# Econometrics - Homework 2

Piche, Alexandre 260478404

October 10, 2014

2.2 Show that the norm of  $\frac{x}{\|x\|}$  is 1.

$$||x|| = \sqrt{x^T x} = \sqrt{x_1^2 + x_2^2 + \ldots + x_n^2} = \text{constant}$$

Thus

$$\|\frac{x}{\|x\|}\| = \frac{1}{\sqrt{x_1^2 + x_2^2 + \dots + x_n^n}} \|x\|$$

Since the denominator  $||x|| = \sqrt{x_1^2 + x_2^2 + \ldots + x_n^n}$  is a constant, we can take it out of the norm. We already know that ||x|| is also equal to  $\sqrt{x_1^2 + x_2^2 + \ldots + x_n^n}$ , it is easy to see that the result is 1.

Compute the norm of the sum and of the difference of  $\boldsymbol{x}$  and  $\boldsymbol{y}$  normalized

The norm of the sum is

$$\sqrt{[\frac{x}{\|x\|} + \frac{y}{\|y\|}]^T[\frac{x}{\|x\|} + \frac{y}{\|y\|}]} = \sqrt{\frac{x^Tx}{\|x\|^2} + \frac{2x^Ty}{\|x\|\|y\|}} + \frac{y^Ty}{\|y\|^2} = \sqrt{1 + \frac{2x^Ty}{\|x\|\|y\|} + 1} = \sqrt{2(1 + \frac{x^Ty}{\|x\|\|y\|})}$$

To take the square root the expression must be non-negative, thus  $\frac{x^Ty}{\|x\|\|y\|} \geq -1$  .

The norm of the difference is given by

$$\sqrt{[\frac{x}{\|x\|} - \frac{y}{\|y\|}]^T[\frac{x}{\|x\|} - \frac{y}{\|y\|}]} = \sqrt{2(1 - \frac{x^Ty}{\|x\|\|y\|)}}$$

Here again  $\frac{x^Ty}{\|x\|\|y\|} \le 1$  for the equation to be non-negative.

Prove the Cauchy-Schwartz inequality and show that this inequality becomes an equality when x and y are parallel.

By taking the last two equations, it is straightforward to obtain the Cauchy-Schwartz inequality  $|x^Ty| \le ||x|| ||y||$ .

If x and y are parallel we can write  $y = \alpha x$  and  $||y|| = \alpha ||x||$ , where  $\alpha$  is a constant. Thus

$$x^T y = x^T \alpha x = \alpha ||x||^2 = \alpha ||x|| ||x|| = ||y|| ||x||$$

Therefore it is an equality when the vectors are parallel.

2.10 Show that  $P = X(W^TX)^{-1}X^T$  is idempotent but not symmetric.

It is easy to show that P is idempotent since

$$PP = X(W^TX)^{-1}W^TX(W^TX)^{-1}W^T = X(W^TX)^{-1}W^T = P$$

We can see that the transpose of P is given by

$$P^T = W(X^T W)^{-1} X^T$$

To prove that they are not equal let's show that they span different subspace. Let y be a n-vector, such that

$$Py = X(W^T X)^{-1} W^T y = Xa$$

where  $a = (W^T X)^{-1} W^T y$  is a k-vector. Thus Py is on S(X).

$$P^T y = W(X^T W)^{-1} X^T y = W b$$

where  $b = (X^T W)^{-1} X^T y$ . Thus  $P^T$  is on S(W). Since  $S(W) \neq S(X), P \neq P^T$  and P is not symmetric.

Characterize the spaces that P and I-P project on to, and show that they are not orthogonal.

Let 
$$P_w=W(W^TW)^{-1}W^T$$
 and  $P_x=X(X^TX)^{-1}X^T$  . We can see that  $(I-P)$  is  $\bot$  to  $S(W)$  since 
$$P_w(I-P)=P_w-W(W^TW)^{-1}W^TX(W^TX)^{-1}W^T=0$$

To prove that P and (I-P) are not orthogonal we need to show that their product is not equal to 0. Since P is on to S(X) and (I-P) is on to  $S^{\perp}(W)$  we can rewrite them as  $P_x$  and  $M_w$  respectively. Thus

$$P_x M_w = P_x (I - P_w) = P_x - P_x P_w$$

where  $P_x P_w = X(X^T X)^{-1} X^T W(W^T W)^{-1} W^T \neq P_x$ . Therefore  $P_x M_w \neq 0$  and S(X) and  $S^{\perp}(W)$  are not  $\perp$ .

#### 2.13 Consider the two regressions

$$y = \beta_1 x_1 + \beta_2 x_2 + \beta_3 x_3 + u$$
, and  $y = \alpha_1 z_1 + \alpha_2 z_2 + \alpha_3 z_3 + u$ ,

where  $z_1 = x_1 - 2x_2$ ,  $z_2 = x_2 + 4x_3$ , and  $z_3 = 2x_1 - 3x_2 + 5x_3$ . Let  $X = [x_1x_2x_3]$  and  $Z = [z_1z_2z_3]$ . Show that the columns of Z can be expressed as linear combinations of the columns of X, that is, that Z = XA, for some  $3 \times 3$  matrix A. Find the elements of this matrix A.

Since Z = XA we can easily show that

$$A = \begin{bmatrix} 1 & 0 & 2 \\ -2 & 1 & -3 \\ 0 & 4 & 5 \end{bmatrix}$$

For A to be invertible we must be able to express X as a linear combination of Z, such that  $X = ZA^{-1}$ 

$$A^{-1} = \begin{bmatrix} 17 & 8 & -2 \\ 10 & 5 & -1 \\ -8 & -4 & 1 \end{bmatrix}$$

Show that the two regressions give the same fitted values and residuals.

$$P_z = Z(Z^T Z)^{-1} Z^T = XA(X^T A^T X A)^{-1} (XA)^T = XAA^{-1}(X^T X)^{-1} (A^T)^{-1} A^T X^T = X(X^T X)^{-1} X^T = P_x$$

Since  $P_z=P_x$ ,  $M_z=M_x$  since they are define by the identity minus their respective fitted value. Therefore they have the same fitted value and residuals.

Precisely how is the OLS estimate  $\beta_1$  related to the OLS estimates  $\alpha_i$ ?

Since  $Z\alpha = y$ ,  $X\beta = y$  and Z = XA we can replace Z in the first equation and get  $XA\alpha = y$ . We can see that  $A\alpha$  must be equal to  $\beta$  for the equality to hold.  $\beta_1 = \alpha_1 + 2\alpha_3$ 

Precisely how is  $\alpha_1$  related to  $\beta_i$ ?

We can rewrite the previous equation as  $\alpha = A^{-1}\beta$ . Therefore  $\alpha_1 = 17\beta_1 + 8\beta_2 - 2\beta_3$ 

## 2.23 For the period 1950:4 to 1996:4, run the regression $\Delta r_t =$

 $\beta_1 + \beta_2 \pi_{t-1} + \beta_3 \Delta y_{t-1} + \beta_4 \Delta r_{t-1} + \beta_5 \Delta r_{t-2} + u_t$ 

Call: lm(formula = dr pi1 + dGDP1 + dr1 + dr2, data = tbills)

Residuals: Min 1Q Median 3Q Max

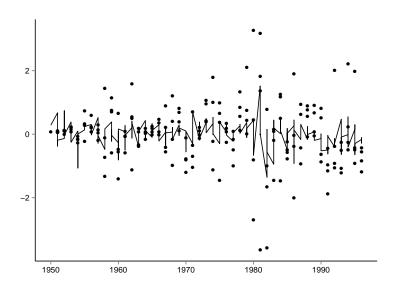


Figure 1: Actual and fitted values

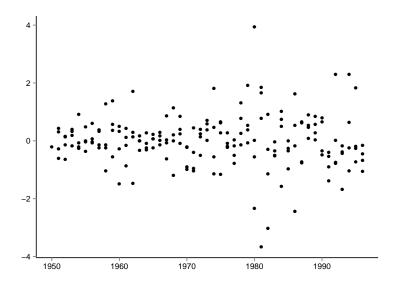


Figure 2: Residuals

### Then regress the residuals on the fitted values and on a constant.

Both coefficient are 0, which is easy to see since the residuals are centered at 0 (for the constant) and the coefficient is also 0 because the fitted values lie on S(X) and the residuals on  $S^{\perp}(X)$ .

```
##
## Call:
## lm(formula = tbills$res ~ tbills$fit)
##
## Residuals:
##
     Min
             1Q Median
                            ЗQ
                                  Max
## -3.660 -0.389 -0.010 0.428
                               3.935
##
## Coefficients:
               Estimate Std. Error t value Pr(>|t|)
##
## (Intercept) -2.84e-18
                           6.50e-02
## tbills$fit 2.12e-16
                           1.88e-01
                                          0
                                                   1
##
## Residual standard error: 0.884 on 183 degrees of freedom
## Multiple R-squared: 1.07e-32,Adjusted R-squared:
## F-statistic: 1.96e-30 on 1 and 183 DF, p-value: 1
```

## Now regress the fitted values on the residuals and on a constant.

We can see that the coefficient of the residuals is 0 since they lie on  $S^{\perp}(X)$  and the intercept is the average of  $\Delta r$ .

```
##
## Call:
## lm(formula = tbills$fit ~ tbills$res)
## Residuals:
##
      Min
               1Q Median
                               3Q
## -1.3741 -0.1802 0.0132 0.2263 1.5055
##
## Coefficients:
              Estimate Std. Error t value Pr(>|t|)
## (Intercept) 0.0134
                        0.0256
                                     0.52
                                           0.6
## tbills$res
                0.0000
                           0.0291
                                     0.00
                                               1.0
##
## Residual standard error: 0.348 on 183 degrees of freedom
## Multiple R-squared: 1.98e-32, Adjusted R-squared:
                                                     -0.00546
## F-statistic: 3.62e-30 on 1 and 183 DF, p-value: 1
```

**2.24** For the same sample period, regress  $\Delta r_t$  on a constant,  $\Delta y_{t-1}$ ,  $\Delta r_{t-1}$ , and  $\Delta r_{t-2}$ . Save the residuals from this regression, and call them  $e_t$ .

```
##
## Call:
## lm(formula = dr ~ dGDP1 + dr1 + dr2, data = tbills)
## Residuals:
##
     Min
            1Q Median
                           3Q
                                 Max
## -3.579 -0.417 -0.031 0.442 4.046
##
## Coefficients:
##
              Estimate Std. Error t value Pr(>|t|)
## (Intercept) -0.1580
                          0.0852
                                   -1.85 0.0654
## dGDP1
               17.4756
                           5.6416
                                     3.10 0.0023
## dr1
                0.2437
                           0.0736
                                     3.31
                                            0.0011
                           0.0720
                                    -2.05
                                          0.0423
## dr2
               -0.1472
## Residual standard error: 0.89 on 181 degrees of freedom
## Multiple R-squared: 0.131, Adjusted R-squared: 0.117
## F-statistic: 9.12 on 3 and 181 DF, p-value: 1.18e-05
```

Regress  $\pi_{t-1}$  on a constant,  $\Delta y_{t-1}$ ,  $\Delta r_{t-1}$ , and  $\Delta r_{t-2}$ . Save the residuals and call them  $v_t$ 

```
##
## lm(formula = pi1 ~ dGDP1 + dr1 + dr2, data = tbills)
## Residuals:
   Min 1Q Median
                          30
                                Max
## -6.528 -2.262 -0.591 1.800 10.366
##
## Coefficients:
##
              Estimate Std. Error t value Pr(>|t|)
## (Intercept) 4.602
                           0.317 14.54 <2e-16
## dGDP1
              -56.346
                          20.953
                                  -2.69 0.0078
## dr1
                0.391
                          0.273
                                  1.43 0.1547
## dr2
                 0.427
                           0.267
                                   1.60 0.1122
##
## Residual standard error: 3.31 on 181 degrees of freedom
## Multiple R-squared: 0.0605, Adjusted R-squared: 0.0449
## F-statistic: 3.89 on 3 and 181 DF, p-value: 0.0101
```

### Now regress $e_t$ on $v_t$ .

We can see that the coefficient is the same than the one for inflation  $(\pi)$  and the residuals are the same than the one in the first regression.

```
##
## Call:
## lm(formula = e ~v + 0)
## Residuals:
            10 Median
   Min
                          3Q
                                Max
## -3.660 -0.389 -0.010 0.428 3.935
##
## Coefficients:
   Estimate Std. Error t value Pr(>|t|)
## v 0.0161 0.0198 0.81 0.42
##
## Residual standard error: 0.881 on 184 degrees of freedom
## Multiple R-squared: 0.00356, Adjusted R-squared: -0.00186
## F-statistic: 0.657 on 1 and 184 DF, p-value: 0.419
```