# MATH-505A: Homework # 3

Due on Friday, September 12, 2014

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# Exercise # 1.7

(1)

Given: Two roads  $r1_{AB}$ ,  $r2_{AB}$  connecting points A and B and  $s1_{BC}$ ,  $s2_{BC}$  connecting B and C.

Let p(AB) denote the probability that path between  $A \longrightarrow B$  is open and let  $p(AB^c)$  denote the probability that there is(are) no path open b/w A and B.

To find: $Y = P(AB|AC^c)$ .

Y is equal to the probability that path between A and B is open AND(given that) the path between A and C is closed  $\implies$  Path between B and C is closed AND between A and B is open

p(AB) = Path b/w A,B is open = 1 - Path b/w A,B is closed = 1 - p \* p

 $p(AC^C|AB) = \text{Path b/w A,C}$  is closed given A,B is open = Path b/w B,C is closed given A,B is open. =  $p(BC^C)$ 

Thus

$$p(AB) = 1 - p^2 \tag{1}$$

Also,

$$p(AB) = p(BC) \tag{2}$$

 $p(AC^C) = 1$  - Probability A,C is open = 1 - Probability AB is open AND BC is open. Thus,

$$p(AC^c) = 1 - p(AB)p(AC) = 1 - (1 - p^2)^2$$
(3)

$$p(AB \cap AC^C) = p(AC^C|AB)p(AB) = p(BC^C)p(AB) = p^2(1-p^2)$$
(4)

$$p(AB|AC^c) = \frac{P(AB \cap AC^c)}{p(AC^c)} = \frac{p(AC^C|AB)p(AB)}{p(AC^c)} = \frac{p^2(1-p^2)}{1-(1-p^2)^2}$$
 (5)

**Part 2:** Additional direct road from A to C. Find  $p(AB|AC^c)$ :

 $p(AC^c|AB)$  = Probability that path b/w A,C is closed given path b/w A,B are open = Probability path b/w A,C(direct) are closed AND path b/w B,C are closed

$$p(AC^C|AB) = p * p(BC^c)p(AB)$$
(6)

where the extra p in 6 as compared to 4 is because the direct path A,C should be blocked too.

$$p(AC^c) = 1 - (p(1-p^2)^2 - (1-p))$$
(7)

where the second term is the probability that direct  $A \longrightarrow C$  is blocked, but  $A \longleftrightarrow B$  and  $B \longrightarrow$  is open(both). The third term represents the probability that  $A \longrightarrow C$  is open directly(does not matter if the other paths are open or blocked),

Thus, for part 2:

$$p(AB|AC^c) = \frac{p^3(1-p^2)}{1-(1-p^2)^2-(1-p)} = \frac{p^2(1-p^2)}{1-(1-p^2)^2}$$
(8)

(2)

$$p(2K \cap 1A) = \frac{\binom{4}{2} * \binom{4}{1} * \binom{52-4-4}{10}}{\binom{52}{13}} \tag{9}$$

 $p(1A|2K) = \frac{p(1A \cap 2K)}{p(2K)}$ 

$$p(2K) = \frac{\binom{4}{2} * \binom{52-4}{11}}{\binom{52}{12}} = \frac{6 * 48! * 13!}{52! * 11!} = \frac{6 * 12 * 13}{49 * 50 * 51 * 52} = 2.52 * 10^{-4}$$
 (10)

Thus,

$$p(1A|2K) = \frac{\binom{4}{2} * \binom{4}{1} * \binom{44}{10}}{\binom{4}{2} * \binom{18}{11}} = \frac{4*11*37!}{46*47*48*34!} = 0.0399$$

(4)

To prove/disprove:  $p(x|C) > p(y|C)ANDp(x|C^c) > p(y|C^c) \implies p(x) > p(y)$ 

$$p(x|C) - p(y|C) > 0 \tag{11}$$

$$p(x|C^c) - p(y|C^c) > 0$$
 (12)

$$p(x) = p(x|C)p(C) + p(x|C^{c})p(C^{c})$$
(13)

Also,

$$p(y) = p(y|C)p(C) + p(y|C^{c})p(C^{c})$$
(14)

Consider p(x) - p(y):

$$p(x) - p(y) = (p(x|C) - p(y|C))p(C) + (p(x|C^c) - p(y|C^c))p(C^c)$$
(15)

From the 12, ?? and 15:

$$p(x) - p(y) > 0 \forall x, y \tag{16}$$

Thus, x is always prefered over y.

(5)

Let  $X_i$  represent the  $i^{th}$  ccard draw

$$p(X_k = m | X_k > X_i) = \frac{p(X_k = m \cap X_k > X_i)}{p(X_k > X_i)} = \frac{p(X_k = m)}{p(X_k > X_i)} = \frac{\frac{1}{m}}{m}$$

Given:  $X_k > X_i$ ,  $\forall i \in [1, k-1] and k \in [1, m]$   $p(X_k = m | X_k > X_i) = \frac{p(X_k = m \cap X_k > X_i)}{p(X_k > X_i)} = \frac{p(X_k = m)}{p(X_k > X_i)} = \frac{\frac{1}{m}}{\frac{1}{k}}$  Where the equality in the last step comes from the fact that the probability of choosing cards such that  $p(X_k > X_i)$  is simply to choose the largest card, i.e. k among the rest i.

Thus 
$$p(X_k = m|X - k > X_i) = \frac{k}{m}$$
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Exercise # 1.8]

# 1

### (a)

Six turns up exactly once The dice on which 6 should appear is choosen in  $\binom{2}{1}$  way and has just one option(6) while the other has 6-1=5 options. Total outcomes =6\*6=36 Thus:

$$p = \frac{1*5}{36} = \frac{5}{36}$$

# (b)

**Both numbers are odd.** 3 out of 6 numbers are odd and the outcome of odd on both dice are independent. Thus,

$$p = \frac{3*3}{36} = \frac{1}{12}$$

# (c)

#### Sum of scores is 4.

Possible configurations: (1,3); (2,2); (3,1)

Thus:

$$p = \frac{3}{36} = \frac{1}{12}$$

# (d)

#### Sum of scores is divisible by 3.

Possible choices for sum to be divisible by 3:

3, 6, 9, 12:

3: (1,2), (2,1)

6: (1,5), (2,4), (3,3), (4,2), (5,1)

9: (3,6), (4,5), (5,4), (6,3)

12: (6,6)

Thus  $p = \frac{2+5+4+1}{36} = \frac{1}{3}$ 

# 2

# (a)

Head appears the first time on  $n^{th}$  throw. Thus a series of n-1 consecutive tails followed by head.  $p = \binom{n}{1} \frac{1}{2^{n-1}} * \frac{1}{2} = \frac{n}{2^n}$ 

(b)

The number of heads and dates to date are equal:

Case1: n is odd.

Clearly the probability of heads and tails being equal in this case is zero!

Case2: n is even.

A series of H,T such that |H| = |T|. Since the probability of occurrence of either a H or a T is  $\frac{1}{2}$ , the result will still be:

$$p = \binom{n}{n/2} \frac{1}{2^n}$$

(c)

Exactly two heads have appeared altogether.

$$p = \binom{n}{2} * \frac{1}{2^2} * \frac{1}{2^{n-2}} = \binom{n}{2} * \frac{1}{2^n}$$

(d)

At least two heads have appeared.

This is same as the one minus probability that 0 or 1 heads have appeared so far:

$$p = 1 - p(0 \ heads) - p(1 \ head \ only) \ p = 1 - \binom{n}{0} \frac{1}{2^0} \frac{1}{2^n} - \binom{n}{1} \frac{1}{2} \frac{1}{2^n - 1} = 1 - \frac{1}{2^n} - \frac{n}{2^n}$$

4

(a)

Biased coin tossed three times:

Sample space: HHH,HHT,HTH,HTT,THH,THT,TTH,TTT

(b)

Balls drawn without replacement from 2U,2V balls

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Sample space: (U1, V1), (U2, V1), (V1, U1), (V1, U2), (U1, V2), (U2, V2), (V2, U1), (V2, U2)
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(c)

Biased coin tossed till H tunrs up

Sample Space:

$$\{H\}, \{T, H\}, \{T, T, H\}, \{T, T, T, H\}, \dots, \{T, T, T, \dots, T, H\}$$

5

To find: Probablity that exactly one of A,B occurs

Probability that exactly one of A,B occurs is equal to the probability that either only A occurs or only **B** occurs. This is simply given by:

 $P(A \cup B) - P(A \cap B)$ . That is either of A,B occurs removing the portion when both A,B occur.

Thus, the probability that exactly one of A,B occurs is :

$$p(A \cup b) - p(A \cap B) = p(A) + p(B) - p(A \cap B) - p(A \cap B) = p(A) + p(B) - 2p(A \cap B)$$

6

$$RHS = 1 - P(A^C|B^C \cap C^C) \frac{P(B^C \cap C^C)}{p(C^C)} p(C^C)$$

Thus, expanding further  $p(A^C|B^C\cap C^C)=\frac{p(A^C\cap B^C\cap C^C)}{p(B^C\cap C^C)}$ 

$$RHS = 1 - \frac{p(A^C \cap B^C \cap C^C)}{p(B^C \cap C^C)} * \frac{P(B^C \cap C^C)}{p(C^C)} * p(C^C)$$

$$RHS = 1 - P(A^C \cap B^C \cap C^C) \text{ which is same as } LHS$$

$$RHS = 1 - P(A^C \cap B^C \cap C^C)$$
 which is same as  $LHS$ 

#### 14

Consider only two events:  $A_i$ .

From the definition of conditional probability:

$$p(A_j|B) = \frac{p(A_j \cap B)}{p(B)} \tag{17}$$

Also:

$$p(B|A_j) = \frac{p(A_j \cap B)}{p(A_j)} \tag{18}$$

Substituing for  $p(A_j \cap B)$  from 17 in 18:

$$p(B|A_j) = \frac{p(A_j|B)p(B)}{p(A_j)} \tag{19}$$

or equivalently,

$$p(A_j|B) = \frac{p(B|A_j)p(A_j)}{p(B)}$$
(20)

where p(B) can be expressed as(using the law of partition):

$$p(B) = p(B|A_1)P(A_1) + P(B|A_2)P(A_2) + \dots + P(B|A_n)P(A_n)$$
(21)

Thus  $p(B) = \sum_{i=1}^{n} p(B|A_i)p(A_i)$ 

Ans hence,

$$p(A_j|B) = \frac{p(B|A_j)p(A_j)}{\sum_{i=1}^{n} p(B|A_i)p(A_i)}$$