Penalized Regression



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MSAN 601: Linear Regression

Plan for this Lecture



- Model Selection
- Regularization in Linear Regression
- Algorithms:
 - Ridge Regression (Tikhanov Regularization)
 - The Lasso
 - Elastic Net

Back to Linear Regression



Model:

$$y_i = \beta_0 + \sum_{j=1}^p \beta_j x_{ij} + \epsilon_i$$

Estimate:

$$\hat{\beta}_{OLS} = \operatorname{argmin}_{\beta} \left(\sum_{i=1}^{n} \left(y_i - \beta_0 - \sum_{j=1}^{p} \beta_j x_{ij} \right)^2 \right)$$

Questions:

- What if we are primarily concerned with variable selection?
- What if p > n? (high dimensional regression)

Best Subset Selection



Algorithm

Given: k predictors $\{\mathbf{x}_1, \dots, \mathbf{x}_p\}$

Loop: for (*k* in 1 to p)

- Fit all $\binom{p}{k}$ models that contain k predictors
- 2 Pick M_k = the "best" among these models

Return: $M^* \in \{M_0, \dots, M_p\}$: $M^* = \operatorname{argmin}_j(S(M_j))$

- $S(M_j)$ = prediction criterion (Mallow's C_p , AIC, BIC, MSPE)

Best Subset Selection



Important Considerations:

- Computational Complexity: must fit 2^p models
- Algorithm is exhaustive: we will find the "best" model
- Often replaced with approximate and less intensive algorithms:
 - Forward stepwise selection
 - Backward stepwise selection
 - Forward-backward stepwise selection

A Nice Alternative: Shrinkage Methods



- Fits all p predictors using a technique that constrains or regularizes the coefficient estimates by optimizing a slightly different objective function
- Equivalently, the techniques shrink coefficient estimates to zero
- Variance of coefficient estimates are reduced as well! Particularly in high dimensional settings!

Ridge Regression



Model:

$$y_i = \beta_0 + \sum_{j=1}^p \beta_j x_{ij} + \epsilon_i$$

Estimate:

$$\hat{\beta}_{Ridge} = \operatorname{argmin}_{\beta} \left(\sum_{i=1}^{n} \left(y_i - \beta_0 - \sum_{j=1}^{p} \beta_j x_{ij} \right)^2 + \lambda \sum_{j=1}^{p} \beta_j^2 \right)$$

 $\lambda \sum_{j=1}^{p} \beta_{j}^{2}$ acts as a shrinkage penalty to standard least squares regression since this value is small when β_{j}^{2} is small.

Note: Also known as Tikhonov regularization



Ridge Regression: regularization perspective



Problem: The variance of \hat{f} for OLS is often high \Leftrightarrow predictions significantly change with small changes in X.

Reason: X^TX is ill-conditioned \Leftrightarrow either $p \approx n$ or variables suffer from multicollinearity:

$$\hat{\beta}_{OLS} = (X^T X)^{-1} X^T y$$

Solution: Ridge regression regularizes (X^TX) :

$$\hat{\beta}_{Ridge} = (X^T X + \lambda I)^{-1} X^T y$$

Ridge Regression



$$\hat{\beta}_{Ridge} = \operatorname{argmin}_{\beta} \left(\sum_{i=1}^{n} \left(y_i - \beta_0 - \sum_{j=1}^{p} \beta_j x_{ij} \right)^2 + \lambda \sum_{j=1}^{p} \beta_j^2 \right)$$

 λ : tuning parameter that adjusts the effect of the penalty

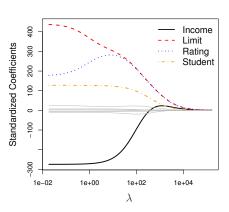
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$$\lambda = 0$$
 \Rightarrow $\hat{\beta}_{OLS} = \hat{\beta}_{Ridge}$

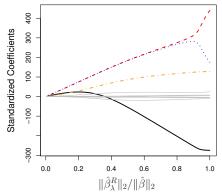
$$\bullet \ \lambda \to \infty \qquad \Rightarrow \qquad \hat{\beta}_{Ridge} \to 0$$

ullet λ chosen using cross validation: amazingly can computationally be determined for all possible values simultaneously!

Example: Comparison of $\hat{\beta}_{OLS}$ and $\hat{\beta}_{Ridge}$







When does Ridge Regression outperform OLS?



- **Omputationally**: Ridge estimates for all values of λ can be determined simultaneously with one fit. Significant advantage over best subset selection that requires 2^p least squares fits.
- **Model Accuracy**: OLS estimates often have high variance but low bias. Increases in λ lead to shrinkage, which subsequently leads to a major decrease in variance and only a slight increase in bias.
- **3** Key is to look across a grid of λ for best MSPE.

When does Ridge Regression outperform OLS?



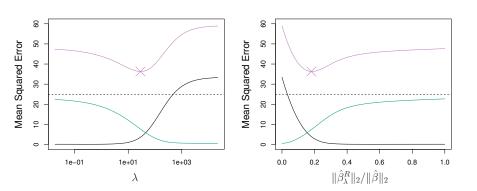


Figure: Squared bias (black), variance (green), and MSPE (purple) for $\hat{\beta}_{Ridge}$ on a simulated data set.

Weaknesses of Ridge Regression



- Requires a user specified tuning parameter λ
- Interpretability of $\hat{\beta}_{Ridge}$
- Subtle but important point: The penalty $\lambda \sum_{j=1}^{p} \beta_j^2$ shrinks β towards 0 but does not set any values exactly to 0.
 - **Exception**: $\lambda = \infty$ here all β_j are exactly 0
 - Consequence: The saturated model is *always* chosen!

Question: Can we shrink some coefficients exactly to zero?



Least absolute shrinkage and selection operator

Model:

$$y_i = \beta_0 + \sum_{j=1}^p \beta_j x_{ij} + \epsilon_i$$

Estimate:

$$\hat{\beta}_{Lasso} = \operatorname{argmin}_{\beta} \left(\sum_{i=1}^{n} \left(y_i - \beta_0 - \sum_{j=1}^{p} \beta_j x_{ij} \right)^2 + \lambda \sum_{j=1}^{p} |\beta_j| \right)$$

 $\lambda \sum_{j=1}^{p} |\beta_j|$ acts as a shrinkage penalty to standard least squares regression since this value is small when $|\beta_j|$ is small.

Historical Note on the Lasso

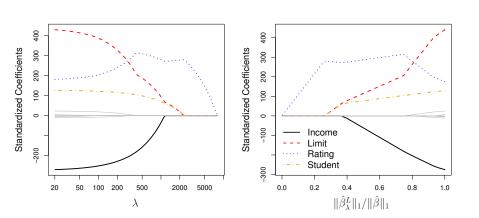


- From the paper "Regression shrinkage via the lasso" (1996) in Journal of the Royal Statistical Society. Series B by Robert Tibshirani (one of the authors of ISL and ESL)
- Considered by many to be the most influential modern statistical method
- Paper currently has 14243 citations! (as of October 27, 2015)
- Website:

http://statweb.stanford.edu/ tibs/lasso.html

Variable Selection Property of Lasso





Note: Changing λ sets various subsets of β to 0! Why?



Re-formulations of Ridge and the Lasso



Both methods can be viewed as optimization problems.

• Ridge Regression:

$$\operatorname{minimize}_{\beta}\left(\sum_{i=1}^{n}\left(y_{i}-\beta_{0}-\sum_{j=1}^{p}\beta_{j}x_{ij}\right)^{2}\right) \qquad \text{subject to} \qquad \sum_{j=1}^{p}\beta_{j}^{2} \leq s$$

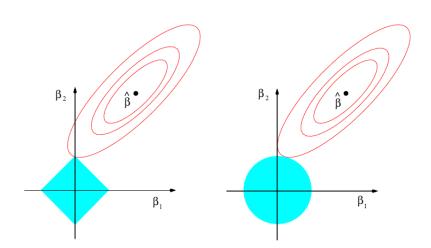
Lasso:

$$\text{minimize}_{\beta} \left(\sum_{i=1}^{n} \left(y_i - \beta_0 - \sum_{j=1}^{p} \beta_j x_{ij} \right)^2 \right) \qquad \text{subject to} \qquad \sum_{j=1}^{p} |\beta_j| \le s$$

Uh, ok so what? Explains the variable selection property of the Lasso!

Comparison of Lasso and Ridge





Often, the Lasso shrinks coefficients exactly to zero!

Comparison of Lasso and Ridge



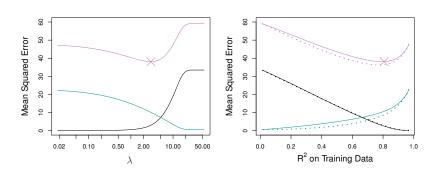


Figure: Squared bias (black), variance (green), and MSPE (purple). Dashed = Ridge, solid = Lasso

Note: Simulated data here included 45 / 45 non-zero coefficients. So, *no* variable selection is needed.

Comparison of Lasso and Ridge



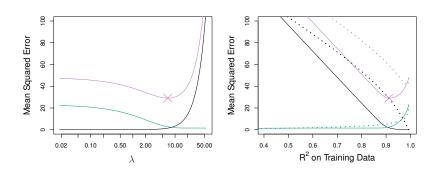


Figure: Squared bias (black), variance (green), and MSPE (purple). Dashed = Ridge, solid = Lasso

Note: Simulated data here included 2 / 45 non-zero coefficients. So, variable selection *is* needed.

A Bayesian Perspective of Ridge and Lasso



Let $\beta = (\beta_0, \beta_1, \dots, \beta_p)$ be the parameters in a linear regression.

Bayesian Framework: Assume that β is a *random vector* with distribution $p(\beta)$. Here,

- $f(y|X,\beta)$ = likelihood of the data (Gaussian if ϵ is Gaussian)
- $p(\beta)$ = prior distribution of β
- $p(\beta|X,y)$ =posterior distrubtion of β given (X,y)

Bayes' Theorem gives us:

$$p(\beta|X,y) \propto f(y|X,\beta)p(\beta)$$

A Bayesian Perspective of Ridge and Lasso



Assumption 1: $p(\beta) = \prod_{i=1}^{p} g(\beta_i)$ (i.e. β_i 's are iid).

Under Assumption 1, the regression model becomes:

$$y = \beta_0 + X_1 \beta_1 + \ldots + X_p \beta_p + \epsilon$$

$$\beta_i \stackrel{iid}{\sim} g(x)$$

A Bayesian Perspective



Properties

• If g(x) is a Gaussian distribution with $\mu = 0$, and $\sigma^2 = h(\lambda)$ then

$$\hat{\beta}_{Ridge} = \mathsf{Mode}(p(\beta|X,y))$$

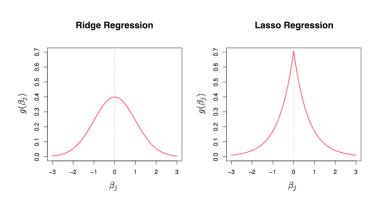
② If g(x) is a Double Exponential distribution with $\mu = 0$, and $\sigma^2 = h(\lambda)$ then

$$\hat{\beta}_{Lasso} = \mathsf{Mode}(p(\beta|X,y))$$

That is, by assuming a certain form of g(x), we find that the Ridge and Lasso estimates are the maximum a posteriori (MAP) estimators for β .

A Bayesian Perspective





Another way of understanding the likelihood of shrinkage!

Elastic Net



In general, the Lasso is best for variable selection / sparse relationships; Ridge for ill-conditioned problems.

Elastic Net: Combines Lasso and Ridge

Model:

$$y_i = \beta_0 + \sum_{j=1}^p \beta_j X_{ij} + \epsilon_i$$

Estimate:

$$\hat{\beta}_{EN} = \operatorname{argmin}_{\beta} \left(\sum_{i=1}^{n} \left(y_i - \beta_0 - \sum_{j=1}^{p} \beta_j x_{ij} \right)^2 + \alpha \lambda \sum_{j=1}^{p} |\beta_j| + \frac{1 - \alpha}{2} \lambda \sum_{j=1}^{p} \beta_j^2 \right)$$

Best of both worlds? – well, this is more difficult to interpret!

Elastic Net Properties



- α = 0: reduces to Ridge Regression
- α = 1: reduces to Lasso
- Has both properties of Ridge and Lasso:
 - Reduces variance
 - Variable selection
- Recently proven that Elastic Net is equivalent to linear support vector machines.

Selecting λ



General Method: Grid search and cross-validation

- Fix a value of λ
- Estimate model and calculate average MSPE from k-fold cross-validation
- **3** Repeat the above procedure across a grid of λ
- **1** Choose λ that leads to smallest MSPE

Important: The above procedure can be done in parallel, easing computation.

Selecting λ

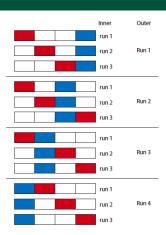


Subtle yet Important Point:

- Contrary to the fitting of a model in standard linear regression which relies upon minimizing MSE in training data only, we choose λ using cross-validation, which relies upon minimizing the MSPE of (cross-validation) test sets. Because of this, we typically hold out a test set initially and then run cross-validation on the training data.
- ② Once λ is chosen, we then evaluate the MSPE on the original held-out test data.

Selection λ

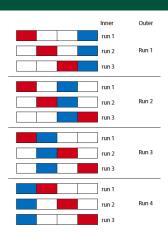




1) First choose a λ from **Validation set**: red (test) + white (training). Here, we will get 4 of them $\lambda_1, \lambda_2, \lambda_3, \lambda_4$

Selection λ

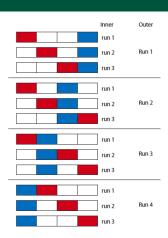




2) Then choose the best λ_i by i) training on **Validation set** and ii) testing on **Held out set**.

Selection λ

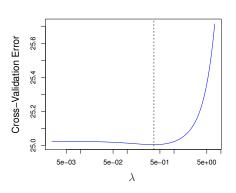


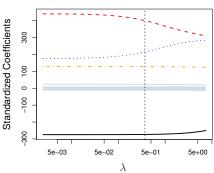


3) Fit final model using entire data set (**Validation** + **Held out**) with best λ_i

Selecting λ : Example with Ridge

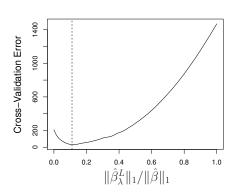


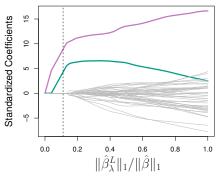




Selecting λ : Example with Lasso







Implementing Shrinkage Methods with R



Now we show how to implement the Lasso, Ridge Regression and Elastic Net in R.