# BigDataHW3

### Alex Fay

## February 2020

### Questions 1

Suppose 
$$\theta$$
 Beta $(a,b)$  such that  $P(\theta;a,b) = \frac{1}{B(a,b)*\theta^{b-1}} = \tau(a)\tau(b)*\theta^{a-1}(1-\theta)^{b-1}$ 

Derive the mean, mode, and variance of theta where B(a,b) is the Beta function and  $\tau(x)$  is the Gamma Function.

Mean: The mean can be used using the mean value theorem.

 $Mean = \frac{1}{a-b}P(\theta)\theta d\theta$  where a is 1 and b is 0 Substituting our knowns above,

$$\begin{array}{l} \int_0^1 (\frac{\theta^{a-1}*((1-\theta)^{b-1}))}{B(a,b)}*\theta d\theta) \\ \text{Since theta} \quad B(a,b) \text{ then we know:} \end{array}$$

$$\tau(a+1*\tau(b)*\tfrac{1}{\tau(a+b+1)}*\tfrac{\tau(a+b)}{\tau(a)\tau(b)}$$
 Thus the mean is  $=\tfrac{a}{a+b}$ 

### Mode:

From finding the gradient vector, we can use this to find the maximum or minimum of the function when the slope is equal to 0.

$$\begin{split} 0 &= \nabla [\theta^{a-1} * (1-\theta^{b-1})] \\ &= (a-1) * \theta^{(a} - 2)(1-\theta)^{b-1} - (b-1) * \theta^{a-1} * (1-\theta)^{b-2} \\ &= (a-1)\theta^{a-2}(1-\theta)^{b-1} = (b-1)\theta^{a-1}(1-\theta)^{b-2} \\ (a-1) &= (a+b-2) * \theta \\ \theta' &= \frac{(a-1)}{(a+b-2)} \end{split}$$

Variance: 
$$Var[\theta] = E[\theta] - E[\theta]^2$$

$$E[\theta] = \frac{a}{a+b} \text{ given above}$$

$$E[\theta]^2 = \frac{a^2}{(a+b)^2}$$

$$\begin{split} &E[\theta^2] = \int_0^1 \theta^2 * P(\theta, a, b) d\theta \\ &= \frac{1}{B(a, b) \int_0^1 \theta^{a+1} (1-\theta)^{b-1} d\theta} \\ &= \frac{B(a+2, b)}{B(a, b)} \end{split}$$

Using similar method to part A: 
$$= \frac{a(a+1)}{(a+b)(a+b+1)}$$
 
$$Var[\theta] = \frac{a(a+1)}{(a+b)(a+b+1)} - \frac{a^2}{(a+b)^2}$$

From simplifying the above we find the Variance is:  $= \frac{ab}{(a+b)^2(a+b+1)}$ 

2. Show that the multinomial distribution is in the exponential family and that the generalized linear model corresponding to this distribution is the same as multinomial log regression.

$$\begin{split} &Cat(x|\mu) = \prod_i = 1_K \mu_i^{xi} \\ &= e^{\log(\prod_i = 1_K \mu_i^{xi})} \\ &= e^{\sigma_{i=1}^K x_i * log(\mu_i)} \\ &\text{Since } \Sigma_{i=1}^K \mu_i = 1 = \Sigma_{1K} x \text{ then:} \\ &= e^{\Sigma_{i=1}^K - 1 x_i log(\mu_i) + (1 - \Sigma_{i=1K} - 1 x_i) * log(\mu_k)} \\ &1 - \sigma_{i=1}^K - 1 x_i \text{ becomes } log\mu_k \\ &\text{Thus from properties of logs:} \\ &e^{\sum_{i=1}^K - 1 x_i * log(\frac{\mu_i}{\mu_k) + log(\mu_K)}} \\ &\text{Since from the above, } \mu_i = \mu_k * e^n thenn = [log(\frac{\mu_i}{\mu_k}); log(\frac{\mu_{k-1}}{\mu_k})] \\ &\text{b(n)} = \mathbf{x} \text{ and } \mathbf{a}(\mathbf{n}) = -log(\Sigma_{i=1}^K - 1e^n) \end{split}$$

Thus, we know the multinomial distribution is part of the exponential family and since the mu is the multinomial log function then it has a multinomial log/softmax regression.