

Solution

1.(a)

$$\int_{-\infty}^{\infty} \int_{-\infty}^{\infty} p(x, y) dx dy = \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} A e^{-(3x+4y)} dx dy = \frac{A}{12} = 1, \\ A = 12.$$

1.(b)

$$F(x, y) = \int_{-\infty}^y \int_{-\infty}^x p(u, v) du dv = \begin{cases} \int_{-\infty}^y \int_{-\infty}^x 12 e^{-(3u+4v)} du dv, \\ 0, \end{cases} \\ = \begin{cases} (1 - e^{-3x})(1 - e^{-4y}), & x > 0, y > 0, \\ 0, & \text{otherwise.} \end{cases}$$

1.(c)

$$P(0 \leq X < 1, 0 \leq Y < 2) = \int_0^2 \int_0^1 12 e^{-(3x+4y)} dx dy = (1 - e^{-3})(1 - e^{-8}) \approx 0.9499.$$

2.

Let $Z = X - Y$. As $\mathbb{E}\{X\} = \mathbb{E}\{Y\} = 0$, we have $\mathbb{E}\{Z\} = 0$.

Because $\mathbb{E}\{Z\} = 0$, and X and Y are independent of each other, we have:

$$\mathbb{E}\{Z^2\} = \text{var}(Z) = \mathbb{E}\{(X - Y)^2\} = \mathbb{E}\{X^2\} + \mathbb{E}\{Y^2\} = 1$$

Then we have

$$\text{var}(|X - Y|) = \text{var}(|Z|) = \mathbb{E}\{|Z|^2\} - [\mathbb{E}\{|Z|\}]^2 = \mathbb{E}\{Z^2\} - [\mathbb{E}\{|Z|\}]^2,$$

Using

$$\mathbb{E}\{|Z|\} = \int_{-\infty}^{+\infty} |z| \frac{1}{\sqrt{2\pi}} e^{-z^2/2} dz = \frac{2}{\sqrt{2\pi}} \int_0^{+\infty} z e^{-z^2/2} dz = \sqrt{\frac{2}{\pi}},$$

Therefore,

$$\text{var}(|X - Y|) = 1 - \frac{2}{\pi}$$

Note that the answer can be validated using MATLAB:

```
>> X=sqrt(0.5).*randn(1,1000000);  
Y=sqrt(0.5).*randn(1,1000000);  
Z=abs(X-Y);  
var(Z)  
ans = 0.3637
```

3.

By observing that $P_{NK}(n, k)$ can be factorized as:

$$P_{NK}(n, k) = \frac{100^n e^{-100}}{n!} \times \binom{100}{k} p^k (1-p)^{100-k}$$

where one corresponds to the Poisson distribution and another corresponds to the binomial distribution, we easily get:

$$P_N(n) = \begin{cases} \frac{100^n e^{-100}}{n!}, & n = 0, 1, \dots, \\ 0, & \text{otherwise} \end{cases}$$

$$P_K(k) = \begin{cases} \binom{100}{k} p^k (1-p)^{100-k}, & k = 0, 1, \dots, 100 \\ 0, & \text{otherwise} \end{cases}$$

4.

$X \sim \mathcal{U}(0, 3)$ and $Y \sim \mathcal{U}(0, 3)$ imply

$$p(x) = \begin{cases} \frac{1}{3}, & 0 \leq x \leq 3, \\ 0, & x < 0, x > 3. \end{cases} \quad p(y) = \begin{cases} \frac{1}{3}, & 0 \leq y \leq 3, \\ 0, & y < 0, y > 3. \end{cases}$$

Because X and Y are independent of each other, the joint PDF is:

$$p(x, y) = \begin{cases} \frac{1}{9}, & 0 \leq x \leq 3, 0 \leq y \leq 3, \\ 0, & x < 0, x > 3, y < 0, y > 3. \end{cases}$$

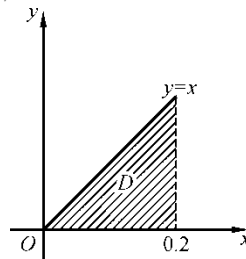
Hence we can easily deduce:

$$P(\max\{X, Y\} \leq 1) = \frac{1}{9}$$

5.

Since random variables X, Y are independent of each other, the joint PDF is:

$$p(x, y) = f_X(x) \cdot f_Y(y) = \begin{cases} \frac{1}{0.2} \cdot 5e^{-5y}, & 0 < x < 0.2 \text{ and } y > 0, \\ 0, & \text{otherwise.} \end{cases}$$



According to the above figure, we have:

$$P(Y \leq X) = \int_0^{0.2} \int_0^x 25e^{-5y} dy dx = \int_0^{0.2} (-5e^{-5x} + 5) dx = e^{-1} = 0.3679.$$

6.(a)

The area of D is $\int_1^{e^2} \frac{1}{x} dx = \ln e^2 - \ln 1 = 2$. Hence the joint PDF is

$$p(x, y) = \begin{cases} \frac{1}{2}, & 1 \leq x \leq e^2 \text{ and } 0 < y \leq \frac{1}{x}, \\ 0, & \text{otherwise.} \end{cases}$$

Hence:

$$p(x) = \begin{cases} \int_0^{1/x} \frac{1}{2} dy = \frac{1}{2x}, & 1 \leq x \leq e^2, \\ 0, & \text{otherwise,} \end{cases}$$

and

$$p(y) = \begin{cases} \int_1^{e^2} \frac{1}{2} dx = \frac{e^2}{2} - \frac{1}{2} & 0 < y \leq \frac{1}{e^2}, \\ \int_1^{1/y} \frac{1}{2} dx = \frac{1}{2y} - \frac{1}{2}, & \frac{1}{e^2} \leq y \leq 1, \\ 0, & \text{otherwise.} \end{cases}$$

6.(b)

It is clear that $p(x, y) \neq p(x)p(y)$. Hence X and Y are not independent.

7.

$$\begin{aligned} \mathbb{E}\{\hat{A}\} &= \mathbb{E}\left\{\frac{1}{N-1} \sum_{n=1}^N r_n\right\} = \frac{1}{N-1} \sum_{n=1}^N \mathbb{E}\{r_n\} = \frac{1}{N-1} \sum_{n=1}^N \mathbb{E}\{A + w_n\} \\ &= \frac{1}{N-1} \sum_{n=1}^N (A + \mathbb{E}\{w_n\}) = \frac{1}{N-1} \sum_{n=1}^N (A + 0) = \frac{N}{N-1} A \end{aligned}$$

$$\begin{aligned} \text{var}(\hat{A}) &= \mathbb{E}\left\{\left(\frac{1}{N-1} \sum_{n=1}^N r_n - \mathbb{E}\{A\}\right)^2\right\} \\ &= \mathbb{E}\left\{\left(\frac{1}{N-1} \sum_{n=1}^N r_n - \frac{N}{N-1} A\right)^2\right\} \\ &= \frac{1}{(N-1)^2} \mathbb{E}\left\{\left(\sum_{n=1}^N w_n + nA - nA\right)^2\right\} = \frac{1}{(N-1)^2} \mathbb{E}\left\{\left(\sum_{n=1}^N w_n\right)^2\right\} \\ &= \frac{1}{(N-1)^2} \sum_{n=1}^N \sum_{m=1}^N \mathbb{E}\{w_n w_m\} = \frac{1}{(N-1)^2} \sum_{n=1}^N \mathbb{E}\{w_n^2\} = \frac{N\sigma_w^2}{(N-1)^2} \end{aligned}$$

We then apply (3.29) to obtain:

$$\text{MSE}(\hat{A}) = \text{var}(\hat{A}) + (A - \mathbb{E}\{A\})^2 = \frac{N\sigma_w^2}{(N-1)^2} + \frac{A^2}{(N-1)^2}$$

8.

Given the event $A = \{\min(X, Y) > 5\}$, we first compute $P(A)$:

$$P(A) = P(X > 5, Y > 5) = \sum_{x=6}^{10} \sum_{y=6}^{10} 0.01 = 0.25$$

Hence $P_{XY|A}(x, y)$ is:

$$P_{XY|A}(x, y) = \begin{cases} 0.04, & x = 6, 7, \dots, 10, y = 6, 7, \dots, 10 \\ 0, & \text{otherwise} \end{cases}$$

9.

Given the event $A = \{X + Y \leq 1\}$, we first compute $P(A)$:

$$P(A) = \int_0^1 \int_0^{1-x} 6e^{-(2x+3y)} dy dx = 1 - 3e^{-2} + 2e^{-3}$$

Hence $P_{XY|A}(x, y)$ is:

$$P_{XY|A}(x, y) = \begin{cases} \frac{6e^{-(2x+3y)}}{1 - 3e^{-2} + 2e^{-3}}, & x + y \leq 1, x \geq 0, y \geq 0 \\ 0, & \text{otherwise} \end{cases}$$

10.

We compute the marginal PDF $P_Y(y)$ first. For $0 \leq y \leq 1$, we have:

$$P_Y(y) = \int_0^1 (x + y) dx = \frac{2y + 1}{2}$$

Hence:

$$P_{X|Y}(x|y) = \frac{P_{XY}(x, y)}{P_Y(y)} = \begin{cases} \frac{2(x + y)}{2y + 1}, & 0 \leq x \leq 1 \\ 0, & \text{otherwise} \end{cases}$$

Due to the symmetry between X and Y , we have:

$$P_{Y|X}(y|x) = \frac{P_{XY}(x, y)}{P_X(x)} = \begin{cases} \frac{2(x + y)}{2x + 1}, & 0 \leq y \leq 1 \\ 0, & \text{otherwise} \end{cases}$$