

Analysis of Variance

Single Factor

Part 2

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Statistics in Computer Engineering Applications
Slide 1



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Outline

- Single-Factor ANOVA:
 - Review: Equal sample size
 - F-test
 - Multiple comparison
 - Unequal Sample Size
 - F-test
 - Multiple comparison
 - Random Effects Model

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Single-Factor ANOVA : Equal Sample Size

- Compare **two or more** populations on one factor

- Let

- μ_1 = mean of treatment (population) 1
- μ_2 = mean of treatment (population) 2
- ...
- μ_I = mean of treatment (population) I

I = number of compared treatments

Two population -> Can perform hypothesis test on two data sets

- Hypothesis

- $H_0: \mu_1 = \mu_2 = \dots = \mu_I$
- H_a : Not all μ_i 's are equal (at least two of the μ_i 's are different)

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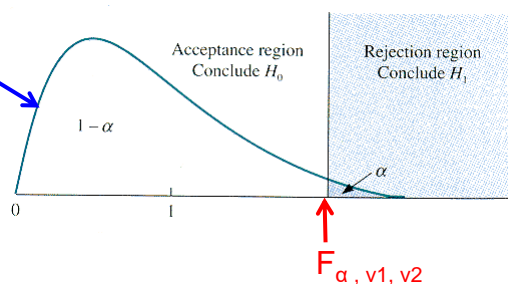
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F-Test

- Test statistic:

$$F_{I-1, I(J-1)} = \frac{\frac{SSTr}{\sigma^2} / (I-1)}{\frac{SSE}{\sigma^2} / (I(J-1))} = \frac{SSTr / (I-1)}{SSE / (I(J-1))} = \frac{MSTr}{MSE}$$

F distribution
for v_1 and v_2



If test statistic $> F_{\alpha, v_1, v_2}$, reject H_0

Image source: <http://www.unc.edu/~nielsen/soci708/m16/m2009.gif>

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Summary : ANOVA

- Test statistic:

$$F_{I-1,I(J-1)} = \frac{\frac{SSTr}{\sigma^2} / (I - 1)}{\frac{SSE}{\sigma^2} / (I(J - 1))} = \frac{SSTr / (I - 1)}{SSE / (I(J - 1))} = \frac{MSTr}{MSE}$$

- When f is large, we are about to reject H₀

	df	Sum of Squares (SS)	Mean Square (MS)	f
Treatment	I-1	SSTr	SSTr/(I-1)	MSTr / MSE
Error	I(J-1)	SSE	SSE/(I(J-1))	
Total	IJ-1	SST		



Summary : ANOVA (cont.)

$$\sum_i \sum_j (x_{ij} - \bar{x})^2 = \sum_i \sum_j (x_{ij} - \bar{x}_i + \bar{x}_i - \bar{x})^2$$

- Test statistic:

$$F_{I-1,I(J-1)} = \frac{\frac{SSTr}{\sigma^2} / (I - 1)}{\frac{SSE}{\sigma^2} / (I(J - 1))} = \frac{SSTr / (I - 1)}{SSE / (I(J - 1))} = \frac{MSTr}{MSE}$$

	df	Sum of Squares (SS)	Mean Square (MS)	f
Treatment	I-1	$J \sum_i (\bar{x}_i - \bar{x})^2$	SSTr/(I-1)	MSTr / MSE
Error	I(J-1)	$\sum_i \sum_j (x_{ij} - \bar{x}_i)^2 = SST - SSTr$	SSE/(I(J-1))	
Total	IJ-1	$\sum_i \sum_j (x_{ij} - \bar{x})^2$		



Summary : ANOVA (cont.)

- Test statistic:

$$F_{I-1, I(J-1)} = \frac{\frac{SSTr}{\sigma^2} / (I-1)}{\frac{SSE}{\sigma^2} / (I(J-1))} = \frac{SSTr / (I-1)}{SSE / (I(J-1))} = \frac{MSTr}{MSE}$$

Another option:
Sample-based
computation

	df	Sum of Squares (SS)	Mean Square (MS)	f
Treatment	I-1	$\frac{1}{J} \sum_{i=1}^I (\sum_{j=1}^J x_{ij})^2 - \frac{1}{IJ} (\sum_{i=1}^I \sum_{j=1}^J x_{ij})^2$	$SSTr / (I-1)$	$MSTr / MSE$
Error	I(J-1)	$SST - SSTr$	$SSE / (I(J-1))$	
Total	IJ-1	$\sum_{i=1}^I \sum_{j=1}^J x_{ij}^2 - \frac{1}{IJ} (\sum_{i=1}^I \sum_{j=1}^J x_{ij})^2$		

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Example 1

- Experiment degree of soiling on 3 mixtures of fabric and polymer
- Prove whether 3 mixture means are the same at $\alpha = 0.01$

Mixture	Degree of soiling				
1:	0.56	1.12	0.90	1.07	0.94
2:	0.72	0.69	0.87	0.78	0.91
3:	0.62	1.08	1.07	0.99	0.93

- Let
 - μ_1 = mean of mixture 1
 - μ_2 = mean of mixture 2
 - μ_3 = mean of mixture 3
 - $I = 3, J = 5$
- Hypothesis
 - $H_0: \mu_1 = \mu_2 = \mu_3$
 - H_a : Not all μ_i 's are equal (at least two of the μ_i 's are different)

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Example 1 (cont.)

	df	Sum of Squares (SS)	Mean Square (MS)	f
Treatment	2	0.0608	0.0304	0.99
Error	12	0.3701 = 0.4309 – 0.0608	0.0308	
Total	14	0.4309		

- Rejection region: given $\alpha = 0.01$
 - $F_{0.01, 2, 12} = 6.93$
- H_0 is not rejected.
- All mixture means are equals.



Example 2

- Experiment consistency lab measurements from 7 labs

Lab 1	Lab 2	Lab 3	Lab 4	Lab 5	Lab 6	Lab 7
4.13	3.86	4.00	3.88	4.02	4.02	4.00
4.07	3.85	4.02	3.88	3.95	3.86	4.02
4.04	4.08	4.01	3.91	4.02	3.96	4.03
4.07	4.11	4.01	3.95	3.89	3.97	4.04
4.05	4.08	4.04	3.92	3.91	4.00	4.10
4.04	4.01	3.99	3.97	4.01	3.82	3.81
4.02	4.02	4.03	3.92	3.89	3.98	3.91
4.06	4.04	3.97	3.90	3.89	3.99	3.96
4.10	3.97	3.98	3.97	3.99	4.02	4.05
4.04	3.95	3.98	3.90	4.00	3.93	4.06



Example 2

- Experiment consistency lab measurements from 7 labs
- Let
 - μ_1 = mean of measurements from lab 1
 - μ_2 = mean of measurements from lab 2
 - ...
 - μ_7 = mean of measurements from lab 7
 - $I = 7, J = 10$
- Hypothesis
 - $H_0: \mu_1 = \mu_2 = \dots = \mu_7$
 - H_a : Not all μ_i 's are equal (at least two of the μ_i 's are different)

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Example 2

- Experiment consistency lab measurements from 7 labs

Lab 1	Lab 2	Lab 3	Lab 4	Lab 5	Lab 6	Lab 7
4.13	3.86	4.00	3.88	4.02	4.02	4.00
4.07	3.85	4.02	3.88	3.95	3.86	4.02
4.04	4.08	4.01	3.91	4.02	3.96	4.03
4.07	4.11	4.01	3.95	3.89	3.97	4.04
4.05	4.08	4.04	3.92	3.91	4.00	4.10
4.04	4.01	3.99	3.97	4.01	3.82	3.81
4.02	4.02	4.03	3.92	3.89	3.98	3.91
4.06	4.04	3.97	3.90	3.89	3.99	3.96
4.10	3.97	3.98	3.97	3.99	4.02	4.05
4.04	3.95	3.98	3.90	4.00	3.93	4.06
\bar{x}_1	\bar{x}_2	\bar{x}_3	\bar{x}_4	\bar{x}_5	\bar{x}_6	\bar{x}_7
4.062	3.997	4.003	3.92	3.957	3.955	3.998

$\bar{x} = 3.985$

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Example 2 (cont.)

7 labs. Each lab collects 10 samples.

	df	Sum of Squares (SS)	Mean Square (MS)	f
Treatment	6	0.125	0.0208	5.66
Error	63	0.356-0.125 = 0.231	0.0037	
Total	69	0.356		

- There is no $F_{\alpha, 6, 63}$ in the f-table
- Examine closest F critical values: 60 vs. 70
 - $F_{0.1, 6, 60} = 1.8747$, $F_{0.05, 6, 60} = 2.2541$, $F_{0.01, 6, 60} = 3.1187$
 - $F_{0.1, 6, 70} = 1.8600$, $F_{0.05, 6, 70} = 2.2312$, $F_{0.01, 6, 70} = 3.0712$
 - For common values of α , $F_{\alpha, 6, 63} < f$
- H_0 is rejected.
- There are differences in 7 means.

```
> qf(0.99, df1=6, df2=63)
[1] 3.102767
> qf(0.95, df1=6, df2=63)
[1] 2.246408
> qf(0.90, df1=6, df2=63)
[1] 1.869819
>
```

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Example 2 (cont.)

7 labs. Each lab collects 10 samples.

	df	Sum of Squares (SS)	Mean Square (MS)	f
Treatment	6	0.125	0.0208	5.66
Error	63	0.356		
Total	69	0.356		

- Reject H_0 for large test statistic -> choose small α
- Do not reject H_0 for small test statistic -> choose large α

- There is no $F_{\alpha, 6, 63}$ in the f-table
- Examine closest F critical values: 60 vs. 70
 - $F_{0.1, 6, 60} = 1.8747$, $F_{0.05, 6, 60} = 2.2541$, $F_{0.01, 6, 60} = 3.1187$
 - $F_{0.1, 6, 70} = 1.8600$, $F_{0.05, 6, 70} = 2.2312$, $F_{0.01, 6, 70} = 3.0712$
 - For common values of α , $F_{\alpha, 6, 63} < f$
- H_0 is rejected.
- There are differences in 7 means.

How to choose α for
rejecting null
hypothesis?
Small or Large?

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Outline

- Single-Factor ANOVA:
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 - Unequal Sample Size
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 - Random Effects Model



Multiple Comparisons

- Question:
 - When H_0 for ANOVA is rejected, how many means are different from each other?
- Procedures:
 - Find confidence interval of pairwise difference $\mu_i - \mu_j$
 - If **confidence interval for** any pairwise difference $\mu_i - \mu_j$ **does not include zero**, we determine that μ_i, μ_j are significantly different from each other



Studentized Range Distribution (cont.)

$$\begin{aligned}
 1 - \alpha &= P\left(\frac{|\bar{X}_i - \bar{X}_j - (\mu_i - \mu_j)|}{\sqrt{MSE/J}} \leq Q_{\alpha, I, I(J-1)} \text{ for all } i, j\right) \\
 &= P\left(-Q_{\alpha, I, I(J-1)} \sqrt{\frac{MSE}{J}} \leq \bar{X}_i - \bar{X}_j - (\mu_i - \mu_j) \leq Q_{\alpha, I, I(J-1)} \sqrt{\frac{MSE}{J}} \text{ for all } i, j\right) \\
 &= P\left(\underbrace{\bar{X}_i - \bar{X}_j - Q_{\alpha, I, I(J-1)} \sqrt{\frac{MSE}{J}} \leq \mu_i - \mu_j \leq \bar{X}_i - \bar{X}_j + Q_{\alpha, I, I(J-1)} \sqrt{\frac{MSE}{J}}}_{\text{Confidence intervals between one pair of true mean difference } \mu_i - \mu_j} \text{ for all } i, j\right)
 \end{aligned}$$

Confidence intervals between one pair of true mean difference $\mu_i - \mu_j$

- There are $\binom{I}{2} = \frac{I(I-1)}{2}$ confidence intervals of $\mu_i - \mu_j$

Studentized Range Distribution (cont.)

$$1 - \alpha = P\left(\underbrace{\bar{X}_i - \bar{X}_j - Q_{\alpha, I, I(J-1)} \sqrt{\frac{MSE}{J}}}_{\text{Lower bound}} \leq \mu_i - \mu_j \leq \underbrace{\bar{X}_i - \bar{X}_j + Q_{\alpha, I, I(J-1)} \sqrt{\frac{MSE}{J}}}_{\text{Upper bound}} \text{ for all } i, j\right)$$

Expect difference between sample mean difference ($\bar{X}_i - \bar{X}_j$) and

true mean difference ($\mu_i - \mu_j$) is not more than this value $Q_{\alpha, I, I(J-1)} \sqrt{\frac{MSE}{J}}$

- The value $w = Q_{\alpha, I, I(J-1)} \sqrt{\frac{MSE}{J}}$ is called
Tukey's honestly significantly difference (HSD)

Example 1

- Experiment degree of soiling on 3 mixtures
- Group mixture means at $\alpha = 0.01$

	df	Sum of Squares (SS)	Mean Square (MS)	f
Treatment	2	0.0608	0.0304	0.99
Error	12	0.3701 = 0.4309 - 0.0608	0.0308	
Total	14	0.4309		

Mixture	Degree of soiling					\bar{x}_i
1:	0.56	1.12	0.90	1.07	0.94	0.918
2:	0.72	0.69	0.87	0.78	0.91	0.794
3:	0.62	1.08	1.07	0.99	0.93	0.938

$$HSD_{\alpha} = q_{\alpha, I, I(J-1)} \sqrt{\frac{MSE}{J}} = q_{0.01, 3, 12} \sqrt{\frac{0.0308}{5}} = 5.05 \sqrt{\frac{0.0308}{5}} = 0.396$$

$\bar{x}_2 \quad \bar{x}_1 \quad \bar{x}_3$

- Sort sample means: 0.794, 0.918, 0.938

0.124 0.020

```
> qtuke(0.01, nmeans=anovaResult$rank,
+       df=anovaResult$df.residual, lower.tail=F)
[1] 5.045934
```

- One group of mixture means

Note that H_0 is not rejected.

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Example 1

- Experiment degree of soiling on 3 mixtures
- Group mixture means at $\alpha = 0.01$

```
> tk <- TukeyHSD(anovaResult, conf.level = 0.99)
> plot(tk)
```

Mixture	Degree of soiling					\bar{x}_i
1:	0.56	1.12	0.90	1.07	0.94	0.918
2:	0.72	0.69	0.87	0.78	0.91	0.794
3:	0.62	1.08	1.07	0.99	0.93	0.938

- Sort sample means:

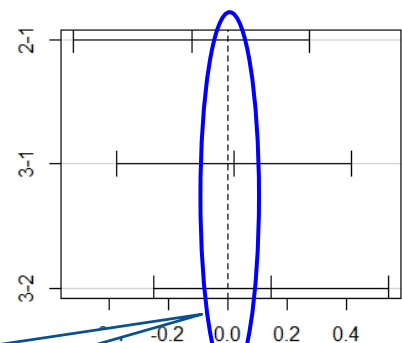
$\bar{x}_2 \quad \bar{x}_1 \quad \bar{x}_3$

0.794, 0.918, 0.938

0.124 0.020

- One group of mixture means

99% family-wise confidence level



Confidence interval of difference ($\bar{x}_i - \bar{x}_j$) includes zero. This means $\bar{x}_i \approx \bar{x}_j$

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Example 3

- For another data set:

$$\bar{x}_1 = 79.28, \bar{x}_2 = 61.54, \bar{x}_3 = 47.92, \bar{x}_4 = 32.76$$

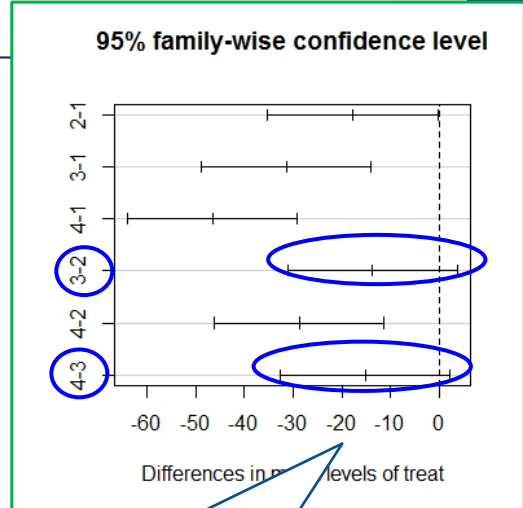
$$HSD_\alpha = 4.05 \sqrt{\frac{932.9625}{5}} = 17.47$$

- Sort sample means:

$\bar{x}_4 \quad \bar{x}_3 \quad \bar{x}_2 \quad \bar{x}_1$
32.76, 47.92, 61.54, 79.28

- 2 groups of means

- $\bar{x}_4, \bar{x}_3, \bar{x}_2$ are not significantly different from each other
- \bar{x}_1 is significantly different from each other



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Confidence intervals of $(\bar{x}_2 - \bar{x}_3)$ and $(\bar{x}_3 - \bar{x}_4)$ include zero. This means $\bar{x}_2 \approx \bar{x}_3 \approx \bar{x}_4$

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More on Multiple Comparison

- Is it possible that H_0 is not rejected but result to multiple group of means?
- Is it also possible that H_0 is rejected but have one group of means?

Measured by MSE

ANOVA tests on ALL means whether they are identical

Multiple comparison tests on PAIRWISE means

Not measured by MSE

ANOVA detects variability among all means

ANOVA test is more sensitive than multiple comparison

More on Multiple Comparison (cont.)

ANOVA tests on ALL means whether they are identical

Multiple comparison tests on PAIRWISE means

ANOVA detects lower variability among all means

ANOVA test is more sensitive than multiple comparison

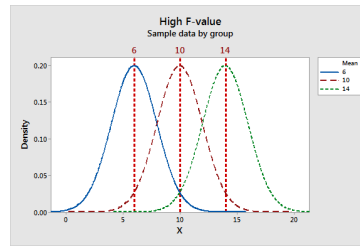
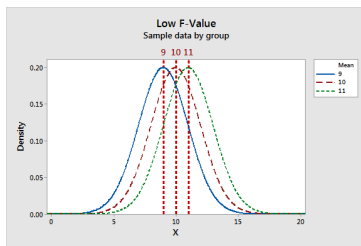


Image source: <http://blog.minitab.com/blog/adventures-in-statistics-2/understanding-analysis-of-variance-anova-and-the-f-test>

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Outline

- Single-Factor ANOVA:
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Example 1

- Measure strength of Mg-based alloys from 3 processes
- Show that alloys from 3 processes are the same at $\alpha = 0.001$

Process	Observations							
Permanent Molding (1)	45.5	45.3	45.4	44.4	44.6	43.9	44.6	44.0
Die casting (2)	44.2	43.9	44.7	44.2	44.0	43.8	44.6	43.1
Plaster molding (3)	46.0	45.9	44.8	46.2	45.1	45.5		

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Single ANOVA with Unequal Size

- ANOVA table:
 - N = Number of all samples = $\sum_{i=1}^I J_i$
 - I = Number of treatments
 - J_i = Number of samples for treatment i , $i = 1, 2, \dots, I$

$$SST_r = \sum_{i=1}^I \sum_{j=1}^{J_i} (\bar{x}_{ij} - \bar{x})^2 = \sum_{i=1}^I \frac{1}{J_i} \left(\sum_{j=1}^{J_i} X_{ij} \right)^2 - \frac{1}{n} \left(\sum_{i=1}^I \sum_{j=1}^{J_i} X_{ij} \right)^2 \quad \text{df} = I-1$$

$$SST = \sum_{i=1}^I \sum_{j=1}^{J_i} (x_{ij} - \bar{x})^2 = \sum_{i=1}^I \sum_{j=1}^{J_i} X_{ij}^2 - \frac{1}{n} \left(\sum_{i=1}^I \sum_{j=1}^{J_i} X_{ij} \right)^2 \quad \text{df} = N-1$$

$$SSE = \sum_{i=1}^I \sum_{j=1}^{J_i} (x_{ij} - \bar{x}_i)^2 = SST - SST_r \quad \text{df} = N-I$$

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Single ANOVA with Unequal Size

- ANOVA table:

- N = Number of all samples = $\sum_{i=1}^I J_i$
- I = Number of treatments
- J_i = Number of samples for treatment i, $i = 1, 2, \dots, I$

	df	Sum of Squares (SS)	Mean Square (MS)	f
Treatment	$I - 1$	$\sum_{i=1}^I \sum_{j=1}^{J_i} (\bar{x}_{ij} - \bar{\bar{x}})^2$	$SSTr/(I-1)$	$MSTr / MSE$
Error	$N - I$	$SSE = \sum_{i=1}^I \sum_{j=1}^{J_i} (x_{ij} - \bar{x}_i)^2$ $= SST - SSTr$	$SSE/(N - I)$	
Total	$N - 1$	$\sum_{i=1}^I \sum_{j=1}^{J_i} (x_{ij} - \bar{\bar{x}})^2$		

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Single ANOVA with Unequal Size

- ANOVA table:

- N = Number of all samples = $\sum_{i=1}^I J_i$
- I = Number of treatments
- J_i = Number of samples for treatment i, $i = 1, 2, \dots, I$

Another option:
Sample-based
computation

	df	Sum of Squares (SS)	Mean Square (MS)	f
Treatment	$I - 1$	$\sum_{i=1}^I \frac{1}{J_i} (\sum_{j=1}^{J_i} x_{ij})^2 - \frac{1}{n} (\sum_{i=1}^I \sum_{j=1}^{J_i} x_{ij})^2$	$SSTr/(I-1)$	$MSTr / MSE$
Error	$N - I$	$SSE = \sum_{i=1}^I \sum_{j=1}^{J_i} (x_{ij} - \bar{x}_i)^2$ $= SST - SSTr$	$SSE/(N - I)$	
Total	$N - 1$	$\sum_{i=1}^I \sum_{j=1}^{J_i} x_{ij}^2 - \frac{1}{n} (\sum_{i=1}^I \sum_{j=1}^{J_i} x_{ij})^2$		

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Example 1

- Measure strength of Mg-based alloys from 3 processes
- Show that alloys from 3 processes are the same at $\alpha = 0.001$

Process	Observations							
Permanent Molding (1)	45.5	45.3	45.4	44.4	44.6	43.9	44.6	44.0
Die casting (2)	44.2	43.9	44.7	44.2	44.0	43.8	44.6	43.1
Plaster molding (3)	46.0	45.9	44.8	46.2	45.1	45.5		

• Let

- μ_1 = mean of alloy 1, μ_2 = mean of alloy 2, μ_3 = mean of alloy 3
- $I = 3, J_1 = J_2 = 8, J_3 = 6, N = 22$

• Hypothesis

- $H_0: \mu_1 = \mu_2 = \mu_3$
- H_a : Not all μ_i 's are equal (at least two of the μ_i 's are different)

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Example 1 (cont.)

Process	Observations								$\sum_{j=1}^{J_i} x_{ij}$
Permanent Molding (1)	45.5	45.3	45.4	44.4	44.6	43.9	44.6	44.0	357.7
Die casting (2)	44.2	43.9	44.7	44.2	44.0	43.8	44.6	43.1	352.5
Plaster molding (3)	46.0	45.9	44.8	46.2	45.1	45.5			273.5

- Fill ANOVA table

	df	Sum of Squares (SS)	Mean Square (MS)	f
Treatment	2	7.93	3.97	12.56
Error	19	13.93 - 7.93 = 6.00	0.32	
Total	21	13.93		

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Example 1 (cont.)

	df	Sum of Squares (SS)	Mean Square (MS)	f
Treatment	2	7.93	3.97	12.56
Error	19	13.93 – 7.93 = 6.00	0.32	
Total	21	13.93		

- Rejection region: given $\alpha = 0.001$
 - $F_{0.001, 2, 19}$
 - H_0 is rejected.
 - Alloys from 3 processes are not the same.

Not exist in f-table.
Use R.

```
> qf (0.999, df1=2, df2=19)
[1] 10.15681
>
```

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Statistics in Computer Engineering Applications
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Outline

- Single-Factor ANOVA:
 - Review: Equal sample size
 - F-test
 - Multiple comparison
 - Unequal Sample Size
 - F-test
 - Multiple comparison
 - Random Effects Model

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Multiple Comparisons : Unequal Sample Size

- When some means are not all equal, how to specify which mean is different from others

- Procedure

- Find Tukey's Honestly Significant Difference (HSD)

$$HSD_{\alpha} = q_{\alpha, I, N-I} \sqrt{\frac{MSE}{2} \left(\frac{1}{J_i} + \frac{1}{J_j} \right)}$$

- $q_{\alpha, I, N-I}$ = q-value from studentized range distribution with 2 degrees of freedom $I, N-I$
- Sort sample means in increasing order
- Underline pairs that differ less than HSD_{α}
- Any pair without underline are considered as significantly different.

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Example 1 : Mg-based Alloy (cont.)

$$HSD_{\alpha, ij} = q_{\alpha, I, N-I} \sqrt{\frac{MSE}{2} \left(\frac{1}{J_i} + \frac{1}{J_j} \right)}$$

- $\bar{x}_1 = 44.71, \bar{x}_2 = 44.06, \bar{x}_3 = 45.58$

$$w_{12} = HSD_{\alpha} = q_{\alpha, I, N-I} \sqrt{\frac{MSE}{2} \left(\frac{1}{J_i} + \frac{1}{J_j} \right)} = q_{0.05, 3, 19} \sqrt{\frac{0.32}{2} \left(\frac{1}{8} + \frac{1}{8} \right)} = 3.59 \sqrt{\frac{0.32}{2} \left(\frac{1}{8} + \frac{1}{8} \right)} = 0.718$$

$$w_{13} = w_{23} = HSD_{\alpha} = q_{\alpha, I, N-I} \sqrt{\frac{MSE}{2} \left(\frac{1}{J_i} + \frac{1}{J_j} \right)} = q_{0.05, 3, 19} \sqrt{\frac{0.32}{2} \left(\frac{1}{8} + \frac{1}{6} \right)} = 3.59 \sqrt{\frac{0.32}{2} \left(\frac{1}{8} + \frac{1}{6} \right)} = 0.775$$

- Sort sample means: $\bar{x}_2 \quad \bar{x}_1 \quad \bar{x}_3$
 $\quad \quad \quad 44.06 \quad 44.71 \quad 45.58$
 $\quad \quad \quad \underline{0.65} \quad \underline{0.87}$

- 2 groups of means

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Example 1 : Mg-based Alloy (cont.)

- $\bar{x}_1 = 44.71, \bar{x}_2 = 44.06, \bar{x}_3 = 45.58$

$$w_{12} = HSD_\alpha = q_{\alpha, I, N-I} \sqrt{\frac{MSE}{2} \left(\frac{1}{j_i} + \frac{1}{j_j} \right)} = q_{0.05, 3, 19} \sqrt{\frac{0.32}{2} \left(\frac{1}{8} + \frac{1}{8} \right)} = 3.50 \sqrt{0.32 \left(\frac{1}{8} + \frac{1}{8} \right)} = 0.719$$

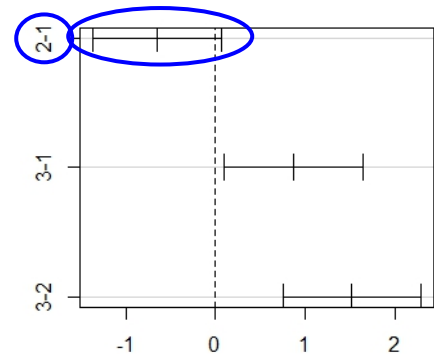
$$w_{13} = w_{23} = HSD_\alpha = q_{\alpha, I, N-I} \sqrt{\frac{MSE}{2} \left(\frac{1}{j_i} + \frac{1}{j_j} \right)} = q_{0.05, 3, 19} \sqrt{\frac{0.32}{2} \left(\frac{1}{8} + \frac{1}{6} \right)}$$

- Sort sample means:

$\bar{x}_2 \quad \bar{x}_1 \quad \bar{x}_3$
44.06, 44.71, 45.58

- 2 groups of means

95% family-wise confidence level



Differences in mean levels of process

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Random Effects Model

- Single factor can be considered as fixed-effects ANOVA model
- The single-factor fixed-effects model is defined as

$$X_{ij} = \mu + \alpha_i + \varepsilon_{ij}$$

$$\sum \alpha_i = 0$$

- X_{ij} = random sample j of treatment i
- μ = overall mean of all treatment i 's
- α_i = effect of treatment i
- ε_{ij} = random error in sample j of treatment i
 - Assumed to be independent and normally distributed with mean = 0, variance = σ^2

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Random Effects Model (cont.)

$$X_{ij} = \mu + \alpha_i + \varepsilon_{ij}$$

$$\sum \alpha_i = 0$$

- Explanations of the single-factor fixed-effects model:
 - Sample is corrupted by random errors
 - Error in one sample is independent from error of other samples
 - Expected response of treatment i

$$E(X_{ij}) = \mu + \alpha_i$$

- If $\alpha_i = 0$, then all treatment i 's have the same response

$$E(X_{ij}) = \mu$$



That's what we try to prove in
single-factor ANOVA

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Null Hypothesis: Single-Factor ANOVA

$$E(X_{ij}) = \mu + \alpha_i$$

- Besides the prior null hypothesis in single-factor ANOVA
 - $H_0: \mu_1 = \mu_2 = \dots = \mu_I$
- Sometimes, the following null hypothesis is also used instead:
 - $H_0: \alpha_1 = \alpha_2 = \dots = \alpha_I = 0$



References

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