Probability and Statistics

5 - Statistics

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Confidence Intervals

Notation (5.13)

Let \overline{x} denote the value of the sample mean of a random sample. Given an $\alpha \in (0,1)$ and intervals $I = I(\overline{x})$ for all such values \overline{x} , such that

$$Pr(\mu \in I) = 1 - \alpha$$

then the intervals are called *confidence intervals* of *confidence level* $1-\alpha$ for μ .

If I is a (symmetric) two-sided confidence interval with respect to \overline{x} , i.e.

$$I = [\overline{x} - \delta, \overline{x} + \delta]$$

for some $\delta \in \mathbb{R}^+$, this may be stated as:

$$\mu = \overline{x} \pm \delta$$
 with a confidence of $100(1-\alpha)\%$

(5.14)(i) Two-sided confidence intervals

$$1 - \alpha = \Pr\left(\mu \in [\overline{X} - \delta, \overline{X} + \delta]\right)$$

$$\iff 1 - \alpha = \Pr\left(|\overline{X} - \mu| \le \frac{\delta}{5/m}\right) = 2\Phi\left(\frac{\delta}{\sigma/\sqrt{n}}\right) - 1$$

$$(el. 4.5.6)$$

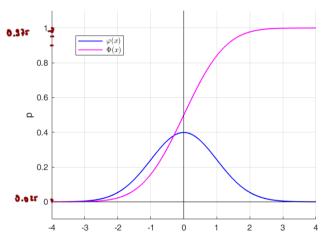
$$\iff \qquad 1 - \frac{\alpha}{2} \;\; = \;\; \Phi\left(\frac{\delta}{\sigma/\sqrt{n}}\right) \qquad \iff \qquad \delta \;\; = \;\; \frac{\sigma}{\sqrt{n}} \cdot \Phi^{-1}\!\!\left(1 - \frac{\alpha}{2}\right)$$

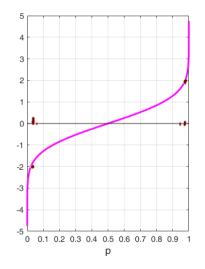
Therefore:
$$\mu = \overline{X} \pm \frac{\sigma}{\sqrt{n}} \cdot \Phi^{-1} \left(1 - \frac{\alpha}{2}\right)$$
 with confidence level $1 - \alpha$

$$\mu = \overline{X} \pm \frac{\sigma}{\sqrt{n}} \cdot \Phi^{-1} \left(1 - \frac{\alpha}{2}\right)$$

with confidence level $1-\alpha$ = 0.5.







(5.14)(ii) One-sided upper confidence intervals

$$1 - \alpha \ = \ \Pr \left(\mu \ \in \ [\overline{X} - \delta, \, \infty] \right) \ = \ \Pr \left(\mu \ \geq \ \overline{X} - \delta \right)$$

$$\iff 1 - \alpha = \Pr\left(\overline{X} - \mu \le \delta\right) = \Pr\left(\frac{\overline{X} - \mu}{\sigma/\sqrt{n}} \le \frac{\delta}{\sigma/\sqrt{n}}\right) = \Phi\left(\frac{\delta}{\sigma/\sqrt{n}}\right)$$

$$\iff \delta = \frac{\sigma}{\sqrt{n}} \cdot \Phi^{-1} (1 - \alpha)$$

Therefore: $\mu \geq \overline{X} - \frac{\sigma}{\sqrt{n}} \cdot \Phi^{-1}(1-\alpha)$ with confidence level $1-\alpha$

(5.14)(iii) One-sided lower confidence intervals

$$1 - \alpha = \Pr\left(\mu \in [-\infty, \overline{X} + \delta]\right) = \Pr\left(\mu \leq \overline{X} + \delta\right)$$

$$\iff 1 - \alpha = \Pr\left(-\delta \leq \overline{X} - \mu\right) = \Pr\left(-\frac{\delta}{\sigma/\sqrt{n}} \leq \frac{\overline{X} - \mu}{\sigma/\sqrt{n}}\right)$$

$$= 1 - \Phi\left(-\frac{\delta}{\sigma/\sqrt{n}}\right) = \Phi\left(\frac{\delta}{\sigma/\sqrt{n}}\right)$$

$$\iff \delta = \frac{\sigma}{\sqrt{n}} \cdot \Phi^{-1}(1 - \alpha)$$

Therefore:

$$\mu \leq \overline{X} + \frac{\sigma}{\sqrt{n}} \cdot \Phi^{-1}(1-\alpha)$$
 with confidence level $1-\alpha$

Confidence intervals for μ with a confidence level of $1-\alpha$ are:

$$\mu = \overline{X} \pm \frac{\sigma}{\sqrt{n}} \cdot \Phi^{-1} \left(1 - \frac{\alpha}{2} \right)$$

$$\mu \geq \overline{X} - \frac{\sigma}{\sqrt{n}} \cdot \Phi^{-1} (1 - \alpha)$$

$$\mu \leq \overline{X} + \frac{\sigma}{\sqrt{n}} \cdot \Phi^{-1} (1 - \alpha)$$

(5.16) Using the sample mean \overline{X} to estimate μ and the sample variance S^2 to estimate σ^2 . confidence intervals for μ can be determined by replacing the standard normal distribution used in (5.14) with the t distribution with n-1 degrees of freedom for the random variable:

$$\frac{\overline{X} - \mu}{S/\sqrt{n}}$$
 $\stackrel{\text{(5.12)}}{\sim}$ t_{n-1}

Two-sided confidence intervals

$$\iff 1 - \frac{\alpha}{2} = F_{t_{n-1}}\left(\frac{\delta}{S/\sqrt{n}}\right) \iff \delta = \frac{S}{\sqrt{n}} \cdot F_{t_{n-1}}^{-1}\left(1 - \frac{\alpha}{2}\right)$$

Therefore:
$$\mu = \overline{X} \pm \frac{S}{\sqrt{n}} \cdot F_{t_{n-1}}^{-1} \left(1 - \frac{\alpha}{2}\right)$$
 with confidence level $1 - \alpha$

(5.16) Using the sample mean \overline{X} to estimate μ and the sample variance S^2 to estimate σ^2 , confidence intervals for μ can be determined by replacing the standard normal distribution used in (5.14) with the t distribution with n-1 degrees of freedom for the random variable:

$$\frac{\overline{X}-\mu}{S/\sqrt{n}}$$
 $\stackrel{(5.12)}{\sim}$ t_{n-1}

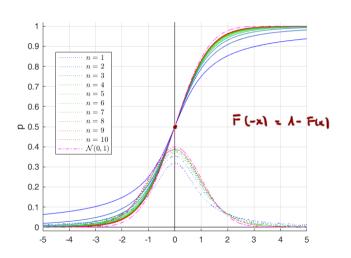
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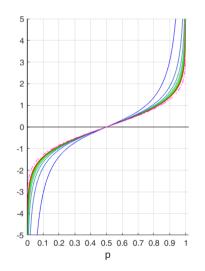
$$\mu = \overline{X} \pm \frac{S}{\sqrt{n}} \cdot F_{t_{n-1}}^{-1} \left(1 - \frac{\alpha}{2}\right)$$

$$\mu \geq \overline{X} - \frac{S}{\sqrt{n}} \cdot F_{t_{n-1}}^{-1} (1 - \alpha)$$

$$\mu \leq \overline{X} + \frac{S}{\sqrt{n}} \cdot F_{t_{n-1}}^{-1} (1 - \alpha)$$

t_n distributions for $n = 1, 2, \dots, 10$

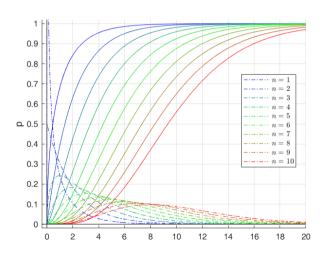


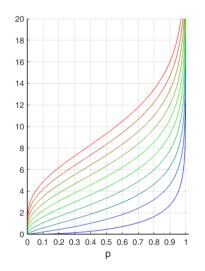


(5.18) Using the sample variance S^2 to estimate σ^2 , confidence intervals for σ^2 can be determined from the random variable:

$$Y_{n-1} := \frac{n-1}{\sigma^2} S^2 \stackrel{(5.11)}{\sim} \chi_{n-1}^2$$

Pdf's and cdf's of $\chi_n^2 = \Gamma\left(\frac{n}{2}, \frac{1}{2}\right)$ distributions for $n = 1, 2, \dots, 10$





(5.18)(i) Two-sided confidence intervals

Given $\alpha \in (0,1)$, put:

$$b_1:=F_{\stackrel{\frown}{Y_{n-1}}}^{-1}\left(rac{lpha}{2}
ight) \qquad ext{ and } \qquad b_2:=F_{\stackrel{\frown}{Y_{n-1}}}^{-1}\left(1-rac{lpha}{2}
ight)$$

$$b_2 := F_{Y_{n-1}}^{-1} \left(1 - \frac{\alpha}{2}\right)$$

Then:

$$\Pr\left(\frac{n-1}{b_2}S^2 \leq \sigma^2 \leq \frac{n-1}{b_1}S^2\right) = \Pr\left(b_1 \leq \frac{n-1}{\sigma^2}S^2 \leq b_2\right) = 1-\alpha$$

$$\sigma^2 \in \left[\frac{n-1}{h_2}S^2, \frac{n-1}{h_1}S^2\right]$$
 with confidence level $1-\alpha$

(5.18)(ii) One-sided lower confidence intervals

Given $\alpha \in (0,1)$, put:

$$b:=F_{Y_{n-1}}^{-1}(\alpha)$$

Then:

$$\Pr\left(\sigma^2 \leq \frac{n-1}{b}S^2\right) = \Pr\left(\frac{n-1}{\sigma^2}S^2 \geq b\right) = 1-\alpha$$

Therefore: $\sigma^2 \leq \frac{n-1}{b} S^2$ with confidence level $1-\alpha$

(5.18)(iii) One-sided upper confidence intervals

Given $\alpha \in (0,1)$, put:

$$b := F_{Y_{n-1}}^{-1} (1 - \alpha)$$

Then:

$$\Pr\left(\sigma^2 \geq \frac{n-1}{b}S^2\right) = \Pr\left(\frac{n-1}{\sigma^2}S^2 \leq b\right) = 1-\alpha$$

Therefore: $\sigma^2 \geq \frac{n-1}{h} S^2$ with confidence level $1-\alpha$

$X_i \sim \mathsf{Bernoulli}(p)$

Remark (5.19)

Let X_1, \ldots, X_n be a random sample with $X_i \sim \text{Bernoulli}(p)$. Then $n\overline{X} \sim \text{binomial}(n, p)$ and for sufficiently large integers n, we get:

$$\frac{\overline{X}-p}{\sqrt{p(1-p)/n}} \stackrel{\approx}{\sim} \mathcal{N}(0,1)$$

(5.20)(i) Two-sided confidence intervals

$$1 - \alpha = \Pr\left(p \in [\overline{X} - \delta, \overline{X} + \delta]\right) = \Pr\left(|\overline{X} - p| \le \delta\right)$$

$$= \Pr\left(-\delta \le \overline{X} - p \le \delta\right)$$

$$= \Pr\left(-\frac{\delta}{\sqrt{p(1-p)/n}} \le \frac{\overline{X} - p}{\sqrt{p(1-p)/n}} \le \frac{\delta}{\sqrt{p(1-p)/n}}\right)$$

$$\approx 2 \cdot \Phi\left(\frac{\delta}{\sqrt{p(1-p)/n}}\right) - 1$$

$$p \in (\mathfrak{I}, \Lambda)$$

$$p \in ($$

$$\iff$$

(5.20)(ii) One-sided lower confidence intervals

$$1 - \alpha = \Pr\left(p \in (-\infty, \overline{X} + \delta]\right) = \Pr\left(\overline{X} - p \ge -\delta\right)$$

$$= \Pr\left(\frac{\overline{X} - p}{\sqrt{p(1 - p)/n}} \ge -\frac{\delta}{\sqrt{p(1 - p)/n}}\right)$$

$$\approx \Phi\left(\frac{\delta}{\sqrt{p(1 - p)/n}}\right)$$

$$\iff \delta \approx \frac{\sqrt{p(1 - p)} \cdot \Phi^{-1}(1 - \alpha)}{\sqrt{n}} \le \frac{\Phi^{-1}(1 - \alpha)}{2\sqrt{n}}$$

(5.20)(iii) One-sided upper confidence intervals

$$1 - \alpha = \Pr\left(p \in [\overline{X} - \delta, \infty)\right) = \Pr\left(\overline{X} - p \le \delta\right)$$

$$= \Pr\left(\frac{\overline{X} - p}{\sqrt{p(1 - p)/n}} \le \frac{\delta}{\sqrt{p(1 - p)/n}}\right)$$

$$\approx \Phi\left(\frac{\delta}{\sqrt{p(1 - p)/n}}\right)$$

$$\iff \delta \approx \frac{\sqrt{p(1 - p)} \cdot \Phi^{-1}(1 - \alpha)}{\sqrt{p}} \le \frac{\Phi^{-1}(1 - \alpha)}{2\sqrt{p}}$$

(5.21) Approximation of n for confidence intervals of given width

Given α , δ and an estimation $\overline{x} \approx p$ (e.g. from a small preliminary sample), an estimation for the sample size n, such that a two-sided confidence interval for p of confidence level $1-\alpha$ has width 2δ , is given by:

$$n \approx \frac{p(1-p)\cdot\left(\Phi^{-1}\left(1-\frac{\alpha}{2}\right)\right)^2}{\delta^2} \approx \frac{\overline{x}(1-\overline{x})\cdot\left(\Phi^{-1}\left(1-\frac{\alpha}{2}\right)\right)^2}{\delta^2}$$

As $p(1-p) \leq rac{1}{4}$ for all $p \in \mathbb{R}$, the estimation

$$n \lesssim \frac{\left(\Phi^{-1}\left(1-\frac{\alpha}{2}\right)\right)^2}{4\delta^2}$$

holds true for all values of p.