MATH-505A: Homework # 6

Due on Friday, October 3, 2014

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Contents

Exercise # 3.1	;
1	;
2(i)	
2 (iii)	
3	
5a	
5b	
5c	
Exercise # 3.2	,
2	
4	
Exercise # 3.3	,
1	
2	
5	

Exercise # 3.1

1

Part a:
$$f(x) = C2^{-x}$$

For f(x) to be a mass function $\sum_{1}^{\infty} C2^{-i} = 1$ $C(\frac{1}{2} + \frac{1}{2^{2}} + \frac{1}{2^{3}} +) = C\frac{1}{2} * \frac{1}{1-\frac{1}{2}} = 1 \implies C = 1$

Part b:
$$f(x) = \frac{C2^{-x}}{x} \sum_{1}^{\infty} \frac{C}{2^{i}i} = 1$$

Notice $ln(1+x) = x + \frac{x^2}{2} + \frac{x^3}{3} + \dots$

Hence
$$C \sum_{1}^{\infty} \frac{(\frac{1}{2})^{i}}{i} = C \ln(1 + 1/2) = 1 \implies C = \frac{1}{\ln 1.5}$$

Part c: $f(x) = Cx^{-2} \sum_{1}^{\infty} \frac{C}{x^{2}} = 1$

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$$f(x) = Cx^{-2} \sum_{1}^{\infty} \frac{C}{x^{2}} = 1$$

$$\textstyle\sum_{1}^{n} \frac{1}{i^2} =$$

Using taylor expansion of sinx and the fact that $\frac{sinx}{x}$ has roots at $x = \pi, 2\pi, 3\pi$:

$$\frac{\sin x}{x} = 1 - \frac{x^2}{3!} + \frac{x^4}{5!} + \dots = (1 - \frac{x}{\pi})(1 - \frac{x}{2\pi})(1 - \frac{x}{3\pi})\dots(1 - \frac{x}{\pi})$$

 $\frac{\sin x}{x} = 1 - \frac{x^2}{3!} + \frac{x^4}{5!} + \dots = (1 - \frac{x}{\pi})(1 - \frac{x}{2\pi})(1 - \frac{x}{3\pi})\dots(1 - \frac{x}{\pi})$ The product of productions of $\frac{\sin x}{x}$ is given the the coefficient of x^2 in the original and hence $-\frac{1}{3!} = \frac{1}{3!}$

Thus
$$\sum_{1}^{\infty} \frac{C}{x^2} = C * \frac{\pi^2}{6}$$
. Thus $C = \frac{6}{\pi^2}$

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Part d: $C2^x/x! \sum_{1}^{\infty} C2^i/i! = C \sum_{1}^{\infty} 2^i/i! = Ce^2 \implies C = \frac{1}{e^2}$

2(i)

Part a
$$P(X > 1) = \sum_{i=1}^{\infty} 2^{-i} = \frac{1}{4} * 2 = \frac{1}{2}$$

Part b
$$P(X > 1) = \sum_{2} 2^{2} = -\frac{1}{4} * 2 = \frac{1}{2}$$

Part b $P(X > 1) = \frac{1}{1.5} \sum_{2}^{\infty} \frac{(\frac{1}{2})^{i}}{i} = 1 - \ln(1.5)/2$
Part c $P(X > 1) = 1 - \frac{6}{\pi^{2}} 1^{-2} = 1 - \frac{6}{\pi^{2}}$
Part d $P(X > 1) = 1 - \frac{1}{e^{2}} 2 = 1 - \frac{2}{e^{2}}$

Part c
$$P(X > 1) = 1 - \frac{6}{\pi^2} 1^{-2} = 1 - \frac{6}{\pi^2}$$

Part d
$$P(X > 1) = 1 - \frac{1}{e^2} = 1 - \frac{2}{e^2}$$

2 (iii)

Probability that X is even = P(X = 2k) for k = 1, 2, 3...

Part a $P(X = 2k) = 2^{-2k}$ Summing up over all k: $P = \frac{1}{2^2} + \frac{1}{2^4} + \frac{1}{2^6} + \dots = \frac{1}{4} \frac{1}{1 - \frac{1}{4}} = \frac{1}{3}$

Part b
$$P(X=2k) = ln(3.5)\frac{\frac{1}{2}^{2k}}{2k} = ln(3.5)\frac{1}{2}\frac{(\frac{1}{4})^k}{k} = \frac{ln(3.5)}{2}e^{1.25}$$

$$P(X = 2k) = \frac{6}{\pi^2} 4k^2 = \frac{3}{2\pi^2} * \frac{\pi^2}{6} = \frac{1}{4}$$

Part d

$$P(X = 2k) = \frac{1}{e^2} \frac{2^{2k}}{(2k)!}$$

3

Since the coin tosses are independent, the choice can be represented by two successive coin tosses with probability of heads being p * p. Thus $P(X = k) = \binom{n}{k} p^{2k} (1 - p^2)^{n-k}$

5a

For Binomial:
$$f(k) = \binom{n}{k} p^k (1-p)^{n-k}$$
 Consider $LHS = f(k-1) * f(k+1)$ $LHS = \binom{n}{k-1} p^{k-1} (1-p)^{n-k-1} + \binom{n}{k+1} p^{k+1} (1-p)^{n-k-1} = \binom{n}{k-1} \binom{n}{k+1} (p^k (1-p)^{n-k})^2$ $RHS = \binom{n}{k}^2$ We now focus on $\binom{n}{k-1} \binom{n}{k+1}$ Let $y = \frac{\binom{n}{k-1} \binom{n}{k+1}}{\binom{n}{k}^2}$ Expanding: $y = \frac{n! n! (n-k)! (n-k)! k! k!}{(k-1)! (k+1)! (n-k+1)!} = \frac{k(n-k)}{(k+1)(n-k+1)}$ $\frac{k}{k+1} \le 1$ and $\frac{n-k}{n-k+1} \le 1$ $\forall k$ Hence $y \le 1$ Thus $LHS \le RHS$ For Poisson $f(k) = e^{-\lambda} \frac{\lambda^k}{k!}$ $LHS = f(k-1)f(k+1) = \frac{e^{-2\lambda} \lambda^{2k}}{(k+1)!(k-1)!}$ Thus $LHS \le RHS$

5b

$$f(k) = \frac{90}{(\pi k)^4} LHS = f(k+1)f(k-1) = \frac{90^2}{\pi^8(k+1)^4(k-1)^4} RHS = f(k+1)^2 = \frac{90^2}{\pi^8(k)^8}$$

$$y = LHS/RHS = \frac{k^8}{(k+1)^4(k-1)^4} = \left(\frac{k}{(k+1)}\frac{k}{(k-1)}\right)^4 = \left(\frac{k^2}{k^2-1}\right)^4 \ge 1$$
Thus $LHS \ge RHS$

5c

Any function of the form $P(x = k) = \frac{1}{n}$ satisfies $f(k)f(k-1) = f(k)^2$ Note: We aren't explicitly talking about countably many case.

Exercise # 3.2

2

Part a
$$P(min(X,Y) \le x) = 1 - P(X > x \cup Y > y) = 1 - P(X > x)P(Y > y) = 1 - 4^{-x}$$

Part b $P(Y > x) = \frac{1}{3}$ by symmetery of $P(X > Y) = P(X < Y) = P(X = Y)$ Part c $P(X = Y) = \frac{1}{3}$ as in (b)
Part d $P(X \ge y) = \frac{1}{3}$ as in (b)
Part e $P(X \text{ divides}Y) = P(Y = kX) = P(Y = kx, X = x) = P(Y = kx)P(X = x) = \sum_{k=1}^{\infty} \frac{1}{k} = \infty \frac{1}{2^{k+1}-1} = \sum_{k=1}^{\infty} \frac{1}{k} = \infty \frac{1}{2^{k}-1} =$

4

Consider three possibilities: 1. A rolls a 6, B,C do not

- 2. A and B roll a 6
- 3. No one rolls 6

$$\begin{array}{l} p = \frac{1}{6}(\frac{5}{6})^2 P(B < C) + \frac{1}{6}\frac{1}{6} + (\frac{5}{6})^3 p \\ P(B < C) = \frac{5}{6}\frac{5}{6}P(B < C) + \frac{1}{6} \implies P(B < C) = \frac{6}{11} \\ p(1 - \frac{125}{216}) = \frac{25}{216}\frac{6}{11} + \frac{6}{216} \\ p(\frac{91}{216}) = \frac{216}{216*11} = \frac{216}{1001} \end{array}$$

Exercise # 3.3

$$E(\frac{1}{X}) = \sum p(x)(\frac{1}{x})$$
$$\frac{1}{E(X)} = \sum p(x)(x)$$

For $E(X) = E(\frac{1}{X})$: $\sum (p(x)(x - \frac{1}{x})) = 0$ The above equation might not true be in general. However one possible case where this is true is for this distribution:

$$p(x) = \begin{cases} 1/2 & x = 1 \text{ or } x = -1 \\ 0 & \text{ otherwise} \end{cases}$$

- a) Given there are c objects with j chosen. The probability to select a new 'distinct' component given j are already selected is $\frac{c-j}{c}$. Thus the distribution is geometric with parameter $p=\frac{c-j}{c}$ and the mean being $\frac{1}{p} = \frac{c}{c-j}$
- b) Mean time required = $\sum 1 \sum_{j=0}^{c-1} \frac{c}{c-j}$

5

$$f(x) = \begin{cases} \frac{1}{x(x+1)} & x = 1, 2, \dots \\ 0 & otherwise \end{cases}$$

 $f(x) = \begin{cases} \frac{1}{x(x+1)} & x = 1, 2, \dots \\ 0 & otherwise \end{cases}$ $E(X^{\alpha}) = \sum \frac{x^{\alpha-1}}{x+1} \text{ For } E(X^{\alpha}) < \infty, \text{ the above sequence should not be diverging. and hence: } E(X^{\alpha}) = \sum \frac{1}{x^{2-\alpha} + x^{(1-\alpha)}} \text{ and hence } x^{2-\alpha} + x^{(1-\alpha)} \text{ should be converging } \Longrightarrow \alpha \le 1$