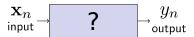
AI4ER 0: Regression

What is regression?

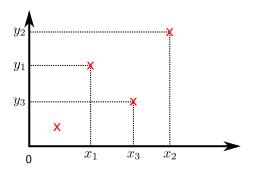
A type of problem in machine learning requiring

- to identify patterns and regularities between input variables and a corresponding continuous output variable,
- from a training data set $\mathcal{D} = \{(\mathbf{x}_n, y_n)\}_{n=1}^N$ formed by pairs of input vectors $\mathbf{x}_n = (x_{n,1}, \dots, x_{n,D})^\mathsf{T}$ and corresponding output values $y_n \in \mathbb{R}$.

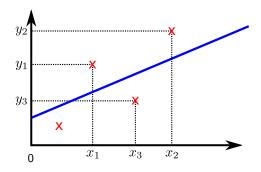


The relationship between inputs \mathbf{x}_n and outputs y_n in $p(y_n|\mathbf{x}_n, \theta)$ is **linear**.

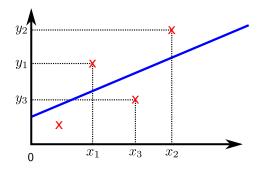
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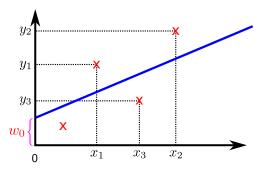


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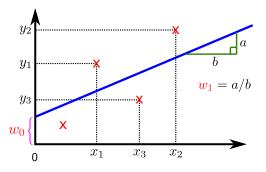
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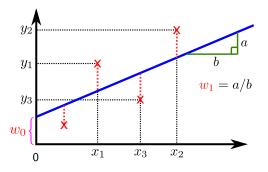
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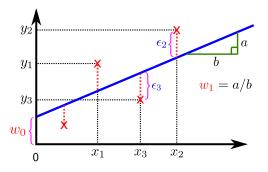
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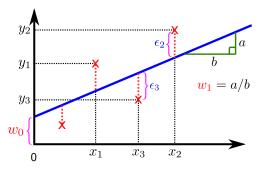
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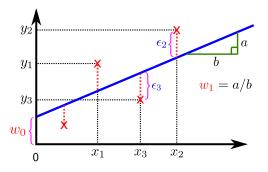
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$$y_n = w_0 + w_1 x_n + \epsilon_n$$

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Why linear? Simple, easy to understand, widely used, easily generalized.

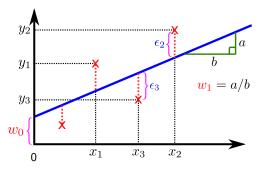


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What should the distribution of ϵ_n be?

The relationship between inputs \mathbf{x}_n and outputs y_n in $p(y_n|\mathbf{x}_n, \theta)$ is **linear**.

Why linear? Simple, easy to understand, widely used, easily generalized.



$$y_n = w_0 + w_1 x_n + \epsilon_n$$

What should the distribution of ϵ_n be? $\epsilon_n \sim \mathcal{N}(0, \sigma^2)$ allows for tractable inference.

Assuming

$$y_n = w_0 + w_1 x_n + \epsilon_n$$
,
 $\epsilon_n \sim \mathcal{N}(0, \sigma^2)$.

What is the form of $p(y_n|\mathbf{x}_n, \theta)$ with $\theta = {\sigma^2, \mathbf{w}_0, \mathbf{w}_1}$?

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What is the form of $p(y_n|\mathbf{x}_n, \boldsymbol{\theta})$ with $\boldsymbol{\theta} = \{\sigma^2, \mathbf{w_0}, \mathbf{w_1}\}$?

We have that

$$\mathbf{E}[y_n] = \mathbf{w}_0 + \mathbf{w}_1 \mathbf{x}_n + \mathbf{E}[\epsilon_n] = \mathbf{w}_0 + \mathbf{w}_1 \mathbf{x}_n,$$

$$\operatorname{Var}[y_n] = \mathbf{E}\left[\left(y_n - \mathbf{E}[y_n]\right)^2\right] = \mathbf{E}\left[\epsilon_n^2\right] = \sigma^2.$$

Assuming

$$y_n = \mathbf{w_0} + \mathbf{w_1} x_n + \epsilon_n$$
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Therefore,

$$p(y_n|\mathbf{x}_n, \boldsymbol{\theta}) = \mathcal{N}(y_n|\mathbf{w}_0 + \mathbf{w}_1 x_n, \sigma^2)$$

$$= \frac{1}{\sqrt{2\pi\sigma^2}} \exp\left\{-\frac{1}{2} \frac{(y_n - \mathbf{w}_0 - \mathbf{w}_1 x_n)^2}{\sigma^2}\right\}.$$

$$y_n = w_0 + w_1 x_{n,1} + \cdots + w_D x_{n,D} + \epsilon_n$$

where
$$|\epsilon_n \sim \mathcal{N}(0, \sigma^2)|$$
 and $\widetilde{\mathbf{x}}_n = (1, \mathbf{x}_n)^\mathsf{T}$, $\boldsymbol{\theta} = \{\sigma^2, \mathbf{w}\}$.

$$y_{n} = w_{0} + w_{1}x_{n,1} + \dots + w_{D}x_{n,D} + \epsilon_{n} = [w_{0}, \dots, w_{D}] \underbrace{\begin{bmatrix} 1 \\ x_{n,1} \\ \vdots \\ x_{n,D} \end{bmatrix}}_{\widetilde{\mathbf{X}}_{n}} + \epsilon_{n}$$

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$$p(y_n|\widetilde{\mathbf{x}}_n, \boldsymbol{\theta}) = \mathcal{N}(y_n|\mathbf{w}^{\mathsf{T}}\widetilde{\mathbf{x}}_n, \sigma^2),$$

For a data set $\mathcal{D} = \{(\mathbf{x}_n, y_n)\}_{n=1}^N$, we assume

$$y_{n} = w_{0} + w_{1}x_{n,1} + \dots + w_{D}x_{n,D} + \epsilon_{n} = [w_{0}, \dots, w_{D}] \underbrace{\begin{bmatrix} 1 \\ x_{n,1} \\ \vdots \\ x_{n,D} \end{bmatrix}}_{\widetilde{\mathbf{X}}_{n}} + \epsilon_{n} = \mathbf{w}^{\mathsf{T}}\widetilde{\mathbf{x}}_{n} + \epsilon_{n},$$

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Jargon for regression:

- \mathbf{x}_n are the inputs, features, covariates, independent variables, etc.
- \mathbf{y}_n are the outputs, responses, targets, dependent variables, etc.
- \mathbf{w} are the coefficients, weights, etc. (w_0 is called the bias or intercept).
- ϵ_n are the errors, disturbances or noise.

Maximize the **likelihood** function $p(y_1, \ldots, y_n | \widetilde{\mathbf{x}}_1, \ldots, \widetilde{\mathbf{x}}_n, \boldsymbol{\theta})$ with respect to $\boldsymbol{\theta}$.

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$$= \sum_{n=1}^N \left\{ -\frac{1}{2} \log(2\pi\sigma^2) - \frac{1}{2\sigma^2} (y_n - \mathbf{w}^{\mathsf{T}} \widetilde{\mathbf{x}}_n)^2 \right\}$$

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$$= -\frac{N}{2} \log(2\pi\sigma^2) - \frac{1}{2\sigma^2} (\mathbf{y} - \widetilde{\mathbf{X}} \mathbf{w})^{\mathsf{T}} (\mathbf{y} - \widetilde{\mathbf{X}} \mathbf{w})$$

Maximize the **likelihood** function $p(y_1, \ldots, y_n | \widetilde{\mathbf{x}}_1, \ldots, \widetilde{\mathbf{x}}_n, \boldsymbol{\theta})$ with respect to $\boldsymbol{\theta}$.

In practice, it is the log-likelihood function what is maximized.

$$\begin{split} \mathcal{L}(\boldsymbol{\theta}) &= \log p(y_1, \dots, y_n | \widetilde{\mathbf{x}}_1, \dots, \widetilde{\mathbf{x}}_n, \boldsymbol{\theta}) \\ &= \log \prod_{n=1}^N p(y_n | \widetilde{\mathbf{x}}_n, \boldsymbol{\theta}) = \sum_{n=1}^N \log \mathcal{N}(y_n | \mathbf{w}^{\mathsf{T}} \widetilde{\mathbf{x}}_n, \sigma^2) \\ &= \sum_{n=1}^N \left\{ -\frac{1}{2} \log(2\pi\sigma^2) - \frac{1}{2\sigma^2} (y_n - \mathbf{w}^{\mathsf{T}} \widetilde{\mathbf{x}}_n)^2 \right\} \\ &= -\frac{N}{2} \log(2\pi\sigma^2) - \frac{1}{2\sigma^2} \left(\mathbf{y} - \widetilde{\mathbf{X}} \mathbf{w} \right)^{\mathsf{T}} \left(\mathbf{y} - \widetilde{\mathbf{X}} \mathbf{w} \right) \end{split}$$

where $\mathbf{y} = (y_1, \dots, y_n)^\mathsf{T}$, $\widetilde{\mathbf{X}} = (\widetilde{\mathbf{x}}_1; \dots; \widetilde{\mathbf{x}}_n)^\mathsf{T}$, and we have used $\mathbf{a}^\mathsf{T} \mathbf{a} = \sum_i a_i^2$.

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where $\mathbf{y} = (y_1, \dots, y_n)^\mathsf{T}$, $\widetilde{\mathbf{X}} = (\widetilde{\mathbf{x}}_1; \dots; \widetilde{\mathbf{x}}_n)^\mathsf{T}$, and we have used $\mathbf{a}^\mathsf{T} \mathbf{a} = \sum_i a_i^2$.

$$\mathbf{a} = \begin{bmatrix} a_0 \\ a_1 \\ \vdots \\ a_D \end{bmatrix}, \qquad \mathbf{w} = \begin{bmatrix} w_0 \\ w_1 \\ \vdots \\ w_D \end{bmatrix}, \qquad \mathbf{a}^\mathsf{T} \mathbf{w} = \sum_{i=0}^D a_i w_i,$$

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$$\frac{d[\mathbf{a}^{\mathsf{T}}\mathbf{w}]}{d\mathbf{w}} = \begin{pmatrix} \frac{d[\mathbf{a}^{\mathsf{T}}\mathbf{w}]}{dw_0} \\ \frac{d[\mathbf{a}^{\mathsf{T}}\mathbf{w}]}{dw_1} \\ \vdots \\ \frac{d[\mathbf{a}^{\mathsf{T}}\mathbf{w}]}{dw_0} \end{pmatrix}$$

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$$\frac{d[\mathbf{a}^{\mathsf{T}}\mathbf{w}]}{d\mathbf{w}} = \begin{vmatrix} \frac{d[\mathbf{a}^{\mathsf{T}}\mathbf{w}]}{dw_0} \\ \frac{d[\mathbf{a}^{\mathsf{T}}\mathbf{w}]}{dw_1} \\ \vdots \\ \frac{d[\mathbf{a}^{\mathsf{T}}\mathbf{w}]}{dw_0} \end{vmatrix} = \mathbf{a}$$

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$$\frac{d[\mathbf{a}^{\mathsf{T}}\mathbf{w}]}{d\mathbf{w}} = \begin{vmatrix} \frac{d[\mathbf{a}^{\mathsf{W}}]}{dw_0} \\ \frac{d[\mathbf{a}^{\mathsf{T}}\mathbf{w}]}{dw_1} \\ \vdots \\ \frac{d[\mathbf{a}^{\mathsf{T}}\mathbf{w}]}{dw} \end{vmatrix} = \mathbf{a} = \frac{d[\mathbf{w}^{\mathsf{T}}\mathbf{a}]}{d\mathbf{w}}.$$

Using the previous result

$$\frac{d[\mathbf{a}^{\mathsf{T}}\mathbf{w}]}{d\mathbf{w}} = \frac{d[\mathbf{w}^{\mathsf{T}}\mathbf{a}]}{d\mathbf{w}} = \mathbf{a},$$

and the product rule of calculus,

$$\frac{d}{dx}[f(x)g(x)] = \underbrace{\left[\frac{d}{dx}f(x)\right]g(x)}_{g(x) \text{ as constant}} + \underbrace{f(x)\left[\frac{d}{dx}g(x)\right]}_{f(x) \text{ as constant}},$$
(1)

we obtain

$$\frac{d[\mathbf{w}^{\mathsf{T}} \mathbf{A} \mathbf{w}]}{d \mathbf{w}} = \underbrace{\mathbf{A}^{\mathsf{T}} \mathbf{w}}_{\mathbf{w}^{\mathsf{T}} \mathbf{A} \text{ as constant}} + \underbrace{\mathbf{A} \mathbf{w}}_{\mathbf{a} \mathbf{w} \text{ as constant}}$$

Vector calculus II

Using the previous result

$$\frac{d[\mathbf{a}^{\mathsf{T}}\mathbf{w}]}{d\mathbf{w}} = \frac{d[\mathbf{w}^{\mathsf{T}}\mathbf{a}]}{d\mathbf{w}} = \mathbf{a},$$

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we obtain

$$\frac{d[\mathbf{w}^{\mathsf{T}} \mathbf{A} \mathbf{w}]}{d\mathbf{w}} = \underbrace{\mathbf{A}^{\mathsf{T}} \mathbf{w}}_{\mathbf{w}^{\mathsf{T}} \mathbf{A} \text{ as constant}} + \underbrace{\mathbf{A} \mathbf{w}}_{\mathbf{a} \mathbf{s} \text{ constant}} = 2\mathbf{A} \mathbf{w} \text{ if } \mathbf{A} \text{ symmetric.}$$

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$$\frac{\partial \mathcal{L}(\boldsymbol{\theta})}{\partial \mathbf{w}} = \frac{\partial}{\partial \mathbf{w}} \left(-\frac{\mathbf{w}^\mathsf{T} \widetilde{\mathbf{X}}^\mathsf{T} \widetilde{\mathbf{X}} \mathbf{w}}{2\sigma^2} + \frac{\mathbf{y}^\mathsf{T} \widetilde{\mathbf{X}} \mathbf{w}}{\sigma^2} \right)$$

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$$= -\frac{\widetilde{\mathbf{X}}^{\mathsf{T}} \widetilde{\mathbf{X}} \mathbf{w}}{\sigma^{2}} + \frac{\widetilde{\mathbf{X}}^{\mathsf{T}} \mathbf{y}}{\sigma^{2}} = 0 \Leftrightarrow$$

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The gradient of the log-likelihood at the maximizer is zero.

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$$\frac{\partial \mathcal{L}(\boldsymbol{\theta})}{\partial \sigma^2} = \frac{\partial}{\partial \sigma^2} \left(-\frac{N}{2} \log(2\pi\sigma^2) - \frac{1}{2\sigma^2} \left(\mathbf{y} - \widetilde{\mathbf{X}} \mathbf{w} \right)^{\mathsf{T}} \left(\mathbf{y} - \widetilde{\mathbf{X}} \mathbf{w} \right) \right)$$

The gradient of the log-likelihood at the maximizer is zero.

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= -\frac{N}{2\sigma^2} + \frac{1}{2\sigma^4} \left(\mathbf{y} - \widetilde{\mathbf{X}} \mathbf{w} \right)^{\mathsf{T}} \left(\mathbf{y} - \widetilde{\mathbf{X}} \mathbf{w} \right) = 0$$

The gradient of the log-likelihood at the maximizer is zero.

$$\begin{split} \frac{\partial \mathcal{L}(\boldsymbol{\theta})}{\partial \mathbf{w}} &= \frac{\partial}{\partial \mathbf{w}} \left(-\frac{\mathbf{w}^\mathsf{T} \widetilde{\mathbf{X}}^\mathsf{T} \widetilde{\mathbf{X}} \mathbf{w}}{2\sigma^2} + \frac{\mathbf{y}^\mathsf{T} \widetilde{\mathbf{X}} \mathbf{w}}{\sigma^2} \right) \\ &= -\frac{\widetilde{\mathbf{X}}^\mathsf{T} \widetilde{\mathbf{X}} \mathbf{w}}{\sigma^2} + \frac{\widetilde{\mathbf{X}}^\mathsf{T} \mathbf{y}}{\sigma^2} = 0 \Leftrightarrow \mathbf{w} = \left(\widetilde{\mathbf{X}}^\mathsf{T} \widetilde{\mathbf{X}} \right)^{-1} \widetilde{\mathbf{X}}^\mathsf{T} \mathbf{y} \,. \end{split}$$

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The gradient of the log-likelihood at the maximizer is zero.

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$$\begin{split} \frac{\partial \mathcal{L}(\boldsymbol{\theta})}{\partial \sigma^2} &= \frac{\partial}{\partial \sigma^2} \left(-\frac{\textit{N}}{2} \log(2\pi\sigma^2) - \frac{1}{2\sigma^2} \left(\mathbf{y} - \widetilde{\mathbf{X}} \mathbf{w} \right)^\mathsf{T} \left(\mathbf{y} - \widetilde{\mathbf{X}} \mathbf{w} \right) \right) \\ &= -\frac{\textit{N}}{2\sigma^2} + \frac{1}{2\sigma^4} \left(\mathbf{y} - \widetilde{\mathbf{X}} \mathbf{w} \right)^\mathsf{T} \left(\mathbf{y} - \widetilde{\mathbf{X}} \mathbf{w} \right) = 0 \Leftrightarrow \\ \sigma^2 &= \frac{1}{\textit{N}} \left(\mathbf{y} - \widetilde{\mathbf{X}} \mathbf{w} \right)^\mathsf{T} \left(\mathbf{y} - \widetilde{\mathbf{X}} \mathbf{w} \right) \,. \end{split}$$

Problems of MLE

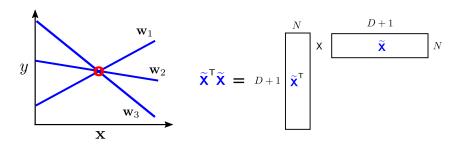
When N < D + 1 the MLE

$$\mathbf{w} = \left(\widetilde{\mathbf{X}}^\mathsf{T}\widetilde{\mathbf{X}}\right)^{-1}\widetilde{\mathbf{X}}^\mathsf{T}\mathbf{y}$$

is not defined. In this case...

Many values of **w** fit the training data equally well, achieving **zero error**.

The matrix $\tilde{\mathbf{X}}^T \tilde{\mathbf{X}}$ is low rank and not invertible:



Non-linear (basis function) regression

Linear regression can model non-linear relationships by replacing \mathbf{x} with some non-linear function of the inputs $\phi(\mathbf{x}) = (\phi_1(\mathbf{x}), \dots, \phi_M(\mathbf{x}))^T$.

Inference does not change, just replace each x_n with the new $\phi(x_n)$.

Example, **polynomials** for 1D data $\phi_m(x) = x^m$, m = 1, ..., M:

$$M = 0, \quad \phi(x) = [1]^{\mathsf{T}},$$

$$M = 1, \quad \phi(x) = [1, x]^{\mathsf{T}},$$

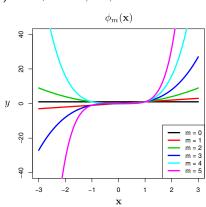
$$M = 2, \quad \phi(x) = [1, x, x^2]^{\mathsf{T}},$$

$$M = 3, \quad \phi(x) = [1, x, x^2, x^3]^{\mathsf{T}},$$

$$M = 4, \quad \phi(x) = [1, x, x^2, x^3, x^4]^{\mathsf{T}},$$

$$M = 5, \quad \phi(x) = [1, x, x^2, x^3, x^4, x^5]^{\mathsf{T}},$$

What should the value of M be?



1D example with polynomials

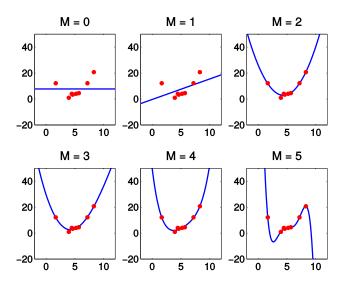
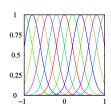


Figure: Z. Ghahramani.

Other basis functions

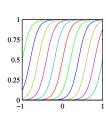
Gaussian radial basis functions with center \mathbf{c}_m and width s:

$$\phi_m(\mathbf{x}) = \exp\left\{-\frac{1}{2}s(\mathbf{x}, \mathbf{c}_m, s)^2\right\}$$
$$s(\mathbf{x}, \mathbf{c}_m, s) = \sqrt{(\mathbf{x} - \mathbf{c}_m)^{\mathsf{T}}(\mathbf{x} - \mathbf{c}_m)/s^2}.$$



Sigmoidal basis functions:

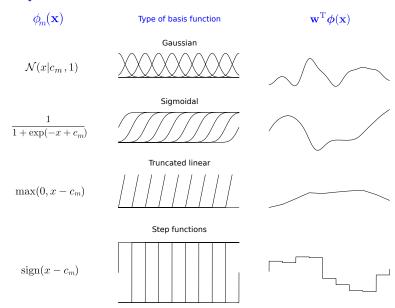
$$\phi_m(\mathbf{x}) = \sigma\left(s(\mathbf{x}, \mathbf{c}_m, s)\right)$$
$$\sigma(\mathbf{x}) = \frac{1}{1 + \exp(-\mathbf{x})}.$$



They are uniformly spread in input space to capture non-linearities everywhere.

Figure: C. Bishop. Pattern Recognition and Machine Learning, 2006.

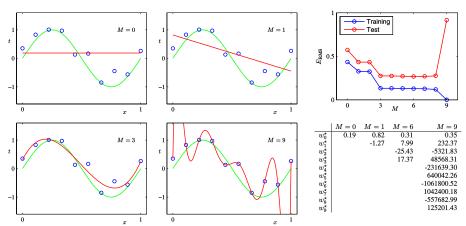
Examples



$$c_1 = -10, c_2 = -8, \ldots, c_9 = 8, c_{10} = 10, s = 1.$$

Overfitting

A large number of basis functions can lead to **over-fitting**: the model fits the **training data** well but it performs poorly on new **test data**.



Solution: use a prior distribution to enforce the entries of \boldsymbol{w} to be small.

Figures and table: C. Bishop. Pattern Recognition and Machine Learning, 2006.