

PwA Cheatsheet

Common Distributions

Discrete				
Name	pmf	cdf	mean	variance
Binomial(n,p)	$\binom{n}{k} p^k (1-p)^{n-k}$	$F(k; n, p) = \Pr(X \leq k) = \sum_{i=0}^{\lfloor k \rfloor} \binom{n}{i} p^i (1-p)^{n-i}$	np	$np(1-p)$
Neg. Binomial(r,p)	$\binom{i-1}{r-1} p^r (1-p)^{i-r}$	-	$\frac{r}{p}$	$r \frac{1-p}{p^2}$
Bernoulli(p)	$\begin{cases} q = (1-p) & \text{for } k=0 \\ p & \text{for } k=1 \end{cases}$	$\begin{cases} 0 & \text{for } k < 0 \\ 1-p & \text{for } 0 \leq k < 1 \\ 1 & \text{for } k \geq 1 \end{cases}$	p	$p(1-p)$
Uniform(a,b)	$\frac{1}{n}, n = b - a + 1$	$\frac{\lfloor k \rfloor - a + 1}{n}$	$\frac{a+b}{2}$	$\frac{(b-a+1)^2 - 1}{12}$
Geometric(p)	$p(1-p)^{i-1}$	$1 - (1-p)^i$	$\frac{1}{p}$	$\frac{1-p}{p^2}$
Hypergeometric(N,K,n) "k successes \subset N, K suc \in N"	$\frac{\binom{K}{k} \binom{N-K}{n-k}}{\binom{N}{n}}$	-	$n \frac{K}{N}$	$n \frac{K}{N} \frac{(N-K)}{N} \frac{N-n}{N-1}$
Poisson(λ)	$\frac{\lambda^k e^{-\lambda}}{k!}$	$e^{-\lambda} \sum_{i=0}^{\lfloor k \rfloor} \frac{\lambda^i}{i!}$	λ	λ

Continuous				
Name	pdf	cdf	mean	variance
Uniform(a,b)	$\begin{cases} \frac{1}{b-a} & \text{for } x \in [a, b] \\ 0 & \text{otherwise} \end{cases}$	$\begin{cases} 0 & \text{for } x < a \\ \frac{x-a}{b-a} & \text{for } x \in [a, b] \\ 1 & \text{for } x \geq b \end{cases}$	$\frac{a+b}{2}$	$\frac{(b-a)^2}{12}$
Normal(μ, ω^2)	$\frac{1}{\sqrt{2\sigma^2\pi}} e^{-\frac{(x-\mu)^2}{2\sigma^2}}$	$\frac{1}{2} \left[1 + \operatorname{erf}\left(\frac{x-\mu}{\sigma\sqrt{2}}\right) \right]$	μ	σ^2
Exponential(λ)	$\lambda e^{-\lambda x}$	$1 - e^{-\lambda x}$	$1/\lambda$	$1/\lambda^2$
Hazard/Failure Rate Functions				
Survival	Hazard	Distribution	Rate	Book
$\bar{F}(t) = 1 - F(t)$	$\lambda(t) = \frac{f(t)}{\bar{F}(t)}$	$F(t) = 1 - \exp\{-\int_0^t \lambda(t)dt\}$	λ	p217

Events	
Sample Space	$S = \{\text{all possible outcomes}\}$
Event	$E \subset S$
Union (either or both)	$E \cup F$
Intersection (both)	$E \cap F$ or EF
Complement	$E^C = S \setminus E \Rightarrow P(E^C) = 1 - P(E)$
Inclusion-Exclusion	$\hookrightarrow P(A \cup B) = P(A) + P(B) - P(A \cap B)$
DeMorgan's Law	<ol style="list-style-type: none"> $(E_1 \cup \dots \cup E_n)^C = E_1^C \cap \dots \cap E_n^C$ $(E_1 \cap \dots \cap E_n)^C = E_1^C \cup \dots \cup E_n^C$ $0 \leq P(E) \leq 1$ $P(S) = 1$ For mutually excl. events $A_i, i \geq 1$: $P(\cup_{i=1}^{\infty} A_i) = \sum_{i=1}^{\infty} P(A_i)$
Axioms	
Finite S, Equal Probability for all point sets:	$P(A) = A \div S $
Odds of Event	$\alpha = \frac{P(A)}{P(A^C)} = \frac{P(A)}{1-P(A)}$

Conditional Probability and Independence I	
Conditional Probability	$P(F E) = \frac{P(F \cap E)}{P(E)}$
Independence if	$P(F \cap E) = P(F)P(E)$
Multiplication Rule	$P(E_1 E_2 \dots E_n) = P(E_1)P(E_2 E_1) \dots P(E_n E_1 \dots E_{n-1})$
Bayes Formula (simple)	$P(A B) = \frac{P(B A)P(A)}{P(B)}$
Bayes Formula (full)	$P(A_i B) = \frac{P(B A_i)P(A_i)}{\sum_j P(B A_j)P(A_j)}$
Conditional pmf (discrete)	$p_{X Y}(x y) = \frac{p(x,y)}{p_Y(y)}$
Conditional pdf (discrete)	$F_{X Y}(x y) = \sum_{a \leq x} p_{X Y}(a y)$
Conditional Density (continuous)	$f_{X Y}(x y) = \frac{f(x,y)}{f_Y(y)}$
Conditional Probabilities (continuous)	$P\{X \in A Y = y\} = \int_A f_{X Y}(x y) dx$

Random Variables (Discrete)	
Distribution Function	$F(x) = P\{X \leq x\}$
Probability Mass Function	$p(x) = PX = x$
Joint Probability Mass Function	$P(X = x \text{ and } Y = y)$ $= P(Y = y X = x) \cdot P(X = x)$ $= P(X = x Y = y) \cdot P(Y = y)$
Expectation	$E[X] = \sum_{x:p(x)>0} x p(x)$
\hookrightarrow note :	$E[g(X)] = \sum_{x:p(x)>0} g(x) p(x)$
Variance	$Var(X) = E[(X - E[X])^2]$ $= E[X^2] - (E[X])^2$
Standard Derivation	$\sigma = \sqrt{Var(X)}$
Covariance	$Cov(X, Y) = E[(X - E[X])(Y - E[Y])]$ $= E[XY] - E[X]E[Y]$
Moment Gener. Function	$M(t) = E[e^{tX}]$ (same for continuous RVs)

Random Variables (Continuous) I	
Probability Density Function	f such that $P\{X \in B\} = \int_B f(x) dx$
Distribution Function	F such that $\frac{d}{dx} F(x) = f(x)$
Expectation	$E[X] = \int_{-\infty}^{\infty} x f(x) dx$
\hookrightarrow note :	$E[g(X)] = \int_{-\infty}^{\infty} g(x) f(x) dx$
Variance	$Var(X) = E[(X - E[X])^2]$ $= E[X^2] - (E[X])^2$
Standard Derivation	$\sigma = \sqrt{Var(X)}$
Covariance	$Cov(X, Y) = E[(X - E[X])(Y - E[Y])]$ $= E[XY] - E[X]E[Y]$
Joint Probability Mass Function	$P\{(X, Y) \in C\} = \iint_{(x,y) \in C} f(x, y) dx dy$ $P\{X \in A, Y \in B\} = \int_B \int_A f(x, y) dx dy$

Random Variables (Continuous) II	
Marginal pmfs	$f_X(x) = \int_{-\infty}^{\infty} f(x, y) dy$ $f_Y(y) = \int_{-\infty}^{\infty} f(x, y) dx$

More on Expectation, Variance, ..	
$E[X + Y] = E[X] + E[Y]$	
$E[aX] = aE[X]$	
$Var(X + a) = Var(X)$	
$Var(aX + b) = a^2 Var(X)$	
$Var(X + Y) = E[(X + Y)^2] - (E[X + Y])^2$	
$= E[X^2 + 2XY + Y^2] - (E[X] + E[Y])^2$	
$= E[X^2] + 2E[XY] + E[Y^2] -$	
$(E[X])^2 - 2E[X]E[Y] - (E[Y])^2$	
$= Var(X) + Var(Y) + 2(E[XY] - E[X]E[Y])$	
$= Var(X) + Var(Y) + 2(Cov(X, Y))$	
$\Rightarrow E[f(X)g(Y)] = E[f(X)]E[g(Y)]$	
$\Rightarrow E[XY] = E[X]E[Y]$	
Independence	$\Rightarrow Cov(X, Y) = 0$
$\Rightarrow Var(X + Y) = Var(X) + Var(Y)$	
Correlation $corr(X, Y) = \rho(X, Y) = \frac{Cov(X, Y)}{\sqrt{Var(X)Var(Y)}}$	
1. $-1 \leq \rho(X, Y) \leq 1$	
2. Independence $\Rightarrow \rho(X, Y) = 0$	
$Y = mX + cm, m \neq 0$ and c :	
3. $m > 0 \Rightarrow \rho(X, Y) = 1$	
$m < 0 \Rightarrow \rho(X, Y) = -1$	
$E[X] = E[E[X Y]]$	
Disc.: $E[X] = \sum_y E[X Y = y] P\{Y = y\}$	
Cont.: $E[X] = \int_{-\infty}^{\infty} E[X Y = y] f_Y(y) dy$	

Combinatorial Analysis

Order matters and k = n	Permutation
Order does matter and k < n	Variation
Order does not matter and k < n	Combination

Counting		
Basic Counting Principle	Experiments $E_1, E_2, ..E_r$ with $n_1, n_2, ..n_r$ possible outcomes. Total outcomes: $\prod_i n_i$	
Permutations (without Repeats)	$n! = n \cdot (n - 1) \cdot \dots \cdot 1$	
Permutations (with Repeats)	$\frac{n!}{k!} = n \cdot (n - 1) \cdot \dots \cdot (k + 1)$	
Variations (without Repeats)	$n \cdot (n - 1) \cdot \dots \cdot (n - k + 1) = \frac{n!}{(n - k)!}$	
Variations (with Repeats)	$\underbrace{n \cdot \dots \cdot n}_k = n^k$	
Combinations (without Repeats)	$\frac{n!}{(n - k)! k!} = \frac{n(n - 1)(n - 2) \dots (n - k + 1)}{k!} = \binom{n}{n - k} = \binom{n}{k}$	"Binomial Coefficient"
Multinomial Coefficient	$\frac{n!}{n_1! n_2! \dots n_r!} = \binom{n!}{n_1, n_2, \dots, n_r}$ "divide n into r non-overlapping subgroups of sizes n1,n2,.."	
Combinations (with Repeats)	$\frac{(n + k - 1)!}{(n - 1)! k!} = \binom{n + k - 1}{k} = \binom{n + k - 1}{n - 1}$	

Limit Theorems	
	$Z_n = \frac{((X_1 + X_2 + \dots + X_n) - n\mu)}{\sigma \sqrt{n}}$
Central Limit Theorem	Then as $n \rightarrow \infty$ $P(Z_n \leq x) \rightarrow \int_{-\infty}^x \frac{1}{\sqrt{2\pi}} \exp(-\frac{1}{2}u^2) du$ $i.e. P(Z_n \leq x) \rightarrow P(Y \leq X)$ where $Y \sim N(0, 1)$ $E[X_i] = \mu \quad Var(X_i) = \sigma^2$
Weak Law of Large Numbers	$s_n = \frac{1}{n}(X_1 + \dots + X_n)$ then for any $\epsilon > 0$ $\lim_{n \rightarrow \infty} P(s_n - \mu > \epsilon) = 0$
Strong Law of Large Numbers	$P\{\lim_{n \rightarrow \infty} (X_1 + X_2 + \dots + X_n) \div n = \mu\} = 1$
Markov's Inequality	$p\{x \geq a\} \leq \frac{E[X]}{a}$
Chebyshev's Inequality	$E[Y^2] < \infty, \forall a > 0.$ $P(Y \geq \frac{1}{a^2} E[Y^2])$
↪	$P(X - \mu \geq a) \leq \frac{\sigma^2}{a^2}$
One-sided Chebyshev (mean 0)	$P\{X \geq a\} \leq \frac{\sigma^2}{\sigma^2 + a^2}$ $P\{X \geq a\} \leq e^{-ta} M(t) \quad t > 0$
Chernoff Bounds	$P\{X \leq a\} \leq e^{-ta} M(t) \quad t < 0$

Markov Chains

Discrete

$P_{i,j} = P(\text{system is in state } j \text{ at time } n + 1 \mid \text{system in state } i \text{ at time } n)$

Transition Matrix: $P = \begin{pmatrix} p_{1,1} & p_{1,2} & \cdots & p_{1,n} \\ p_{2,1} & p_{2,2} & \cdots & p_{2,n} \\ \vdots & \vdots & \ddots & \vdots \\ p_{n,1} & p_{n,2} & \cdots & p_{n,n} \end{pmatrix}$

Probability vector $\pi^{(n)}$: Probabilities that we are in state i at n.

$\pi^{(n+1)} = \pi^{(n)} P$
 $\pi^{(n)} = \pi^{(0)} P^n$

A **Markov Chain** is **ergodic**(aperiodic and irreducible) iff there exists $n \in \mathbb{N}^+$ such that P^n has no zero entries. It then has a Steady State Probability vector $\pi = \lim_{n \rightarrow \infty} \pi^{(n)}$ independent of $\pi^{(0)}$.

- 1. $\pi_0 + \pi_1 + \dots + \pi_N - 1 = 1$
- 2. $\pi = \pi P$

Continuous

Poisson

$P(\tilde{N}(t) = k) = \frac{(\lambda t)^k}{k!} e^{-\lambda t}$ if

- 1. For any fixed t , $\tilde{N}(t)$ is a discrete RV
- 2. $\tilde{N}(0) = 0$
- 3. # of events in disjoint intervals are independent
- 4. $\tilde{N}(t + h) - \tilde{N}(t) = \#$ of events in $[t, t + h]$ for $h \rightarrow 0$
- 5. $P(\tilde{N}(h) = 1) = P(\text{event occurs in } [t, t + h]) = \lambda h + E(h)$
 $(E(h)/h \rightarrow 0, \text{ as } h \rightarrow 0)$
- 6. $\frac{1}{h} P(\tilde{N}(h) \geq 2) \rightarrow 0, \text{ as } h \rightarrow 0$

Birth-Death

Birth Rates $\lambda_{i,i+1} = b_i$
Death Rates $\lambda_{i,i-1} = d_i$
 $\lambda_{i,j} = 0$, otherwise

→ Have steady state prob. vector if b_i s and d_i s are non-zero and we have a finite number of states.

- 1. $\pi_0 + \pi_1 + \dots + \pi_N - 1 = 1$
- 2. $\pi_j = \frac{b_0 \cdots b_{j-1}}{d_0 \cdots d_{j-1}} \pi_0$

M/M/S Queue

Customers arrive with Poisson Process rate λ , S servers, Service time exponentially distributed with mean $\frac{1}{\mu}$. State j = j customers in queue,

$b_j = \lambda, d_j = \begin{cases} j\mu, & j = 1, 2, \dots, S \\ S\mu, & j \geq S \end{cases}$

→ Has steady state prob. vector if $\lambda < S\mu$.

M/M/1 Queue

$\pi_j = (1 - \frac{\lambda}{\mu})(\frac{\lambda}{\mu})^j$, Mean Queue Length $E[J] = \frac{\lambda}{\mu - \lambda}$

Surprise, Uncertainty & Entropy	
Entropy	$H(X) := -\sum_k p_X(x_k) \log_2 p_X(x_k)$ $(0 \log_2(0) := 0)$
Surprise	$S(X = x_k) = -\log_2 p_X(x_k)$ 1. $S(1) = 0 \neq S(0)$ (which is undefined) 2. S decreases: $p < q \Rightarrow S(q) < S(p)$ 3. $S(pq) = S(p) + S(q)$ If S is continuous and these are satisfied, $\exists \mathcal{C} > 0. \forall p \in [0, 1], S(p) = -\mathcal{C} \log(p)$
↪ Properties	
Average Uncertainty	$H(X, Y) := -\sum_j \sum_k p_{X,Y}(x_j, y_k) \log_2 p_{X,Y}(x_j, y_k)$ $H_{Y=y_k}(X) :=$
Uncertainty of X given Y	$-\sum_j p_{X (Y=y_k)}(x_j) \log_2 p_{X (Y=y_k)}(x_j)$
Conditional Entropy	$H_Y(X) := \sum_k H_{Y=y_k}(X) p_Y(y_k)$

Coding Theory							
Code \mathcal{C}	A map from $\{x_k\} \subset \mathbb{R}$ into sequences of 0's and 1's. Sequences are called code words.						
Code Word length	$x_k \mapsto 0111 \Rightarrow n + k =$						
Expected Length of Code \mathcal{C}	$E[\mathcal{C}] = \sum_k n_k p_k = \sum_k n_k P(X = x_k)$ No code word extends another one:						
Acceptable Code	<table><tr><td>✗</td><td>✓</td></tr><tr><td>$x_1 \mapsto 0$</td><td>$x_1 \mapsto 0$</td></tr><tr><td>$x_2 \mapsto 00$</td><td>$x_2 \mapsto 10$</td></tr></table>	✗	✓	$x_1 \mapsto 0$	$x_1 \mapsto 0$	$x_2 \mapsto 00$	$x_2 \mapsto 10$
✗	✓						
$x_1 \mapsto 0$	$x_1 \mapsto 0$						
$x_2 \mapsto 00$	$x_2 \mapsto 10$						
Noiseless Theorem	For any acceptable code assigning n_k bits to x_k the following holds: $E[\mathcal{C}] = \sum_k n_k p_k \geq H(X) = -\sum_k p_k \log_2 p_k$ Where $p_k = P(X = x_k)$ and n_k length of a codeword associated with x_k For any discrete RV X there exists an acceptable code with the expected length $E[\mathcal{C}] = L$ such that \hookrightarrow Thrm						
	$H(X) \leq L < H(X) + 1$						

Algorithm for finding an acceptable code with expected length $H(X) \leq E[\mathcal{C}] = L < H(X) + 1$ for discrete RV X:

- 1. Let n_j be the integer satisfying $-\log_2 p_j \geq n_j > -\log_2 p_j + 1$
- 2. Find any acceptable code assigning n_j bits to x_j

There is no unique nearly-optimal code in general. Optimal or nearly-optimal coding depends on the pmg of X.

Common Moment Generating Functions M(t)	
Binomial	$(pe^t + 1 - p)^n$
Neg. Binomial	$[(pe^t) \div (1 - (1 - p)e^t)]^r$
Poisson	$\exp(\lambda(e^t - 1))$
Uniform	$(e^{tb} - e^{ta}) \div (t(b - a))$
Exponential	$\lambda \div (\lambda - t)$
Normal	$\exp(\mu t + ((\sigma^2 t^2) \div 2)$