

Lecture 4, September 10, 2020

geometric distribution

$$0 < p < 1$$

$$q = 1 - p$$

$$\Xi \sim G(p)$$

$$\Xi \sim \begin{pmatrix} 1 & 2 & 3 & \dots & k & \dots \\ p & qp & q^2p & \dots & q^{k-1}p & \dots \end{pmatrix}$$

$$p + qp + q^2p + \dots = \frac{p}{1-q} = 1$$

$$E\Xi = \sum_{k=1}^{\infty} k q^{k-1} p = \frac{p}{(1-q)^2} = \frac{1}{p}$$

$$\sum_{k=1}^{\infty} q^k = \frac{q}{1-q} \quad -1 < q < 1$$

$$\checkmark \sum_{k=1}^{\infty} k q^{k-1} = \frac{1-q + q}{(1-q)^2} = \frac{1}{(1-q)^2}$$

$$\text{Var } \Xi = E\Xi^2 - (E\Xi)^2 = \frac{1+q}{p^2} - \left(\frac{1}{p}\right)^2 = \frac{q}{p^2}$$

$$E\Xi^2 = \sum_{k=1}^{\infty} k^2 q^{k-1} p = \frac{1+q}{(1-q)^2} = \frac{1+q}{p^2} \rightarrow \sum_{k=1}^{\infty} k^2 q^{k-1} = \frac{q+1}{(1-q)^3}$$

$$\sum_{k=1}^{\infty} k(k-1) q^{k-2} = \frac{2}{(1-q)^3}$$

$$\checkmark \sum_{k=1}^{\infty} (k^2 - k) q^{k-1} = \frac{2q}{(1-q)^3}$$

$$\xi \sim G(p)$$

$$P(\xi > k+n \mid \xi > n) \stackrel{=}{=} \quad k, n \in \mathbb{N}$$

$$k, n \in \mathbb{N}$$

$$\begin{aligned} P(\xi > n) &= P(\xi = n+1) + P(\xi = n+2) + \dots = \\ &= pq^n + pq^{n+1} + pq^{n+2} + \dots = \frac{pq^n}{1-q} = q^n \end{aligned}$$

$$\stackrel{=}{=} \frac{P(\xi > k+n \ \& \ \xi > n)}{P(\xi > n)} = \frac{P(\xi > k+n)}{P(\xi > n)} =$$

$$= \frac{q^{k+n}}{q^n} = q^k = P(\xi > k)$$

Independent random variables

X, Y

$$\forall A, B \rightarrow P(X \in A, Y \in B) = P(X \in A) \cdot P(Y \in B)$$

$$P(X=a, Y=b) = P(X=a) \cdot P(Y=b) \quad \forall a, b \in \mathbb{R}$$

$$A = \{a_1, a_2, \dots, a_n\}$$

$$P(X=a_1, Y=b) = P(X=a_1) \cdot P(Y=b)$$

$$+ \quad \frac{P(X=a_n, Y=b)}{P(X=a_n)} = P(Y=b)$$

$$P(X \in A, Y=b) = P(X \in A) \cdot P(Y=b)$$

$$P(X_1 \in A_1, X_2 \in A_2, \dots, X_n \in A_n) = \\ = P(X_1 \in A_1) \cdot P(X_2 \in A_2) \cdot \dots \cdot P(X_n \in A_n)$$

$$\forall A_1, A_2, \dots, A_n$$

$E(XY) = EX \cdot EY$ if X and Y are independent.

$$\hookrightarrow \sum_{i,j} x_i y_j P(X=x_i, Y=y_j) \stackrel{\text{indep.}}{=}$$

$$= \sum_{i,j} x_i y_j P(X=x_i) P(Y=y_j) =$$

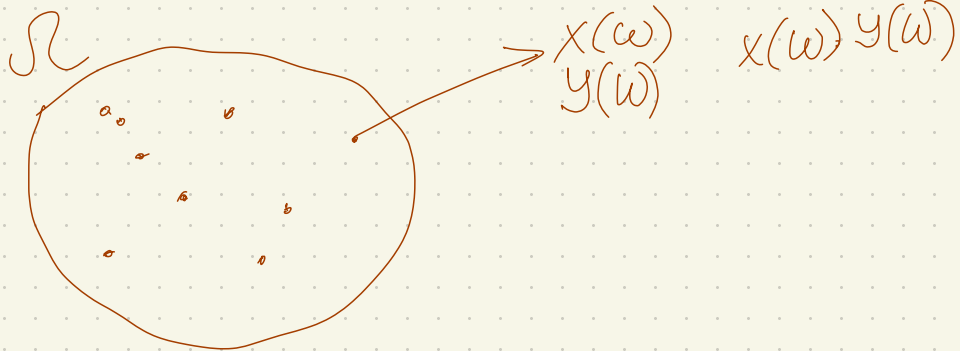
$$= \sum_i x_i P(X=x_i) \cdot \sum_j y_j P(Y=y_j) = EX \cdot EY$$

$X \backslash Y$	0	1
0	0,5	0,2
1	0,1	0,2

$$P(X=0, Y=1) = 0,2$$

$$P(X=0) = 0,7$$

$$P(Y=1) = 0,4$$



$$\sum_{i=1}^n a_i \cdot \sum_{j=1}^k b_j = \sum_{\substack{i \in [1, n] \\ j \in [1, k]}} a_i b_j$$

$$(a_1 + a_2 + \dots + a_n)(b_1 + b_2 + \dots + b_k)$$

$$E(x^2)$$

$$\text{Cov}(X, Y) = E((X - EX)(Y - EY)) = \text{covariance}$$

$$\begin{aligned} &= E(XY - YEX - XEY + EXEY) = \\ &= E(XY) - E(YEX) - E(XEY) + E(EXEY) = \\ &= E(XY) - EXEY - EXEY + EXEY \end{aligned}$$

If X and Y are independent then $\text{Cov}(X, Y) = 0$.

$$\text{Cov}(X, X) = E((X - EX)^2) = \text{Var } X.$$

$$\begin{aligned} \text{Var}(X \pm Y) &= E((X \pm Y)^2) - (E(X \pm Y))^2 = \\ &= E(X^2 \pm 2XY + Y^2) - (EX \pm EY)^2 = \underline{EX^2} \pm \\ &\pm 2E(XY) + \underline{EY^2} - \underline{(EX)^2} \pm 2EXEY - \underline{(EY)^2} = \\ &= \text{Var } X + \text{Var } Y \pm 2\text{Cov}(X, Y) \end{aligned}$$

If X, Y are independent then $\text{Var}(X \pm Y) =$
 $= \text{Var } X + \text{Var } Y$

Probability generating functions

$g_X(t) = E t^X$ (probability generating function of random variable X)

X takes nonnegative integer values only

$$X \sim \begin{pmatrix} 0 & 1 & 3 & 6 \\ 1/7 & 1/7 & 2/7 & 3/7 \end{pmatrix}$$

$$t^X \sim \begin{pmatrix} 1 & t & t^3 & t^6 \\ 1/7 & 1/7 & 2/7 & 3/7 \end{pmatrix} \quad g_X(t) = \frac{1}{7} + \frac{1}{7}t + \frac{2}{7}t^3 + \frac{3}{7}t^6$$

polynomial / power series, the sum of all coefficients is equal to 1

$$g'_X(t) = E(X t^{X-1})$$

$$g''_X(t) = E(X(X-1)t^{X-2})$$

$$t=1 \Rightarrow g'_X(1) = EX$$

$$t=1 \Rightarrow g''_X(1) = E(X^2 - X)$$

X and Y are independent r.v.

$$\begin{aligned} g_{X+Y}(t) &= E t^{X+Y} = E(t^X \cdot t^Y) = E t^X \cdot E t^Y = \\ &= g_X(t) \cdot g_Y(t) \end{aligned}$$

Poisson distribution

$X \sim \text{Po}(\lambda)$ if $X \in \mathbb{N} \cup \{0\}$, $\lambda > 0$

$$P(X=k) = \frac{\lambda^k}{k!} e^{-\lambda}$$

$$\sum_{k=0}^{\infty} \frac{\lambda^k}{k!} e^{-\lambda} = e^{-\lambda} \sum_{k=0}^{\infty} \frac{\lambda^k}{k!} = e^{-\lambda} \cdot e^{\lambda} = 1$$

$$g_X(t) = E t^X = \sum_{k=0}^{\infty} t^k \cdot \frac{\lambda^k}{k!} e^{-\lambda} = e^{-\lambda} \sum_{k=0}^{\infty} \frac{(t\lambda)^k}{k!} =$$

$$= e^{-\lambda} \cdot e^{t\lambda} = e^{\lambda(t-1)}$$

$$g'_X(t) = \lambda e^{\lambda(t-1)}, \quad g''_X(t) = \lambda^2 e^{\lambda(t-1)}$$

$$g'_X(1) = \lambda = EX, \quad g''_X(1) = \lambda^2 = E(X^2 - X)$$

$$\lambda^2 + \lambda = EX^2$$

$$\text{Var } X = EX^2 - (EX)^2 = \lambda^2 + \lambda - \lambda^2 = \lambda$$

$$EX = \text{Var } X = \lambda$$

$X \sim \text{Po}(\lambda)$, $Y \sim \text{Po}(\mu)$, X and Y are independent

$X+Y \sim ?$

$$g_{X+Y}(t) = g_X(t) \cdot g_Y(t) = e^{\lambda(t-1)} \cdot e^{\mu(t-1)} =$$

$$= e^{(\lambda+\mu)(t-1)} \Rightarrow X+Y \sim \text{Po}(\lambda+\mu)$$

Continuous distributions

$$\underbrace{F_X(x)}_{\parallel} = \underbrace{\int_{-\infty}^x f_X(t) dt}_{\text{probability density}}, \quad x \in \mathbb{R}$$

$P(X < x)$

$$F'_X(x) = f_X(x) \geq 0$$



$$\begin{aligned} P(a \leq x < b) &= P(x < b) - P(x < a) = F_X(b) - F_X(a) = \\ &= \int_{-\infty}^b f_X(t) dt - \int_{-\infty}^a f_X(t) dt = \int_a^b f_X(t) dt \end{aligned}$$



$$EX = \int_{-\infty}^{+\infty} x \underbrace{f_X(x) dx}$$

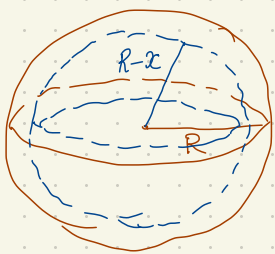
$$\int_{-\infty}^{+\infty} f_X(x) dx = 1$$

$$E h(x) = \int_{-\infty}^{+\infty} h(x) f_X(x) dx$$

$$E(CX) = C EX$$

$$E(X \pm Y) = EX \pm EY$$

$$E(C) = C$$



A random point (M) is taken inside the ball. ξ is equal to the distance from M to the sphere.
 $E\xi = ?$ $\text{Var}\xi = ?$

$$F_{\xi}(x) = P(\xi < x) = \begin{cases} 0, & x \leq 0 \\ \frac{R^3 - (R-x)^3}{R^3}, & 0 < x \leq R \\ 1, & x > R \end{cases}$$

$$f_{\xi}(x) = \begin{cases} 0, & x \leq 0 \\ \frac{3(R-x)^2}{R^3}, & 0 < x \leq R \\ 0, & x > R \end{cases} \quad f'_{\xi}(x) = \frac{3(R-x)^2}{R^3} I_{0 < x \leq R}$$

$$E\xi = \int_{-\infty}^{+\infty} x f_{\xi}(x) dx = \int_{-\infty}^{+\infty} x \cdot \frac{3(R-x)^2}{R^3} I_{0 < x \leq R} dx =$$

$$= \int_0^R \frac{3}{R^3} (R^2 x - 2Rx^2 + x^3) dx = \frac{3}{R^3} \left(\frac{R^2 x^2}{2} - \frac{2Rx^3}{3} + \frac{x^4}{4} \right) \Big|_0^R =$$

$$= \frac{3}{R^3} \left(\frac{R^4}{2} - \frac{2R^4}{3} + \frac{R^4}{4} \right) = \frac{R}{4}$$

$$E\xi^2 = \int_{-\infty}^{+\infty} x^2 f_{\xi}(x) dx = \int_0^R \frac{3}{R^3} (R^2 x^2 - 2Rx^3 + x^4) dx =$$

$$= \frac{3}{R^3} \left(\frac{R^2 x^3}{3} - \frac{Rx^4}{2} + \frac{x^5}{5} \right) \Big|_0^R = \frac{3}{R^3} \left(\frac{R^5}{3} - \frac{R^5}{2} + \frac{R^5}{5} \right) =$$

$$= \frac{R^2}{10}$$

$$\text{Var}\xi = \frac{R^2}{10} - \left(\frac{R}{4} \right)^2 = \frac{3R^2}{80}$$

$$I_A = \begin{cases} 1, & A \text{ has happened} \\ 0, & A \text{ has not happened} \end{cases}$$

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$$I(A)$$