

Exponential Families

- Turns out all the classical ^{univariate} distributions that we have covered can be written as special cases of a more general construction.
↓
Exponential Family

Example:
→ Poisson pdf

$$f_{\mu}(x) = \frac{e^{-\mu} \mu^x}{x!}$$

Consider another Poisson pdf, with fixed parameter μ_0

$$f_{\mu_0}(x) = \frac{e^{-\mu_0} \mu_0^x}{x!}$$

if we divide them, we get

$$\frac{f_{\mu}(x)}{f_{\mu_0}(x)} = e^{-(\mu - \mu_0)} \left(\frac{\mu}{\mu_0}\right)^x$$

$$\rightarrow f_{\mu}(x) = e^{-(\mu - \mu_0)} \left(\frac{\mu}{\mu_0}\right)^x \cdot f_{\mu_0}(x)$$

→ we're writing $\text{Poi}(\mu)$ pdf in terms of the $\text{Poi}(\mu_0)$ pdf. we're "tilting" $f_{\mu_0}(x)$ by

We can rewrite this as

$$f_{\mu}(x) = e^{\alpha x - \psi(\alpha)} \cdot f_{\mu_0}(x)$$

where $\alpha = \log\left(\frac{\mu}{\mu_0}\right)$ and $\psi(\alpha) = \mu_0(e^{\alpha} - 1)$

$$\exp\left\{\log\left(\frac{\mu}{\mu_0}\right) \cdot x - \mu_0(e^{\log(\mu/\mu_0)} - 1)\right\}$$

$$\exp\left\{\log\left[\left(\frac{\mu}{\mu_0}\right)^x\right] - \mu_0\left(\frac{\mu}{\mu_0} - 1\right)\right\}$$

$$\exp\left\{\log\left[\left(\frac{\mu}{\mu_0}\right)^x\right] - (\mu - \mu_0)\right\} \quad \checkmark$$

Then, we can take any one Poisson Distribution $f_{\mu_0}(x)$ and for any value $\mu > 0$, let $\alpha = \log\left(\frac{\mu}{\mu_0}\right)$ and

calculate

$$\tilde{f}_{\mu}(x) = e^{\alpha x} f_{\mu_0}(x)$$

and divide by $e^{\psi(\alpha)}$ to get Poisson density $f_{\mu}(x)$.

Exponential tilt

$\psi(\cdot)$ not a function of x .

Then $e^{-\psi(\alpha)}$ is your renormalization constant.

In the same way that we write the Poisson pmf in exponential family form, we can write other distributions with more than 1 parameter as exponential families.

$$f_{\alpha}(x) = e^{\alpha^T y - \psi(\alpha)} \cdot f_0(x) \quad \text{for } x \in A$$

$A \subseteq \mathbb{R}^p$

" α is a p -dimensional vector of parameters".

$y = t(x)$: sufficient statistic vector.

The normalizing function $e^{-\psi(\alpha)}$

$$e^{-\psi(\alpha)} = \int_{\mathcal{X}} e^{\alpha^T y} f_0(x) dx$$

" $e^{-\psi(\alpha)}$ is the value of this integral, and we must divide over it for the pdf to integrate to 1 (or the pmf to sum to 1)."

For instance in the previous Poisson example, we got

$$e^{\eta(\alpha)} = \sum_0^{\infty} e^{\alpha x} f_{\mu_0}(x).$$

because x is discrete.

A 2-parameter example would be the Gamma distribution

$$f(x) = \frac{x^{v-1} e^{-x/\sigma}}{\sigma^v \Gamma(v)}$$

$$\sigma > 0, v > 0$$

$$\alpha = (\alpha_1, \alpha_2) = \left(-\frac{1}{\sigma}, v\right)$$

$$(y_1, y_2) = (x, \log x)$$

$$\begin{aligned} \text{and } \eta(\alpha) &= v \log \sigma + \log \Gamma(v) \\ &= -\alpha_2 \log \{-\alpha_1\} + \log \{\Gamma(\alpha_2)\} \end{aligned}$$

and do the algebra yourself...

→ Just like we can write 1 or 2-dim classical families in exponential family form, we can extend the dimensionality of α to get more general / flexible parametric distributions

Having said this, under repeated sampling,
we can take the exponential family pdf

$$f_{\alpha}(x) = e^{\alpha^T y - \psi(\alpha)} \cdot f_0(x_i) \quad (\text{density})$$

$$f_{\alpha}(\underline{x}) = \prod_{i=1}^n e^{\alpha^T y_i - \psi(\alpha)} f_0(x_i) \quad (\text{likelihood})$$

recall
 $y = t(x)$

$$= e^{\alpha^T \sum_{i=1}^n y_i - n\psi(\alpha)} f_0(x_i)$$

$$= e^{n(\alpha^T \bar{y} - \psi(\alpha))} f_0(x_i)$$

→ So then we have that once we have
defined $y = t(x)$, we only need its mean \bar{y}
to obtain its likelihood.

$$\bar{y} = \frac{\sum_{i=1}^n y_i}{n}$$

Argument 1

You can write many of
the classical pdfs/
pmfs

as exponential families

→ need clever substituting
transformations, etc.

Argument 2

→ you can come up
with almost any $t(x)$
you would like and
use likelihood to

fit for data.
the exp. family

Example of how to use the exponential family construction to fit data.

- use gfr data.

gfr

x_1

x_2

x_3

\vdots

x_n

(1) Define our sufficient statistic vector

$$y(x) = (x, x^2, x^3, \dots, x^7)$$

(2) This will allow us to fit a 7-parameter exponential family!
(Apply to each observation).