Causal Inference in Python: A Vignette

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This document illustrates the use of CausalInference with a simple simulated data set. We begin with some basic definitions.

1 Setting and Notation

As is standard in the literature, we work within the framework of Rubin's potential outcome model (Rubin, 1974).

Let Y(0) denote the potential outcome of a subject in the absence of treatment, and let Y(1) denote the unit's potential outcome when it is treated. Let D denote treatment status, with D=1 indicating treatment and D=0 indicating control, and let X be a K-column vector of covariates or individual characteristics.

For unit i, i = 1, 2, ..., N, the observed outcome can be written as

$$Y_i = (1 - D_i)Y_i(0) + D_iY_i(1).$$

The set of observables (Y_i, D_i, X_i) , i = 1, 2, ..., N, forms the basic input data set for CausalInference. CausalInference is appropriate for settings in which treatment can be said to be *strongly ignorable*, as defined in Rosenbaum and Rubin (1983). That is, for all x in the support of X, we have

- (i) Unconfoundedness: D is independent of (Y(0), Y(1)) conditional on X = x;
- (ii) Overlap: c < P(D = 1|X = x) < 1 c, for some c > 0.

In the following, we illustrate the typical flow of a causal analysis using the tools of CausalInference and a simulated data set. In simulating the data, we specified a constant treatment effect of 10 for simplicity, and incorporated systematic overlap issues and nonlinearities to highlight a number of tools in the package. We focus mostly on illustrating the use of CausalInference; for details on methodology please refer to Imbens and Rubin (2015).

2 Initialization

The main object of interest in CausalInference is the class CausalModel, which we can import with

>>> from causalinference import CausalModel

CausalModel takes as inputs three NumPy arrays: Y, an N-vector of observed outcomes; D, an N-vector of treatment status indicators; and X, an N-by-K matrix of covariates. To initialize a CausalModel instance, simply run:

Once an instance of the class CausalModel has been created, it will contain a number of attributes and methods that are relevant for conducting a causal analysis. Tables 1 and 2 contain a brief description of these attributes and methods.

CausalModel is *stateful*. As we employ some of the methods to be discussed subsequently, the instance causal will mutate, with new data being added or existing data being modified or dropped. Running

will return causal to its initial state.

3 Summary Statistics

Once an instance of the class CausalModel has been created, basic summary statistics will be computed and stored in the attribute summary_stats. We can display it by running:

>>> print causal.summary_stats

Summary Statistics

	Controls	(N_c=392)	Treated	(N_t=608)	
Variable	Mean	S.d.	Mean	S.d.	Raw-diff
Y	43.097	31.353	90.911	41.815	47.814
	Controls	(N_c=392)	Treated	(N_t=608)	
Variable	Mean	S.d.	Mean	S.d.	Nor-diff
ХО	3.810	2.950	5.762	2.566	0.706
X1	3.436	2.848	5.849	2.634	0.880

The attribute summary_stats is in reality just a dictionary-like object with special method defined to enable the display of the above table. In many situations it is more convenient to simply access the relevant statistic directly. To retrieve the vector of covariate mean for the treatment group, for example, we simply run:

```
>>> causal.summary_stats['X_t_mean']
array([ 5.76232357,  5.8489734 ])
```

Since summary_stats behaves like a dictionary, it is equipped with the usual Python dictionary methods. To list the dictionary keys, for instance, we go:

```
>>> causal.summary_stats.keys()
['Y_c_mean', 'X_t_sd', 'N_t', 'K', 'ndiff', 'N', 'Y_t_sd', 'rdiff', 'Y_t_mean',
'X_c_mean', 'X_t_mean', 'Y_c_sd', 'X_c_sd', 'N_c']
```

Here rdiff refers to the difference in average observed outcomes between treatment and control groups. ndiff, on the other hand, refers to the normalized differences in average covariates, defined as

$$\frac{\bar{X}_{k,t} - \bar{X}_{k,c}}{\sqrt{\left(s_{k,t}^2 + s_{k,c}^2\right)/2}},$$

where $\bar{X}_{k,t}$ and $s_{k,t}$ are the sample mean and sample standard deviation of the kth covariate of the treatment group, and $\bar{X}_{k,c}$ and $s_{k,c}$ are the analogous statistics for the control group.

The normalized differences in average covariates provide a way to measure the covariate balance between the treatment and the control groups. Unlike the t-statistic, its absolute magnitude does not increase (in expectation) as the sample size increases.

4 Least Squares Estimation

One of the simplest treatment effect estimators is the ordinary least squares (OLS) estimator. Causal-Inference provides several common regression specification.

By default, the method est_via_ols will run the following regression:

$$Y_i = \alpha + \beta D_i + \gamma'(X_i - \bar{X}) + \delta' D_i(X_i - \bar{X}) + \varepsilon_i.$$

To inspect any treatment effect estimates produced, we can simply invoke print on the attribute estimates, as in below:

```
>>> causal.est_via_ols()
>>> print causal.estimates
```

Treatment Effect Estimates: OLS

 	Est.	S.e.	z	P> z	[95% Con	f. int.]
ATE	3.672	0.906	4.051	0.000	1.895	5.449
ATC	-0.227	0.930	-0.244	0.807	-2.050	1.596
ATT	6.186	1.067	5.799	0.000	4.095	8.277

Here ATE, ATC, and ATT stand for, respectively, average treatment effect, average treatment effect for the controls, and average treatment effect for the treated. Like summary_stats, the attribute estimates is a dictionary-like object that contains the estimation results.

Including interaction terms between the treatment indicator D and covariates X implies that treatment effects can differ across individuals. In some instances we may want to assume a constant treatment effect, and only run

$$Y_i = \alpha + \beta D_i + \gamma'(X_i - \bar{X}_i) + \varepsilon_i.$$

This can be achieved by supplying a value of 1 in est_via_ols to the optional parameter adj (its default value is 2). To compute the raw difference in average outcomes between treatment and control groups, we can set adj=0.

In this example, the least squares estimates are radically different from the true treatment effect of 10. This is the result of the nonlinearity and non-overlap issues intentionally introduced into the data simulation process. As we shall see, several other tools exist in CausalInference that can better deal with a lack of overlap and that will allow us to obtain estimates that are less sensitive to functional form assumptions.

5 Propensity Score Estimation

The probability of getting treatment conditional on the covariates, $p(X_i) = P(D_i = 1|X_i)$, also known as the propensity score, plays a central role in much of what follows. Two methods, est_propensity and est_propensity_s, are provided for propensity score estimation. Both involve running a logistic regression of the treatment indicator D on functions of the covariates. est_propensity allows the user to specify the covariates to include linearly and/or quadratically, while est_propensity_s will make this choice automatically based on a sequence of likelihood ratio tests.

In the following, we run est_propensity_s and display the estimation results. In this example, the specification selection algorithm decided to include both covariates and all the interaction and quadratic terms.

- >>> causal.est_propensity_s()
- >>> print causal.propensity

Estimated Parameters of Propensity Score

	Coef.	S.e.	z	P> z	[95% C	Conf. int.]
Intercept	-2.839	0.526	-5.401	0.000	-3.870	-1.809
X1	0.486	0.153	3.178	0.001	0.186	0.786
XO	0.466	0.155	3.011	0.003	0.163	0.770
X1*X0	0.080	0.015	5.391	0.000	0.051	0.109
X0*X0	-0.045	0.012	-3.579	0.000	-0.069	-0.020
X1*X1	-0.045	0.013	-3.542	0.000	-0.070	-0.020

The propensity attribute is again another dictionary-like container of results. The dictionary keys of propensity can be found by running:

```
>>> causal.propensity.keys()
['coef', 'lin', 'qua', 'loglike', 'fitted', 'se']
```

The selected linear and quadratic terms are contained in the lists causal.propensity['lin'] and causal.propensity['qua']. Though we won't make direct calls to it, most of the propensity-based techniques discussed subsequently are based on causal.propensity['fitted'], the vector of estimated propensity scores.

6 Improving Covariate Balance

When there is indication of covariate imbalance, we may wish to construct a sample where the treatment and control groups are more similar than the original full sample. One way of doing so is by dropping units with extreme values of propensity score. For these subjects, their covariate values are such that the probability of being in the treatment (or control) group is so overwhelmingly high that we cannot reliably find comparable units in the opposite group. We may wish to forego estimating treatment effects for such units since nothing much can be credibly said about them.

A good rule-of-thumb is to drop units whose estimated propensity score is less than $\alpha = 0.1$ or greater than $1 - \alpha = 0.9$. By default, once the propensity score has been estimated by running either est_propensity or est_propensity_s, a value of 0.1 will be set for the attribute cutoff:

```
>>> causal.cutoff
0.1
```

Calling causal.trim() at this point will drop every unit that has propensity score outside of the $[\alpha, 1-\alpha]$ interval. Alternatively, a procedure exists that will estimate the optimal cutoff that minimizes the asymptotic sampling variance of the trimmed sample. The method trims will perform this calculation, set the cutoff to the optimal α , and then invoke trim to construct the subsample. For our example, the optimal α was estimated to be slightly less than 0.1:

- >>> causal.trim_s()
- >>> causal.cutoff
- 0.0954928016329

The complexity of this cutoff selection algorithm is only O(N), i.e., linear in the sample size, so in practice there is very little reason to not employ it.

If we now print summary_stats again to view the summary statistics of the trimmed sample, we see that the normalized differences in average covariates has fallen noticeably.

>>> print causal.summary_stats

Summary Statistics

	Controls	(N_c=371)	Treated	(N_t=363)	
Variable	Mean	S.d.	Mean	S.d.	Raw-diff
Y	41.331	29.608	66.067	28.108	24.736
	Controls	(N_c=371)	Treated	(N_t=363)	
Variable	Mean	S.d.	Mean	S.d.	Nor-diff
				B.u.	NOI GIII
ХО	3.709	2.872	4.658	2.522	0.351

7 Stratifying the Sample

With the propensity score estimated, one may wish to stratify the sample into blocks that have units that are more similar in terms of their covariates. This makes the treatment and control groups within each propensity bin more comparable, and therefore treatment effect estimates more credible.

CausalInference provides two methods for subclassification based on propensity score. The first, stratify, splits the sample based on what is specified in the attribute blocks. The default value of blocks is set to 5, which means that stratify will split the sample into 5 equal-sized bins. In contrast, the second method, stratify_s, will use a data-driven procedure for selecting both the number of blocks and their boundaries, with the expectation that the number of blocks should increase with the sample size. Operationally this method is a divide-and-conquer algorithm that recursively divides the sample into two until there is no significant advantage of doing so. The algorithm runs in $O(N \log N)$ time, so costs relatively little to use.

To inspect the results of the stratification, we can invoke print on the attribute strata to display some summary statistics, as follows:

```
>>> causal.stratify_s()
```

Stratification Summary

		Propensi	ty Score	Sa	mple Size	Ave. P	ropensity	Outcome
	Stratum	Min.	Max.	Controls	Treated	Controls	Treated	Raw-diff
-								
	1	0.095	0.265	157	28	0.188	0.187	11.885
	2	0.266	0.474	111	72	0.360	0.367	12.025
	3	0.477	0.728	70	113	0.598	0.601	11.696
	4	0.728	0.836	23	69	0.781	0.787	10.510
	5	0.838	0.904	10	81	0.865	0.873	3.405

Under the hood, the attribute strata is actually a list-like object that contains, as each of its elements, a full instance of the class CausalModel, with the input data being those that correspond to the units that are in the propensity bin. We can thus, for example, access each stratum and inspect its summary_stats attribute, or as the following illustrates, loop through strata and estimate within-bin treatment effects using least squares.

```
>>> for stratum in causal.strata:
...     stratum.est_via_ols(adj=1)
...
>>> [stratum.estimates['ols']['ate'] for stratum in causal.strata]
[10.379170390195197, 9.2918973715823707, 9.67876709257445, 9.6722830043583023,
9.2239596078238222]
```

Note that these estimates are much more stable and closer to the true value of 10 than the within-bin raw differences in average outcomes that were reported in the stratification summary table, highlighting the virtue of further controlling for covariates even within blocks.

Taking the sample-weighted average of the above within-bin least squares estimates results in a propensity score matching estimator that is commonly known as the subclassification estimator or blocking estimator. However, instead of manually looping through the strata attribute, estimating within-bin treatment effects, and then averaging appropriately to arrive at an overall estimate, we can also simply call est_via_blocking, which will perform these operations and collect the results in the attribute estimates. We will report these estimates in the next section along with estimates obtained from other, alternative estimators.

>>> print causal.strata

8 Treatment Effect Estimation

In addition to least squares and the blocking estimator described in the last section, CausalInference provides two alternative treatment effect estimators. The first is the nearest neighborhood matching estimator of Abadie and Imbens (2006). Instead of relying on the propensity score, this estimator pairs treatment and control units by matching directly on the covariate vectors themselves. More specifically, each unit i in the sample is matched with a unit m(i) in the opposite group, where

$$m(i) = \operatorname*{argmin}_{j:D_j \neq D_i} \|X_j - X_i\|,$$

and $||X_j - X_i||$ is some measure of distance between the covariate vectors X_j and X_i . The method est_via_matching implements this estimator, as well as several extensions that can be invoked through optional arguments.

The last estimator is a version of the Horvitz-Thompson weighting estimator, modified to further adjust for covariates. Mechanically, this involves running the following weight least squares regression:

$$Y_i = \alpha + \beta D_i + \gamma' X_i + \varepsilon_i,$$

where the weight for unit i is $1/\hat{p}(X)$ if i is in the treatment group, and $1/(1-\hat{p}(X))$ if i is in the control group. This estimator is also sometimes called the doubly-robust estimator, referring to the fact that this estimator is consistent if either the specification of the propensity score is correct, or the specification of the regression function is correct. We can invoke it by calling est_via_weighting. Note that under this specification the treatment effect does not differ across units, so the ATC and the ATT are both equal to the overall ATE.

In the following we invoke each of the four estimators (including least squares, since the input data has changed now that the sample has been trimmed), and print out the resulting estimates.

- >>> causal.est_via_ols()
- >>> causal.est_via_weighting()
- >>> causal.est_via_blocking()
- >>> causal.est_via_matching(bias_adj=True)
- >>> print causal.estimates

Treatment Effect Estimates: OLS

	Est.	S.e.	Z	P> z	[95% Con	f. int.]
 ATE	2.913	0.803	3.627	0.000	1.339	4.487
ATC	2.435	0.824	2.956	0.003	0.820	4.049
ATT	3.401	0.885	3.843	0.000	1.667	5.136

Treatment Effect Estimates: Weighting

		Est.	S.e.	z	P> z	[95% Cor	nf. int.]
	ATE	17.821	1.684	10.585	0.000	14.521	21.121
Treatment	Effect	Estimates:	Blocking				
		Est.	S.e.	z	P> z	[95% Cor	nf. int.]
	ATE	9.702	0.381	25.444	0.000	8.954	10.449
	ATC	9.847	0.527	18.701	0.000	8.815	10.879
	ATT	9.553	0.332	28.771	0.000	8.903	10.204

Treatment Effect Estimates: Matching

	Est.	S.e.	z	P> z	[95% (Conf. int.]
ATE	9.624	0.245				10.103
ATC ATT	9.642 9.606	0.270 0.318	35.776 30.159	0.000		10.170 10.230

As we can see above, despite the trimming the least squares estimates are still severely biased, as is the weighting estimator (since neither the propensity score or the regression function is correctly specified). The blocking and matching estimators, on the other hand, are less sensitive to specification assumptions, and thus result in estimates that are closer to the true average treatment effects.

References

- Abadie, A., & Imbens, G. (2006). Large sample properties of matching estimators for average treatment effects. *Econometrica*, 74, 235-267.
- Crump, R., Hotz, V. J., Imbens, G., & Mitnik, O. (2009). Dealing with limited overlap in estimation of average treatment effects. *Biometrika*, 96, 187-199.
- Imbens, G. W., & Rubin, D. B. (2015). Causal inference in statistics, social, and biomedical sciences:

 An introduction. Cambridge University Press.
- Rosenbaum, P. R., & Rubin, D. B. (1983). The central role of the propensity score in observational studies for causal effects. *Biometrika*, 70, 41-55.
- Rubin, D. B. (1974). Estimating Causal Effects of Treatments in Randomized and Nonrandomized Studies. *Journal of Educational Psychology*, 66, 688-701.

Attribute	Description
$summary_stats$	Dictionary-like object containing summary statistics for the
	outcome variable and the covariates.
propensity	Dictionary-like object containing propensity score data,
	including estimated logistic regression coefficients, predicted
	propensity score, maximized log-likelihood, and the lists of the
	linear and quadratic terms that are included in the regression.
cutoff	Floating point number specifying the cutoff point for trimming
	on propensity score.
blocks	Either an integer indicating the number of equal-sized blocks to
	stratify the sample into, or a list of ascending numbers specifying
	the boundaries of each stratum.
strata	List-like object containing the list of stratified propensity bins.
estimates	Dictionary-like object containing treatment effect estimates for

Table 1: Attributes of the class CausalModel. Invoking print on any of the dictionary- or list-like attribute above yields customized summary tables.

each estimator used.

Method	Description
reset	Reinitializes data to original inputs, and drops any estimated results.
${\sf est_propensity}$	Estimates via logistic regression the propensity score using specified
	linear and quadratic terms.
${\sf est_propensity_s}$	Estimates via logistic regression the propensity score using the
	covariate selection algorithm of Imbens and Rubin (2015).
trim	Trims data based on propensity score using the threshold specified
	by the attribute cutoff.
trim_s	Trims data based on propensity score using the cutoff selected by
	the procedure of Crump, Hotz, Imbens, and Mitnik (2009).
stratify	Stratifies the sample based on propensity score as specified by
	the attribute blocks.
${ t stratify_s}$	Stratifies the sample based on propensity score using the bin
	selection procedure suggested by Imbens and Rubin (2015).
${\sf est_via_ols}$	Estimates average treatment effects using least squares.
${\sf est_via_weighting}$	Estimates average treatment effects using the Horvitz-Thompson
	weighting estimator modified to incorporate covariates.
${\sf est_via_blocking}$	Estimates average treatment effects using regression within blocks.
${\sf est_via_matching}$	Estimates average treatment effects using matching with replacement.

Table 2: Methods of the class CausalModel. Invoke help on any of the above methods for more detailed documentation.