

Multiple linear regression

Labor Economics

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Model:

$$Y = \beta_0 + \beta_1 X_1 + \cdots + \beta_p X_p + \epsilon$$

We interpret β_j as the **average effect** of a unit increase in X_j on Y , **holding all other predictors fixed** ("ceteris paribus").

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In the advertising example:

$$\text{sales} = \beta_0 + \beta_1 \cdot \text{TV} + \beta_2 \cdot \text{radio} + \beta_3 \cdot \text{newspaper} + \epsilon$$

Interpreting regression coefficients

The ideal scenario is when all predictors are **uncorrelated**:

- An increase in the value of X_1 does not affect the value of X_2
- A trial/experiment design is called a **balanced design** in that case, and each coefficient can be estimated and tested separately

Interpretations such as “*a unit change in X_j is associated with a β_j change in Y , while all the other variables stay fixed*”, are possible.

Interpreting regression coefficients

- **Correlations amongst predictors** cause problems:
 1. The variance of all coefficients tends to increase, sometimes dramatically
 2. Interpretations become hazardous — when X_j changes, everything else changes.
- Claims of *causality* should be avoided for *observational data*
 - Identifying causality is a big topic in economics.
We'll touch on that topic later in this course.

The woes of (interpreting) regression coefficients

- The regression coefficient β_j estimates the expected change in Y per unit change in X_j , *with all other predictors held fixed*.
- When predictors are correlated, they *change together*!

Two Examples

1. Y = total amount of paper money in your pocket;
 X_1 = # of papers; X_2 = # of 10- and 20- RMB papers.
By itself, regression coefficients of Y on X_2 will be > 0 .
But how about with X_1 in the model?

Two Examples

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But how about with X_1 in the model?
2. Y = number of tackles by a football player in a season;
 W and H are his weight and height. Fitted regression model is
 $\hat{Y} = b_0 + 0.5W - 0.10H$. How do we interpret $\hat{\beta}_2 < 0$?

Interpreting regression coefficients

“The only way to find out what will happen when a complex system is disturbed is to disturb the system, not merely to observe it passively.” -- Fred Mosteller and John Tukey

- These are said by statisticians. What can we (economists) do to deal with correlations between X_i 's?

Interpreting regression coefficients

“The only way to find out what will happen when a complex system is disturbed is to disturb the system, not merely to observe it passively.” -- Fred Mosteller and John Tukey

Economists Esther and Banerjee won Nobel Prize in 2019 for their usage of **‘randomised control trials’ (RCT)** in economics.



Estimation and Prediction for Multiple Regression

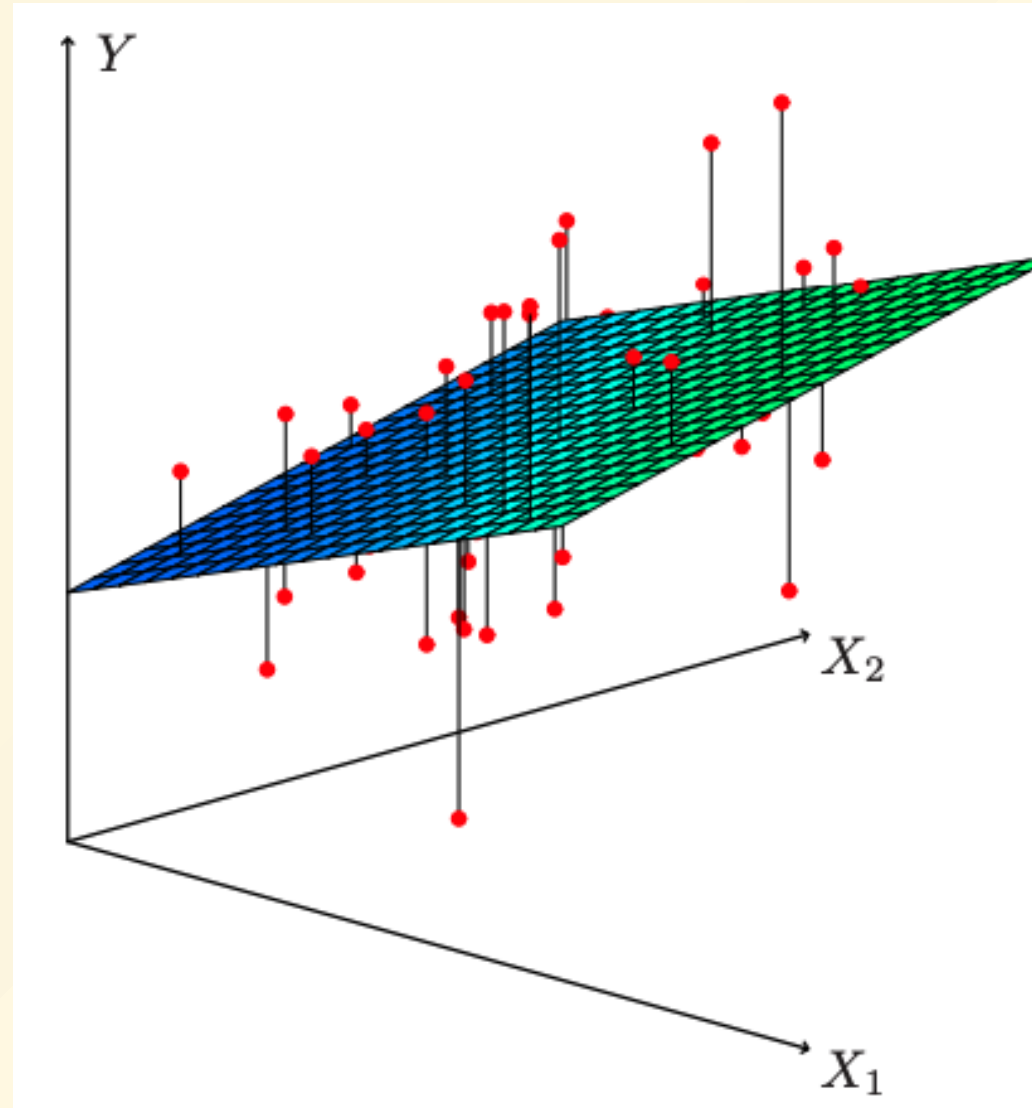
- Given estimates $\hat{\beta}_0, \hat{\beta}_1, \dots, \hat{\beta}_p$, we can make predictions using:

$$\hat{y} = \hat{\beta}_0 + \hat{\beta}_1 x_1 + \dots + \hat{\beta}_p x_p$$

- LS estimators are obtained by minimizing the RSS:

$$\min_{\hat{\beta}_0, \dots, \hat{\beta}_p} \text{RSS} = \sum_{i=1}^n (y_i - \hat{y}_i)^2$$

Example: $p = 2$.



Results for advertising data

| | Coefficient | Std. Error | t-statistic | p-value |
|-----------|-------------|------------|-------------|----------|
| Intercept | 2.939 | 0.3119 | 9.42 | < 0.0001 |
| TV | 0.046 | 0.0014 | 32.81 | < 0.0001 |
| radio | 0.189 | 0.0086 | 21.89 | < 0.0001 |
| newspaper | -0.001 | 0.0059 | -0.18 | 0.8599 |

| | Correlations: | | | |
|-----------|---------------|--------|-----------|--------|
| | TV | radio | newspaper | sales |
| TV | 1.0000 | 0.0548 | 0.0567 | 0.7822 |
| radio | | 1.0000 | 0.3541 | 0.5762 |
| newspaper | | | 1.0000 | 0.2283 |
| sales | | | | 1.0000 |

Some important questions

1. Is *at least one of the predictors* X_1, X_2, \dots, X_p useful in predicting the response?
2. Do *all the predictors* help to explain Y , or only *part of the predictors* useful?
3. How well does the linear model fit the data?

Q1: Is *at least one of the predictors* useful in predicting Y ?

- We can use the F-statistic:

$$F = \frac{(TSS - RSS)/p}{RSS/(n - p - 1)} \sim F_{p, n-p-1}$$

| Quantity | Value |
|-------------------------|-------|
| Residual Standard Error | 1.69 |
| R^2 | 0.897 |
| F-statistic | 570 |

Q2: Do *all the predictors* help to explain Y , or only *part of the predictors* useful?

- Essentially, this is to **decide on the important variables**.
- The most direct approach is called all subsets or best subsets regression:
 - we compute the least squares fit for all possible subsets;
 - then choose between them based on some criterion that balances training error with model size.
- However, usually we cannot examine all possible models. For example, when $p = 40$, there are $2^p \geq$ a billion models!

Forward selection

- Begin with the *null model* — a model that contains an intercept but no predictors.
- Fit p simple linear regressions and add to the null model the variable that results in the **lowest RSS**.
- Add to that model the variable that results in the **lowest RSS** amongst all two-variable models.
- Continue until **some stopping rule** is satisfied, for example when all remaining variables have a p-value above some threshold

Backward selection

- Start with *all variables* in the model.
- Remove the variable with the **largest p-value** — that is, the variable that is the least statistically significant.
- The new $(p - 1)$ -variable model is fit, and the variable with the **largest p-value** is removed.
- Continue until a **stopping rule** is reached. For instance, we may stop when all remaining variables have a **significant p-value** defined by some significance threshold.

Model selection

- Forward and Backward selections are two specialized cases of *model selection*.
- There are more systematic criteria for choosing an “optimal” member in the path of models produced by forward or backward stepwise selection.
 - Especially for time-series data.
- These include **Mallow's C_p** , **Akaike information criterion (AIC)**, **Bayesian information criterion (BIC)**, **adjusted R^2** and **Cross-validation (CV)**.

Q3. How well does the model fit the data?

- R^2 fails to be a good judge: adding more predictor variables always increases R^2 !
- We need to have **some punishments** for those high R^2 cases with high p .

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$$R_{adj}^2 = 1 - \frac{(1 - R^2)(n - 1)}{n - p}$$

- where n is # of observations and p is # of predictors.

Other Considerations in the Regression Model

- Some predictors are not **quantitative** but are *qualitative*.
- Also called *categorical/factor* predictors:
gender, student/martial status, ethnicity,

Motivating example: a credit card company (say Bank of China) has the following data about its clients:

- Balance, Age, Cards, Education, Income, Limit, Rating

Qualitative Predictors — continued

Example: investigate differences in credit card balance between males and females, ignoring the other variables.

A *dummy variable* for gender:

$$x_i = \begin{cases} 1 & \text{if i-th person is female} \\ 0 & \text{if i-th person is male} \end{cases}$$

Model: $y_i = \beta_0 + \beta_1 x_i + \epsilon_i$.

Interpretation?

Results for gender model:

| | Coefficient | Std. Error | t-statistic | p-value |
|----------------|-------------|------------|-------------|---------|
| Intercept | 509.80 | 33.13 | 15.389 | <0.0001 |
| gender[Female] | 19.73 | 46.05 | 0.429 | 0.6690 |

Qualitative predictors with more than two levels

- With more than two levels, we create additional dummy variables.
- For example, for the ethnicity variable (Asian/African/American) we create two dummy variables:
 - $x_{i1} = 1$ if i-th person is Asian, or 0 otherwise;
 - $x_{i2} = 1$ if i-th person is African, or 0 otherwise;
- There will always be one fewer dummy variable than the number of levels. The level with no dummy variable —American in this example — is known as the **baseline**.

Results for ethnicity

| | Coefficient | Std. Error | t-stat. | p-value |
|--------------------|-------------|------------|---------|----------|
| Intercept | 531.00 | 46.32 | 11.464 | < 0.0001 |
| ethnicity[Asian] | -18.69 | 65.02 | -0.287 | 0.7740 |
| ethnicity[African] | -12.50 | 56.68 | -0.221 | 0.8260 |

Extensions of the Linear Model

Extensions of the Linear Model

- In the ads example, we have assumed that the effect on sales of increasing *one advertising medium* is independent of the amount spent on *the other media*. (**No Interactions**)
- However, suppose that spending money on radio advertising actually increases the effectiveness of TV advertising, so that the slope term for TV should increase as radio increases.
 - In econ (marketing), this is known as a complementary (synergy) effect
 - in statistics it is referred to as an **interaction effect**.

Modelling interactions — Advertising data

$$\begin{aligned}\text{sales} &= \beta_0 + \beta_1 \times \text{TV} + \beta_2 \times \text{radio} + \beta_3 \times (\text{radio} \times \text{TV}) + \epsilon \\ &= \beta_0 + (\beta_1 + \beta_3 \times \text{radio}) \times \text{TV} + \beta_2 \times \text{radio} + \epsilon\end{aligned}$$

- Implications: when the expenses on radio ads get higher, the **marginal benefit** of expenses on TV ads get higher!

Modelling interactions — Advertising data

Results as below. **Interpretation?**

| | Coefficient | Std. Error | t-statistic | p-value |
|-----------|-------------|------------|-------------|---------|
| Intercept | 6.7502 | 0.248 | 27.23 | <0.0001 |
| TV | 0.0191 | 0.002 | 12.70 | <0.0001 |
| radio | 0.0289 | 0.009 | 3.24 | 0.0014 |
| TV×radio | 0.0011 | 0.000 | 20.73 | <0.0001 |

Interpretation

Interactions are important:

- The p-value for the interaction term TV×radio is extremely low, indicating that there is strong evidence for $H_A : \beta_3 \neq 0$.
- The R^2 for the interaction model is 96.8%, compared to only 89.7% for the model that predicts sales using TV and radio without an interaction term.

Interpretation — continued

- This means that $(96.8 - 89.7)/(100 - 89.7) = 69\%$ of the variability in sales that remains after fitting the additive model has been explained by the interaction term.
- The coefficient estimates in the table suggest that an increase in TV advertising of \$1,000 is associated with increased sales of $(\hat{\beta}_1 + \hat{\beta}_3 \times \mathbf{radio}) \times 1000 = 19 + 1.1 \times \mathbf{radio}$ units.
- An increase in radio advertising of \$1, 000 will be associated with an increase in sales of $(\hat{\beta}_2 + \hat{\beta}_3 \times \mathbf{TV}) \times 1000 = 29 + 1.1 \times \mathbf{TV}$ units.

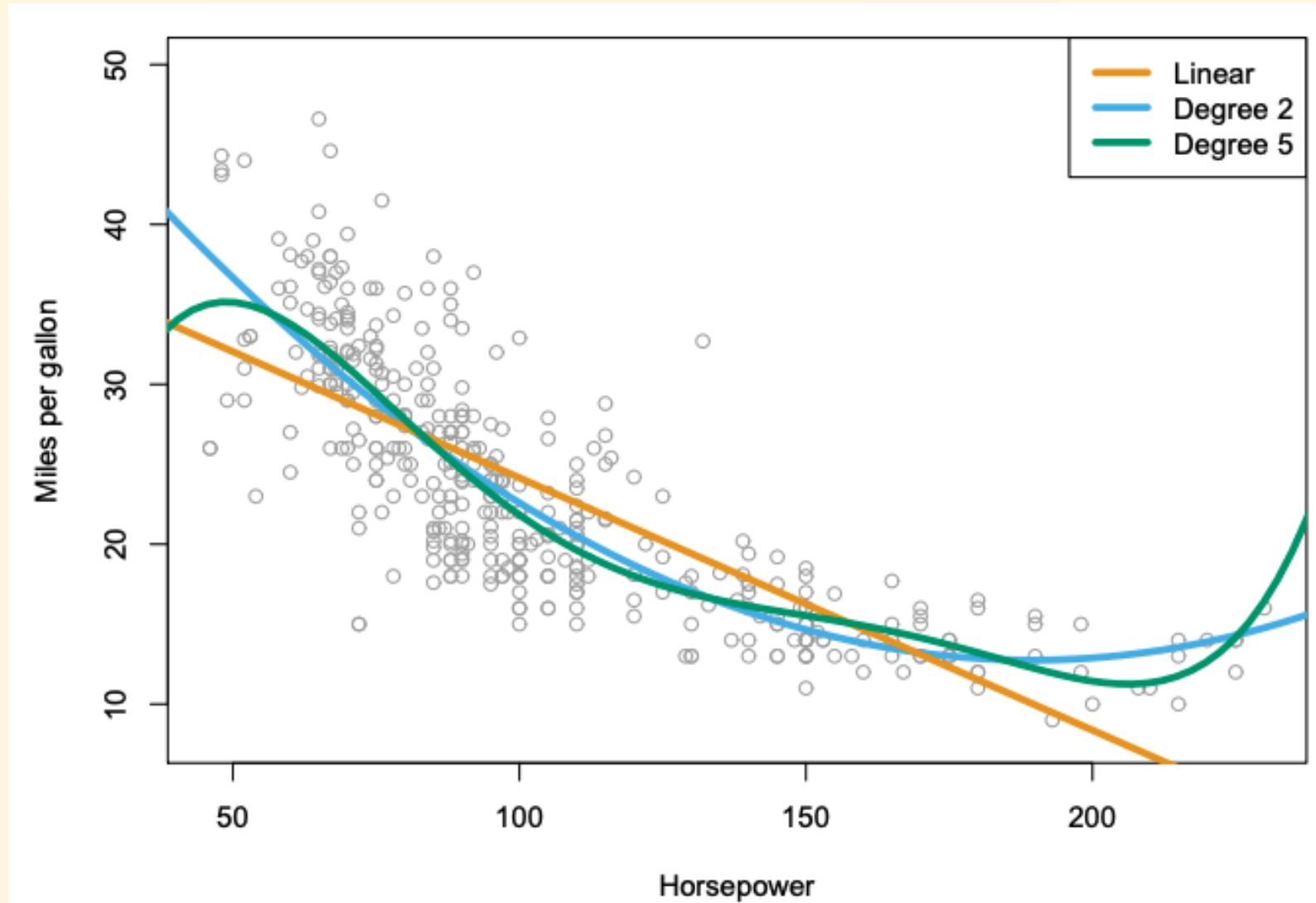
Hierarchy

- Sometimes it is the case that an interaction term has a very small p-value, but the associated main effects (in this case, TV and radio) do not.
- The **hierarchy principle**:
If we include an interaction in a model, we should also include the main effects, even if the p-values associated with their coefficients are not significant.

Hierarchy

- The rationale for this principle is that interactions are hard to interpret in a model without main effects — their meaning is changed.
- Specifically, the interaction terms also contain main effects, if the model has no main effect terms.

Another extension: Incorporating non-linear effects



The figure suggests that

- $\text{mpg} = \beta_0 + \beta_1 \times \text{horsepower} + \beta_2 \times \text{horsepower}^2 + \epsilon$

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| | Coefficient | Std. Error | t-statistic | p-value |
|-------------------------|-------------|------------|-------------|---------|
| Intercept | 56.9001 | 1.8004 | 31.6 | <0.0001 |
| horsepower | -0.4662 | 0.0311 | -15.0 | <0.0001 |
| horsepower ² | 0.0012 | 0.0001 | 10.1 | <0.0001 |

Generalizations of the Linear Model

- **Classification problems:** logistic regression, support vector machines
- **Non-linearity:** kernel smoothing, splines and generalized additive models, nearest neighbor methods.
- **Interactions:** Tree-based methods, bagging, random forests and boosting (these also capture non-linearities)
- **Regularized fitting:** Ridge regression and lasso