

Analysis of Variance Single Factor

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Outline

- Introduction
- Single-Factor ANOVA:
 - Equal sample size
 - F-test
 - Multiple comparison

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Introduction : Analysis of Variance

- Analysis of quantitative responses
- In short, ANOVA
- Simplest ANOVA
 - Single-factor or One-way
 - Factorial or Multiple-way
- Examples
 - Study of five brands of gasoline on car efficiency
 - Study of four types of sugar on bacteria growth



Outline

- Introduction
- Single-Factor ANOVA:
 - Equal sample size
 - F-test
 - Multiple comparison



Single-Factor ANOVA : Equal Sample Size

- Compare two or more populations on one factor
 - Let
 - μ_1 = mean of treatment (population) 1
 - μ_2 = mean of treatment (population) 2
 - ...
 - μ_I = mean of treatment (population) I
- I = number of compared treatments
- Hypothesis
 - $H_0: \mu_1 = \mu_2 = \dots = \mu_I$
 - H_a : Not all μ_i 's are equal (at least two of the μ_i 's are different)

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Notation

- Let
 - X_{ij} = Random variable for measurement j of treatment i
 - x_{ij} = Sample value for measurement j of treatment i
 - J = Number of samples in one treatment
 - I = Number of treatments

- (Treatment) sample mean: $\bar{X}_i = \frac{\sum_{j=1}^J X_{ij}}{J}$

Divided by number of samples
in one treatment

- Grand mean: $\bar{X} = \frac{\sum_{i=1}^I \bar{X}_i}{I} = \frac{\sum_{i=1}^I \sum_{j=1}^J X_{ij}}{IJ}$

Divided by number of samples
from treatments

- (Treatment) sample variance: $S_i^2 = \frac{\sum_{j=1}^J (X_{ij} - \bar{X}_i)^2}{J-1}$

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Example

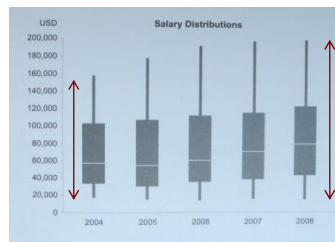
- Experiment testing strength of 4 shipping boxes

Type	Compression Strength (lb)						Sample Mean	Sample SD
1	655.5	788.3	734.3	721.4	679.1	699.4	713.00	46.55
2	789.2	772.5	786.9	686.1	732.1	774.8	756.93	40.34
3	737.1	639.0	696.3	671.7	717.2	727.1	698.07	37.20
4	535.1	628.7	542.4	559.0	586.9	520.0	562.02	39.87

$$\text{Grand Mean} = (713.00 + 756.93 + 698.07 + 562.02) / 4 = 682.50$$

Box Plots

- ✓ Analyze distribution within series
- ✓ Analyze change of distributions over time



- To analyze distribution change within series or over time
- Feature
 - Length represents distributed amount

Image source: Figure 7.23, [2]

Box Plots (cont.)

• Example

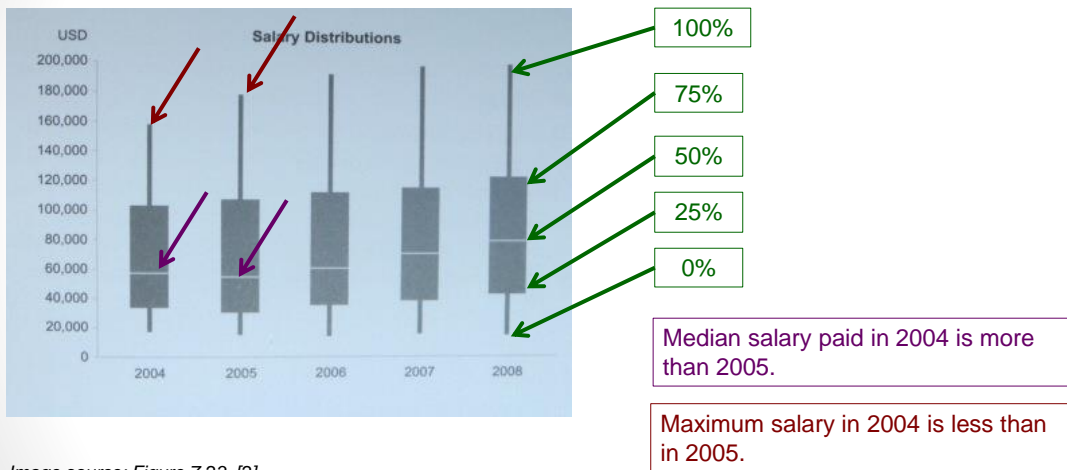


Image source: Figure 7.23, [2]

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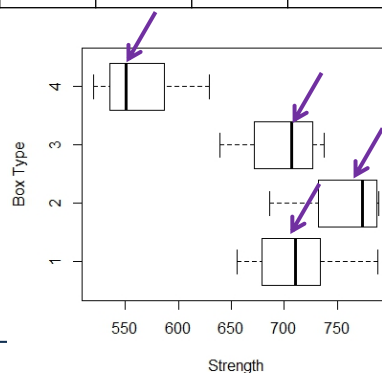
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Example (cont.)

Type	Compression Strength (lb)						Sample Mean	Sample SD
1	655.5	788.3	734.3	721.4	679.1	699.4	713.00	46.55
2	789.2	772.5	786.9	686.1	732.1	774.8	756.93	40.34
3	737.1	639.0	696.3	671.7	717.2	727.1	698.07	37.20
4	535.1	628.7	542.4	559.0	586.9	520.0	562.02	39.87



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Basic Assumption

- Assume
 - Distribution of each population is normal with the same variance = σ^2
- Therefore, each sample X_{ij} comes from normal distribution with
 - $E(X_{ij}) = \mu_i$
 - $V(X_{ij}) = \sigma^2$
- Hypothesis
 - $H_0: \mu_1 = \mu_2 = \dots = \mu_I$
 - H_a : Not all μ_i 's are equal (at least two of the μ_i 's are different)



Sum of Squares

- When H_0 is true, all sample means ($\bar{x}_1, \bar{x}_2, \dots, \bar{x}_I$) should be the same
- Therefore, test statistics will be measured from differences of sample means
- Treatment sum of squares (SST_r): Difference between different treatments
 - Sum of difference between each sample mean and grand mean
 - $$SST_r = J(\bar{X}_1 - \bar{X})^2 + J(\bar{X}_2 - \bar{X})^2 + \dots + J(\bar{X}_I - \bar{X})^2$$
$$= J \sum_i (\bar{X}_i - \bar{X})^2$$
- Error sum of squares (SSE): Difference within the same treatment
 - Sum of differences between samples and sample mean
 - $$SSE = \sum_i \sum_j (X_{ij} - \bar{X}_i)^2$$



Review: Sample Variance Distribution

- Let X_1, X_2, \dots, X_n be random sample from a normal distribution with mean value = μ and standard deviation = σ .
- Then, sample variance S^2 has distribution to be a chi-square distribution with degree of freedom = n-1

$$S^2 \approx \sigma^2 \frac{\chi_{n-1}^2}{(n-1)} \Rightarrow \frac{(n-1)S^2}{\sigma^2} \approx \chi_{(n-1)}^2$$

Sum of Squares (cont.)

- Treatment sum of squares (SST_r) :

$$\begin{aligned} SST_r &= J \sum_i (\bar{X}_i - \bar{X})^2 \\ &= J \sum_i (Y_i - \bar{Y})^2 \\ &= J \frac{\sum_i (Y_i - \bar{Y})^2}{(I-1)} \end{aligned}$$

$$= (I-1)J S_Y^2$$

$$\begin{aligned} \frac{SST_r}{\sigma^2} &= \frac{(I-1)J S_Y^2}{\sigma^2/J} \\ &= \frac{(I-1)}{(\sigma^2/J)} \cdot \frac{\sigma_Y^2 \chi_{I-1}^2}{I-1} \end{aligned}$$

$$= \frac{\cancel{(I-1)} \cancel{(\sigma^2/J)} \chi_{I-1}^2}{\cancel{(\sigma^2/J)} \cancel{(I-1)}} \approx \chi_{I-1}^2$$

$$X_i \sim N(\mu, \sigma^2) \rightarrow Y_i = \bar{X}_i \sim N\left(\mu, \frac{\sigma^2}{J}\right)$$

$$S^2 = \frac{\sum_{i=1}^n (x - \bar{x})^2}{n-1}$$

$$S_X^2 \approx \sigma_X^2 \frac{\chi_{n-1}^2}{(n-1)}$$

Sum of Squares (cont.)

- Error sum of squares (SSE)

- $$SSE = \sum_i \sum_j (X_{ij} - \bar{X}_i)^2$$

$$= \sum_j (X_{1j} - \bar{X}_1)^2 + \sum_j (X_{2j} - \bar{X}_2)^2 + \dots + \sum_j (X_{Ij} - \bar{X}_I)^2$$

$$= (J-1)S_1^2 + (J-1)S_2^2 + \dots + (J-1)S_I^2$$

$$= (J-1)[S_1^2 + S_2^2 + \dots + S_I^2]$$

Each X_i has the same variance

$$= I(J-1)S^2$$

$$S^2 = \frac{\sum_{i=1}^n (x_i - \bar{x})^2}{n-1}$$

$$X_i \sim N(\mu, \sigma^2)$$

$$\frac{SSE}{\sigma^2} = \frac{I(J-1)S^2}{\sigma^2} \approx \chi_{I(J-1)}^2$$

$$\frac{(n-1)S^2}{\sigma^2} \approx \chi_{(n-1)}^2$$

F distribution

- Let X_1 and X_2 be independent *chi-squared* random variables with v_1 and v_2 degrees of freedom

$$F_{v_1, v_2} = \frac{X_1/v_1}{X_2/v_2}$$

- Generally, sample variance has sampling distribution in term of chi-squared distribution with degree of freedom = $n-1$

$$S^2 \approx \sigma^2 \frac{\chi_{n-1}^2}{(n-1)}$$

- Let S_1^2 and S_2^2 be sample variances with chi-squared distribution

$$\frac{(m-1)S_1^2}{\sigma_1^2} \approx \chi_{(m-1)}^2 \quad \text{and} \quad \frac{(n-1)S_2^2}{\sigma_2^2} \approx \chi_{(n-1)}^2$$

$$F_{m-1, n-1} = \frac{\frac{(m-1)S_1^2/\sigma_1^2}{m-1}}{\frac{(n-1)S_2^2/\sigma_2^2}{n-1}} = \frac{S_1^2/\sigma_1^2}{S_2^2/\sigma_2^2}$$

F distribution (cont.)

- Let X_1 and X_2 be independent chi-squared random variables with v_1 and v_2 degrees of freedom

$$F_{v_1, v_2} = \frac{X_1/v_1}{X_2/v_2}$$

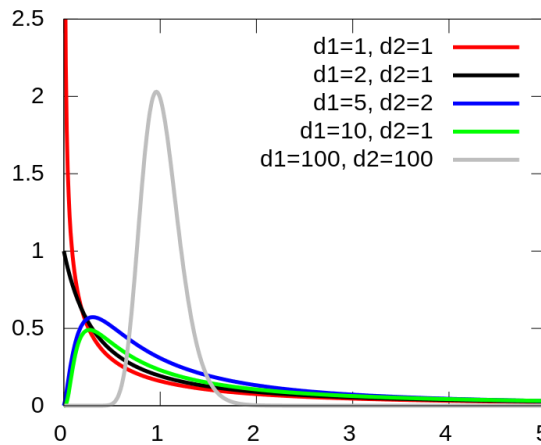


Image source: <http://en.wikipedia.org/wiki/F-distribution>

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Sum of Squares (cont.)

- Treatment sum of squares (SST_r) :

$$\frac{SST_r}{\sigma^2} = \frac{(I-1)S_Y^2}{\sigma^2/J} = \frac{(I-1)(\sigma^2/J)\chi_{I-1}^2}{(\sigma^2/J)(I-1)} \approx \chi_{I-1}^2$$

- Error sum of squares (SSE)

$$\frac{SSE}{\sigma^2} = \frac{I(J-1)S^2}{\sigma^2} \approx \chi_{I(J-1)}^2$$

Use as test statistic

$$\frac{\frac{SST_r}{\sigma^2} / (I-1)}{\frac{SSE}{\sigma^2} / I(J-1)} = \frac{SST_r / (I-1)}{SSE / I(J-1)} = \frac{\chi_{I-1}^2 / (I-1)}{\chi_{I(J-1)}^2 / I(J-1)} \approx F_{I-1, I(J-1)}$$

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To Perform F-Test

Mean of chi-square distribution = degree of freedom

- Test statistic:

$$f = F_{I-1, I(J-1)} = \frac{SSTr/(I-1)}{SSE/(I(J-1))}$$

$$= \frac{MSTr}{MSE}$$

$$E\left(\frac{SSTr}{\sigma^2}\right) = I - 1$$

$$E\left(\frac{SSE}{\sigma^2}\right) = I(J - 1)$$

$$E\left(\frac{SSTr}{I - 1}\right) = \sigma^2$$

$$E\left(\frac{SSE}{I(J - 1)}\right) = \sigma^2$$

$$E(MSTr) = \sigma^2$$

$$E(MSE) = \sigma^2$$

- If H_0 is true (\bar{X}_i' s are about the same),
MSTr and MSE are unbiased estimates of $\sigma^2 \Rightarrow$ test statistic $f \sim 1$ (f is small)
- If H_0 is false (\bar{X}_i' s are not the same),
 $E(MSTr) > \sigma^2 \Rightarrow$ test statistic f is large.
- Hence, rejection region covers large test statistic f

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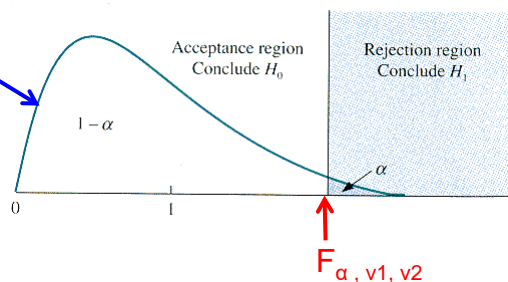
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F-Test

- Test statistic:

$$F_{I-1, I(J-1)} = \frac{\frac{SSTr}{\sigma^2} / (I - 1)}{\frac{SSE}{\sigma^2} / (I(J - 1))} = \frac{SSTr / (I - 1)}{SSE / (I(J - 1))} = \frac{MSTr}{MSE}$$

F distribution
for v_1 and v_2



If test statistic $> F_{\alpha, v_1, v_2}$, reject H_0

Image source: <http://www.unc.edu/~nielsen/soci708/m16/m2009.gif>

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Summary : ANOVA

- Test statistic:

$$F_{I-1, I(J-1)} = \frac{\frac{SSTr}{\sigma^2} / (I-1)}{\frac{SSE}{\sigma^2} / (I(J-1))} = \frac{SSTr / (I-1)}{SSE / (I(J-1))} = \frac{MSTr}{MSE}$$

- When f is large, we are about to reject H_0

	df	Sum of Squares (SS)	Mean Square (MS)	f
Treatment	I-1	SSTr	SSTr/(I-1)	MSTr / MSE
Error	I(J-1)	SSE	SSE/(I(J-1))	
Total	IJ-1	SST		

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Summary : ANOVA (cont.)

- $SST = SSTr + SSE$

$$\sum_i \sum_j (x_{ij} - \bar{x})^2 = \sum_i \sum_j (x_{ij} - \bar{x}_i + \bar{x}_i - \bar{x})^2$$

$$= \sum_i \sum_j (x_{ij} - \bar{x}_i)^2 + \sum_i \sum_j (\bar{x}_i - \bar{x})^2 - 2 \sum_i \sum_j (x_{ij} - \bar{x}_i)(\bar{x}_i - \bar{x})$$

$$= SSE + SSTr - 2 \sum_i \sum_j (x_{ij} - \bar{x}_i)(\bar{x}_i - \bar{x}) \quad 0$$

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Summary : ANOVA (cont.)

- Test statistic:

$$F_{I-1, I(J-1)} = \frac{\frac{SSTr}{\sigma^2} / (I-1)}{\frac{SSE}{\sigma^2} / (I(J-1))} = \frac{SSTr / (I-1)}{SSE / (I(J-1))} = \frac{MSTr}{MSE}$$

Another option:
Sample-based
computation

	df	Sum of Squares (SS)	Mean Square (MS)	f
Treatment	I-1	$\frac{1}{J} \sum_{i=1}^I (\sum_{j=1}^J x_{ij})^2 - \frac{1}{IJ} (\sum_{i=1}^I \sum_{j=1}^J x_{ij})^2$	SSTr/(I-1)	MSTr / MSE
Error	I(J-1)	SST - SSTr	SSE/(I(J-1))	
Total	IJ-1	$\sum_{i=1}^I \sum_{j=1}^J x_{ij}^2 - \frac{1}{IJ} (\sum_{i=1}^I \sum_{j=1}^J x_{ij})^2$		

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Example 1

- Experiment degree of soiling on 3 mixtures of fabric and polymer
- Prove whether 3 mixture means are the same at $\alpha = 0.01$

Mixture	Degree of soiling				
1:	0.56	1.12	0.90	1.07	0.94
2:	0.72	0.69	0.87	0.78	0.91
3:	0.62	1.08	1.07	0.99	0.93

- Let
 - μ_1 = mean of mixture 1
 - μ_2 = mean of mixture 2
 - μ_3 = mean of mixture 3
 - $I = 3, J = 5$
- Hypothesis
 - $H_0: \mu_1 = \mu_2 = \mu_3$
 - H_a : Not all μ_i 's are equal (at least two of the μ_i 's are different)

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Example 1 (cont.)

- Experiment degree of soiling on 3 mixtures

Mixture	Degree of soiling					\bar{x}_i
1:	0.56	1.12	0.90	1.07	0.94	0.918
2:	0.72	0.69	0.87	0.78	0.91	0.794
3:	0.62	1.08	1.07	0.99	0.93	0.938

- Fill ANOVA table

$$\bar{x} = 0.883$$

	df	Sum of Squares (SS)	Mean Square (MS)	f
Treatment	2	0.0608	0.0304	0.99
Error	12	0.3701 = 0.4309 – 0.0608	0.0308	
Total	14	0.4309		

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Example 1 (cont.)

	df	Sum of Squares (SS)	Mean Square (MS)	f
Treatment	2	0.0608	0.0304	0.99
Error	12	0.3701 = 0.4309 – 0.0608	0.0308	
Total	14	0.4309		

- Rejection region: given $\alpha = 0.01$
 - $F_{0.01, 2, 12} = 6.93$
- H_0 is not rejected.
- All mixture means are equals.

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Outline

- Introduction
- Single-Factor ANOVA:
 - Equal sample size
 - F-test
 - Multiple comparison



Multiple Comparisons

- Question:
 - When H_0 for ANOVA is rejected, how many means are different from each other?
- Procedures:
 - Find confidence interval of pairwise difference $\mu_i - \mu_j$
 - If **confidence interval for** any pairwise difference $\mu_i - \mu_j$ **does not include zero**, we determine that μ_i, μ_j are significantly different from each other



Studentized Range Distribution

- Let Z_1, Z_2, \dots, Z_m be m independent *standard normal* random variables
- Let W be a *chi-squared* random variable with *degree of freedom* = v , and independent of the Z_i 's
- Then, *Q distribution*, called *studentized range distribution* is defined as

$$Q = \frac{\max |Z_i - Z_j|}{\sqrt{\frac{W}{v}}} \quad \text{where } W = \frac{SSE}{\sigma^2} = \chi^2_{I(J-1)} = \frac{I(J-1)MSE}{\sigma^2}$$

- This Q distribution has 2 parameters: m and v
 - Hence, it is denoted by $Q_{\alpha, m, v}$



Studentized Range Distribution (cont.)

$$X_i \sim N(\mu, \sigma^2) \rightarrow \bar{X}_i \sim N(\mu, \frac{\sigma^2}{J})$$

$$Z_i = \frac{\bar{X}_i - \mu_i}{\sigma/\sqrt{J}} \sim N(0, 1)$$

- From

$$Z_i = \frac{\bar{X}_i - \mu_i}{\sigma/\sqrt{J}}, \quad W = \frac{SSE}{\sigma^2} = \chi^2_{I(J-1)} = \frac{I(J-1)MSE}{\sigma^2}, \quad m = I, \quad v = I(J-1)$$

$$Q = \frac{\max |Z_i - Z_j|}{\sqrt{\frac{W}{v}}} = \frac{\max \left| \frac{\bar{X}_i - \mu_i}{\sigma/\sqrt{J}} - \frac{\bar{X}_j - \mu_j}{\sigma/\sqrt{J}} \right|}{\sqrt{\frac{I(J-1)MSE}{\sigma^2} \frac{1}{I(J-1)}}} = \frac{\max |\bar{X}_i - \bar{X}_j - (\mu_i - \mu_j)|}{\sqrt{MSE/J}}$$

$$1 - \alpha = P \left(\frac{\max |\bar{X}_i - \bar{X}_j - (\mu_i - \mu_j)|}{\sqrt{MSE/J}} \leq Q_{\alpha, I, I(J-1)} \right)$$

$$= P \left(\frac{|\bar{X}_i - \bar{X}_j - (\mu_i - \mu_j)|}{\sqrt{MSE/J}} \leq Q_{\alpha, I, I(J-1)} \text{ for all } i, j \right)$$

Probability that maximum difference between sample mean difference ($\bar{X}_i - \bar{X}_j$) true mean difference ($\mu_i - \mu_j$) is less than $Q_{\alpha, I, I(J-1)} \sqrt{MSE/J}$ = $1 - \alpha$



Studentized Range Distribution (cont.)

$$\begin{aligned}
 1 - \alpha &= P\left(\frac{|\bar{X}_i - \bar{X}_j - (\mu_i - \mu_j)|}{\sqrt{MSE/J}} \leq Q_{\alpha, I, I(J-1)} \text{ for all } i, j\right) \\
 &= P\left(-Q_{\alpha, I, I(J-1)} \sqrt{\frac{MSE}{J}} \leq \bar{X}_i - \bar{X}_j - (\mu_i - \mu_j) \leq Q_{\alpha, I, I(J-1)} \sqrt{\frac{MSE}{J}} \text{ for all } i, j\right) \\
 &= P\left(\underbrace{\bar{X}_i - \bar{X}_j - Q_{\alpha, I, I(J-1)} \sqrt{\frac{MSE}{J}} \leq \mu_i - \mu_j \leq \bar{X}_i - \bar{X}_j + Q_{\alpha, I, I(J-1)} \sqrt{\frac{MSE}{J}}}_{\text{Confidence intervals between one pair of true mean difference } \mu_i - \mu_j} \text{ for all } i, j\right)
 \end{aligned}$$

Confidence intervals between one pair of true mean difference $\mu_i - \mu_j$

- There are $\binom{I}{2} = \frac{I(I-1)}{2}$ confidence intervals of $\mu_i - \mu_j$

Studentized Range Distribution (cont.)

$$1 - \alpha = P\left(\bar{X}_i - \bar{X}_j - Q_{\alpha, I, I(J-1)} \sqrt{\frac{MSE}{J}} \leq \mu_i - \mu_j \leq \bar{X}_i - \bar{X}_j + Q_{\alpha, I, I(J-1)} \sqrt{\frac{MSE}{J}} \text{ for all } i, j\right)$$

Expect difference between sample mean difference ($\bar{X}_i - \bar{X}_j$) and

true mean difference ($\mu_i - \mu_j$) is not more than this value $Q_{\alpha, I, I(J-1)} \sqrt{\frac{MSE}{J}}$

- The value $w = Q_{\alpha, I, I(J-1)} \sqrt{\frac{MSE}{J}}$ is called
Tukey's honestly significantly difference (HSD)

Multiple Comparisons : Equal Sample Size

- When some means are not all equal, how to specify which mean is different from others
- Procedure
 - Find Tukey's Honestly Significant Difference (HSD)

$$HSD_{\alpha} = q_{\alpha, I, I(J-1)} \sqrt{\frac{MSE}{J}}$$

- $q_{\alpha, I, I(J-1)}$ = q-value from studentized range distribution with 2 degrees of freedom $I, I(J-1)$
- Sort sample means in increasing order
- Underline pairs that differ less than HSD_{α}
- Any pair without underline are considered as significantly different.

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Example 1

- Experiment degree of soiling on 3 mixtures
- Group mixture means at $\alpha = 0.01$

	df	Sum of Squares (SS)	Mean Square (MS)	f
Treatment	2	0.0608	0.0304	0.99
Error	12	0.3701 = 0.4309 - 0.0608	0.0308	
Total	14	0.4309		

Mixture	Degree of soiling					\bar{x}_i
1:	0.56	1.12	0.90	1.07	0.94	0.918
2:	0.72	0.69	0.87	0.78	0.91	0.794
3:	0.62	1.08	1.07	0.99	0.93	0.938

$$HSD_{\alpha} = q_{\alpha, I, I(J-1)} \sqrt{\frac{MSE}{J}} = q_{0.01, 3, 12} \sqrt{\frac{0.0308}{5}} = 5.05 \sqrt{\frac{0.0308}{5}} = 0.396$$

- Sort sample means: 0.794, 0.918, 0.938

0.124 0.020

```
> qtukekey(0.01, nmeans=anovaResult$rank,
+ df=anovaResult$df.residual, lower.tail=F)
[1] 5.045934
```

- One group of mixture means

Note that H_0 is not rejected.

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Example 1

- Experiment degree of soiling on 3 m
- Group mixture means at $\alpha = 0.01$

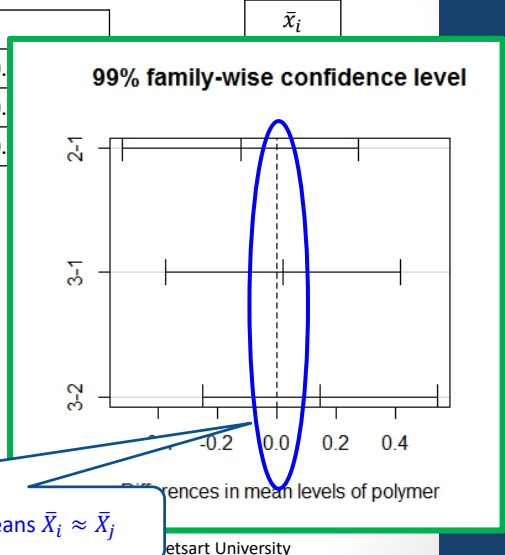
```
> tk <- TukeyHSD(anovaResult, conf.level = 0.99)
> plot(tk)
```

Mixture	Degree of soiling				
1:	0.56	1.12	0.90	1.07	0.
2:	0.72	0.69	0.87	0.78	0.
3:	0.62	1.08	1.07	0.99	0.

- Sort sample means:

\bar{x}_2 \bar{x}_1 \bar{x}_3
 0.794, 0.918, 0.938
 0.124 0.020

- One group of mixture means



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Confidence interval of difference ($\bar{X}_i - \bar{X}_j$) includes zero. This means $\bar{X}_i \approx \bar{X}_j$

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Example 2

- Test on 5 brands of automobile oil filters
 - Use 9 samples for each brands
- $\bar{x}_1 = 14.5$, $\bar{x}_2 = 13.8$, $\bar{x}_3 = 13.3$, $\bar{x}_4 = 14.3$, $\bar{x}_5 = 13.1$

	df	Sum of Squares (SS)	Mean Square (MS)	f
Treatment	4	13.32	3.33	37.84
Error	40	3.53	0.088	
Total	44	16.85		

- Rejection region:
 - $F_{0.05, 4, 40} = 2.61$
 - H_0 is rejected
 - Find Tukey's HSD to see mean differences

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Example 2 (cont.)

- $\bar{x}_1 = 14.5$, $\bar{x}_2 = 13.8$, $\bar{x}_3 = 13.3$, $\bar{x}_4 = 14.3$, $\bar{x}_5 = 13.1$

	df	Sum of Squares (SS)	Mean Square (MS)	f
Treatment	4	13.32	3.33	37.84
Error	40	3.53	0.088	
Total	44	16.85		

$$HSD_{\alpha} = q_{\alpha, I, I(J-1)} \sqrt{\frac{MSE}{J}} = q_{0.05, 5, 40} \sqrt{\frac{0.088}{9}} = 4.04 \sqrt{\frac{0.088}{9}} = 0.399$$

$\bar{x}_5 \quad \bar{x}_3 \quad \bar{x}_2 \quad \bar{x}_4 \quad \bar{x}_1$

- Sort sample means: 13.1, 13.3, 13.8, 14.3, 14.5
- 3 groups of means:
 - \bar{x}_5, \bar{x}_3 are not significantly different from each other
 - \bar{x}_4, \bar{x}_1 are not significantly different from each other
 - \bar{x}_2 is significantly different from \bar{x}_5, \bar{x}_3 and \bar{x}_4, \bar{x}_1

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Example 2 (cont.)

- If use another value for sample mean and same HSD_{α} :
 - $\bar{x}_1 = 14.5$, $\bar{x}_2 = 14.15$, $\bar{x}_3 = 13.3$, $\bar{x}_4 = 14.3$, $\bar{x}_5 = 13.1$

$$HSD_{\alpha} = q_{0.05, 5, 40} \sqrt{\frac{0.088}{9}} = 4.04 \sqrt{\frac{0.088}{9}} = 0.399$$

$\bar{x}_5 \quad \bar{x}_3 \quad \bar{x}_2 \quad \bar{x}_4 \quad \bar{x}_1$

- Sort sample means: 13.1, 13.3, 14.15, 14.3, 14.5
- 2 groups of means
 - \bar{x}_5, \bar{x}_3 are not significantly different from each other
 - $\bar{x}_2, \bar{x}_4, \bar{x}_1$ are not significantly different from each other

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Example 3

- For another data set:

- $\bar{x}_1 = 79.28, \bar{x}_2 = 61.54,$
 $\bar{x}_3 = 47.92, \bar{x}_4 = 32.76$

- $I = 4, J = 5$

$$HSD_{\alpha} = q_{\alpha, I, I(J-1)} \sqrt{\frac{MSE}{J}} = q_{0.05, 4, 16} \sqrt{\frac{92.9625}{5}} = 4.05 \sqrt{\frac{92.9625}{5}} = 17.47$$

$\bar{x}_4 \quad \bar{x}_3 \quad \bar{x}_2 \quad \bar{x}_1$

- Sort sample means: 32.76, 47.92, 61.54, 79.28

Although \bar{x}_4, \bar{x}_2 are different from each other, they are not different from \bar{x}_3

- 2 groups of means

- $\bar{x}_4, \bar{x}_3, \bar{x}_2$ are not significantly different from each other
- \bar{x}_1 is significantly different from each other

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Example 3

- For another data set:

- $\bar{x}_1 = 79.28, \bar{x}_2 = 61.54, \bar{x}_3 = 47.92, \bar{x}_4 = 32.76$

$$HSD_{\alpha} = 4.05 \sqrt{\frac{92.9625}{5}} = 17.47$$

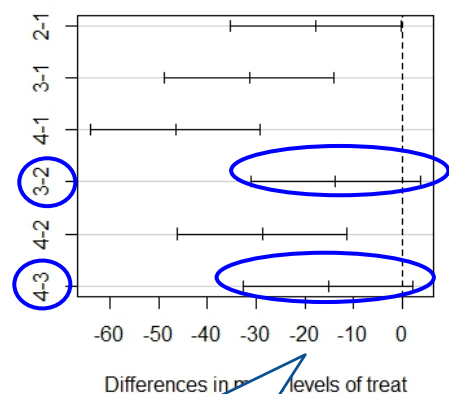
- Sort sample means:

$\bar{x}_4 \quad \bar{x}_3 \quad \bar{x}_2 \quad \bar{x}_1$
32.76, 47.92, 61.54, 79.28

- 2 groups of means

- $\bar{x}_4, \bar{x}_3, \bar{x}_2$ are not significantly different from each other
- \bar{x}_1 is significantly different from each other

95% family-wise confidence level



Confidence intervals of $(\bar{x}_2 - \bar{x}_3)$ and $(\bar{x}_3 - \bar{x}_4)$ include zero. This means $\bar{x}_2 \approx \bar{x}_3 \approx \bar{x}_4$

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References

1. J.L. Devore and K.N.Berk, Modern Mathematical Statistics with Applications, Springer, 2012.
2. S. Few, Now You See It: Simple Visualization Techniques for Quantitative Analysis, Analytics Press, 2009

