

# Probability and Statistics

## 6 – Hypothesis Testing

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# Null Hypothesis, Tests

The objective of statistical hypothesis tests, is to check whether a hypothesis is consistent with data from a random sample.

## Notation (6.1)

A hypothesis about a parameter  $\theta$  of a random distribution is called null hypothesis  $H_0$ .

If  $H_0$  completely determines the distribution (e.g.  $H_0 : \theta = 0$ ),  $H_0$  is called a *simple hypothesis*, otherwise (e.g.  $H_0 : \theta > 0$ ),  $H_0$  is called a *composite hypothesis*.

A *test of a null hypothesis* is always based on a random sample  $X_1, \dots, X_n$  and the (a priori) definition of a so-called *critical region*  $C \subseteq \mathbb{R}^n$ . A particular test consists of the observation of  $n$  values  $x_1, \dots, x_n$  from the random variables  $X_1, \dots, X_n$  and:

- $H_0$  is *accepted* if  $(x_1, \dots, x_n) \notin C$ .
- $H_0$  is *rejected* if  $(x_1, \dots, x_n) \in C$ .

## Example (Ross, Chapter 8, Exc. 4)

In a certain chemical process, it is very important that a particular solution that is to be used as a reactant have a pH of exactly 8.20. A method for determining pH that is available for solutions of this type is known to give measurements that are normally distributed with a mean equal to the actual pH and with a standard deviation of 0.02. Suppose 10 independent measurements yielded the following pH values:

8.18, 8.17, 8.16, 8.15, 8.17, 8.21, 8.22, 8.16, 8.19, 8.18

$$X_i \sim \mathcal{N}(\mu, \sigma), \sigma = 0.02$$

$$H_0: \mu = 8.20$$

$$\alpha = 0.05$$

$$\Phi^{-1}\left(1 - \frac{\alpha}{2}\right) \approx 1.96$$

$$\bar{x} = 8.179$$

$$\mu \pm \frac{\sigma}{\sqrt{n}} \cdot \Phi^{-1}\left(1 - \frac{\alpha}{2}\right)$$

$$\mu \pm \frac{0.02}{\sqrt{10}} \cdot 1.96 \approx 0.012 \quad \text{with confidence } 95\%$$

# Error Types, Significance Level

## Notation (6.1 contd.)

In a test, two different types of errors may occur:

- Errors of *type I*:  $H_0$  is rejected (i. e.  $(x_1, \dots, x_n) \in C$ ), although  $H_0$  is correct.
- Errors of *type II*:  $H_0$  is accepted (i. e.  $(x_1, \dots, x_n) \notin C$ ), although  $H_0$  is false.

A test has significance level  $\alpha$  if the probability of false rejection of  $H_0$  (i.e. of a type I error) is at most  $\alpha$ .

# Tests concerning the mean of a normal distribution

Let  $X_1, \dots, X_n$  be a random sample with:

$$X_i \sim \mathcal{N}(\mu, \sigma)$$

$$(1) \sigma \text{ known} \rightarrow \bar{X} \sim \mathcal{N}\left(\mu, \frac{\sigma}{\sqrt{n}}\right) \quad \text{cdf: } \Phi$$

$$(2) \sigma \text{ unknown} \rightarrow \tilde{X}, S \quad \text{cdf: } F_{t_{n-1}}$$

## Two-sided tests concerning $\mu$

$$H_0 : \mu = \mu_0$$

As the sample mean  $\bar{X}$  is a suitable estimator for  $\mu$ , the hypothesis should be rejected if the mean value  $\bar{x}$  of a set of sampled data  $x_1, \dots, x_n$  differs significantly from  $\mu_0$ . A critical region is therefore given by

$$C := \{(x_1, \dots, x_n) \mid |\bar{x} - \mu_0| > c\}$$

for a suitable constant  $c$ .

Given a significance level  $\alpha$ ,  $c$  has to be determined, such that:

$$\Pr_{H_0}(|\bar{X} - \mu_0| > c) := \Pr(|\bar{X} - \mu_0| > c \mid \mu = \mu_0) = \alpha$$

↑ conditional probability

## $X_i \sim \mathcal{N}(\mu, \sigma)$ : Two-sided tests concerning $\mu$

$$(1) \quad \Pr(|\bar{X} - \mu_0| \leq c \mid \mu = \mu_0) = \alpha$$

$$\Pr(-c < \bar{X} - \mu < c) = 1 - \alpha$$

$$\Pr\left(-\frac{c}{\sigma/\sqrt{n}} < \underbrace{\frac{\bar{X} - \mu}{\sigma/\sqrt{n}}}_{\sim \mathcal{N}(0,1)} < \frac{c}{\sigma/\sqrt{n}}\right) = 1 - \alpha \iff \frac{c}{\sigma/\sqrt{n}} = \Phi^{-1}\left(1 - \frac{\alpha}{2}\right)$$

$$\iff c = \frac{\sigma}{\sqrt{n}} \cdot \Phi^{-1}\left(1 - \frac{\alpha}{2}\right)$$

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$$(2) \quad \Pr\left(-\frac{c}{S/\sqrt{n}} < \underbrace{\frac{\bar{X} - \mu}{S/\sqrt{n}}}_{\sim t_{n-1}} < \frac{c}{S/\sqrt{n}}\right) = 1 - \alpha \iff c = \frac{S}{\sqrt{n}} \cdot F_{t_{n-1}}^{-1}\left(1 - \frac{\alpha}{2}\right)$$

# $X_i \sim \mathcal{N}(\mu, \sigma)$ : Two-sided tests concerning $\mu$

- $H_0$  is rejected if:

$$|\bar{x} - \mu_0| > c \iff \frac{\sqrt{n}}{\sigma} \cdot |\bar{x} - \mu_0| > \underbrace{\Phi^{-1}\left(1 - \frac{\alpha}{2}\right)}_{\hat{=}} \hat{=} \Phi^{-1}\left(1 - \frac{\alpha}{2}\right)$$

$$\frac{\sqrt{n}}{S} \cdot |\bar{x} - \mu_0| > F_{t_{n-1}}^{-1}\left(1 - \frac{\alpha}{2}\right)$$

*Handwritten notes:  $\Phi^{-1}(1 - \frac{\alpha}{2}) = 1 - \frac{\alpha}{2} \Leftrightarrow \alpha = 2 \cdot (1 - \Phi^{-1})$*

- $H_0$  is accepted if:

$$|\bar{x} - \mu_0| < c \iff \frac{\sqrt{n}}{\sigma} \cdot |\bar{x} - \mu_0| < \Phi^{-1}\left(1 - \frac{\alpha}{2}\right)$$

$$\frac{\sqrt{n}}{S} \cdot |\bar{x} - \mu_0| < F_{t_{n-1}}^{-1}\left(1 - \frac{\alpha}{2}\right)$$



$X_i \sim \mathcal{N}(\mu, \sigma)$ : Two-sided tests concerning  $\mu$ 

Starting with a set of sampled data,

$$v := \frac{\sqrt{n}}{\sigma} \cdot |\bar{X} - \mu_0|$$

can be calculated and the maximal value for  $\alpha$  can be determined, such that  $H_0$  will be accepted on the basis of the data sampled. This value

$$\alpha_{\bar{X}} := 2(1 - \Phi(v)) \quad \checkmark$$

is called the *p-value* of the sample.