval]	
t	0 (omitted)					
_cons	.474359
sigma_u	.50074571					
sigma_e	0					
rho	1	(fraction of variance due to u_i)				

F test that all u_i=0: F(134, 254) = . Prob > F = .

```

.
. predict E_res
(option xb assumed; fitted values)

```

Question 2b(ii)

```
. reg hrsemp_res grant_res d88_res E_res
note: d88_res omitted because of collinearity
note: E_res omitted because of collinearity
```

Source	SS	df	MS	Number of obs =	390
Model	8923.22499	1	8923.22499	F(1, 388) =	.
Residual	0	388	0	Prob > F =	.
Total	8923.22499	389	22.9388817	R-squared =	1.0000
				Adj R-squared =	1.0000
				Root MSE =	0

hrsemp_res	Coef.	Std. Err.	t	P> t	[95% Conf. Interval]
grant_res	53.57414
d88_res	0 (omitted)				
E_res	0 (omitted)				
_cons	6.862733

Interpretation:

Generally, the result shows that when we considering firm-specific variables and trend for the three years, the most important residual is the one coming from grant. However, other residuals are not important to the firm. In conclusion, estimating fixed effect model using this approach is inefficient. We present the STATA output in Appendix 8,9,10,11, and 12. Also, the programming codes are shown in Appendix 14.