

A 32 | a) $-F(0, b) = N_{\text{eff}} \ln(b) + N_{\text{on}} \ln(\alpha b) - (1+\alpha)b - \ln(N_{\text{eff}}!) - \ln(N_{\text{on}}!)$

$$-\frac{dF(0, b)}{db} = \frac{N_{\text{eff}}}{b} + \frac{N_{\text{on}}}{b} - 1 - \alpha \stackrel{!}{=} 0$$

$$\Rightarrow b_0 = \frac{N_{\text{eff}} + N_{\text{on}}}{1 + \alpha}$$

$$\sigma^2 = \left(\frac{d^2 F(0, b)}{db^2} \right)^{-1} \Big|_{b=b_0} = \left(\frac{N_{\text{eff}} + N_{\text{on}}}{b_0^2} \right)^{-1} = \frac{N_{\text{eff}} + N_{\text{on}}}{(1 + \alpha)^2}$$

$$b) \lambda = \frac{L_0}{\hat{L}} = \exp\left(\ln\left(\frac{L_0}{\hat{L}}\right)\right) = \exp\left(\ln(L_0) - \ln(\hat{L})\right) = \exp(-F_0 - (-F))$$

$$= \exp\left[N_{\text{eff}} \ln\left(\frac{b_0}{\hat{L}}\right) + N_{\text{on}} \ln\left(\frac{\alpha b_0}{\hat{L} + \alpha \hat{L}}\right) - (1 + \alpha)(\hat{L} - b_0) + \hat{L}\right]$$

$$= \exp\left[N_{\text{eff}} \ln\left(\frac{1 + \frac{N_{\text{on}}}{N_{\text{eff}}}}{1 + \alpha}\right) + N_{\text{on}} \ln\left(\alpha \frac{N_{\text{eff}} + N_{\text{on}}}{1 + \alpha} \cdot \frac{1}{N_{\text{on}} - \alpha N_{\text{eff}} + \alpha N_{\text{eff}}}\right) - (1 + \alpha)\left(\frac{N_{\text{eff}} + N_{\text{on}}}{1 + \alpha} - N_{\text{eff}}\right) + N_{\text{on}} - \alpha N_{\text{eff}}\right]$$

$$= \exp\left[N_{\text{eff}} \ln\left(\frac{1 + \frac{N_{\text{on}}}{N_{\text{eff}}}}{1 + \alpha}\right) + N_{\text{on}} \ln\left(\alpha \cdot \frac{1 + \frac{N_{\text{eff}}}{N_{\text{on}}}}{1 + \alpha}\right) - N_{\text{eff}} - N_{\text{on}} + N_{\text{eff}} + \alpha N_{\text{eff}} + N_{\text{on}} - \alpha N_{\text{eff}}\right]$$

$$= \exp\left[N_{\text{eff}} \ln\left(\frac{1 + \frac{N_{\text{on}}}{N_{\text{eff}}}}{1 + \alpha}\right) + N_{\text{on}} \ln\left(\alpha \cdot \frac{1 + \frac{N_{\text{eff}}}{N_{\text{on}}}}{1 + \alpha}\right)\right]$$

c) u^2 ist $\chi^2(1)$ verteilt

$$\Rightarrow u^2 \sim \chi^2 \quad \wedge \quad D \sim \chi^2$$

$$\Rightarrow \sqrt{D} \sim |u| \sim |\mathcal{N}(0, 1)|$$

$\Rightarrow \sqrt{D} \equiv$ Konfidenzschranke in Einheiten von σ

d) i) $\sqrt{D} \approx 1,8363 \sigma$

Signifikanz Level = p-Value = p

$$p = 1 - 2 \int_0^{\sqrt{D}} \frac{1}{\sqrt{2\pi}} e^{-\frac{x^2}{2}} dx \approx 0,0663 = 6,63\%$$

ii) $\sqrt{D} \approx 4,3724 \sigma$

$$p \approx 0,0000120 = 0,00120\%$$

A33) $\chi^2 = \sum_{i=1}^n \frac{(x_i - \mu)^2}{\sigma_i^2}$

a) $\chi^2 = \frac{(37,6 \text{ meV} - 37,3 \text{ meV})^2}{(0,5 \text{ meV})^2} + \frac{(32,2 \text{ meV} - 37,3 \text{ meV})^2}{(0,5 \text{ meV})^2} + \dots$

$\approx 6,08$

$\chi^2(p=0,05, \text{DoF}=7) = 14,067 > 6,08$

$\Rightarrow H_0$ kann nicht abgelehnt werden

b) $\chi^2 \approx 27,9$

$\chi^2(p=0,05, \text{DoF}=8) = 14,067 < 27,9$

$\Rightarrow H_0$ wird abgelehnt

A34/ a)

$$\Gamma = \frac{\sup_{\theta \in \theta_0} L(\theta|X)}{\sup_{\theta \in \theta} L(\theta|X)}$$

$$= \frac{\sup_{\sigma^2} \prod_{i=1}^N \frac{1}{\sqrt{2\pi}\sigma} \exp\left[-\frac{1}{2\sigma^2}(x_i - \mu_0)^2\right]}{\sup_{\sigma^2, \mu} \prod_{i=1}^N \frac{1}{\sqrt{2\pi}\sigma} \exp\left[-\frac{1}{2\sigma^2}(x_i - \mu)^2\right]}$$

$\Gamma < k_\alpha \Rightarrow H_0$ verwerfen

$\Gamma \geq k_\alpha \Rightarrow H_0$ annehmen

b) $L = \left(\frac{1}{\sqrt{2\pi}\sigma}\right)^N \exp\left[-\frac{1}{2\sigma^2} \sum_i (x_i - \mu)^2\right]$

$$-\ln(L) = F = N \ln(\sqrt{2\pi}) + N \underbrace{\ln(\sigma)}_{\frac{1}{2} \ln(\sigma^2)} + \frac{1}{2\sigma^2} \sum_i (x_i - \mu)^2$$

$$\frac{\partial F}{\partial \sigma^2} = \frac{N}{2\sigma^2} - \frac{1}{2\sigma^4} \sum_i (x_i - \mu)^2 \stackrel{!}{=} 0$$

$$\Leftrightarrow \sigma^2 = \frac{1}{N} \sum_i (x_i - \mu)^2$$

$$\frac{\partial F}{\partial \mu} = -\frac{1}{\sigma^2} \sum_i (x_i - \mu) \stackrel{!}{=} 0$$

$$\Leftrightarrow \sum_i x_i = N\mu$$

$$\Leftrightarrow \mu = \frac{1}{N} \sum_i x_i = \bar{x}$$

c)
$$\Gamma = \frac{\left(\frac{1}{N} \sum_i (x_i - \mu_0)^2\right)^{-\frac{N}{2}}}{\left(\frac{1}{N} \sum_i (x_i - \bar{x})^2\right)^{-\frac{N}{2}}} \exp\left[-\frac{1}{2} \left(\frac{1}{N} \sum_i (x_i - \mu_0)^2\right)^{-1} \sum_i (x_i - \mu_0)^2 - \left(\frac{1}{N} \sum_i (x_i - \bar{x})^2\right)^{-1} \sum_i (x_i - \bar{x})^2\right]$$

$$= \left(\frac{\sum_i (x_i - \mu_0)^2}{\sum_i (x_i - \bar{x})^2}\right)^{-\frac{N}{2}} \exp\left[-\frac{1}{2} (N - N)\right]$$

$$= \left(\frac{\frac{1}{N-1} \sum_i (x_i - \mu_0)^2}{s^2}\right)^{-\frac{N}{2}}$$

$$\Rightarrow \Gamma^{-\frac{2}{N}} = \frac{1}{N-1} \sum_i (x_i - \mu_0)^2 / s^2$$

$$= \frac{1}{s^2(N-1)} \sum_i (x_i^2 - 2x_i\mu_0 + \mu_0^2)$$

$$= \frac{1}{s^2(N-1)} \left(\sum_i x_i^2 - \underbrace{2N\bar{x}\mu_0 + N\mu_0^2}_{N(\mu_0 - \bar{x})^2} + \underbrace{N\bar{x}^2}_{N\bar{x} \frac{1}{N} \sum_i x_i} \right)$$

$$\begin{aligned}
\Gamma^{-\frac{2}{N}} &= \frac{1}{s^2(N-1)} \left(\sum_i (x_i^2) - \underbrace{2N\bar{x}\mu_0 + N\mu_0^2 + N\bar{x}^2}_{N(\mu_0 - \bar{x})^2} - \underbrace{N\bar{x}^2}_{N\bar{x} \frac{1}{N} \sum_i x_i} \right) \\
&= \frac{\sum_i (x_i^2 - \bar{x}x_i) + N(\mu_0 - \bar{x})^2}{(N-1) \cdot s^2} \\
&= \frac{\sum_i (x_i^2 - \bar{x}x_i - \bar{x}x_i + \bar{x}^2) + N(\mu_0 - \bar{x})^2}{(N-1) s^2} \\
&= \frac{\sum_i (x_i^2 - 2\bar{x}x_i) + N\bar{x}^2 + N(\mu_0 - \bar{x})^2}{(N-1) s^2} \\
&= \frac{\sum_i (x_i^2 - 2\bar{x}x_i + \bar{x}^2) + N(\mu_0 - \bar{x})^2}{(N-1) s^2} \\
&= \frac{\sum_i (x_i - \bar{x})^2}{(N-1) s^2} + \frac{N(\mu_0 - \bar{x})^2}{(N-1) s^2} \\
&= 1 + \frac{T^2}{N-1}
\end{aligned}$$

$$\Leftrightarrow \sqrt{(N-1)(\Gamma^{-\frac{2}{N}} - 1)} = T$$

$$T \stackrel{!}{\geq} k_\alpha \Leftrightarrow T \stackrel{!}{\leq} k_\alpha^*$$

$$(T \stackrel{!}{\leq} k_\alpha \Leftrightarrow T \stackrel{!}{\geq} k_\alpha^*)$$

k_α^* kann in Tabellen nachgeschlagen werden

$$d) T = \frac{\sqrt{25} (205 \text{ ml} - 200 \text{ ml})}{10 \text{ ml}} = 2,5$$

$\alpha = 5\%$; DoF = 24; einseitiger Test

$\Rightarrow k_\alpha^* = 1,711 < T \Rightarrow H_0$ wird nicht abgelehnt

zweiseitiger Test:

$\Rightarrow k_\alpha^* = 2,064 < |T| \Rightarrow H_0$ wird nicht abgelehnt

A35/ Bayes Theorem:

$$p(H_i | D, I) = \frac{p(H_i | I) p(D | H_i, I)}{p(D | I)}$$

$$p(H_\pi | I) = 0,8$$

$$p(H_k | I) = 0,1$$

$$p(H_p | I) = 0,1$$

$$\begin{aligned} a) p(D | I) &= \sum_i p(H_i | I) p(D | H_i, I) \\ &= 0,8 \cdot 0,13 + 0,1 \cdot 1,5 + 0,1 \cdot 0,5 = 0,304 \end{aligned}$$

$$\Rightarrow p(H_\pi | D, I) = \frac{0,8 \cdot 0,13}{0,304} \approx 34,2\%$$

$$p(H_k | D, I) = \frac{0,1 \cdot 1,5}{0,304} \approx 49,3\%$$

$$p(H_p | D, I) = \frac{0,1 \cdot 0,5}{0,304} \approx 16,4\%$$

$$b) p(H_\pi | D, I) = 36,7\%$$

$$p(H_k | D, I) = 3,02\%$$

$$p(H_p | D, I) = 0,302\%$$

$$c) p(H_\pi | D, I) = 23,7\%$$

$$p(H_k | D, I) = 27,2\%$$

$$p(H_p | D, I) = 55,2\%$$