Notes of [1]

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1 Problem Setup

There is an *tabular episodic* MDP $\mathcal{M} = (\mathcal{S}, \mathcal{A}, \mathbb{P}, R, H)$ where only the transition probability \mathbb{P} is *unknown*. For simplicity, we also assume the reward function R is *deterministic*. We want to find a policy such that the regret incurred by this policy after K episodes is minimized. Given a policy $\pi = (\pi_1, \dots, \pi_K)$, the regret incurred by this policy is defined by

$$\mathcal{R}_K^{\pi} \coloneqq \mathbb{E}\left[\sum_{k=1}^K (V_1^* - V_1^{\pi_k})(x_{k,1})\right],$$

where the initial state for each episode can be either randomized or adversarial.

Remark 1. There exists an optimal policy which is Markov and deterministic (may depend on time $t \in [H]$).

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\{1,2,\ldots,n\}
[n]
\mathcal{A}
                               action space
A
                               |\mathcal{A}|
\mathcal{S}
                               state space
S
                               |\mathcal{S}|
H
                               horizon
K
                               # of episodes
T
                               HK
R: \mathcal{S} \times \mathcal{A} \rightarrow [0, 1]
                               known reward function
\mathbb{P}: \mathcal{S} \times \mathcal{A} \to \Delta(\mathcal{S})
                               transition probability of the underlying MDP
\pi = (\pi_1, \dots, \pi_K)
                               an arbitrary policy where \pi_k is the policy in the kth episode for k \in [K]
Q_h^{\pi_k}: \mathcal{S} \times \mathcal{A} \to \mathbb{R}V_h^{\pi_k}: \mathcal{S} \to \mathbb{R}
                               Q-value function of policy \pi_k at time h
                               value function of policy \pi_k at time h
Q_h^*
                               Q-value function of the optimal policy
V_h^*
                               value function of the optimal policy
                               initial state of the kth episode
x_{k,1}
(x_{k,h},a_{k,h})
                               state-action pair in the kth episode and at the hth step
                               history before the kth episode (x_{1,1}, a_{1,1}, \dots, x_{1,H+1}, \dots, x_{k-1,1}, a_{k-1,1}, \dots, x_{k-1,H+1})
\mathcal{H}_k
n_k: \mathcal{S} \times \mathcal{A} \to \mathbb{N}
                               number of hits of state-action pair before the kth episode
n_k(y \mid x, a)
                               number of hits of state y when taking action a at state x before the kth episode
\widehat{\mathbb{P}}_k
                               empirical transition probability using \mathcal{H}_k
\widetilde{Q}_{k,h}
                               optimistic estimate of the optimal Q-value function in the kth episode
\widetilde{V}_{k,h}
                               optimistic estimate of the optimal value function in the kth episode
                               an arbitrary transition probability
V
                               an arbitrary value function
(\rho V)(x,a)
                               \sum_{y \in \mathcal{S}} \rho(y \mid x, a) V(y)
                               regret incurred by policy \pi
\mathcal{R}^{\pi}_{K}
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2 Notations and Definitions

3 Algorithm

Algorithm 1: UCBVI-CH ([1])

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1 initialization: \widetilde{Q}_{1,h}(x,a) = H - h + 1 for every (h,x,a) \in [H] \times \mathcal{S} \times \mathcal{A}
2 for episode \ k = 1 to K do
3 | if k > 1 then
4 | call Algorithm 2 to compute \widetilde{Q}_{k,h}
5 | for step \ h = 1 to H do
6 | observe state x_{k,h}
7 | take action a_{k,h} = \operatorname{argmax}_{a \in \mathcal{A}} \widetilde{Q}_{k,h}(x_{k,h},a).
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Algorithm 2: Computation of $\widetilde{Q}_{k,h}$

Algorithm 3: Computation of $b_k(x, a)$

$$b_k(x,a) = c_1 H \sqrt{\frac{\ln(SAT/\delta)}{n_k(x,a)}}$$

Here $\overline{c_1}$ is a constant which will be defined when event \mathcal{E}_1 is defined.

4 Proofs

4.1 Favorable Events

4.1.1 \mathcal{E}_1

Given any $(x, a, t) \in \mathcal{S} \times \mathcal{A} \times [T]$, define *i.i.d.* random variables $X_1^{x,a}, \dots, X_t^{x,a}$ following the distribution $\mathbb{P}(x, a)$. Let

$$\mathcal{E}_1 \coloneqq \left\{ \forall (h, x, a, t) \in [H] \times \mathcal{S} \times \mathcal{A} \times [T], \left| \frac{\sum_{i=1}^t V_h^*(X_i^{x, a})}{t} - \sum_{y \in \mathcal{S}} \mathbb{P}(y \mid x, a) V_h^*(y) \right| \leq c_1 H \sqrt{\frac{\ln(SAT/\delta)}{t}} \right\},$$

where c_1 is a constant which will be defined later.

By Hoeffding's inequality (Lemma 9) and a union bound, there exists a constant c_1 such that $\mathbf{Pr}(\mathcal{E}_1) \ge 1 - \delta/4$.

4.1.2 \mathcal{E}_2

Given any $(x, a, y, t) \in \mathcal{S} \times \mathcal{A} \times \mathcal{S} \times [T]$, suppose *i.i.d.* random variables $X_1^{x,a,y}, \dots, X_t^{x,a,y}$ follow the Bernoulli distribution $\mathcal{B}(\mathbb{P}(y \mid x, a))$. Let

$$\mathcal{E}_2 \coloneqq \bigg\{ \forall (x,a,y,t) \in \mathcal{S} \times \mathcal{A} \times \mathcal{S} \times [T] \text{ satisfying } \mathbb{P}(y \mid x,a)t \geq c_2H^2 \ln(SAT/\delta), \\ \frac{\sum_{i=1}^t X_i^{x,a,y}}{t} \leq (1+1/H)\mathbb{P}(y \mid x,a) \bigg\},$$

where c_2 is a constant which will be defined later.

By Multiplicative Chernoff bound (Lemma 10) and a union bound, there exists a constant c_2 such that $\mathbf{Pr}(\mathcal{E}_2) \geq 1 - \delta/4$.

4.1.3 \mathcal{E}_3

Given any $(x, a, y, t) \in \mathcal{S} \times \mathcal{A} \times \mathcal{S} \times [T]$, suppose *i.i.d.* random variables $X_1^{x,a,y}, \dots, X_t^{x,a,y}$ follow the Bernoulli distribution $\mathcal{B}(\mathbb{P}(y \mid x, a))$. Let

$$\mathcal{E}_3 \coloneqq \left\{ \forall (x, a, y, t) \in \mathcal{S} \times \mathcal{A} \times \mathcal{S} \times [T] \text{ satisfying } \mathbb{P}(y \mid x, a)t \leq c_2 H^2 \ln(SAT/\delta), \\ \frac{\sum_{i=1}^t X_i^{x, a, y}}{t} \leq \frac{c_3 H \ln(SAT/\delta)}{t} \right\},$$

where c_3 is a constant which will be defined later.

By Bernstein's inequality (Lemma 11) and a union bound, there exists a constant c_3 such that $\mathbf{Pr}(\mathcal{E}_3) \ge 1 - \delta/4$.

4.2 Main Theorem

Theorem 2. With probability at least $1 - \delta$, the regret incurred by Algorithm 1 is bounded by

$$O\left(H\sqrt{SAT\ln(SAT/\delta)} + H^2S^2A\ln\left(\frac{T}{SA}\right)\ln(SAT/\delta)\right).$$

Remark 3. When T is large, the regret is bounded by $\widetilde{O}(H\sqrt{SAT})$.

Proof. The following arguments are conditioned on event $\mathcal{E} := \mathcal{E}_1 \wedge \mathcal{E}_2 \wedge \mathcal{E}_3 \wedge \mathcal{E}_4$, where \mathcal{E}_4 is defined later. And for simplicity, we use $\pi = (\pi_1, \dots, \pi_K)$ to represent Algorithm 1.

We first prove that the estimated Q-value function $Q_{k,h}$ is optimistic.

Lemma 4. For every $(k, h, x, a) \in [K] \times [H] \times S \times A$, it holds that

$$\widetilde{Q}_{k,h}(x,a) \ge Q_h^*(x,a).$$

Corollary 5. For every $(k, h, x) \in [K] \times [H] \times S$, it holds that $\widetilde{V}_{k,h}(x) \geq V_h^*(x)$.

Proof. Fix (k, h, x, a) and note that

$$\begin{aligned} \widetilde{Q}_{k,h}(x,a) - Q_h^*(x,a) &= (\widehat{\mathbb{P}}_k \widetilde{V}_{k,h+1})(x,a) - (\mathbb{P}V_{h+1}^*)(x,a) + b_k(x,a) \\ &= (\widehat{\mathbb{P}}_k (\widetilde{V}_{k,h+1} - V_{h+1}^*))(x,a) + ((\widehat{\mathbb{P}}_k - \mathbb{P})V_{h+1}^*)(x,a) + b_k(x,a) \end{aligned}$$

By event \mathcal{E}_1 , we have $|(\widehat{\mathbb{P}}_k - \mathbb{P})V_{h+1}^*(x,a)| \leq b_k(x,a)$. Using mathematical induction, we prove this lemma.

With optimistic guarantee, we can give a direct upper bound of \mathcal{R}_K^{π} . Note that

$$\mathcal{R}_{K}^{\pi} = \mathbb{E}\left[\sum_{k=1}^{K} (V_{1}^{*} - V_{1}^{\pi_{k}})(x_{k,1})\right]$$

$$\leq \mathbb{E}\left[\sum_{k=1}^{K} (\widetilde{V}_{k,1} - V_{1}^{\pi_{k}})(x_{k,1})\right]$$

$$= \sum_{k=1}^{K} \mathbb{E}\widetilde{\delta}_{k,1}.$$

where we have defined $\widetilde{\delta}_{k,h} \coloneqq (\widetilde{V}_{k,h} - V_h^{\pi_k})(x_{k,h})$

The next step idea is to bound $\widetilde{\delta}_{k,h}$ using $\widetilde{\delta}_{k,h+1}$. We first show

Lemma 6.

$$\widetilde{\delta}_{k,h} = ((\widehat{\mathbb{P}}_k - \mathbb{P})\widetilde{V}_{k,h+1})(x_{k,h}, a_{k,h}) + ((\mathbb{P}(\widetilde{V}_{k,h+1} - V_{h+1}^{\pi_k}))(x_{k,h}, a_{k,h}) - \widetilde{\delta}_{k,h+1}) + \widetilde{\delta}_{k,h+1} + b_k(x_{k,h}, a_{k,h}) + ((\mathbb{P}(\widetilde{V}_{k,h+1} - V_{h+1}^{\pi_k}))(x_{k,h}, a_{k,h}) - \widetilde{\delta}_{k,h+1}) + \widetilde{\delta}_{k,h+1} + b_k(x_{k,h}, a_{k,h}) + ((\mathbb{P}(\widetilde{V}_{k,h+1} - V_{h+1}^{\pi_k}))(x_{k,h}, a_{k,h}) - \widetilde{\delta}_{k,h+1}) + \widetilde{\delta}_{k,h+1} + b_k(x_{k,h}, a_{k,h}) + ((\mathbb{P}(\widetilde{V}_{k,h+1} - V_{h+1}^{\pi_k}))(x_{k,h}, a_{k,h}) - \widetilde{\delta}_{k,h+1}) + \widetilde{\delta}_{k,h+1} + b_k(x_{k,h}, a_{k,h}) + ((\mathbb{P}(\widetilde{V}_{k,h+1} - V_{h+1}^{\pi_k}))(x_{k,h}, a_{k,h}) - \widetilde{\delta}_{k,h+1}) + \widetilde{\delta}_{k,h+1} + b_k(x_{k,h}, a_{k,h}) + ((\mathbb{P}(\widetilde{V}_{k,h+1} - V_{h+1}^{\pi_k}))(x_{k,h}, a_{k,h}) - \widetilde{\delta}_{k,h+1}) + \widetilde{\delta}_{k,h+1} + b_k(x_{k,h}, a_{k,h}) + \widetilde{\delta}_{k,h+1} + \widetilde{\delta}_{k,h+$$

The idea to write in this way is that the expectation of $(\mathbb{P}(\widetilde{V}_{k,h+1}-V_{h+1}^{\pi_k}))(x_{k,h},a_{k,h})-\widetilde{\delta}_{k,h+1}$ is 0 conditioned on history \mathcal{H}_k .

Proof. Just note that

$$\begin{split} \widetilde{\delta}_{k,h} &= \widetilde{V}_{k,h}(x_{k,h}) - V_h^{\pi_k}(x_{k,h}) \\ &= (\widehat{\mathbb{P}}_k \widetilde{V}_{k,h+1})(x_{k,h}, a_{k,h}) - (\mathbb{P}V_{h+1}^{\pi_k})(x_{k,h}, a_{k,h}) + b_k(x_{k,h}, a_{k,h}) \\ &= ((\widehat{\mathbb{P}}_k - \mathbb{P})\widetilde{V}_{k,h+1})(x_{k,h}, a_{k,h}) + ((\mathbb{P}(\widetilde{V}_{k,h+1} - V_{h+1}^{\pi_k}))(x_{k,h}, a_{k,h}) - \widetilde{\delta}_{k,h+1}) + \widetilde{\delta}_{k,h+1} + b_k(x_{k,h}, a_{k,h}) \end{split}$$

We next focus on bounding

$$((\widehat{\mathbb{P}}_k - \mathbb{P})\widetilde{V}_{k,h+1})(x_{k,h}, a_{k,h}) \tag{1}$$

and show

Lemma 7 (One Step Transition Probability Error).

$$(1) \leq \frac{1}{H} \widetilde{\delta}_{k,h+1} + c_1 H \sqrt{\frac{\ln(SAT/\delta)}{n_k(x_{k,h}, a_{k,h})}} + \frac{1}{H} \left((\mathbb{P}(\widetilde{V}_{k,h+1} - V_{h+1}^{\pi_k}))(x_{k,h}, a_{k,h}) - \widetilde{\delta}_{k,h+1} \right) + \frac{\max\{c_2, c_3\}H^2 S \ln(SAT/\delta)}{n_k(x_{k,h}, a_{k,h})}.$$

Remark 8. There exists an easier way to bound (1) which just rewrite

$$(1) \le \left\| (\widehat{\mathbb{P}}_k - \mathbb{P})(x_{k,h}, a_{k,h}) \right\|_1 \cdot \left\| \widetilde{V}_{k,h+1} \right\|_{\infty}$$

and then uses the inequality in [3] to bound the ℓ_1 -norm deviation of the transition probability. Using this method will lead to an extra \sqrt{S} in the final regret.

Proof. Rewrite (1) we have

$$(1) = \underbrace{((\widehat{\mathbb{P}}_k - \mathbb{P})V_{h+1}^*)(x_{k,h}, a_{k,h})}_{(I)} + \underbrace{((\widehat{\mathbb{P}}_k - \mathbb{P})(\widetilde{V}_{k,h+1} - V_{h+1}^*))(x_{k,h}, a_{k,h})}_{(II)}$$
(2)

Consider (I) first. Note that

$$(I) = \sum_{y \in \mathcal{S}} \left(\widehat{\mathbb{P}}_{k}(y \mid x_{k,h}, a_{k,h}) - \mathbb{P}(y \mid x_{k,h}, a_{k,h}) \right) V_{h+1}^{*}(y)$$

$$= \left(\sum_{y \in \mathcal{S}} \widehat{\mathbb{P}}_{k}(y \mid x_{k,h}, a_{k,h}) V_{h+1}^{*}(y) \right) - \left(\sum_{y \in \mathcal{S}} \mathbb{P}(y \mid x_{k,h}, a_{k,h}) V_{h+1}^{*}(y) \right)$$
(3)

The first part of (3) can be seen as the empirical mean of $\sum_{y \in \mathcal{S}} \mathbb{P}(y \mid x_{k,h}, a_{k,h}) V_{h+1}^*(y)$ after $n_k(x_{k,h}, a_{k,h})$ trials. By event \mathcal{E}_1 , we conclude that

$$|(I)| \le c_1 H \sqrt{\frac{\ln(SAT/\delta)}{n_k(x_{k,h}, a_{k,h})}}.$$
(4)

We now take care of (II). Note that

$$(II) = \sum_{y \in \mathcal{S}} \left(\widehat{\mathbb{P}}_k(y \mid x_{k,h}, a_{k,h}) - \mathbb{P}(y \mid x_{k,h}, a_{k,h}) \right) (\widetilde{V}_{k,h+1} - V_{h+1}^*)(y). \tag{5}$$

Let S' be the set of states such that

$$\mathbb{P}(y \mid x_{k,h}, a_{k,h}) n_k(x_{k,h}, a_{k,h}) \ge c_2 H^2 \ln(SAT/\delta).$$

Rewrite (5) we get

$$(II) \leq \frac{1}{H} \widetilde{\delta}_{k,h+1} + \underbrace{\sum_{y \in \mathcal{S}'} \left(\widehat{\mathbb{P}}_{k}(y \mid x_{k,h}, a_{k,h}) - \mathbb{P}(y \mid x_{k,h}, a_{k,h})\right) (\widetilde{V}_{k,h+1} - V_{h+1}^{\pi_{k}})(y) - \frac{1}{H} \widetilde{\delta}_{k,h+1}}_{(III)} + \underbrace{\sum_{y \in (\mathcal{S} - \mathcal{S}')} \left(\widehat{\mathbb{P}}_{k}(y \mid x_{k,h}, a_{k,h}) - \mathbb{P}(y \mid x_{k,h}, a_{k,h})\right) (\widetilde{V}_{k,h+1} - V_{h+1}^{*})(y),}_{(IV)}$$

$$(6)$$

where we have used $V_{h+1}^{\pi_k}(y) \leq V_{h+1}^*(y)$. Due to event \mathcal{E}_2 , we have

$$(III) \leq \frac{1}{H} \left(\sum_{y \in \mathcal{S}'} \mathbb{P}(y \mid x_{k,h}, a_{k,h}) (\widetilde{V}_{k,h+1} - V_{h+1}^{\pi_k})(y) - \widetilde{\delta}_{k,h+1} \right)$$

$$\leq \frac{1}{H} \left((\mathbb{P}(\widetilde{V}_{k,h+1} - V_{h+1}^{\pi_k})) (x_{k,h}, a_{k,h}) - \widetilde{\delta}_{k,h+1} \right)$$
(7)

Next we upper bound (IV). By event \mathcal{E}_3 and plugging in inequality $\mathbb{P}(y \mid x_{k,h}, a_{k,h}) \leq \frac{c_2 H^2 \ln(SAT/\delta)}{n_k(x_{k,h}, a_{k,h})}$, we have

$$(IV) \le \frac{c_3 H S \ln(SAT/\delta)}{n_k(x_{k,h}, a_{k,h})} + \frac{c_2 H^2 S \ln(SAT/\delta)}{n_k(x_{k,h}, a_{k,h})}$$

$$\le \frac{\max\{c_2, c_3\} H^2 S \ln(SAT/\delta)}{n_k(x_{k,h}, a_{k,h})}$$

Plugging in upper bounds of (I), (III), (IV) to (1), we prove this lemma.

Combining Lemma 6 and Lemma 7, we get

$$\begin{split} \widetilde{\delta}_{k,h} & \leq \left(1 + \frac{1}{H}\right) \widetilde{\delta}_{k,h+1} + \left(1 + \frac{1}{H}\right) (\mathbb{P}(\widetilde{V}_{k,h+1} - V_{h+1}^{\pi_k})(x_{k,h}, a_{k,h}) - \widetilde{\delta}_{k,h+1}) \\ & + 2c_1 H \sqrt{\frac{\ln(SAT/\delta)}{n_k(x_{k,h}, a_{k,h})}} + \frac{\max\{c_2, c_3\} H^2 S \ln(SAT/\delta)}{n_k(x_{k,h}, a_{k,h})}. \end{split}$$

Hence

$$\sum_{k=1}^{K} \widetilde{\delta}_{k,1} \leq H + \left(1 + \frac{1}{H}\right)^{H} \left[\underbrace{\sum_{k=2}^{K} \sum_{h=1}^{H} (\mathbb{P}(\widetilde{V}_{k,h+1} - V_{h+1}^{\pi_{k}})(x_{k,h}, a_{k,h}) - \widetilde{\delta}_{k,h+1})}_{(*)} + 2c_{1}H \underbrace{\sum_{k=2}^{K} \sum_{h=1}^{H} \sqrt{\frac{\ln(SAT/\delta)}{n_{k}(x_{k,h}, a_{k,h})}}}_{(**)} + \max\{c_{2}, c_{3}\} \underbrace{\sum_{k=2}^{K} \sum_{h=1}^{H} \frac{H^{2}S \ln(SAT/\delta)}{n_{k}(x_{k,h}, a_{k,h})}}_{(***)} \right]$$

$$\lesssim H + (*) + H(**) + (***). \tag{8}$$

(*) can be seen as a martingale with (K-1)H r.v.'s and each r.v. is bounded by [-H,H]. By Azuma's inequality (Lemma 12), with probability at least $(1-\delta/4)$, it holds that

$$|(*)| \lesssim H\sqrt{\ln(1/\delta)T}.\tag{9}$$

And this defines event \mathcal{E}_4 . Rewrite (**), we have

$$(**) \le \sum_{(x,a)\in\mathcal{S}\times\mathcal{A}} \sum_{t=1}^{n_K(x,a)} \sqrt{\frac{\ln(SAT/\delta)}{t}} \lesssim \sqrt{\ln(SAT/\delta)} \cdot \sum_{(x,a)\in\mathcal{S}\times\mathcal{A}} \sqrt{n_K(x,a)} \le \sqrt{SAT\ln(SAT/\delta)},$$
(10)

where the last inequality is due to Cauchy-Schwarz inequality. Using a similar way, we get

$$(***) \leq \sum_{(x,a)\in\mathcal{S}\times\mathcal{A}} \sum_{t=1}^{n_K(x,a)} \frac{H^2 S \ln(SAT/\delta)}{t}$$

$$\lesssim H^2 S \ln(SAT/\delta) \sum_{(x,a)\in\mathcal{S}\times\mathcal{A}} \ln(n_K(x,a)) \leq H^2 S^2 A \ln\left(\frac{T}{SA}\right) \ln(SAT/\delta), \quad (11)$$

where the last inequality is due to Jensen's inequality applied to $\ln(\cdot)$ function. Putting back (9), (10), and (11) into (8), we prove this theorem.

5 Probability Tools

The following lemma states Hoeffding's inequality.

Lemma 9. Let X_1, X_2, \ldots, X_t be independent random variables bounded by [0, M]. Let $X = \sum_{i=1}^t X_i$. For every $\epsilon \geq 0$, it holds that

$$\mathbf{Pr}\left(|X - \mathbb{E}X| \ge \epsilon\right) \le 2\exp\left(-\frac{2\epsilon^2}{tM^2}\right).$$

The following lemma states a weak Multiplicative Chernoff bound.

Lemma 10. Let X_1, X_2, \ldots, X_t be independent random variables bounded by [0,1]. Let $X = \sum_{i=1}^t X_i$. For every $\epsilon \in [0,1]$, it holds that

$$\mathbf{Pr}\left(X \ge (1+\epsilon)\mathbb{E}X\right) \le \exp\left(-\frac{\epsilon^2\mathbb{E}X}{3}\right).$$

The following lemma states Bernstein's inequality.

Lemma 11. Let $X_1, X_2, ..., X_t$ be zero-mean independent random variables bounded by [-M, M]. Let $X = \sum_{i=1}^t X_i$. For every $\epsilon \ge 0$, it holds that

$$\mathbf{Pr}(X > \epsilon) \le \exp\left(-\frac{\frac{1}{2}\epsilon^2}{\sum_{i=1}^t \mathbb{E}[X_i^2] + \frac{1}{3}M\epsilon}\right).$$

Assuming $X_0 = 0$, a martingale (X_1, \dots, X_t) is c-Lipschitz if $|X_i - X_{i-1}| \le c_i$ where $c = (c_1, \dots, c_t)$. The following lemma states Azuma's inequality.

Lemma 12. ([2]) If a martingale $(X_1, ..., X_t)$ is **c**-Lipschitz, define $X = \sum_{i=1}^t X_i$, then for every $\epsilon \geq 0$, it holds that

$$\mathbf{Pr}(|X - \mathbb{E}X| \ge \epsilon) \le 2 \exp\left(-\frac{\epsilon^2}{2\sum_{i=1}^t c_i^2}\right),$$

where $c = (c_1, \ldots, c_t)$.

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