Survey on approximation methods for discrete scan statistics: a software illustration

Alexandru Amărioarei, Cristian Preda

Laboratoire de Mathématiques Paul Painlevé Département de Probabilités et Statistique Université de Lille 1, INRIA/Modal Team, France

International Workshop on Applied Probability 16-19 June, 2014, Antalya, Turkey



- Introduction
 - Framework
 - Problem
- Approximations and bounds for scan statistics
 - One dimensional scan statistics
 - Two dimensional scan statistics
 - Three dimensional scan statistics
- Scan statistics and 1-dependent sequences
 - Methodology
 - Numerical examples
- 4 Some dependent models for scan statistics
 - Model and discussion
 - Example (d = 1): Longest monotone run
- Seferences



- Introduction
 - Framework
 - Problem
- - One dimensional scan statistics
 - Two dimensional scan statistics
 - Three dimensional scan statistics
- - Methodology
 - Numerical examples
- - Model and discussion
 - Example (d=1): Longest monotone run





The d-dimensional discrete scan statistics

Let T_1, T_2, \ldots, T_d be positive integers, with $d \geq 1$

- The rectangular region, $\mathcal{R}_d = [0, T_1] \times [0, T_2] \times \cdots \times [0, T_d]$
- The r.v.'s $X_{s_1,s_2,...,s_d}$, $1 \le s_j \le T_j$, $j \in \{1,2,...,d\}$

Let $2 \le m_i \le T_i$, $1 \le j \le d$, be positive integers

• Define for $1 \le i_l \le T_l - m_l + 1$, $1 \le l \le d$,

$$Y_{i_1,i_2,\dots,i_d} = \sum_{s_1=i_1}^{i_1+m_1-1} \sum_{s_2=i_2}^{i_2+m_2-1} \cdots \sum_{s_d=i_d}^{i_d+m_d-1} X_{s_1,s_2,\dots,s_d}$$

• The d-dimensional discrete scan statistic,

$$S_{\mathbf{m}}(\mathsf{T}) = \max_{\substack{1 \leq i_j \leq T_j - m_j + 1 \ j \in \{1, 2, \dots, d\}}} Y_{i_1, i_2, \dots, i_d}$$

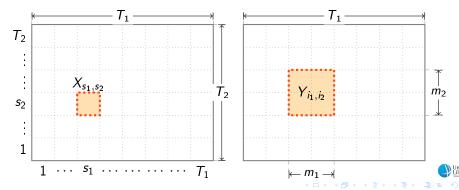
with ${\bf m}=(m_1,m_2,\ldots,m_d)$ and ${\bf T}=(T_1,T_2,\ldots,T_d)$



Example: two dimensional scan statistics (d = 2)

We have for d=2

$$Y_{i_1,i_2} = \sum_{s_1=i_1}^{i_1+m_1-1} \sum_{s_2=i_2}^{i_2+m_2-1} X_{s_1,s_2}, \quad S_{m_1,m_2}(T_1,T_2) = \max_{\substack{1 \leq i_1 \leq T_1-m_1+1 \\ 1 \leq i_2 \leq T_2-m_2+1}} Y_{i_1,i_2}$$



- Introduction
 - Framework
 - Problem
- Approximations and bounds for scan statistics
 - One dimensional scan statistics
 - Two dimensional scan statistics
 - Three dimensional scan statistics
- Scan statistics and 1-dependent sequences
 - Methodology
 - Numerical examples
- 4 Some dependent models for scan statistics
 - Model and discussion
 - Example (d = 1): Longest monotone run
- 6 References





Problem

Goal

Find a good estimate for the distribution of d-dimensional discrete scan statistic

$$Q_{\mathbf{m}}(\mathsf{T}) = \mathbb{P}\left(S_{\mathbf{m}}(\mathsf{T}) \leq \tau\right)$$

The distribution of $S_{\mathbf{m}}(\mathsf{T})$ is used for testing the null hypotheses of randomness against the alternative hypothesis of clustering.

Example: Bernoulli model

 H_0 : The r.v.'s $X_{s_1,s_2,...,s_d}$ are i.i.d. $\mathcal{B}(p)$

 H_1 : There exists $\mathcal{R}(i_1, i_2, \dots, i_d) = [i_1 - 1, i_1 + m_1 - 1] \times \dots \times [i_d - 1, i_d + m_d - 1] \subset \mathcal{R}_d$ where the r.v.'s $X_{s_1,s_2,...,s_d} \sim \mathcal{B}(p')$, p' > p and $X_{s_1,s_2,...,s_d} \sim \mathcal{B}(p)$ outside $\mathcal{R}(i_1, i_2, \dots, i_d)$

- Introduction
 - Framework
 - Problem
- 2 Approximations and bounds for scan statistics
 - One dimensional scan statistics
 - Two dimensional scan statistics
 - Three dimensional scan statistics
- Scan statistics and 1-dependent sequences
 - Methodology
 - Numerical examples
- Some dependent models for scan statistics
 - Model and discussion
 - Example (d = 1): Longest monotone run
- 6 References





Exact methods (Bernoulli case)

There are three main approaches used for investigating the exact distribution of the one dimensional discrete scan statistics over a sequence of binary trials:

- Combinatorial method: [Naus, 1974], [Naus, 1982]
- Finite Markov chain imbedding technique: [Fu, 2001],
 [Balakrishnan and Koutras, 2002], [Fu and Lou, 2003], [Wu, 2013],
 [Fu et al., 2012] etc.
- Conditional generating function method: [Ebneshahrashoob and Sobel, 1990], [Ebneshahrashoob et al., 2005], [Shinde and Kotwal, 2008] etc.



Product Type Approximation

Considering that $T_1=L_1m_1$, [Naus, 1982] gave the following approximation

$$\mathbb{P}\left(S_{m_1}(T_1) \leq \tau\right) \approx Q(2m_1) \left[\frac{Q(3m_1)}{Q(2m_1)}\right]^{\frac{T_1}{m_1}-2},$$

where $Q(2m_1)=\mathbb{P}\left(S_{m_1}(2m_1)\leq au\right)$ and $Q(3m_1)=\mathbb{P}\left(S_{m_1}(3m_1)\leq au\right)$.

$$\begin{split} Q(2m_1) &= F^2(\tau; m_1, \rho) - \tau b(\tau + 1; m_1, \rho) F(\tau - 1; m_1, \rho) + m_1 \rho b(\tau + 1; m_1, \rho) F(\tau - 2, m_1 - 1), \\ Q(3m_1) &= F^3(\tau; m_1, \rho) - A_1 + A_2 + A_3 - A_4, \\ A_1 &= 2b(\tau + 1; m_1, \rho) F(\tau; m_1, \rho) [\tau F(\tau - 1; m_1, \rho) - m_1 \rho F(\tau - 2; m_1 - 1, \rho)], \\ A_2 &= 0.5b^2(\tau + 1; m_1, \rho) \left[\tau(\tau - 1) F(\tau - 2; m_1, \rho) - 2(\tau - 1) m_1 F(\tau - 3; m_1 - 1, \rho) + m_1(m_1 - 1)\rho^2 F(\tau - 4; m_1 - 2, \rho)\right], \\ A_3 &= \sum_{r=1}^{\tau} b(2(\tau + 1) - r; m_1, \rho) F^2(r - 1; m_1, \rho), \\ A_4 &= \sum_{r=2}^{\tau} b(2(\tau + 1) - r; m_1, \rho) b(r + 1; m_1, \rho) [r F(r - 1; m_1, \rho) - m_1 \rho F(r - 2; m_1 - 1, \rho)]. \end{split}$$





Product Type Approximation for Binomial and Poisson

If $X_i \sim Bin(n, p)$ or $X_i \sim Pois(\lambda)$, we have the approximation

$$\mathbb{P}\left(S_{m_{1}}(T_{1}) \leq \tau\right) \approx Q(2m_{1}) \left[\frac{Q(3m_{1})}{Q(2m_{1})}\right]^{\frac{T_{1}}{m_{1}}-2}, T_{1} \geq 3m_{1}$$

$$\approx Q(2m_{1}-1) \left[\frac{Q(2m_{1})}{Q(2m_{1}-1)}\right]^{T_{1}-2m_{1}+1}, T_{1} \geq 2m_{1}$$

where $Q(2m_1-1)$, $Q(2m_1)$ and $Q(3m_1)$ are computed by [Karwe and Naus, 1997] recurrence.

▶ Karwe Naus algorithm for $Q(2m_1 - 1)$ and $Q(2m_1)$



Bounds

[Glaz and Naus, 1991] developed a variety of tight bounds:

Lower Bounds

$$\mathbb{P}\left(S_{m_{1}}(T_{1}) \leq \tau\right) \leq \frac{Q(2m_{1})}{\left[1 + \frac{Q(2m_{1}-1) - Q(2m_{1})}{Q(2m_{1}-1)Q(2m_{1})}\right]^{T_{1}-2m_{1}}}, T_{1} \geq 2m_{1}$$

$$\leq \frac{Q(3m_{1})}{\left[1 + \frac{Q(2m_{1}-1) - Q(2m_{1})}{Q(3m_{1}-1)}\right]^{T_{1}-3m_{1}}}, T_{1} \geq 3m_{1}$$

Upper Bounds

$$\mathbb{P}\left(S_{m_1}(T_1) \leq \tau\right) \leq Q(2m_1)\left[1 - Q(2m_1 - 1) + Q(2m_1)\right]^{T_1 - 2m_1}, \ T_1 \geq 2m_1$$

$$\leq Q(3m_1)\left[1 - Q(2m_1 - 1) + Q(2m_1)\right]^{T_1 - 3m_1}, \ T_1 \geq 3m_1$$

The values $Q(2m_1-1)$, $Q(2m_1)$, $Q(3m_1-1)$, $Q(3m_1)$ are computed using [Karwe and Naus, 1997] algorithm.

- Introduction
 - Framework
 - Problem
- 2 Approximations and bounds for scan statistics
 - One dimensional scan statistics
 - Two dimensional scan statistics
 - Three dimensional scan statistics
- Scan statistics and 1-dependent sequences
 - Methodology
 - Numerical examples
- Some dependent models for scan statistics
 - Model and discussion
 - Example (d = 1): Longest monotone run
- 6 References





Product Type Approximation Bernoulli Case

[Boutsikas and Koutras, 2000] using Markov Chain Imbedding approach proposed the approximation

$$\mathbb{P}\left(S_{m_1,m_2}(T_1,T_2) \leq \tau\right) \approx \frac{Q(m_1,m_2)(T_1-m_1-1)(T_2-m_2-1)}{Q(m_1,m_2+1)(T_1-m_1-1)(T_2-m_2)} \frac{Q(m_1+1,m_2+1)(T_1-m_1)(T_2-m_2)}{Q(m_1+1,m_2)(T_1-m_1)(T_2-m_2-1)}$$

Where,

$$\begin{split} Q(m_1,m_2) &= F\left(\tau; m_1 m_2, p\right) \\ Q(m_1+1,m_2) &= \sum_{s=0}^{\tau} F^2\left(\tau-s; m_2, p\right) b\left(s; (m_1-1) m_2, p\right) \\ Q(m_1+1,m_2+1) &= \sum_{s_1,s_2=0}^{\tau} \sum_{t_1,t_2=0}^{\tau} \sum_{i_1,i_2,i_3,i_4=0}^{1} b(s_1; m_1-1, p) b(s_2; m_1-1, p) b(t_1; m_2-1, p) \times \\ b(t_2; m_2-1, p) p^{\sum_{s_1} i_1} \left(1-p\right)^{4-\sum_{s_1} i_2} F(u; (m_1-1) (m_2-1), p) \\ u &= \min\left\{\tau-s_1-t_1-i_1, \tau-s_2-t_1-i_2, \tau-s_1-t_2-i_3, \tau-s_2-t_2-i_4\right\} \\ b(s; n, p) &= \binom{n}{s} p^s (1-p)^{n-s} \\ F(s; n, p) &= \sum_{i=0}^{s} b(i; n, p) \end{split}$$

Iniversité ille1

Bounds for the Bernoulli Case

The following bounds were established by [Boutsikas and Koutras, 2003]

Lower Bound

$$LB = (1 - Q_1)^{(T_1 - m_1)(T_2 - m_2)} (1 - Q_2)^{T_1 - m_1} (1 - Q_3)^{T_2 - m_2} (1 - Q_4)$$

Upper Bound

$$\begin{split} \mathit{UB} &= (1 - \mathit{Q}_1) \left(1 - q^{(m_1 - 1)(3m_2 - 2) + (2m_1 - 1)(m_2 - 1)} \mathit{Q}_1 \right)^{(T_1 - m_1 - 1)(T_2 - m_2 - 1)} \left(1 - q^{m_1(m_2 - 1)} \mathit{Q}_1 \right)^{T_2 - m_2 - 1} \\ &\times \left(1 - q^{(m_1 - 1)(2m_2 - 1) + (m_1 - 1)(m_2 - 1)} \mathit{Q}_1 \right)^{T_1 - m_1 - 1} \left(1 - q^{(m_1 - 1)(2m_2 - 1) + m_1(m_2 - 1)} \mathit{Q}_2 \right)^{T_1 - m_1} \\ &\times \left(1 - q^{(m_1 - 1)(3m_2 - 2) + m_1(m_2 - 1) + (m_1 - 1)(m_2 - 1)} \mathit{Q}_3 \right)^{T_2 - m_2} \left(1 - q^{(m_1 - 1)(2m_2 - 1) + m_1(m_2 - 1)} \mathit{Q}_4 \right). \end{split}$$

Where $X_{ij} \sim B(p)$, q=1-p and

$$\begin{aligned} Q_1 &= F_{\tau+1,m_1m_2}^c - q^{m_2} F_{\tau+1,(m_1-1)m_2}^c - q^{m_1} F_{\tau+1,m_1(m_2-1)}^c + q^{m_1+m_2-1} F_{\tau+1,(m_1-1)(m_2-1)}^c, \\ Q_2 &= F_{\tau+1,m_1m_2}^c - q^{m_2} F_{\tau+1,(m_1-1)m_2}^c, \quad Q_3 &= F_{\tau+1,m_1m_2}^c - q^{m_1} F_{\tau+1,m_1(m_2-1)}^c, \quad Q_4 &= F_{\tau+1,m_1m_2}^c, \\ F_{i,m}^c &= 1 - F(i-1;m,p). \end{aligned}$$



Product Type Approximation Binomial and Poisson

For $X_{ij} \sim Bin(n,p)$ or $X_{ij} \sim Pois(\lambda)$, [Chen and Glaz, 2009] proposed the product type approximation

$$\mathbb{P}\left(S_{m_1,m_2}(T_1,T_2) \leq \tau\right) \approx \frac{Q(m_1+1,m_2+1)^{(T_1-m_1)(T_2-m_2)}}{Q(m_1+1,m_2)^{(T_1-m_1)(T_2-m_2-1)}} \times \frac{Q(m_1,2m_2-1)^{(T_1-m_1-1)(T_2-2m_2)}}{Q(m_1,2m_2)^{(T_1-m_1-1)(T_2-2m_2+1)}}$$

Where,

$$Q(m_1, 2m_2 - 1) = \mathbb{P}(S_{m_1, m_2}(m_1, 2m_2 - 1) \le \tau)$$

$$Q(m_1, 2m_2) = \mathbb{P}(S_{m_1, m_2}(m_1, 2m_2) \le \tau)$$

To compute the unknown variables we use

- ullet $Q(m_1,2m_2-1)$ and $Q(m_1,2m_2)$ adaptation of [Karwe and Naus, 1997] algorithm
- $Q(m_1 + 1, m_2)$ and $Q(m_1 + 1, m_2 + 1)$ conditioning

ightharpoonup Formulas for $Q(m_1+1,m_2)$ and $Q(m_1+1,m_2+1)$





Lower Bound for Binomial and Poisson

[Chen and Glaz, 1996] gave a lower bound applying Hoover Bonferroni type inequality of order $r \geq 3$,

$$\begin{split} \mathbb{P}\left(S_{m_{1},m_{2}}(T_{1},T_{2}) \geq \tau\right) &= \mathbb{P}\left(\bigcup_{i_{1}=1}^{T_{1}-m_{1}+1} \bigcup_{i_{2}=1}^{T_{2}-m_{2}+1} A_{i_{1},i_{2}}\right) \leq \sum_{i_{1}=1}^{T_{1}-m_{1}+1} \sum_{i_{2}=1}^{T_{2}-m_{2}+1} \mathbb{P}\left(A_{i_{1},i_{2}}\right) \\ &- \sum_{i_{1}=1}^{T_{1}-m_{1}+1} \sum_{i_{2}=1}^{T_{2}-m_{2}} \mathbb{P}\left(A_{i_{1},i_{2}} \cap A_{i_{1},i_{2}+1}\right) - \sum_{i_{1}=1}^{T_{1}-m_{1}} \mathbb{P}\left(A_{i_{1},1} \cap A_{i_{1}+1,1}\right) \\ &- \sum_{i_{1}=1}^{T_{1}-m_{1}+1} \sum_{l=2}^{r-1} \sum_{i_{2}=1}^{T_{2}-m_{2}+1-l} \mathbb{P}\left(A_{i_{1},i_{2}} \cap A_{i_{1},i_{2}+1}^{c} \cdots A_{i_{1},i_{2}+l-1}^{c} \cap A_{i_{1},i_{2}+l}\right) \end{split}$$

Where $A_{i_1,i_2} = \{Y_{i_1,i_2} \geq \tau\}$ and for r = 4,

Lower Bound

$$\mathbb{P}\left(S_{m_1,m_2}(T_1,T_2) \leq \tau\right) \geq (T_1 - m_1)\left(Q(m_1 + 1, m_2) - 2Q(m_1, m_2)\right) - (T_1 - m_1 + 1)(T_2 - m_2 - 3) \times Q(m_1, m_2 + 2) + (T_1 - m_1 + 1)(T_2 - m_2 - 2)Q(m_1, m_2 + 3).$$

• $Q(m_1 + 1, m_2)$, $Q(m_1, m_2)$, $Q(m_1, m_2 + 2)$, $Q(m_1, m_2 + 3)$ - [Karwe and Naus, 1997] algorithm (variant)



Upper Bound for Binomial and Poisson

For the upper bound we adapt the inequality of [Kuai et al., 2000] to the two dimensional framework:

$$\begin{split} \mathbb{P}\left(S_{\pmb{m_1},\pmb{m_2}}\ (\pmb{T_1},\pmb{T_2}) \leq \tau\right) &= 1 - \mathbb{P}\left(\bigcup_{\pmb{i_1}=1}^{\pmb{T_1}-\pmb{m_1}+1} \bigcup_{\pmb{i_2}=1}^{\pmb{T_2}-\pmb{m_2}+1} A_{\pmb{i_1},\pmb{i_2}}\right) \\ &\leq 1 - \sum_{\pmb{i_1}=1}^{\pmb{T_1}-\pmb{m_1}+1} \sum_{\pmb{i_2}=1}^{\pmb{T_2}-\pmb{m_2}+1} \left[\frac{\theta_{\pmb{i_1},\pmb{i_2}}\mathbb{P}(A_{\pmb{i_1},\pmb{i_2}})^2}{\Sigma(\pmb{i_1},\pmb{i_2}) + (1-\theta_{\pmb{i_1},\pmb{i_2}})\mathbb{P}(A_{\pmb{i_1},\pmb{i_2}})} + \frac{(1-\theta_{\pmb{i_1},\pmb{i_2}})\mathbb{P}(A_{\pmb{i_1},\pmb{i_2}})^2}{\Sigma(\pmb{i_1},\pmb{i_2}) - \theta_{\pmb{i_1},\pmb{i_2}}\mathbb{P}(A_{\pmb{i_1},\pmb{i_2}})^2}\right] \end{split}$$

where

$$\Sigma(i_1,i_2) = \sum_{j_1=1}^{T_1-m_1+1} \sum_{j_2=1}^{T_2-m_2+1} \mathbb{P}\left(A_{j_1,j_2} \cap A_{j_1,j_2}\right) \text{ and } \theta_{i_1,i_2} = \frac{\Sigma(i_1,i_2)}{\mathbb{P}(A_{j_1,j_2})} - \left\lfloor \frac{\Sigma(i_1,i_2)}{\mathbb{P}(A_{j_1,j_2})} \right\rfloor.$$

We have

$$\mathbb{P}\left(A_{\pmb{i_1},\pmb{i_2}} \cap A_{\pmb{j_1},\pmb{j_2}}\right) = \left\{ \begin{array}{l} [1-Q(m_1,m_2)]^2 \text{, if } |i_1-j_1| \geq m_1 \text{ or } |i_2-j_2| \geq m_2, \\ 1-2Q(m_1,m_2) + \mathbb{P}\left(Y_{\pmb{i_1},\pmb{i_2}} \leq \tau, Y_{\pmb{j_1},\pmb{j_2}} \leq \tau\right) \text{, otherwise} \end{array} \right. \\ \mathbb{P}\left(Y_{\pmb{i_1},\pmb{i_2}} \leq \tau, Y_{\pmb{j_1},\pmb{j_2}} \leq \tau\right) = \sum_{k=0}^{\tau} \mathbb{P}(Z=k) \mathbb{P}(Y_{\pmb{i_1},\pmb{i_2}} - Z \leq \tau-k)^2, \\ Z = \sum_{\substack{k=0 \\ k \neq k}} \frac{(i_1+m_1-1) \wedge (j_1+m_1-1) (i_2+m_2-1) \wedge (j_2+m_2-1)}{\sum_{\substack{k=0 \\ k \neq k}} (i_k+m_1-1) \wedge (i_2+m_2-1) \wedge (i_2+m_2-1)} X_{\pmb{s},\pmb{t}}. \end{array}$$

niversité

18 / 46

- Introduction
 - Framework
 - Problem
- Approximations and bounds for scan statistics
 - One dimensional scan statistics
 - Two dimensional scan statistics
 - Three dimensional scan statistics
- Scan statistics and 1-dependent sequences
 - Methodology
 - Numerical examples
- Some dependent models for scan statistics
 - Model and discussion
 - Example (d = 1): Longest monotone run
- 6 References





Product Type Approximation

Glaz et al. proposed in [Guerriero et al., 2010] the product type approximation

$$\begin{split} & \mathbb{P}\left(S_{m_1,m_2,m_3}(T_1,T_2,T_3) \leq \tau\right) \approx \\ & \frac{Q(m_1+1,m_2+1,m_3+1)(T_1-m_1)(T_2-m_2)(T_3-m_3)}{Q(m_1,m_2,m_3)(T_1-m_1)(T_2-m_2-1)(T_3-m_3-1)} \times \\ & \frac{Q(m_1,m_2,m_3)(T_1-m_1-1)(T_2-m_2-1)(T_3-m_3-1)}{Q(m_1,m_2+1,m_3)(T_1-m_1-1)(T_2-m_2)(T_3-m_3-1)} \times \\ & \frac{Q(m_1,m_2+1,m_3)(T_1-m_1-1)(T_2-m_2)(T_3-m_3-1)}{Q(m_1+1,m_2,m_3+1)(T_1-m_1-1)(T_2-m_2)(T_3-m_3-1)} \\ & \frac{Q(m_1+1,m_2,m_3+1)(T_1-m_1-1)(T_2-m_2)(T_3-m_3-1)}{Q(m_1+1,m_2,m_3+1)(T_1-m_1-1)(T_2-m_2)(T_3-m_3-1)} \end{split}$$

Where,

$$Q(N_1, N_2, N_3) = \mathbb{P}(S_{m_1, m_2, m_3}(N_1, N_2, N_3) \leq \tau)$$

- The approximation also works for binomial and Poisson distribution
- Three Poisson Type Approximation

Lower and upper bounds for the distribution of three dimensional scan statistics were proposed by [Akiba and Yamamoto, 2004].



- Introduction
 - Framework
 - Problem
- 2 Approximations and bounds for scan statistics
 - One dimensional scan statistics
 - Two dimensional scan statistics
 - Three dimensional scan statistics
- 3 Scan statistics and 1-dependent sequences
 - Methodology
 - Numerical examples
- 4 Some dependent models for scan statistics
 - Model and discussion
 - Example (d = 1): Longest monotone run
- 6 References





Haiman Type Approximation: Key Idea

Haiman proposed in [Haiman, 2000] a different approach

Main Observation

The scan statistic r.v. can be viewed as a maximum of a sequence of 1-dependent stationary r.v..

- The idea:
 - discrete and continuous one dimensional scan statistic: [Haiman, 2000], [Haiman, 2007]
 - discrete and continuous two dimensional scan statistic: [Haiman and Preda, 2002], [Haiman and Preda, 2006]
 - discrete three dimensional scan statistic: [Amărioarei and Preda, 2013]



Expressing $S_{\mathbf{m}}(\mathsf{T})$ as a maximum of a 1-dependent sequence

Let $L_j = rac{T_j}{m_j - 1}$, $j \in \{1, 2, \ldots, d\}$, be positive integers

ullet Define for each $k_1 \in \{1,2,\ldots,L_1-1\}$ the random variables

$$Z_{k_1} = \max_{\substack{(k_1-1)(m_1-1)+1 \leq i_1 \leq k_1(m_1-1) \\ 1 \leq i_j \leq (L_j-1)(m_j-1) \\ j \in \{2,...,d\}}} Y_{i_1,i_2,...,i_d}$$

• $(Z_j)_j$ is 1-dependent and stationary

Example: d = 1

$$\underbrace{X_1, X_2, \dots, X_{m_1-1}, \overbrace{X_{m_1}, \dots, X_{2(m_1-1)}}^{Z_2}, \underbrace{X_{2m_1-1}, \dots, X_{3(m_1-1)}}_{Z_3}, X_{3m_1-2}, \dots, X_{4(m_1-1)}}_{Z_3}$$

Observe

$$S_{\mathbf{m}}(\mathbf{T}) = \max_{1 \leq k_1 \leq L_1 - 1} Z_{k_1}$$



The main tool

Let $(Z_j)_{j\geq 1}$ be a strictly stationary 1-dependent sequence of r.v.'s and let $q_m=q_m(x)=\mathbb{P}(\max(Z_1,\ldots,Z_m)\leq x)$, with $x<\sup\{u|\mathbb{P}(Z_1\leq u)<1\}$.

Theorem [Amărioarei, 2012]

For x such that $\mathbb{P}(Z_1>x)=1-q_1<0.1$ and m>3 we have

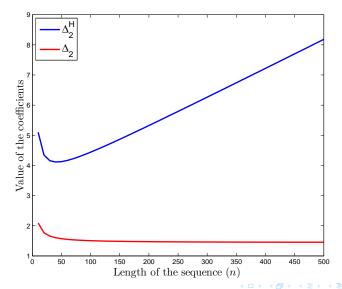
$$\left|q_m - rac{2q_1 - q_2}{\left[1 + q_1 - q_2 + 2(q_1 - q_2)^2
ight]^m}
ight| \leq \Delta_2 (1 - q_1)^2$$

- $\Delta_2 = mF(q_1, m) = m \left[1 + \frac{3}{m} + K(\alpha)(1 q_1) + \frac{\Gamma(\alpha)(1 q_1)}{m} \right].$
- Increased range of applicability
- Sharp bounds values (ex. $\alpha = 0.025$: $561 \rightarrow 145$ and $88 \rightarrow 17.5$)

ightharpoonup Selected values for K(lpha) and $\Gamma(lpha)$



Difference between the results: $1 - q_1 = 0.025$







Approximation process

Define for $t_1 \in \{2, 3\}$,

$$Q_{\mathbf{t_{1}}} = Q_{\mathbf{t_{1}}}(\tau) = \mathbb{P}\left(\bigcap_{k_{1}=1}^{t_{1}-1} \left\{Z_{k_{1}} \leq \tau\right\}\right) = \mathbb{P}\left(\bigcap_{\substack{1 \leq i_{1} \leq (t_{1}-1)(m_{1}-1) \\ 1 \leq i_{j} \leq (t_{j}-1)(m_{j}-1) \\ j \in \left\{2, \dots, d\right\}}} Y_{i_{1}}, i_{2}, \dots, i_{d} \leq \tau\right)$$

If $1-Q_2 \leq 0.1$ then

$$\left|\,Q_{\mathsf{m}}(\mathsf{T}) - \frac{2\,Q_2 - Q_3}{[1 + Q_2 - Q_3 + 2(Q_2 - Q_3)^2]^{L_1 - 1}}\,\right| \leq (L_1 \, - \, 1) F(Q_2, L_1 \, - \, 1) (1 - \, Q_2)^2$$

Example: d = 1

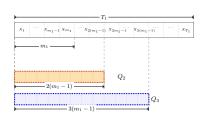
• The approximation

$$\mathbb{P}(S_{\boldsymbol{m_{1}}}(T_{1}) \leq \tau) \approx H(Q_{2}, Q_{3}, L_{1})$$

$$H(x, y, m) = \frac{2x-y}{[1+x-y+2(x-y)^2]^{m-1}}$$

Approximation error, about

$$(L_1-1)F(Q_2,L_1-1)(1-Q_2)^2$$



Approximation process 2

The approximation of $S_{\mathbf{m}}(\mathsf{T})$ is an iterative process. The s step, $1 \leq s \leq d$, becomes:

Let

$$Q_{t_1,t_2,...,t_s} = Q_{t_1,t_2,...,t_s}(\tau) = \mathbb{P} \begin{pmatrix} \max_{\substack{1 \leq i_j \leq (t_j-1)(m_j-1) \\ l \in \{1,...,s\} \\ 1 \leq i_j \leq (l_j-1)(m_j-1) \\ j \in \{s+1,...,d\}}} Y_{i_1,i_2,...,i_d} \leq \tau \end{pmatrix}$$

ullet Define for $t_I \in \{2,3\},\ I \in \{1,\ldots,s-1\}$ and $k_s \in \{1,2,\ldots,L_s-1\}$

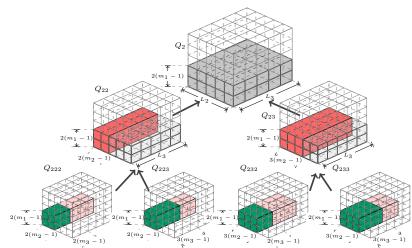
$$Z_{k_s}^{(t_1,t_2,...,t_{s-1})} = \max_{\substack{1 \leq i_j \leq (t_j-1)(m_j-1) \\ i \in \{1,2,...,s-1\} \\ (k_s-1)(m_s-1)+1 \leq i_s \leq k_s(m_s-1) \\ 1 \leq i_j \leq (t_j-1)(m_j-1) \\ j \in \{s+1,...,d\}}} Y_{i_1,i_2,...,i_d}$$

- $\bullet \ \left\{Z_1^{(t_1,t_2,\ldots,t_{s-1})},\ldots,Z_{L_s-1}^{(t_1,t_2,\ldots,t_{s-1})}\right\} \ \text{forms a 1-dependent stationary sequence}$
- The approximation

$$\left|Q_{t_1,...,t_{s-1}} - H\left(Q_{t_1,...,t_{s-1},2},Q_{t_1,...,t_{s-1},3},L_s\right)\right| \leq (L_s - 1)F(Q_{t_1,...,t_{s-1},2},L_s - 1)(1 - Q_{t_1,...,t_{s-1},2})^2$$



Illustration of the approximation process: d = 3







- Introduction
 - Framework
 - Problem
- 2 Approximations and bounds for scan statistics
 - One dimensional scan statistics
 - Two dimensional scan statistics
 - Three dimensional scan statistics
- 3 Scan statistics and 1-dependent sequences
 - Methodology
 - Numerical examples
- 4 Some dependent models for scan statistics
 - Model and discussion
 - Example (d = 1): Longest monotone run
- 6 References





Comparison between methods: d=1

Table 1:
$$n = 1, p = 0.005, m_1 = 10, T_1 = 1000, lt_{App} = 10^4$$

τ	Exact	Glaz and Naus Product type	Haiman Approximation	Approximation Error	Lower Bound	Upper Bound
1	0.810209	0.810216	0.810404	0.001111	0.809903	0.810439
2	0.995764	0.995764	0.995764	3×10^{-7}	0.995764	0.995764
3	0.999950	0.999950	0.999950	$4\times \mathbf{10^{-11}}$	0.999950	0.999950

Table 2:
$$n = 5, p = 0.05, m_1 = 25, T_1 = 500, lt_{App} = 10^4, lt_{Sim} = 10^3$$

τ	$\hat{\mathbb{P}}(S \leq au)$	Glaz and Naus Product type	Haiman Approximation	Total Error	Lower Bound	Upper Bound
13	0.712750	0.705787	0.714699	0.039308	0.697431	0.706948
14	0.867498	0.862184	0.865029	0.012502	0.859543	0.862407
15	0.946912	0.943329	0.946177	0.004169	0.942552	0.943362
16	0.980230	0.978959	0.979822	0.001354	0.978733	0.978963
17	0.993486	0.992821	0.993134	0.000433	0.992756	0.992822
18	0.997802	0.997726	0.997849	0.000127	0.997708	0.997726
19	0.999362	0.999327	0.999358	3×10^{-5}	0.999322	0.999327
20	0.999819	0.999813	0.999825	9×10^{-6}	0.999812	0.999813
21	0.999954	0.999951	0.999953	$2 imes 10^{-6}$	0.999951	0.999951





Comparison between methods: d = 2

Table 3: $n = 1, p = 0.005, m_1 = m_2 = 6, T_1 = T_2 = 30, It_{App} = 10^3, It_{Sim} = 10^3$

τ	$\hat{\mathbb{P}}(S \leq \tau)$	Glaz and Naus Product type	Haiman Approximation	Total Error(App+Sim)	Lower Bound	Upper Bound
2	0.915903	0.914013	0.920211	0.041483	0.901935	0.945623
3	0.994292	0.994395	0.994578	0.000803	0.993785	0.996638
4	0.999747	0.999757	0.999760	2×10^{-5}	0.999737	0.999858
5	0.999992	0.999992	0.999992	7×10^{-7}	0.999992	0.999995

Table 4:
$$n = 5, p = 0.002, m_1 = 5, m_2 = 10, T_1 = 50, T_2 = 80, It_{App} = 10^4, It_{Sim} = 10^3$$

au	$\hat{\mathbb{P}}(S \leq \tau)$	Glaz and Naus Product type	Haiman Approximation	Total Error(App+Sim)	Lower Bound	Upper Bound
4	0.894654	0.873256	0.893724	0.037136	0.803422	0.944318
5	0.988003	0.986249	0.988144	0.002125	0.981418	0.993451
6	0.998963	0.998847	0.998963	0.000152	0.998543	0.999401
7	0.999926	0.999919	0.999925	9×10^{-6}	0.999903	0.999955
8	0.999995	0.999995	0.999995	5×10^{-7}	0.999994	0.9999971

Comparison between methods: d = 3

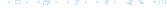
Table 5:
$$n = 1, p = 0.00005, m_1 = m_2 = m_3 = 5, T_1 = T_2 = T_3 = 60, lt_{App} = 10^5$$

τ	$\hat{\mathbb{P}}(S \leq \tau)$	Glaz et al. Product type	Haiman Approximation	Approximation Error	Simulation Error	Total Error
1	0.841806	0.841424	0.851076	0.011849	0.064889	0.076738
2	0.999119	0.999142	0.999192	0.00000	0.000170	0.000170
3	0.999997	0.999998	0.999997	0.00000	3×10^{-7}	3×10^{-7}

Table 6:
$$n = 1, p = 0.0001, m_1 = m_2 = m_3 = 5, T_1 = T_2 = T_3 = 60, It_{App} = 10^5$$

τ	$\hat{\mathbb{P}}(S \leq \tau)$	Glaz et al. Product type	Haiman Approximation	Approximation Error	Simulation Error	Total Error
2	0.993294	0.993241	0.993192	0.000010	0.001367	0.001377
3	0.999963	0.999964	0.999963	0.00000	0.000005	0.000005
4	0.999999	0.999999	0.999999	0.00000	2×10^{-9}	2×10^{-9}





- Introduction
 - Framework
 - Problem
- 2 Approximations and bounds for scan statistics
 - One dimensional scan statistics
 - Two dimensional scan statistics
 - Three dimensional scan statistics
- Scan statistics and 1-dependent sequences
 - Methodology
 - Numerical examples
- 4 Some dependent models for scan statistics
 - Model and discussion
 - Example (d = 1): Longest monotone run
- 6 References



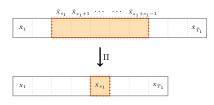


Introducing the model d=1

Let $m_1, c_1, \, ilde{\mathcal{T}}_1$ be positive integers such that $c_1 \geq 1$ and $2 \leq m_1 \leq \mathcal{T}_1$

- ullet Consider $\left(ilde{X}_i
 ight)_{1 \leq i \leq ilde{T}_1}$ be a sequence of i.i.d. r.v's
- ullet Take $\Pi:\mathbb{R}^{c_1} o\mathbb{R}$ to be a measurable function
- ullet Define $\mathcal{T}_1= ilde{\mathcal{T}}_1-c_1+1$ and for $1\leq s_1\leq \mathcal{T}_1$

$$X_{s_1} = \Pi\left(\tilde{X}_{s_1}, \dots, \tilde{X}_{s_1+c_1-1}\right)$$



• The sequence $(X_{s_1})_{1 \leq s_1 \leq T_1}$ is $(c_1 - 1)$ -dependent



Approximation and bounds

Assume that $L_1 = \frac{\tilde{T}_1}{m_1 + c_1 - 2}$ is a positive integer

• Define for $k_1 \in \{1, 2, \dots, L_1 - 1\}$

$$Z_{k_1} = \max_{(k_1-1)(m_1+c_1-2)+1 \le i_1 \le k_1(m_1+c_1-2)} Y_{i_1}$$

- \bullet $(Z_j)_j$ is 1-dependent and stationary
- We have the approximation

$$\mathbb{P}\left(S_{m_1}(T_1) \le \tau\right) \approx \frac{2Q_2 - Q_3}{[1 + Q_2 - Q_3 + 2(Q_2 - Q_3)^2]^{L_1 - 1}}$$

Error bounds

$$\begin{split} E_{app}(1) &= (L_1 - 1) F\left(Q_2, L_1 - 1\right) (1 - Q_2)^2 \,, \\ E_{sapp}(1) &= (L_1 - 1) F\left(\hat{Q}_2, L_1 - 1\right) \left(1 - \hat{Q}_2 + \beta_2\right)^2 \,, \\ E_{sf}(1) &= (L_1 - 1) (\beta_2 + \beta_3) \end{split}$$



- Introduction
 - Framework
 - Problem
- Approximations and bounds for scan statistics
 - One dimensional scan statistics
 - Two dimensional scan statistics
 - Three dimensional scan statistics
- 3 Scan statistics and 1-dependent sequences
 - Methodology
 - Numerical examples
- 4 Some dependent models for scan statistics
 - Model and discussion
 - Example (d = 1): Longest monotone run
- 6 References





Longest increasing run

Let \tilde{X}_1 , \tilde{X}_2 , ..., $\tilde{X}_{\tilde{T}_1}$ be a sequence of i.i.d. r.v.'s with the common distribution G.

A subsequence $(\tilde{X}_k,\ldots,\tilde{X}_{k+l-1})$ forms an *increasing run* (or ascending run) of length $l\geq 1$, starting at position $k\geq 1$, if it verifies the following relation

$$\tilde{X}_{k-1} > \tilde{X}_k < \tilde{X}_{k+1} < \cdots < \tilde{X}_{k+l-1} > \tilde{X}_{k+l}$$

Denote the length of the longest increasing run among the first T_1 random variables by $M_{\tilde{T}_1}$.

The asymptotic distribution was studied

- G continuous distribution: [Pittel, 1981], [Révész, 1983], [Grill, 1987], [Novak, 1992], etc.
- G discrete distribution: [Csaki and Foldes, 1996], [Grabner et al., 2003], [Eryilmaz, 2006], etc.





Longest increasing run 2

Let $c_1=2$, $\mathcal{T}_1=\tilde{\mathcal{T}}_1-1$ and define $\Pi:\mathbb{R}^2 \to \mathbb{R}$ by

$$\Pi(x,y) = \begin{cases} 1, & \text{if } x < y \\ 0, & \text{otherwise} \end{cases}$$

Thus our block-factor model becomes

$$X_{s_1} = 1_{\tilde{X}_{s_1} < \tilde{X}_{s_1+1}}$$

Let L_{T_1} be the distribution of the length of the longest run of ones, among the first T_1 observations.

$$\mathbb{P}\left(M_{\widetilde{\mathcal{T}}_{1}} \leq m_{1}
ight) = \mathbb{P}\left(L_{\mathcal{T}_{1}} < m_{1}
ight) = \mathbb{P}\left(\mathcal{S}_{m_{1}}\left(\mathcal{T}_{1}
ight) < m_{1}
ight)$$
, for $m_{1} \geq 1$

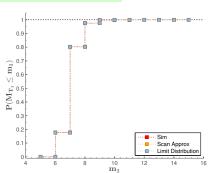


Longest increasing run: numerical results

For $ilde{X}_{s_1} \sim \mathcal{U}\left([0,1]\right)$, [Novak, 1992] showed that

$$\max_{1 \leq m_1 \leq T_1} \left| \mathbb{P}\left(L_{T_1} < m_1 \right) - \mathrm{e}^{-T_1 \frac{m_1 + 1}{(m_1 + 2)!}} \right| = \mathcal{O}\left(\frac{|\mathsf{n} \ T_1}{T_1} \right)$$

_				
m_1	Sim	АррН	$E_{total}(1)$	LimApp
5	0.00000700	0.00000733	0.14860299	0.00000676
6	0.17567262	0.17937645	0.01089628	0.17620431
7	0.80257424	0.80362353	0.00110990	0.80215088
8	0.97548510	0.97566460	0.00011579	0.97550345
9	0.99749821	0.99751049	0.00001114	0.99749792
10	0.99977074	0.99977183	0.00000098	0.99977038
11	0.99998075	0.99998083	0.00000008	0.99998073
12	0.99999851	0.99999851	0.00000001	0.99999851
13	0.99999989	0.99999989	0.00000000	0.99999989
14	0.99999999	0.99999999	0.00000000	0.99999999
15	1.00000000	1.00000000	0.00000000	1.00000000





Longest increasing run: numerical results

For $\tilde{X}_{s_1} \sim \textit{Geom}(p)$, [Louchard and Prodinger, 2003] showed that

$$\mathbb{P}\left(M_{T_1} \leq m_1\right) \sim \exp\left(-\exp\eta\right),$$

$$\eta = \frac{m_1(m_1+1)}{2}\log\frac{1}{1-p} + m_1\log\frac{1}{p} - \log T_1 - \log p + \log D(m_1),$$

$$D(m_1) = \prod_{k=1}^{m_1} \left[1 - (1-p)^k\right] \left[1 - (1-p)^{m_1+2}\right]$$

m_1	Sim	AppH	$E_{total}(1)$	LimApp
6	0.56445934	0.56997462	0.00255592	0.56810748
7	0.95295406	0.95325180	0.00018554	0.95294598
8	0.99658057	0.99659071	0.00001214	0.99657969
9	0.99979460	0.99979550	0.00000068	0.99979435
10	0.99998950	0.99998950	0.00000003	0.99998947

We used $T_1 = 10000$, p = 0.1 and $Iter = 10^5$.





Akiba, T. and Yamamoto, H. (2004).

Upper and lower bounds for 3-dimensional r-within n consecutive $-(r_1, r_2, r_3)$ -out-of $-(n_1, n_2, n_3)$: F-system. In Advanced Reliability Modelling.



Amărioarei, A. (2012).

Approximation for the distribution of extremes of one dependent stationary sequences of random variables.

arXiv:1211.5456v1, submitted.



Amărioarei, A. and Preda, C. (2013).

Approximation for the distribution of three-dimensional discrete scan statistic.

Methodol Comput Appl Probab, DOI 10.1007/s11009-013-9382-3.



Balakrishnan, N. and Koutras, M. V. (2002).

Runs and scans with applications.

Wiley Series in Probability and Statistics. Wiley-Interscience [John Wiley & Sons], New York.





Boutsikas, M. V. and Koutras, M. V. (2000).

Reliability approximation for Markov chain imbeddable systems.

Methodol. Comput. Appl. Probab., 2:393-411.



Bounds for the distribution of two-dimensional binary scan statistics.

Probab. Eng. Inform. Sci., 17:509-525.



Two-dimensional discrete scan statistics.

Statist. Probab. Lett., 31:59-68.



Scan statistics, chapter 5, Approximations for two-dimensional variable window scan statistics., pages 109–128.

Birkhäuser Boston, Inc., Boston.



On the length of theh longest monnotone block.

Studio Scientiarum Mathematicarum Hungarica, 31:35-46.





Ebneshahrashoob, M., Gao, T., and Wu, M. (2005).

An efficient algorithm for exact distribution of discrete scan statistics. Methodol. Comput. Appl. Probab., 7:1423-1436.



Ebneshahrashoob, M. and Sobel, M. (1990).

Sooner and later waiting time problems for Bernoulli trials: frequency and run quotas.

Statist. Probab. Lett., 9:5–11.



Eryilmaz, S. (2006).

A note on runs of geometrically distributed random variables.

Discrete Mathematics, 306:1765–1770.



Fu, J. (2001).

Distribution of the scan statistic for a sequence of bistate trials.

J. Appl. Probab., 38:908-916.



Fu, J., Wu, T.-L., and Lou, W. (2012).

Continuous, discrete, and conditional scan statistics.

J. Appl. Prob., 49:199–209.





Fu, J. C. and Lou, W. (2003).

Distribution theory of runs and patterns and its applications. A finite Markov chain imbedding approach.

World Scientific Publishing Co., Inc., River Edge, NJ.



Glaz, J. and Naus, J. (1991).

Tight bounds and approximations for scan statistic probabilities for discrete data.

Annals of Applied Probability, 1:306-318.



Grabner, P., Knopfmacher, A., and Prodinger, H. (2003).

Combinatorics of geometrically distributed random variables: run statistics.

Theoret. Comput. Sci., 297:261-270.



📄 Grill, K. (1987).

Erdos-Révész type bounds for the length of the longest run from a stationary mixing sequence.

Probab. Theory Relat. Fields, 75:169-179.





Guerriero, M., Glaz, J., and Sen, R. (2010).

Approximations for a three dimensional scan statistic.

Methodol. Comput. Appl. Probab., 12:731-747.



Haiman, G. (2000).

Estimating the distributions of scan statistics with high precision.

Extremes, 3:349-361.



Haiman, G. (2007).

Estimating the distribution of one-dimensional discrete scan statistics viewed as extremes of 1-dependent stationary sequences.

J. Statist. Plann. Inference, 137:821-828.



Haiman, G. and Preda, C. (2002).

A new method for estimating the distribution of scan statistics for a two-dimensional Poisson process.

Methodol. Comput. Appl. Probab., 4:393-407.



Haiman, G. and Preda, C. (2006).



Estimation for the distribution of two-dimensional discrete scan statistics.

Methodol. Comput. Appl. Probab., 8:373–381.

🚺 Karwe, V. and Naus, J. (1997).

New recursive methods for scan statistic probabilities.

Computational Statistics & Data Analysis, 17:389-402.

廜 Kuai, H., Alajaji, F., and Takahara, G. (2000).

A lower bound on the probability of a finite union of events.

Discrete Mathematics, 215:147-158.

📘 Louchard, G. and Prodinger, H. (2003).

Ascending runs of sequences of geometrically distributed random variables: a probabilistic analysis.

Theoret. Comput. Sci., 304:59-86.

Naus, J. (1974).

Probabilities for a generalized birthday problem.

Journal of American Statistical Association, 69:810-815.





Naus, J. (1982).

Approximations for distributions of scan statistics.

Journal of American Statistical Association, 77:177-183.



Longest runs in a sequence of *m*-dependent random variables.

Probab. Theory Relat. Fields, 91:269-281.



Limiting behavior of a process of runs.

Ann. Probab., 9:119-129.



Three problems on the llength of increasing runs.

Stochastic Process. Appl., 5:169–179.



Distributions of scan statistics in a sequence of Markov Bernoulli trials. *International Journal of Statisites*, LXVI:135–155.



Wu, T.-L. (2013).



On finite Markov chain imbedding technique.

Methodol Comput Appl Probab, 15:453-465.



Computation of $Q(m_1 + 1, m_2)$ and $Q(m_1 + 1, m_2 + 1)$

Consider $X_{ij} \sim B(n,p)$ and the notation $Y_{i_1,i_2}^{j_1,j_2} = \sum_{i_1=1}^{j_1} \sum_{i_2=1}^{j_2} X_{ij}$,

$$\begin{split} \mathbb{P}\left(S_{m_{1},m_{2}}(m_{1}+1,m_{2}) \leq k\right) &= \sum_{y=0}^{k \wedge (m_{1}-1)m_{2}n} \mathbb{P}^{2}\left(Y_{1,1}^{1,m_{2}} \leq k - y\right) \mathbb{P}\left(Y_{2,1}^{m_{1},m_{2}} = y\right) \\ \mathbb{P}\left(S_{m_{1},m_{2}}(m_{1}+1,m_{2}+1) \leq k\right) &= \sum_{y_{1}=0}^{k \wedge (m_{1}-1)(m_{2}-1)n} (k-y_{1}) \wedge (m_{2}-1)n (k-y_{1}) \wedge (m_{2}-1)n} \sum_{y_{3}=0} \sum_{y_{3}=0}^{k \wedge (m_{1}-1)(m_{2}-1)n} \mathbb{P}\left(Y_{1,1}^{1,1} \leq a_{1}\right) \\ &= \sum_{y_{4}=0} \sum_{y_{4}=0} \sum_{y_{5}=0} \mathbb{P}\left(Y_{1,1}^{m_{1}-1,1} \leq a_{1}\right) \mathbb{P}\left(Y_{1,1}^{1,1} \leq a_{1}\right) \\ &\times \mathbb{P}\left(Y_{1,m_{2}+1}^{1,m_{2}+1} \leq a_{2}\right) \mathbb{P}\left(Y_{m_{1}+1,1}^{m_{1}+1,1} \leq a_{3}\right) \mathbb{P}\left(Y_{m_{1}+1,m_{2}+1}^{m_{1}+1} \leq a_{4}\right) \\ &\times \mathbb{P}\left(Y_{2,1}^{m_{1},m_{2}} = y_{1}\right) \mathbb{P}\left(Y_{1,2}^{1,m_{2}} = y_{2}\right) \mathbb{P}\left(Y_{m_{1}+1,2}^{m_{1}+1,m_{2}+1} \leq a_{4}\right) \\ &\times \mathbb{P}\left(Y_{2,1}^{m_{1},1} = y_{4}\right) \mathbb{P}\left(Y_{2,m_{2}+1}^{m_{1},m_{2}+1} = y_{5}\right) \\ &a_{1} = k - y_{1} - y_{2} - y_{4}, \quad a_{2} = k - y_{1} - y_{2} - y_{5}, \\ &a_{3} = k - y_{1} - y_{3} - y_{4}, \quad a_{4} = k - y_{1} - y_{3} - y_{5}, \end{split}$$



IWAP 2014 Conference

Karwe Naus recursive methods

Define

$$b_{2(m)}^{k}(y) = \mathbb{P}\left(S_{m}(2m) \leq k, Y_{m+1}(m) = y\right)$$
$$f(y) = \mathbb{P}(X_{1} = y)$$
$$Q_{2m}^{k} = \mathbb{P}\left(S_{m}(2m) \leq k\right)$$

We have the recurrence relations

$$\begin{aligned} b_{2(1)}^k(y) &= \left(\sum_{j=0}^k f(j)\right) f(y) \\ b_{2(m)}^k(y) &= \sum_{\eta=0}^y \sum_{\nu=0}^{k-y+\eta} b_{2(m-1)}^{k-\nu}(y-\eta) f(\nu) f(\eta) \\ Q_{2m}^k &= \sum_{y=0}^k b_{2(m)}^k(y) \\ Q_{2m-1}^k &= \sum_{v=0}^k f(x) Q_{2(m-1)}^{k-x} \end{aligned}$$





Selected Values for $K(\alpha)$ and $\Gamma(\alpha)$

α	$K(\alpha)$	$\Gamma(\alpha)$
0.1	38.63	480.69
0.05	21.28	180.53
0.025	17.56	145.20
0.01	15.92	131.43

Table 7 : Selected values for $K(\alpha)$ and $\Gamma(\alpha)$

◆ Return



Error bounds

Let $\gamma_{t_1,...,t_d}=Q_{t_1,...,t_d}$, with $t_j\in\{2,3\}$, $j\in\{1,\ldots,d\}$, and define for $2\leq s\leq d$

$$\gamma_{t_1,\ldots,t_{s-1}} = H\left(\gamma_{t_1,\ldots,t_{s-1},2},\gamma_{t_1,\ldots,t_{s-1},3},L_s\right)$$

Denote with $\hat{Q}_{t_1,...,t_d}$ the estimated value of $Q_{t_1,...,t_d}$ and define for $2 \leq s \leq d$

$$\hat{Q}_{\boldsymbol{t_1},...,\boldsymbol{t_{s-1}}} = H\left(\hat{Q}_{\boldsymbol{t_1},...,\boldsymbol{t_{s-1}},2},\hat{Q}_{\boldsymbol{t_1},...,\boldsymbol{t_{s-1}},3},L_s\right)$$

We observe that

$$\left|Q_{\mathsf{m}}(\mathsf{T}) - H\left(\hat{Q}_{2}, \hat{Q}_{3}, L_{1}\right)\right| \leq \underbrace{\left|Q_{\mathsf{m}}(\mathsf{T}) - H\left(\gamma_{2}, \gamma_{3}, L_{1}\right)\right|}_{E_{\mathit{alpp}}(d)} + \underbrace{\left|H\left(\gamma_{2}, \gamma_{3}, L_{1}\right) - H\left(\hat{Q}_{2}, \hat{Q}_{3}, L_{1}\right)\right|}_{E_{\mathit{sim}}(d)}$$

∢ Return





Error bounds: approximation and simulation errors 1

Approximation error

$$E_{app}(d) = \sum_{s=1}^{d} (L_1 - 1) \cdots (L_s - 1) \sum_{t_1, \dots, t_{s-1} \in \{2,3\}} F_{t_1, \dots, t_{s-1}} \left(1 - \gamma_{t_1, \dots, t_{s-1}, 2} + B_{t_1, \dots, t_{s-1}, 2}\right)^2,$$

where for $2 \le s \le d$

$$\begin{split} &F_{t_{1},...,t_{s-1}} = F\left(Q_{t_{1},...,t_{s-1},2},L_{s}-1\right), \ F = F\left(Q_{2},L_{1}-1\right), \\ &B_{t_{1},...,t_{s-1}} = \left(L_{s}-1\right) \left[F_{t_{1},...,t_{s-1}}\left(1-\gamma_{t_{1},...,t_{s-1},2}+B_{t_{1},...,t_{s-1},2}\right)^{2}+\sum_{t_{s}\in\{2,3\}}B_{t_{1},...,t_{s}}\right], \\ &B_{t_{1},...,t_{d-1}} = \left(L_{d}-1\right)F_{t_{1},...,t_{d-1}}\left(1-\gamma_{t_{1},...,t_{d-1},2}+B_{t_{1},...,t_{d-1},2}\right)^{2}, \ B_{t_{1},...,t_{d}} = 0, \end{split}$$

and for
$$s=1$$
: $\sum_{\mathbf{t_1}, \mathbf{t_0} \in \{2,3\}} x = x$, $F_{\mathbf{t_1}, \mathbf{t_0}} = F$, $\gamma_{\mathbf{t_1}, \mathbf{t_0}, 2} = \gamma_2$ and $B_{\mathbf{t_1}, \mathbf{t_0}, 2} = B_2$.





Error bounds: approximation and simulation errors 2

Simulation errors

$$\begin{split} E_{sf}(d) &= (L_1 - 1) \dots (L_d - 1) \sum_{t_1, \dots, t_d \in \{2, 3\}} \beta_{t_1, \dots, t_d} \\ E_{sapp}(d) &= \sum_{s=1}^d (L_1 - 1) \dots (L_s - 1) \sum_{t_1, \dots, t_{s-1} \in \{2, 3\}} F_{t_1, \dots, t_{s-1}} \left(1 - \hat{Q}_{t_1, \dots, t_{s-1}, 2} + A_{t_1, \dots, t_{s-1}, 2} + C_{t_1, \dots, t_{s-1}, 2}\right)^2 \end{split}$$

where for 2 < s < d

$$\begin{aligned} A_{t_{1},...,t_{s-1}} &= (L_{s}-1)...(L_{d}-1) \sum_{t_{s},...,t_{d} \in \{2,3\}} \beta_{t_{1},...,t_{d}}, A_{t_{1},...,t_{d}} &= \beta_{t_{1},...,t_{d}} \\ C_{t_{1},...,t_{s-1}} &= (L_{s}-1) \left[F_{t_{1},...,t_{s-1}} \left(1 - \hat{Q}_{t_{1},...,t_{s-1},2} + A_{t_{1},...,t_{s-1},2} + C_{t_{1},...,t_{s-1},2} \right)^{2} \right. \\ &\left. + \sum_{t_{s} \in \{2,3\}} C_{t_{1},...,t_{s}} \right] \end{aligned}$$



