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Advanced Probabilistic Machine Learning

Lecture 2 - Bayesian linear regression



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Summary of lecture 1 (I/IV)

Conditional probability is defined as

$$p(\mathbf{x}|y) = \frac{p(\mathbf{x},y)}{p(y)}$$
 where $p(y) \neq 0$

Marginalization is defined as

$$p(\mathbf{x}) = \sum_{y} p(\mathbf{x}, y)$$
 or $p(\mathbf{x}) = \int_{y} p(\mathbf{x}, y) dy$

Much of the probability theory can be derived from these two rules.

Bayes' theorem is derived by using the def. of conditional probability twice

$$p(\mathbf{x}|y) = \frac{p(y|\mathbf{x})p(\mathbf{x})}{p(y)}$$



Summary of lecture 1 (II/IV)

In this course we solve problems using Bayes' theorem

$$p(\boldsymbol{\theta}|\mathcal{D}) = \frac{p(\mathcal{D}|\boldsymbol{\theta})p(\boldsymbol{\theta})}{p(\mathcal{D})}$$

- D: observed data
- ullet parameters of some model explaining the data
- $p(\theta)$: **prior** belief of parameters before we collected any data
- $p(\theta|\mathcal{D})$: **posterior** belief of parameters after inferring data
- $p(\mathcal{D}|\theta)$: **likelihood** of the data in view of the parameters
- $p(\mathcal{D})$: The marginal likelihood



Summary of lecture 1 (III/IV)

If we view the quantities as functions of θ , we can disregard the normalization constant p(y).

$$\underbrace{p(\boldsymbol{\theta}|\mathcal{D})}_{\text{posterior}} \propto \underbrace{p(\mathcal{D}|\boldsymbol{\theta})}_{\text{likelihood}} \underbrace{p(\boldsymbol{\theta})}_{\text{prior}}$$

Conjugate prior A prior ensuring that the posterior and the prior belong to the same probability distribution family.

Example: Beta-Binomial

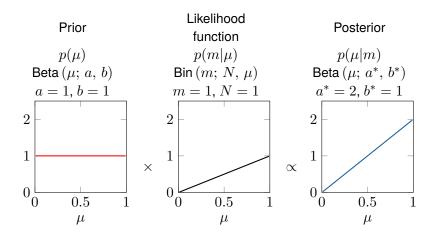
$$\underbrace{\operatorname{Beta}\left(\mu;\,a^{*},\,b^{*}\right)}_{\operatorname{posterior}} \propto \underbrace{\operatorname{Bin}\left(m;\,N,\,\mu\right)}_{\operatorname{likelihood}} \underbrace{\operatorname{Beta}\left(\mu;\,a,\,b\right)}_{\operatorname{prior}} \quad a^{*} = a + m$$

$$b^* = b + N - m$$

Beta distribution is a conjugate prior to the binomial likelihood.



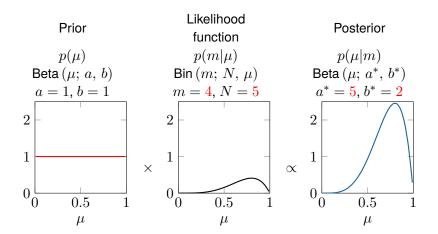
Summary of lecture 1 (IV/IV)



Assume you get N=1 data point, of which m=1 is head, $\mathcal{D}=\{1\}$.



Summary of lecture 1 (IV/IV)



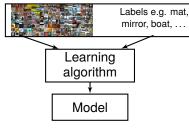
Assume you get N=5 data points, of which m=4 are heads, $\mathcal{D}=\{1,0,1,1,1\}.$



Supervised machine learning

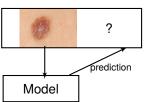
Learning a model from labeled data.

Training data



Predicting output of new data based on this model.

Unseen data



How do we rephrase supervised machine learning as a within the probabilistic methodology?



Supervised machine learning – Probabilistic perspective

Given: Data of inputs & outputs $\mathcal{D} = \{(\mathbf{x}_1, y_1), \dots, (\mathbf{x}_N, y_N)\}.$

Task: Predict the output y_* for a new unseen input x_* .

Solution:

Likelihood Define the likelihood

$$p(\mathbf{y}|\theta, \mathbf{X})$$

2. Prior Define the prior $p(\theta)$

- $\mathbf{X} = \begin{bmatrix} \mathbf{x}_1 \\ \vdots \\ \mathbf{x}^T \end{bmatrix}, \quad \mathbf{y} = \begin{bmatrix} y_1 \\ \vdots \\ \vdots \end{bmatrix}$
- 3. Learning Do inference by applying Bayes' theorem

$$p(\theta|\mathbf{y}, \mathbf{X}) \propto p(\mathbf{y}|\theta, \mathbf{X})p(\theta)$$

4. **Prediction** Compute **predictive distribution** by marginalizing

$$p(y_*|x_*, \mathbf{y}, \mathbf{X}) = \int p(y_*|\theta, x_*)p(\theta|\mathbf{y}, \mathbf{X})d\theta$$



Example: Linear regression model

- Recall the linear regression from lecture 2 in the SML course
- Now we introduce a prior over the parameter w

Linear regression model

$$y_n = \mathbf{w}^\mathsf{T} \mathbf{x}_n + \varepsilon_n, \qquad \varepsilon_n \sim \mathcal{N}(0, \sigma^2), \qquad n = 1, ..., N$$

 $\mathbf{w} \sim p(\mathbf{w}).$

Present assumptions:

- 1. y_n observed **random** variable.
- 2. w unknown deterministic
- 3. \mathbf{x}_n known **deterministic** variable.
- 4. ε_n unknown random variable.
- 5. σ known **deterministic** variable.



Linear regression: Maximum likelihood

Two equivalent ways of expressing the linear regression model:

1.
$$y_n = \mathbf{w}^\mathsf{T} \mathbf{x}_n + \varepsilon_n$$
, $\varepsilon_n \sim \mathcal{N}(0, \sigma^2)$

2.
$$p(y_n | \mathbf{w}) = \mathcal{N}(y_n; \mathbf{w}^\mathsf{T} \mathbf{x}_n, \sigma^2)$$
.

The **likelihood** $p(\mathbf{y} \mid \mathbf{w})$ is given by

$$p(\mathbf{y} \mid \mathbf{w}) = \prod_{n=1}^{N} p(y_n \mid \mathbf{w}) = \prod_{n=1}^{N} \mathcal{N}\left(y_n; \mathbf{w}^{\mathsf{T}} \mathbf{x}_n, \sigma^2\right)$$
$$= \mathcal{N}\left(\mathbf{y}; \mathbf{X}\mathbf{w}, \sigma^2 \mathbf{I}_N\right). \quad \mathbf{X} = \begin{bmatrix} \mathbf{x}_1^{\mathsf{T}} \\ \vdots \\ \mathbf{x}_N^{\mathsf{T}} \end{bmatrix}, \quad \mathbf{y} = \begin{bmatrix} y_1 \\ \vdots \\ y_N \end{bmatrix}$$

The solution if found by maximizing the likelihood

$$\hat{\mathbf{w}} = \arg\max p(\mathbf{y} \,|\, \mathbf{w})$$



Example: Linear regression model

- Recall the linear regression from lecture 2 in the SML course
- Now we introduce a prior over the parameter w

Bayesian linear regression model

$$y_n = \mathbf{w}^\mathsf{T} \mathbf{x}_n + \varepsilon_n, \qquad \varepsilon_n \sim \mathcal{N}(0, \sigma^2), \qquad n = 1, ..., N$$

 $\mathbf{w} \sim p(\mathbf{w}).$

Present assumptions:

- 1. y_n observed **random** variable.
- 2. w unknown random variable. (difference from SML)
- 3. \mathbf{x}_n known **deterministic** variable.
- 4. ε_n unknown random variable.
- 5. σ known **deterministic** variable.



Bayesian linear regression model

Remember Bayes' theorem

$$p(\mathbf{w} \mid \mathbf{y}) = \frac{p(\mathbf{y} \mid \mathbf{w})p(\mathbf{w})}{p(\mathbf{y})}$$

- Prior distribution: $p(\mathbf{w})$ describes the knowledge we have about \mathbf{w} before observing any data.
- **Likelihood:** $p(\mathbf{y} \mid \mathbf{w})$ described how "likely" the observed data is for a particular parameter value.
- Posterior distribution: $p(\mathbf{w} \mid \mathbf{y})$ summarize all our knowledge about \mathbf{w} from the observed data and the model.

In Bayesian linear regression we use a Gaussian distribution as prior

$$p(\mathbf{w}) = \mathcal{N}(\mathbf{w}; \mathbf{m}_0, \Sigma_0)$$

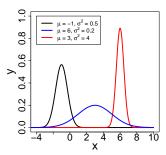


Scalar Gaussian (Normal) distribution

For a scalar variable x, the Gaussian distribution can be written on the form

$$\mathcal{N}\left(x;\mu,\,\sigma^{2}\right) = \underbrace{\frac{1}{\sqrt{2\pi\sigma^{2}}}}_{Z} e^{-\frac{(x-\mu)^{2}}{2\sigma^{2}}}$$

- μ is the mean (expected value of the distribution)
- \bullet σ is the standard deviation
- σ^2 is the variance
- Z is the normalization constant



What if **x** is a vector
$$\mathbf{x} = \begin{pmatrix} x_1 & x_2 & \cdots & x_D \end{pmatrix}^\mathsf{T}$$
?



Multivariate Gaussian

For a D-dimensional vector \mathbf{x} , the **multivariate** Gaussian distribution can be written on the form

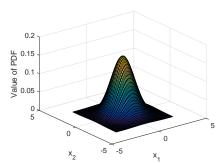
$$\mathcal{N}\left(\mathbf{x}; \boldsymbol{\mu}, \boldsymbol{\Sigma}\right) = \underbrace{\frac{1}{(2\pi)^{D/2}\sqrt{\det \boldsymbol{\Sigma}}}}_{Z} \exp\left(-\frac{1}{2}\underbrace{(\mathbf{x} - \boldsymbol{\mu})^{\mathsf{T}}\boldsymbol{\Sigma}^{-1}(\mathbf{x} - \boldsymbol{\mu})}_{\text{quadratic form}}\right).$$

- μ is the mean vector
- \(\Sigma \) is the covariance matrix
- Z is the normalization constant

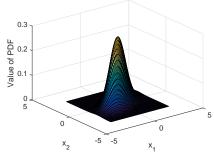
Gaussian $\propto e^{\mathrm{quadratic\ form}}$



Multivariate Gaussian



$$\mathbf{\Sigma} = \begin{pmatrix} 1 & 0 \\ 0 & 1 \end{pmatrix}$$



$$\mathbf{\Sigma} = \begin{pmatrix} 1 & 0.4 \\ 0.4 & 0.5 \end{pmatrix}$$



Partitioned Gaussian - marginalization

Partition the Gaussian random vector $\mathbf{x} \sim \mathcal{N}(\boldsymbol{\mu}, \boldsymbol{\Sigma})$, where $\mathbf{x} \in \mathbb{R}^n$ into two sets of random variables $\mathbf{x}_a \in \mathbb{R}^{n_a}$ and $\mathbf{x}_b \in \mathbb{R}^{n_b}$,

$$\mathbf{x} = \begin{pmatrix} \mathbf{x}_a \\ \mathbf{x}_b \end{pmatrix}, \qquad \boldsymbol{\mu} = \begin{pmatrix} \boldsymbol{\mu}_a \\ \boldsymbol{\mu}_b \end{pmatrix}, \qquad \boldsymbol{\Sigma} = \begin{pmatrix} \boldsymbol{\Sigma}_{aa} & \boldsymbol{\Sigma}_{ab} \\ \boldsymbol{\Sigma}_{ba} & \boldsymbol{\Sigma}_{bb} \end{pmatrix}.$$

Task: Compute the marginal distribution $p(\mathbf{x}_a)$,

$$p(\mathbf{x}_a) = \int p(\mathbf{x}_a, \mathbf{x}_b) d\mathbf{x}_b.$$



Partitioned Gaussian – marginalization

Theorem 1 (Marginalization)

Partition the Gaussian random vector $\mathbf{x} \sim \mathcal{N}(\boldsymbol{\mu}, \boldsymbol{\Sigma})$ according to

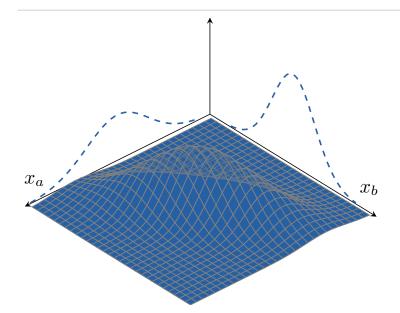
$$\mathbf{x} = egin{pmatrix} \mathbf{x}_a \ \mathbf{x}_b \end{pmatrix}, \qquad oldsymbol{\mu} = egin{pmatrix} oldsymbol{\mu}_a \ oldsymbol{\mu}_b \end{pmatrix}, \qquad oldsymbol{\Sigma} = egin{pmatrix} oldsymbol{\Sigma}_{aa} & oldsymbol{\Sigma}_{ab} \ oldsymbol{\Sigma}_{ba} & oldsymbol{\Sigma}_{bb} \end{pmatrix}.$$

The marginal distribution $p(\mathbf{x}_a)$ is then given by

$$p(\mathbf{x}_a) = \mathcal{N}(\mathbf{x}_a; \boldsymbol{\mu}_a, \boldsymbol{\Sigma}_{aa}).$$

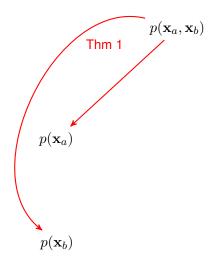


Partitioned Gaussian – marginalization





Partitioned Gaussian – Theorems





Partitioned Gaussian - conditioning

Theorem 2 (Conditioning)

Partition the Gaussian random vector $\mathbf{x} \sim \mathcal{N}(\boldsymbol{\mu}, \boldsymbol{\Sigma})$ according to

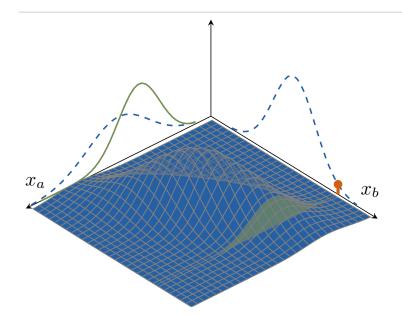
$$\mathbf{x} = \begin{pmatrix} \mathbf{x}_a \\ \mathbf{x}_b \end{pmatrix}, \qquad \boldsymbol{\mu} = \begin{pmatrix} \boldsymbol{\mu}_a \\ \boldsymbol{\mu}_b \end{pmatrix}, \qquad \boldsymbol{\Sigma} = \begin{pmatrix} \boldsymbol{\Sigma}_{aa} & \boldsymbol{\Sigma}_{ab} \\ \boldsymbol{\Sigma}_{ba} & \boldsymbol{\Sigma}_{bb} \end{pmatrix}.$$

The conditional distribution $p(\mathbf{x}_a \mid \mathbf{x}_b)$ is then given by

$$p(\mathbf{x}_{a} | \mathbf{x}_{b}) = \mathcal{N}\left(\mathbf{x}_{a}; \boldsymbol{\mu}_{a | b}, \boldsymbol{\Sigma}_{a | b}\right),$$
$$\boldsymbol{\mu}_{a | b} = \boldsymbol{\mu}_{a} + \boldsymbol{\Sigma}_{ab} \boldsymbol{\Sigma}_{bb}^{-1} (\mathbf{x}_{b} - \boldsymbol{\mu}_{b}),$$
$$\boldsymbol{\Sigma}_{a | b} = \boldsymbol{\Sigma}_{aa} - \boldsymbol{\Sigma}_{ab} \boldsymbol{\Sigma}_{bb}^{-1} \boldsymbol{\Sigma}_{ba}.$$

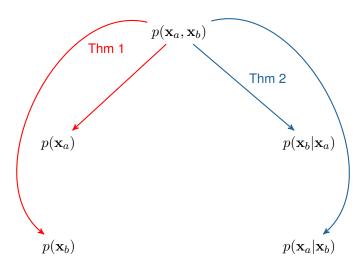


Partitioned Gaussian - conditioning





Partitioned Gaussian – Theorems





Affine transformation of multivar. Gauss

We can also do the opposite: compute $p(\mathbf{x}_a, \mathbf{x}_b)$ based on $p(\mathbf{x}_b \mid \mathbf{x}_a)$ and $p(\mathbf{x}_a)$

Theorem 3 (Affine transformation)

Assume that \mathbf{x}_a , as well as \mathbf{x}_b conditioned on \mathbf{x}_a , are Gaussian distributed according to

$$\begin{split} p(\mathbf{x}_a) &= \mathcal{N}\left(\mathbf{x}_a; \boldsymbol{\mu}_a, \, \boldsymbol{\Sigma}_a\right), \\ p(\mathbf{x}_b \,|\, \mathbf{x}_a) &= \mathcal{N}\left(\mathbf{x}_b; \mathbf{M} \mathbf{x}_a, \, \boldsymbol{\Sigma}_{b \,|\, a}\right). \end{split}$$

Then the joint distribution of x_a and x_b is

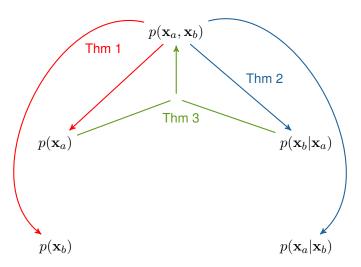
$$p(\mathbf{x}_a, \mathbf{x}_b) = \mathcal{N}\left(\begin{bmatrix} \mathbf{x}_a \\ \mathbf{x}_b \end{bmatrix}; \begin{bmatrix} \boldsymbol{\mu}_a \\ \mathbf{M}\boldsymbol{\mu}_a \end{bmatrix}, \mathbf{R}\right)$$

with

$$\mathbf{R} = egin{bmatrix} \mathbf{\Sigma}_a & \mathbf{\Sigma}_a \mathbf{M}^\mathsf{T} \ \mathbf{M} \mathbf{\Sigma}_a & \mathbf{\Sigma}_{b+a} + \mathbf{M} \mathbf{\Sigma}_a \mathbf{M}^\mathsf{T} \end{bmatrix}$$



Partitioned Gaussian – Theorems





Bayesian linear regression model

Bayesian linear regression model:

$$y_n = \mathbf{w}^\mathsf{T} \mathbf{x}_n + \varepsilon_n,$$
 $\varepsilon_n \sim \mathcal{N}(0, \beta^{-1}),$ $\mathbf{w} \sim p(\mathbf{w}).$ $\beta = \sigma^{-2} \text{ is called}$ the precision

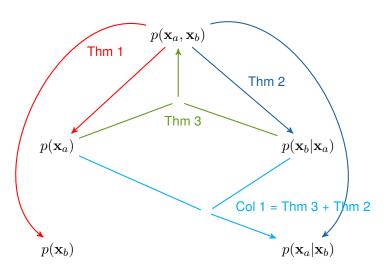
The probabilistic model is given by:

$$p(\mathbf{y} \mid \mathbf{w}) = \mathcal{N}(\mathbf{y}; \mathbf{X}\mathbf{w}, \beta^{-1}\mathbf{I}_N),$$
 likelihood $p(\mathbf{w}) = \mathcal{N}(\mathbf{w}; \mathbf{m}_0, \mathbf{S}_0),$ prior distribution

Task: Compute the posterior distribution: $p(\mathbf{w} \mid \mathbf{y})$.



Partitioned Gaussian – Theorems





Affine transformation of multivar. Gauss

By combining **Theorem 3** and **Theorem 2** we get

Corollary 1 (Affine transformation – conditional)

Assume that x_a , as well as x_b conditioned on x_a , are Gaussian distributed according to

$$\begin{split} p(\mathbf{x}_a) &= \mathcal{N}\left(\mathbf{x}_a; \boldsymbol{\mu}_a, \, \boldsymbol{\Sigma}_a\right), \\ p(\mathbf{x}_b \,|\, \mathbf{x}_a) &= \mathcal{N}\left(\mathbf{x}_b; \mathbf{M}\mathbf{x}_a + \mathbf{b}, \, \boldsymbol{\Sigma}_{b \,|\, a}\right). \end{split}$$

Then the conditional distribution of x_a given x_b is

$$p(\mathbf{x}_a \mid \mathbf{x}_b) = \mathcal{N}\left(\mathbf{x}_a; \boldsymbol{\mu}_{a \mid b}, \boldsymbol{\Sigma}_{a \mid b}\right),$$

with

$$\begin{split} \boldsymbol{\mu}_{a\,|\,b} &= \boldsymbol{\Sigma}_{a\,|\,b} \left(\boldsymbol{\Sigma}_{a}^{-1} \boldsymbol{\mu}_{a} + \mathbf{M}^{\mathsf{T}} \boldsymbol{\Sigma}_{b\,|\,a}^{-1} (\mathbf{x}_{b} - \mathbf{b}) \right), \\ \boldsymbol{\Sigma}_{a\,|\,b} &= \left(\boldsymbol{\Sigma}_{a}^{-1} + \mathbf{M}^{\mathsf{T}} \boldsymbol{\Sigma}_{b\,|\,a}^{-1} \mathbf{M} \right)^{-1}. \end{split}$$



The probabilistic model is given by:

$$\begin{split} p(\mathbf{y} \,|\, \mathbf{w}) &= \mathcal{N}\left(\mathbf{y}; \mathbf{X} \mathbf{w}, \, \beta^{-1} \mathbf{I}_N\right), \quad \text{likelihood} \\ p(\mathbf{w}) &= \mathcal{N}\left(\mathbf{w}; \mathbf{m}_0, \, \mathbf{S}_0\right), \qquad \text{prior distribution} \end{split}$$

Task: Compute the posterior distribution: $p(\mathbf{w} \mid \mathbf{y})$.

Solution: Identify

$$\mathbf{x}_a = \mathbf{w}, \quad \mathbf{x}_b = \mathbf{v},$$

With **Corollary 1** we get the posterior distribution

$$p(\mathbf{w}|\mathbf{v}) = \mathcal{N}(\mathbf{w}; \mathbf{m}_N, \mathbf{S}_N)$$

where

$$\mathbf{m}_N = \mathbf{S}_N (\mathbf{S}_0^{-1} \mathbf{m}_0 + \beta \mathbf{X}^\mathsf{T} \mathbf{y}),$$

$$\mathbf{S}_N^{-1} = \mathbf{S}_0^{-1} + \beta \mathbf{X}^\mathsf{T} \mathbf{X},$$



Consider the problem of fitting a straight line to noisy measurements.

Let the model be $(y_n \in \mathbb{R}, x_n \in \mathbb{R})$

$$y_n = \underbrace{w_0 + w_1 x_n}_{\mathbf{w}^\mathsf{T} \mathbf{x}_n} + \varepsilon_n, \qquad n = 1, \dots, N.$$

where

$$\mathbf{x}_n = \begin{bmatrix} 1 \\ x_n \end{bmatrix}, \quad \mathbf{w} = \begin{bmatrix} \mathbf{w}_0 \\ \mathbf{w}_1 \end{bmatrix}$$

$$\varepsilon_n \sim \mathcal{N}(0, \beta^{-1}), \quad \beta = 5^2.$$

Furthermore, let the prior be

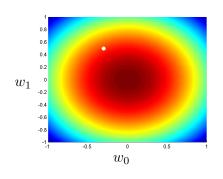
$$p(\mathbf{w}) = \mathcal{N} \begin{pmatrix} \mathbf{w} \mid \begin{pmatrix} 0 & 0 \end{pmatrix}^\mathsf{T}, \alpha^{-1} \mathbf{I}_2 \end{pmatrix},$$

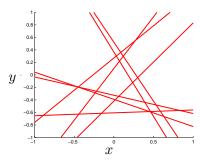
where

$$\alpha = 2$$
.



Plot of the situation before any data arrives.





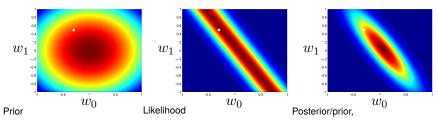
Prior,

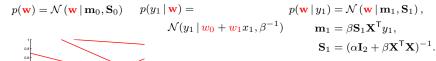
$$p(\mathbf{w}) = \mathcal{N} \begin{pmatrix} \mathbf{w} \mid \begin{pmatrix} 0 & 0 \end{pmatrix}^\mathsf{T}, \frac{1}{2} \mathbf{I}_2 \end{pmatrix}$$

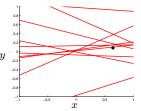
Example of a few realizations from the prior.



Plot of the situation after one measurement has arrived.



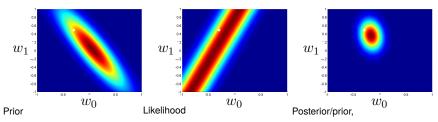




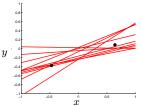
Example of a few realizations from the posterior and the first measurement (black circle).



Plot of the situation after **two** measurements have arrived.



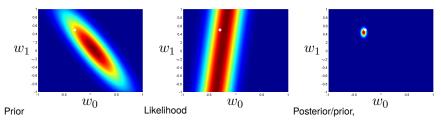
 $p(\mathbf{w} \mid y_1) = \mathcal{N}(\mathbf{w} \mid \mathbf{m}_1, \mathbf{S}_1) \quad p(y_2 \mid \mathbf{w}) = \qquad p(\mathbf{w} \mid y_2) = \mathcal{N}(\mathbf{w} \mid \mathbf{m}_2, \mathbf{S}_2),$ $\mathcal{N}(y_2 \mid \mathbf{w}_0 + \mathbf{w}_1 x_2, \beta^{-1}) \qquad \mathbf{m}_2 = \beta \mathbf{S}_2 \mathbf{X}^{\mathsf{T}} \mathbf{y},$ $\mathbf{S}_2 = (\alpha \mathbf{I}_2 + \beta \mathbf{X}^{\mathsf{T}} \mathbf{X})^{-1}.$



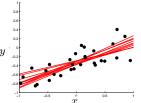
Example of a few realizations from the posterior and the first measurement (black circle).



Plot of the situation after 30 measurements have arrived.



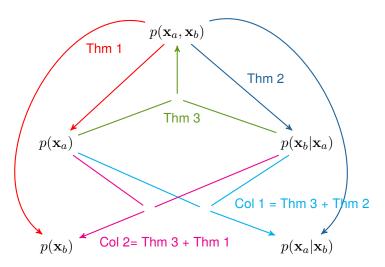
$$\begin{split} p(\mathbf{w} \,|\, y_2) &= \mathcal{N}\left(\mathbf{w} \,|\, \mathbf{m}_2, \mathbf{S}_2\right) \ p(y_3 \,|\, \mathbf{w}) = \\ &\qquad \qquad \mathcal{N}(y_3 \,|\, \mathbf{w}_0 + \mathbf{w}_1 x_3, \beta^{-1}) \\ &\qquad \qquad \mathbf{m}_3 = \beta \mathbf{S}_3 \mathbf{X}^\mathsf{T} \mathbf{y}, \\ &\qquad \qquad \mathbf{S}_3 = (\alpha \mathbf{I}_2 + \beta \mathbf{X}^\mathsf{T} \mathbf{X})^{-1}. \end{split}$$



Example of a few realizations from the posterior and the first measurement (black circle).



Partitioned Gaussian – Theorems





Affine transformation of multivar. Gauss

By combining Theorem 3 and Theorem 1 we get

Corollary 2 (Affine transformation – Marginalization)

Assume that \mathbf{x}_a , as well as \mathbf{x}_b conditioned on \mathbf{x}_a , are Gaussian distributed according to

$$p(\mathbf{x}_a) = \mathcal{N}(\mathbf{x}_a; \boldsymbol{\mu}_a, \boldsymbol{\Sigma}_a),$$
$$p(\mathbf{x}_b | \mathbf{x}_a) = \mathcal{N}(\mathbf{x}_b; \mathbf{M}\mathbf{x}_a + \mathbf{b}, \boldsymbol{\Sigma}_{b|a}).$$

Then the marginal distribution of x_b is then given by

$$p(\mathbf{x}_b) = \mathcal{N}(\mathbf{x}_b; \boldsymbol{\mu}_b, \boldsymbol{\Sigma}_b),$$

where

$$egin{aligned} oldsymbol{\mu}_b &= \mathbf{M} oldsymbol{\mu}_a + \mathbf{b}, \ oldsymbol{\Sigma}_b &= oldsymbol{\Sigma}_{b \mid a} + \mathbf{M} oldsymbol{\Sigma}_a \mathbf{M}^\mathsf{T}. \end{aligned}$$



Predictive distribution

For a new data point (y_*, \mathbf{x}_*) , we have:

$$\begin{split} p(y_*|\mathbf{w}) &= \mathcal{N}\left(y_*; \mathbf{x}_*^\mathsf{T} \mathbf{w}, \, \beta^{-1}\right), \quad \text{likelihood} \\ p(\mathbf{w}|\mathbf{y}) &= \mathcal{N}\left(\mathbf{w}; \mathbf{m}_N, \, \mathbf{S}_N\right) \quad \text{posterior} \end{split}$$

Identify

$$\mathbf{x}_a = \mathbf{w}, \quad \mathbf{x}_b = y_*,$$

With Corollary 2 we get the predictive distribution

$$p(y_*|\mathbf{y}) = \mathcal{N}(y_*; m_*, s_*)$$

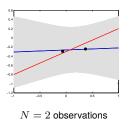
where

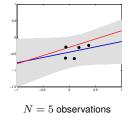
$$m_* = \mathbf{x}_*^\mathsf{T} \mathbf{m}_N$$
$$s_* = \beta^{-1} + \mathbf{x}_*^\mathsf{T} \mathbf{S}_N \mathbf{x}_*$$

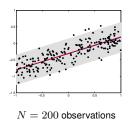


ex) Predictive distribution

Investigating the predictive distribution for the example above







• Gray shaded area: One standard deviation of the predictive distribution as function of x^*

$$p(y_*|\mathbf{y}) = \mathcal{N}\left(y_*; \mathbf{x}_*^\mathsf{T} \mathbf{m}_N, \, \beta^{-1} + \mathbf{x}_*^\mathsf{T} \mathbf{S}_N \mathbf{x}_*\right) \quad \text{where} \quad \mathbf{x}_* = \begin{bmatrix} 1 \\ x_* \end{bmatrix}$$

- Blue line: Mean of predictive distribution
- Black circles: Observations
- Red line: true model



Conjugate priors (I/II)

The probabilistic model with unknown w is given by:

$$\begin{split} p(\mathbf{w}) &= \mathcal{N}\left(\mathbf{w}; \mathbf{m}_0, \, \mathbf{S}_0\right) & \text{prior distribution} \\ p(\mathbf{y} \, | \, \mathbf{w}) &= \mathcal{N}\left(\mathbf{y}; \mathbf{X}\mathbf{w}, \, \beta^{-1}\mathbf{I}_N\right) & \text{likelihood} \end{split}$$

which gives the posterior

$$p(\mathbf{w}|\mathbf{y}) = \mathcal{N}(\mathbf{w}; \mathbf{m}_N, \mathbf{S}_N)$$
 posterior

Note that, using a Gaussian prior gives a Gaussian posterior

$$\underbrace{p(\mathbf{w}|\mathbf{y})}_{\text{Gaussian}} \propto \underbrace{p(\mathbf{y}\,|\,\mathbf{w})}_{\text{Ikelihood}} \underbrace{p(\mathbf{w})}_{\text{p(\mathbf{w})}}$$

Hence, the Gaussian prior is a **conjugate prior** for the Gaussian likelihood unknown **w**.

Q: What if also precision β is unknown?



Conjugate prior (II/II)

The probabilistic model with unknown \mathbf{w} and $\boldsymbol{\beta}$ is given by:

$$p(\mathbf{w}, \boldsymbol{\beta}) = \mathcal{N}(\mathbf{w}; \mathbf{m}_0, \boldsymbol{\beta}^{-1} \mathbf{S}_0) \operatorname{Gam}(\boldsymbol{\beta}; a_0, b_0)$$
 prior
$$p(\mathbf{y} | \mathbf{w}) = \mathcal{N}(\mathbf{y}; \mathbf{X} \mathbf{w}, \boldsymbol{\beta}^{-1} \mathbf{I}_N)$$
 likelihood

which gives the posterior

$$p(\mathbf{w}, \boldsymbol{\beta}|\mathbf{y}) = \mathcal{N}(\mathbf{w}; \mathbf{m}_N, \boldsymbol{\beta}^{-1}\mathbf{S}_N) \operatorname{Gam}(\boldsymbol{\beta}; a_N, b_N)$$
 posterior

Using a Gauss-Gamma prior gives a Gauss-Gamma posterior

$$\underbrace{p(\mathbf{w}, \boldsymbol{\beta} | \mathbf{y})}_{\text{Gauss-Gamma}} \propto \underbrace{p(\mathbf{y} | \mathbf{w}, \boldsymbol{\beta})}_{\text{Iikelihood}} \underbrace{p(\mathbf{w}, \boldsymbol{\beta})}_{\text{p}(\mathbf{w}, \boldsymbol{\beta})}$$

Hence, the Gauss-Gamma prior is a **conjugate prior** for the Gaussian likelihood with unknown \mathbf{w} and unknown precision β .

See further in Exercise 2.11



Non-conjugate priors

In the first two lectures we could solve Bayes' theorem analytically since we used conjugate priors

$$p(\mathbf{w} \mid \mathbf{y}) = \frac{p(\mathbf{y} \mid \mathbf{w})p(\mathbf{w})}{p(\mathbf{y})}$$

However, often you have a personal belief incompatible with the conjugacy.

For example:

- Likelihood with heavy tails
- Multi modal distributions

We have to use approximative inference methods. In this course will discuss two methods

- Monte carlo (lecture 4)
- Variational inference (lecture 6)



A few concepts to summarize lecture 2

Prior distribution: $p(\mathbf{w})$ The representation we have about the unknown parameters \mathbf{w} before we have considered any data.

Likelihood distribution: $p(\mathbf{y} \mid \mathbf{w})$ describes how likely the measurements are for a particular parameter value.

Posterior distribution: $p(\mathbf{w} \mid \mathbf{y})$ summarize our knowledge about the parameters \mathbf{w} based on the information we have from the measurements \mathbf{y} and the model.

Predictive distribution: $p(y_{\star} | \mathbf{y})$ the distribution of unobserved observations y_{\star} conditional on the observed data \mathbf{y} .