

1. The distribution of weights (in pounds) of male students at a college is approximately normal with an unknown mean μ and with an unknown standard deviation σ .

- (a) (1.5 pts) Assume that the average weight of a random sample of 20 male students in the college is 152.33 lb with a sample standard deviation of 15.24 lb. Determine a two-sided confidence interval for μ at 95% level of confidence using two-decimal places of rounding. Justify your steps including your choice of the interval estimation procedure (e.g. t-interval, z-interval etc.).

Answer. Since σ is unknown and a random sample from an approximate normal population is used, a t-interval procedure applies with $\bar{x} = 152.33$, $s = 15.24$, $1 - \alpha = 0.95$, $n = 20$ and $df = 19$. We also use $t_{19,0.025} =$ as the t-score. Then, the margin of error is $E = 2.093 \frac{15.24}{\sqrt{20}} \simeq 7.1325$ lb, and the interval is 152.33 ± 7.1325 which is approximately $[145.20, 159.46]$ lb, meaning we are 95% certain that the actual mean weight is between 145.20 lb and 159.46 lb.

- (b) (1 pt) Assume that the population standard deviation of the weights of the male students is 15 lb. Based on this, determine the minimum sample size n needed to ensure that the margin of error at 98% level of confidence is not larger than 5 lb. Use an integer end result with proper units.

Answer. If $\sigma = 15$ is given, we can use a z-interval approach with $z_{0.01} = 2.326$ (the level of confidence has increased here). Then, $n \simeq (2.326 \frac{15}{5})^2 = 48.692$ and so we need to use at least $n = 49$ male students in the sample.

2. In a random sample of 64 cars registered in a certain state, 11 of them were found to have emission levels that exceed the state standard. Let p denote the actual (unknown) proportion of all cars registered in that state whose emission levels exceed the state standards.

- (a) (1.5 pts) Obtain a lower-bound confidence interval (CI) for p at 95% confidence level (use three decimal places). Justify the procedure that you apply.

Answer. Since both $x = 11$ and $n - x = 53$ are larger than 10, the minimum criteria for a large-sample z-interval procedure applies (though the Agresti-Coull interval with $\tilde{p} = \frac{x+2}{n+4} = \frac{13}{68}$ is also an option) at level $1 - \alpha = 0.95$. With $\hat{p} = \frac{x}{n} = \frac{11}{64} = 0.17188$ and $z_{0.05} = 1.645$, the lower bound is

$$LB = \hat{p} - z_{0.05} \sqrt{\frac{\hat{p}(1-\hat{p})}{n}} = \frac{11}{64} - (1.645) \sqrt{\frac{\frac{11}{64} \frac{53}{64}}{64}} \simeq 0.0943 \simeq 0.094.$$

So, the lower-bound CI for p is $[0.094, 1)$.

- (b) (0.5 pts) Based on the interval above, can you reasonably conclude that p is larger than 0.1? Briefly explain.

Answer. The CI constructed above also covers proportions between 0.094 and 0.10 as plausible values of p . So, we can NOT conclude that p is larger than 0.1.

3. (1 pt) Let $X \sim \text{Bin}(16, p)$ where p is unknown. Consider the estimator $\hat{\theta} = \frac{X+2}{20}$ for p . Determine $\text{Bias}(\hat{\theta}) = p - E[\hat{\theta}]$, $\text{Var}(\hat{\theta})$ and $\text{MSE}(\hat{\theta})$ in terms of the values of p .

Answer. Using the mean and variance of $\text{Bin}(n = 16, p)$ distribution as well as the properties of linear functions, we get

$$\begin{aligned} E[\hat{\theta}] &= \frac{E[X] + 2}{20} = \frac{16p + 2}{20} = \frac{4}{5}p + \frac{1}{10} \text{ and} \\ \text{Var}(\hat{\theta}) &= \frac{\text{Var}(X + 2)}{20^2} = \frac{\text{Var}(X)}{400} = \frac{16}{400}p(1-p) = \frac{p(1-p)}{25}. \end{aligned}$$

Then, $\text{Bias}(\hat{\theta}) = p - (\frac{4}{5}p + \frac{1}{10}) = \frac{1}{5}p - \frac{1}{10}$ or $\frac{2p-1}{10}$ (which is 0 if $p = 0.5$) and

$$\text{MSE}(\hat{\theta}) = (\frac{2p-1}{10})^2 + \frac{p(1-p)}{25} = \frac{1}{100}.$$