Willingness-to-pay for Warnings: Main Tables

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Research Question

- How much do people value alerts (signals) about potential preventable threats?
- How do signal's probabilistic characteristics affect the willingness-to-pay for it and the welfare gains from using it?
- Applications:
 - Natural disaster warnings (tornados, floods, earthquakes)
 - Medical tests for treatable conditions
 - Investing in research on likelihood of catastrophic events (rogue Al, global warming, pandemics)
- Note: most real-life applications provide little practice with using the signal



Overview of the Experiment

- An insurance experiment:
 - ullet Two states of the world: bad $(\omega=1)$ and good $(\omega=0)$
 - Probability of a bad state is $P(\omega = 1) = \pi$
 - Bad state \implies loss of \$L
 - ullet A perfectly protective insurance can be purchased for $\$
- Subject can purchase a signal s before purchasing the insurance:
 - A signal is characterized by its true-positive ($P(s=1|\omega=1)$) and true-negative rates ($P(s=0|\omega=0)$)

Research objective

How do signal characteristics affect the WTP?



- Theoretically, what should be the WTP for a signal?
- If bad states are a priori rare ($\pi L << c$) \implies never protect without a signal
- The theoretical WTP b for an expected utility maximizer given a signal s is a solution b^* to the following:

$$P(s=1)u(Y_0 - b^* - c) + \pi P(s=0|\omega=1)u(Y_0 - b^* - L) +$$
$$+(1-\pi)P(s=0|\omega=0)u(Y_0 - b^*) = (1-\pi)u(Y_0) + \pi u(Y_0 - L)$$

A risk-neutral agent then pays:

$$b^* = \pi(1 - P(s = 0 | \omega = 1))L - P(s = 1)c$$

- The formulas become more complicated if subjects can protect without a signal (bad state are not rare enough)
- The theoretical WTP b for an expected utility maximizer given a signal s is a solution b^* to the following:

$$P(s=1)u(Y_0 - b^* - c) + \pi P(0|1)u(Y_0 - b^* - L) + (1 - \pi)P(0|0)u(Y_0 - b^*) =$$

$$= \min[(1 - \pi)u(Y_0) + \pi u(Y_0 - L), u(Y_0 - c)]$$

A risk-neutral agent then pays:

$$b^* = \min[\pi L, c] - \pi(1 - P(s = 0 | \omega = 1))L - P(s = 1)c$$

Hypotheses

- Conditional on the signal's value for risk-neutral subjects, false positive and false negative rates reduce the perceived value of the signal (WTP)
 - The opposite is true: subjects underreact to false positive and false negative rates and overpay for bad signals
- 2 Conditional on the signal's value for risk-neutral subjects, false positive and false negative rates increase expected costs
 - No: FP and FN rates have no significant effects on costs besides their predicted theoretical effect
- Extra: how much of these disrepancies result from belief updating issues or risk aversion?



Risk Aversion Measurement

- Measure risk aversion based on blind protection choices:
 - Exclude obs from subjects switching back and forth
 - \bullet The lowest probability for which a subject chooses to protect is π^*
 - ullet Calculate their coefficient of relative risk aversion heta as the solution to the following equation:

$$\pi^* u(Y_0 - L; \theta) + (1 - \pi^*) u(Y_0; \theta) = u(Y_0 - c; \theta)$$

• Where u() is the CRRA utility function:

$$u(x;\theta) = \frac{x^{1-\theta} - 1}{1 - \theta}$$

• Note: risk lovers have $\theta < 0$



CRRA Estimates

• Most subjects are moderately risk averse:

| Probability (π^*) | θ | N |
|-----------------------|----------|----|
| Always protect | >2 | 1 |
| 0.1 | 2 | 2 |
| 0.15 | 1.216 | 7 |
| 0.2 | 0.573 | 17 |
| 0.25 | 0 | 7 |
| 0.3 | -0.539 | 5 |
| Never protect | <-0.539 | 8 |
| | | |

Note:

- There are 18 subjects (out of 65) switching multiple times
- Can use more sophisticated methods to measure risk aversion for those



WTP for the Signal

Theoretical value of signal for risk-neutral subject:

$$b^* = \underbrace{\min[\pi L, c]}_{\text{BP costs}} - \underbrace{\pi(1 - P(s = 0 | \omega = 1))L}_{\text{False neg. costs}} - \underbrace{P(s = 1)c}_{\text{Protection costs}}$$

- Two potential approaches:

$$V - b^* = \alpha_0 + \alpha_1 \text{FN costs} + \alpha_2 \text{Prot. costs} + \epsilon$$

Regress WTP directly on its components and account for censoring at 0:

$$V = \min[0, \beta_0 + \beta_1 \mathsf{FN} \; \mathsf{costs} + \beta_2 \mathsf{Prot}. \; \mathsf{costs} - \beta_3 \mathsf{BP} \; \mathsf{costs} + \gamma]$$

Note: protection costs include costs due to false positive signals



WTP for the Signal (Approach 1)

 Coefficient sign. different from zero is an anomaly: people overpay for bad signals
 Figure: WTP for Information (Discrepancy)

| rigule. Will for information (Discrepancy) | | | | |
|--------------------------------------------|-------------|-------------|-------------|-----------|
| | (1) (2) (3) | | (3) | (4) |
| | All | Risk-averse | Risk-loving | Switchers |
| Prot. costs | .205** | .372** | 0232 | .183 |
| | (2.1) | (2.4) | (-0.1) | (1.0) |
| False neg. costs | .36*** | .259** | .519*** | .341** |
| | (4.4) | (2.0) | (3.1) | (2.6) |
| Constant | 543*** | 58* | 54 | 441 |
| | (-2.8) | (-1.9) | (-1.5) | (-1.2) |
| Observations | 390 | 156 | 126 | 108 |
| Adjusted R^2 | 0.05 | 0.05 | 0.07 | 0.04 |

t statistics in parentheses



 $^{^{\}ast}$ p < 0.10 , ** p < 0.05 , *** p < 0.01

WTP for the Signal (Approach 2, Tobit Estimation)

Coefficient should **differ from one** in abs. value to show an anomaly (ignore stars for now)
 Figure: WTP for Information (Tobit Estimation)

| Figure: WTP for Information (Tobit Estimation) | | | | | |
|------------------------------------------------|---------|-------------|-------------|-----------|--|
| | (1) | (2) (3) | | (4) | |
| | All | Risk-averse | Risk-loving | Switchers | |
| model | | | | | |
| BP costs | .562*** | .63*** | .557*** | .453** | |
| | (5.8) | (3.9) | (3.3) | (2.5) | |
| Prot. costs | 342** | 199 | 561** | 321 | |
| | (-2.5) | (-0.9) | (-2.3) | (-1.3) | |
| False neg. costs | 398*** | 533*** | 214 | 395** | |
| | (-3.9) | (-3.4) | (-1.0) | (-2.5) | |
| Constant | .141 | 115 | .0424 | .704 | |
| | (0.4) | (-0.2) | (0.1) | (1.2) | |
| sigma | | | | | |
| Constant | 1.78*** | 1.81*** | 1.84*** | 1.59*** | |
| | (24.8) | (15.5) | (12.9) | (13.4) | |
| Observations | 390 | 156 | 126 | 108 | |
| AIC | 1413.41 | 575.80 | 452.26 | 392.50 | |

t statistics in parentheses

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WTP for the Signal (Risk Aversion)

 Explaining the discrepancy between WTP and value with risk aversion:

| I igure. VV I | rigure: vv i P for information (different risk aversion) | | | | |
|-------------------------|----------------------------------------------------------|----------------|----------------|----------------|--|
| | (1) | (2) | (3) | (4) | |
| | Heterogeneous | $\theta = 0.5$ | $\theta = 1.0$ | $\theta = 1.5$ | |
| BP costs | 284*** | 39*** | 157** | .165** | |
| | (-2.7) | (-5.3) | (-2.1) | (2.2) | |
| Prot. costs | .726*** | .766*** | .686*** | .515*** | |
| | (4.8) | (6.8) | (6.1) | (4.6) | |
| False neg. costs | .602*** | .667*** | .743*** | .775*** | |
| | (5.1) | (7.7) | (8.6) | (8.9) | |
| Constant | 54 | 277 | -1.15*** | -2.21*** | |
| | (-1.5) | (-1.1) | (-4.7) | (-9.2) | |
| Observations | 228 | 390 | 390 | 390 | |
| Adjusted \mathbb{R}^2 | 0.14 | 0.17 | 0.22 | 0.30 | |

t statistics in parentheses



 $^{^{\}ast}$ p < 0.10 , ** p < 0.05 , *** p < 0.01

Actual Costs vs Theoretical Costs

- Calculate actual costs based on decisions made in the Informed Protection treatment and actual posterior probabilities of losses.
- Each reported participant's strategy s is a tuple of numbers (r_w, r_b) representing protection responses correspondingly to white and black hints
- Then the expected cost of each decision are:

$$EC(s) = \pi(P(0|1)(1 - r_w) + P(1|1)(1 - r_b))L + P(s = 1)c$$
$$+ (P(s = 0)r_w + P(s = 1)r_b)c$$

 Regress expected costs on minimal theoretical costs and other signal characteristics



Actual Costs vs Theoretical Costs

Figure: Actual Exp. Costs vs Theoretical Costs

| | (1) | (2) | (3) | (4) | (5) | (6) |
|-------------------------|----------|---------|--------|---------|----------|---------|
| | ÒĽS | ÒĽŚ | ÒĽŚ | ÈÉ | ÈÉ | ÈÉ |
| Optimal exp. costs | .958*** | .967*** | .956 | .987*** | 1.04*** | .938*** |
| | (9.2) | (8.2) | (1.6) | (17.5) | (6.3) | (11.4) |
| Prior prob. | | .102 | .0367 | | .632 | |
| | | (0.1) | (0.0) | | (0.3) | |
| False neg. rate | | | 106 | | | 111 |
| | | | (-0.0) | | | (-0.1) |
| False pos. rate | | | .0229 | | | 841 |
| | | | (0.0) | | | (-0.4) |
| Constant | -1.07*** | -1.08** | -1.07 | -1*** | -1.06*** | 999*** |
| | (-2.8) | (-2.2) | (-1.6) | (-7.9) | (-4.9) | (-8.4) |
| Observations | 390 | 390 | 390 | 390 | 390 | 390 |
| Adjusted \mathbb{R}^2 | 0.10 | 0.10 | 0.10 | 0.11 | 0.11 | 0.11 |

t statistics in parentheses



 $^{^{\}ast}$ p<0.10, ** p<0.05, *** p<0.01

Additional Complementary Tables

- Factors affecting informed protection responses
- The effect of beliefs on informed protection
- 4 How accurate are their beliefs?
- Oecomposition of belief updating: priors vs signals

Informed Protection: Determinants

| Figure: Informed Protection | | | | |
|-----------------------------|---------|----------|---------|---------|
| | (1) | (2) | (3) | (4) |
| | All | All | Smart | Smart |
| Informed protection | | | | |
| Posterior prob. | 2.14*** | .668** | 2.29*** | .708** |
| | (14.0) | (2.5) | (12.7) | (2.4) |
| Prior prob. | | 1.32*** | | 1.34*** |
| | | (3.4) | | (3.0) |
| Gremlin says Black | | 1.32*** | | 1.43*** |
| | | (6.3) | | (6.2) |
| Constant | 702*** | -1.12*** | 761*** | -1.2*** |
| | (-11.0) | (-8.6) | (-10.8) | (-8.2) |

780

736.93

780

774.36

Observations

AIC



636

573.32

636

608.24

t statistics in parentheses

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Informed Protection: Do Subject's Beliefs Matter?

| Figure: Informed Protection: Response to Reported Beliefs | | | | |
|-----------------------------------------------------------|---------|---------|----------|--|
| | (1) | (2) | (3) | |
| | All | All | Smart | |
| Informed protection | | | | |
| Belief | 2.49*** | 1.57*** | 1.71*** | |
| | (13.8) | (7.0) | (6.8) | |
| Posterior prob. | | 1.22*** | 1.23*** | |
| | | (6.9) | (5.8) | |
| Constant | 887*** | 985*** | -1.07*** | |
| | (-12.0) | (-12.8) | (-12.7) | |
| Observations | 780 | 780 | 636 | |
| AIC | 767.11 | 720.08 | 558.84 | |

t statistics in parentheses



 $^{^{\}ast}$ p<0.10 , ** p<0.05 , *** p<0.01

Belief Updating: Correlation

| Figure: Belief Elicitation: Belief vs Posterior | | | | |
|-------------------------------------------------|---------|---------------|-----------|--|
| | (1) | (2) | (3) | |
| | All | Not_honest | Good quiz | |
| Posterior prob. | .669*** | .711*** | .523*** | |
| | (29.0) | (29.1) | (15.7) | |
| Constant | .151*** | .147*** | .226*** | |
| | (16.2) | (14.1) | (17.8) | |
| Observations | 780 | 636 | 520 | |
| Adjusted ${\cal R}^2$ | 0.57 | 0.61 | 0.39 | |

t statistics in parentheses



 $^{^{\}ast}$ p < 0.10 , ** p < 0.05 , *** p < 0.01

Belief Updating: Decomposition

• Posterior probability $\mu = P(B|S=x)$ that the ball is black conditional on a hint S=x can be written as:

$$\ln\left(\frac{\mu}{1-\mu}\right) = \lambda_0 + S_B + S_W$$

- With $\lambda_0 \equiv \ln(p/(1-p))$ representing (transformed) prior beliefs
- And S_B , S_W describing the effect of new evidence:

$$S_B \equiv I(S = B) \ln(P(s = B|B)/P(s = B|W))$$

 $S_W \equiv I(S = W) \ln((1 - P(s = B|B))/(1 - P(s = B|W))$

Belief Updating: Decomposition

| Figure: Belief Elicitation: Decomposition | | | | |
|-------------------------------------------|--------|---------|-----------|--|
| | (1) | (2) | (3) | |
| | OLS | FE | Smart, FE | |
| lt_prior | .162** | .132** | .141** | |
| | (2.0) | (2.6) | (2.3) | |
| signalB | .477** | .937*** | 1.04*** | |
| | (2.2) | (5.0) | (4.6) | |
| signalW | .46** | 0 | 0 | |
| | (2.5) | (.) | (.) | |
| Constant | 0878 | 595*** | 637*** | |
| | (-0.4) | (-5.3) | (-4.8) | |
| Observations | 172 | 172 | 140 | |
| Adjusted \mathbb{R}^2 | 0.26 | 0.31 | 0.31 | |

t statistics in parentheses



^{*} p < 0.10, ** p < 0.05, *** p < 0.01