

Assignment 2

1. M

2. M

3. Marginal

$$\begin{aligned}
f(x) &= \left(\frac{1}{\sqrt{2\pi D}} \exp\left(-\frac{(x-\mu)^2}{2D}\right) \right) * \left(\frac{1}{\sqrt{2\pi A}} \exp\left(-\frac{(\mu-M)^2}{2A}\right) \right) \\
&= \exp\left(-\frac{(x-\mu)^2}{2D} - \frac{(\mu-M)^2}{2A}\right) \\
&= \exp\left(-\frac{1}{2DA}(x^2 A - 2x\mu A + \mu^2 A + \mu^2 A - 2\mu MD + M^2 D)\right) \\
&= \exp\left(-\frac{1}{2DA}\mu^2(D+A) - 2\mu(xA + MD) + (x^2 A + \mu^2 D)\right) \\
&= \exp\left(-\frac{D+A}{2DA}\left(\mu^2 - \frac{2M(xA + MD)}{D+A} + \frac{x^2 A + M^2 D}{D+A}\right)\right) \\
&= \exp\left(-\frac{D+A}{2DA}\left(\mu - \frac{xA + MD}{D+A}\right)^2\right) * \exp\left(-\frac{D+A}{2DA}\left(\mu - \frac{xA + MD}{D+A}\right)^2\right)
\end{aligned}$$

4. Marginal density is $\alpha = (1-p)/p$ and $\beta = n$ The conditional density is $\mu \sim \text{gamma}$

5. The F distribution in this case relates to the Beta distribution with the equation

$$\frac{x_1}{x_1 + x_2}$$

6. By transforming it into a multivariate distribution making $x_2 \sim \text{Bin}(20, 0.5 * \theta)$ so we define

$$f(20, \theta_1, \theta_2) = \frac{20!}{x_1! x_2!} \theta_1^{x_1} \theta_2^{x_2}$$

We obtain the fisher information for both θ 's their combination an the hessian matrix

$$\begin{aligned}
I_{\theta} &= -E_{\theta} \left\{ \frac{\partial^2}{\partial \theta^2} \log \left(\frac{20!}{x_1! x_2!} \theta_1^{x_1} \theta_2^{x_2} \right) \right\} \\
&= -E_{\theta} \left\{ \frac{\partial^2}{\partial \theta^2} \log(20!) - \log(x_1!) - \log(x_2!) + \log(\theta_1^{x_1}) + \log(\theta_2^{x_2}) \right\} \\
&= -E_{\theta} \left\{ \frac{\partial^2}{\partial \theta^2} \log(\theta_1^{x_1}) + \log(\theta_2^{x_2}) \right\}
\end{aligned}$$

$$I_{\theta_1} = -E_{\theta_1} \left\{ \frac{\partial^2}{\partial \theta_1^2} x_1 \log(\theta_1) \right\} = -E_{\theta_1} \left\{ \frac{\partial^2}{\partial \theta_1^2} \frac{x_1}{\theta_1} \right\} = E_{\theta_1} \left\{ \frac{x_1}{\theta_1^2} \right\} = \frac{1}{\theta_1}$$

$$I_{\theta_2} = -E_{\theta_2} \left\{ \frac{\partial^2}{\partial \theta_2^2} x_2 \log(\theta_2) \right\} = -E_{\theta_2} \left\{ \frac{\partial^2}{\partial \theta_2^2} \frac{x_2}{\theta_2} \right\} = E_{\theta_2} \left\{ \frac{x_2}{\theta_2^2} \right\} = \frac{1}{\theta_2}$$

$$\begin{aligned} I_{\theta_{12}} = I_{\theta_{21}} &= -E_{\theta_{21}} \left\{ \frac{\partial^2}{\partial \theta_2 \partial \theta_1} \log(\theta_1^{x_1}) + \log(\theta_2^{x_2}) \right\} \\ &= -E_{\theta_{12}} \left\{ \frac{\partial^2}{\partial \theta_2 \partial \theta_1} \log(\theta_1^{x_1}) + \log(\theta_2^{x_2}) \right\} \\ &= -E_{\theta_{21}} \left\{ \frac{\partial^2}{\partial \theta_2 \partial \theta_1} x_1 \log(\theta_1) + \log x_2(\theta_2) \right\} \\ &= -E_{\theta_{12}} \left\{ \frac{\partial^2}{\partial \theta_2 \partial \theta_1} x_1 \log(\theta_1) + \log x_2(\theta_2) \right\} = -E_{\theta_{21}} \left\{ \frac{\partial}{\partial \theta_2} \left(\frac{x_1}{\theta_1} \right) \right\} \\ &= -E_{\theta_{21}} \left\{ \frac{\partial}{\partial \theta_1} \left(\frac{x_2}{\theta_2} \right) \right\} = -E_{\theta} \{0\} = 0 \end{aligned}$$

$$I_{\theta} = \begin{bmatrix} \frac{1}{\theta_1} & 0 \\ 0 & \frac{1}{\theta_2} \end{bmatrix}$$

Giving

$$\text{var}(\theta) \geq \frac{1}{n * I_{\theta}}$$

7. The relationship between them is that when the flat prior density $g(\theta) = 1$ then the posterior expectation of θ will be the same as the MLE θ

8. T

$$\begin{aligned} 9. \frac{f_{x_1}(x)}{f_{x_2}(x)} &= \frac{\frac{n!}{x_1!(n-x_1)!} x_1^x (1-x_1)^{n-x}}{\frac{n!}{x_2!(n-x_2)!} x_2^x (1-x_2)^{n-x}} = \frac{x_1^x (1-x_1)^{n-x}}{x_2^x (1-x_2)^{n-x}} = \left(\frac{x_1}{x_2} \right)^x * \frac{(1-x_1)^{n-x}}{(1-x_2)^{n-x}} = \left(\frac{x_1}{x_2} \right)^x * \frac{\left(\frac{1-x_1}{1-x_2} \right)^n}{\left(\frac{1-x_1}{1-x_2} \right)^x} = \\ &\quad \left(\frac{x_1(1-x_2)}{x_2(1-x_1)} \right)^x * \left(\frac{1-x_1}{1-x_2} \right)^n \end{aligned}$$

$$\begin{aligned} \alpha &= \log \left(\frac{x_1(1-x_2)}{x_2(1-x_1)} \right) & \Psi(\alpha) &= \left(\left(\frac{x_1}{x_2} \right) * \left(\frac{1}{e^{\alpha}} \right) \right)^n \\ \alpha &= \log \left(\frac{x_1(1-x_2)}{x_2(1-x_1)} \right) & \Psi(\alpha) &= \left(\left(\frac{x_1}{x_2} \right) * \left(\frac{1}{e^{\alpha}} \right) \right)^n \end{aligned}$$

10. With a Bayesian approach the way to estimate theta would be to first use the known probability of 0.5 heads and 0.5 tails as a prior in which case the conjugate prior of the binomial is the beta

$$P(\theta) = \frac{1}{B(h_1, t_1)} \theta^{h_1-1} (1-\theta)^{t_1-1}$$

Where h_1 is the times that the coin landed in heads and t_1 the times that it landed on tails.

For the probability of the event happening because we have a likelihood with binomial distribution and a beta prior we obtain that:

$$P(D) = {}_{h_2+t_2}C_{h_2} \frac{B(h_2 + h_1, t_2 + t_1)}{B(h_1, t_1)}$$

After obtaining this we join it to obtain the probability of obtaining heads that would be

$$P(\theta|D) = \frac{1}{B(h_2 + h_1, t_2 + t_1)} \theta^{h_2-1} (1 - \theta)^{t_2-1}$$