

In the real world you see $x = \langle 0, 0, 1, 0, 1, 0 \rangle$, the "data".

Then, you pick a curly-F (i.e. a parametric model).

But you don't know theta! So you have to guess theta. This guessing is "inference". There are typically three goals of "statistical inference":

- (1) Point estimation. Give me your best guess of theta (one value).
- (2) Confidence sets. Give me a range of likely theta's.
- (3) Theory testing. Evaluate a theory about the value of theta.

Assume curly-F = Bernoulli iid. Once you make an assumption of the parametric model, you can compute the JMF or JDF:

$$p(x; \theta) = \prod_{i=1}^6 p(x_i; \theta)$$

$$p(\langle 0, 0, 1, 0, 1, 0 \rangle; \theta) = (\theta^0 (1-\theta)^{1-0}) (\theta^0 (1-\theta)^{1-0}) (\theta^1 (1-\theta)^{1-1}) \dots = \theta^2 (1-\theta)^4$$

$$\begin{aligned} \text{if } \theta = 0.5 &\Rightarrow p(x; \theta) = 0.5^2 (1-0.5)^4 = 0.0156 \\ \text{if } \theta = 0.25 &\Rightarrow p(x; \theta) = 0.25^2 (1-0.25)^4 = 0.0118 \end{aligned}$$

=> theta = 0.25 seems "more likely" than theta = 0.5.

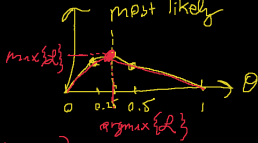
$$\mathcal{L}(\theta; x) = p(x; \theta) \quad \leftarrow \begin{array}{l} \text{probability of the data} \\ \text{with theta known} \end{array}$$

$\sum_{\theta} \mathcal{L}(\theta; x) = 1$ likelihood function, probability of theta given x known or the likelihood of "seeing" the parameter at a certain value.

$$\int_{\Theta} \mathcal{L}(\theta; x) d\theta = \text{no rule} \quad \sum_{\text{supp}[x]} p(x; \theta) = 1$$

Maximum Likelihood Estimator / Estimate

"hat" means "estimator"

$$\text{Define } \hat{\theta}_{MLE} := \underset{\theta \in \Theta}{\operatorname{argmax}} \{ \mathcal{L}(\theta; x) \}$$


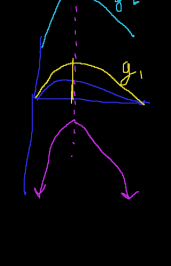
Let g be a strictly increasing function.

$$= \underset{\theta \in \Theta}{\operatorname{argmax}} \{ g(\mathcal{L}(\theta; x)) \}$$

let $g = \ln$ log-likelihood

$$= \underset{\theta \in \Theta}{\operatorname{argmax}} \{ \ln(\mathcal{L}(\theta; x)) \}$$

Define $\ell(\theta; x) := \ln(\mathcal{L}(\theta; x))$

$$= \underset{\theta \in \Theta}{\operatorname{argmax}} \{ \ell(\theta; x) \}$$


$$\ell(\theta; x) = \ln(p(x; \theta)) \stackrel{\text{iid}}{=} \ln\left(\prod_{i=1}^n p(x_i; \theta)\right) = \sum_{i=1}^n \ln(p(x_i; \theta))$$

In our example of $x = \langle 0, 0, 1, 0, 1, 0 \rangle \dots$

$$\begin{aligned} \ell(\theta; x) &= \sum_{i=1}^6 \ln(\theta^{x_i} (1-\theta)^{1-x_i}) = \sum_{i=1}^6 (x_i \ln(\theta) + (1-x_i) \ln(1-\theta)) \\ &= (\sum x_i) \ln(\theta) + (6 - \sum x_i) \ln(1-\theta) \quad \text{Note: } \bar{x} := \frac{1}{n} \sum x_i \Rightarrow \sum x_i = n \bar{x} \\ &= 6 \bar{x} \ln(\theta) + (6 - 6 \bar{x}) \ln(1-\theta) = 6 \left(\bar{x} \ln(\theta) + (1-\bar{x}) \ln(1-\theta) \right) \end{aligned}$$

We need to find the argmax of this function...

$\hat{\theta}_{MLE}$ = this represents estimate (a realization from the estimator)
take derivative of the log likelihood wrt theta and set = 0 and solve.

$$\frac{d}{d\theta} [\ell(\theta; x)] = 6 \left(\frac{\bar{x}}{\theta} - \frac{1-\bar{x}}{1-\theta} \right) \stackrel{\text{set}}{=} 0 \Rightarrow \frac{\bar{x}}{\theta} = \frac{1-\bar{x}}{1-\theta}$$

$$\Rightarrow \bar{x}(1-\theta) = (1-\bar{x})\theta \Rightarrow \bar{x} - \bar{x}\theta = \theta - \bar{x}\theta \Rightarrow \hat{\theta}_{MLE} = \bar{x} = \frac{2}{6} = 0.33 \dots$$

The estimator, $\hat{\theta}_{MLE} = \bar{X}$ is a rv whose realizations are the estimates. This rv has nice properties:

- ① $\hat{\theta}_{MLE}$ is "consistent". This means that this estimator can provide arbitrary precision on theta given enough n.
- ② $\hat{\theta}_{MLE} \sim N(\theta, \text{SE}[\hat{\theta}_{MLE}]^2)$ asymptotic normality
standard error
- ③ "Efficiency" means that among all consistent estimators, it has minimum variance.

Consider $X \sim \text{Geom}(\theta) := (1-\theta)^X \theta \Rightarrow \text{supp}[X] = \{0, 1, 2, \dots\}, \Theta = (0, 1)$

Consider a sequence of iid Bernoulli thetas. This rv tells you the number of failures (realizations of zero) before the first success (realizations of one).

If $\theta = 1\%$

$$\frac{0}{1^{.99}}, \frac{0}{2^{.99}}, \frac{0}{3^{.99}}, \dots, \frac{0}{41^{.99}}, \frac{1}{50^{.99}} \Rightarrow x=49 \quad P(X=49; \theta=0.01) = 0.99^{49} \cdot 0.01$$

$\mathcal{F} = \text{iid Geometric}$. n realizations

$$\begin{aligned} \mathcal{L}(\theta; x) &= \prod_{i=1}^n (1-\theta)^{x_i} \theta = (1-\theta)^{\sum x_i} \theta^n \\ \ell(\theta; x) &= \ln(\mathcal{L}(\theta; x)) = (\sum x_i) \ln(1-\theta) + n \ln(\theta) = n \bar{x} \ln(1-\theta) + n \ln(\theta) \\ &= n (\bar{x} \ln(1-\theta) + \ln(\theta)) \end{aligned}$$

Let's find the MLE. We take the derivative of the log-likelihood wrt theta and set it equal to zero and solve.

$$\begin{aligned} \frac{d}{d\theta} [\ell] &= n \left(-\frac{\bar{x}}{1-\theta} + \frac{1}{\theta} \right) \stackrel{\text{set}}{=} 0 \Rightarrow \frac{1}{\theta} = \frac{\bar{x}}{1-\theta} \Rightarrow \frac{1-\theta}{\theta} = \bar{x} \\ &\Rightarrow \frac{1}{\theta} - 1 = \bar{x} \Rightarrow \frac{1}{\theta} = \bar{x} + 1 \Rightarrow \hat{\theta}_{MLE} = \frac{1}{\bar{x} + 1} \end{aligned}$$

$$\text{Consider } \bar{x} = 49 \Rightarrow \hat{\theta}_{MLE} = \frac{1}{49+1} = 2\%$$

Let's examine MLE property #2:

$$\hat{\theta}_{MLE} \sim N(\theta, \text{SE}[\hat{\theta}_{MLE}]^2) = N\left(\theta, \sqrt{\frac{\theta(1-\theta)}{n}}\right)^2$$

In the curly-F: iid Bernoulli case,

$$\hat{\theta}_{MLE} = \bar{X}, \quad \text{SE}[\hat{\theta}_{MLE}] = \text{SE}[\bar{X}] = \sqrt{\text{Var}[\bar{X}]} = \sqrt{\frac{\sigma^2}{n}} = \sqrt{\frac{\theta(1-\theta)}{n}}$$

In the curly-F: iid Geometric case,

$$\hat{\theta}_{MLE} = \frac{1}{\bar{X} + 1}, \quad \text{SE}\left[\frac{1}{\bar{X} + 1}\right] = ? \quad \text{difficult without more mathematics...}$$

We now use property 2 to attack the other goals of inference:

Confidence Sets: we use a method called the "confidence interval":

$$CI_{\theta, 1-\alpha} := \left[\hat{\theta}_{MLE} \pm z_{\frac{\alpha}{2}} \text{SE}[\hat{\theta}_{MLE}] \right]$$

parameter, level of confidence, standard normal quantile at alpha/2

For the iid Bernoulli case,

$$CI_{\theta, 1-\alpha} = \left[\bar{x} \pm z_{\frac{\alpha}{2}} \sqrt{\frac{\bar{x}(1-\bar{x})}{n}} \right]$$

Letting $1-\alpha = 95\% \Rightarrow \alpha = 5\%$

$$CI_{\theta, 95\%} = \left[\bar{x} \pm 1.96 \sqrt{\frac{\bar{x}(1-\bar{x})}{n}} \right]$$

