



ETC3580: Advanced Statistical Modelling

Week 11: Additive models

Outline

1 Additive models

2 Generalized additive models

Additive models

Avoid curse of dimensionality by assuming additive surface:

$$y = \beta_0 + \sum_{j=1}^p f_j(x_j) + \varepsilon$$

where $\varepsilon \sim N(0, \sigma^2)$.

- Restricts complexity but a much richer class of surfaces than parametric models.
- Need to estimate p one-dimensional functions instead of one p -dimensional function.
- Usually set each f_j to have zero mean.
- Some f_j may be linear.

Additive models

- Up to p different bandwidths to select.
- Generalization of multiple regression model

$$y = \beta_0 + \sum_{j=1}^p \beta_j x_j + \varepsilon$$

which is also additive in its predictors.

- Estimated functions, f_j , are analogues of coefficients in linear regression.
- Interpretation easy with additive structure.

Additive models

- Categorical predictors: fit constant for each level as for linear models.
- Allow interaction between two continuous variables x_j and x_k by fitting a bivariate surface $f_{j,k}(x_j, x_k)$.
- Allow interaction between factor x_j and continuous x_k by fitting separate functions $f_{j,k}(x_k)$ for each level of x_j .

Additive models in R

- gam package: more smoothing approaches, uses a backfitting algorithm for estimation.
- mgcv package: simplest approach, with automated smoothing selection and wider functionality.
- gss package: smoothing splines only

Estimation

Back-fitting-algorithm (Hastie and Tibshirani, 1990)

- 1 Set $\beta_0 = \bar{y}$.
- 2 Set $f_j(x) = \hat{\beta}_j x$ where $\hat{\beta}_j$ is OLS estimate.
- 3 For $j = 1, \dots, p, 1, \dots, p, 1, \dots, p, \dots$

$$f_j(x) = S(x_j, y - \beta_0 - \sum_{i \neq j} f_i(x_i))$$

where $S(x, u)$ means univariate smooth of u on x .

Iterate step 3 until convergence.

Estimation

Back-fitting-algorithm (Hastie and Tibshirani, 1990)

- 1 Set $\beta_0 = \bar{y}$.
- 2 Set $f_j(x) = \hat{\beta}_j x$ where $\hat{\beta}_j$ is OLS estimate.
- 3 For $j = 1, \dots, p, 1, \dots, p, 1, \dots, p, \dots$

$$f_j(x) = S(x_j, y - \beta_0 - \sum_{i \neq j} f_i(x_i))$$

where $S(x, u)$ means univariate smooth of u on x .

Iterate step 3 until convergence.

- S could be *any* univariate smoother.
- $y - \beta_0 - \sum_{i \neq j} f_i(x_i)$ is a “partial residual”

Estimation

Regression splines

No need for iterative back-fitting as the model can be written as a linear model.

Penalized regression splines

No need for iterative back-fitting as the model can be written as a linear mixed-effects model.

Inference for Additive Models

Each fitted function can be written as a linear smoother $\hat{\mathbf{f}}_j = \mathbf{S}_j \mathbf{y}$ for some $n \times n$ matrix \mathbf{S}_j .

$\hat{\mathbf{f}}(\mathbf{x})$ is a linear smoother. Denote smoothing matrix as \mathbf{S} :

$$\hat{\mathbf{f}}(\mathbf{x}) = \mathbf{S} \mathbf{y} = \beta_0 \mathbf{1} + \sum_{j=1}^p \mathbf{S}_j \mathbf{y}$$

where $\mathbf{1} = [1, 1, \dots, 1]^T$. Then $\mathbf{S} = \sum_{j=0}^p \mathbf{S}_j$ where \mathbf{S}_0 is such that $\mathbf{S}_0 \mathbf{y} = \beta_0 \mathbf{1}$.

Thus all inference results for linear smoothers may be applied to additive model.

Outline

1 Additive models

2 Generalized additive models

Generalised additive models

Generalized Linear Model (GLM)

- Distribution of y
- Link function g
- $E(y \mid x_1, \dots, x_p) = \mu$ where $g(\mu) = \beta_0 + \sum_{j=1}^p \beta_j x_j$.

Generalised additive models

Generalized Linear Model (GLM)

- Distribution of y
- Link function g
- $E(y \mid x_1, \dots, x_p) = \mu$ where $g(\mu) = \beta_0 + \sum_{j=1}^p \beta_j x_j$.

Generalised Additive Model (GAM)

- Distribution of y
- Link function g
- $E(y \mid x_1, \dots, x_p) = \mu$ where $g(\mu) = \beta_0 + \sum_{j=1}^p f_j(x_j)$.

Generalised additive models

Examples:

- Y binary and $g(\mu) = \log[\mu(1 - \mu)]$. This is a logistic additive model.
- Y normal and $g(\mu) = \mu$. This is a standard additive model.

Estimation

Hastie and Tibshirani describe method for fitting GAMs using a method known as “local scoring” which is an extension of the Fisher scoring procedure.

- Revision lecture on Monday
- No lecture on Tuesday

Real world projects: Business analytics with imperfect data

Presenter: Matt McInnes,
Managing Director - Asia Pacific, Lynxx Consulting

Thursday, 18 October 2018, 5:00pm to 6:30pm

Location: Learning and Teaching Building, Room G58