

Modern Sampling Methods

Class 4: Treatment and Policy Choice

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Outline

- ▶ Setup
- ▶ Example
- ▶ CES Rules and Minmax Regret
- ▶ Local Asymptotic Optimality
- ▶ Empirical Welfare Maximization

Basic Setup

Based on Manski (2004) and Dehejia (2005).

$\mathcal{T} = \{0, 1\}$: set of possible treatments.

$Y(0), Y(1)$: potential outcomes

$X \in \mathcal{X}$: background characteristics.

Let θ be parameters associated with potential outcomes:

$$X \sim F_X(\cdot)$$

$$Y(0)|X = x \sim F_0(\cdot|x, \theta)$$

$$Y(1)|X = x \sim F_1(\cdot|x, \theta)$$

Treatment Assignment Rules and Social Welfare

A treatment assignment rule selects treatment based on X :

$$\delta : \mathcal{X} \rightarrow \{0, 1\}.$$

(Could also allow for randomization.)

Suppose we want to maximize average outcomes

$$W(\theta, \delta, x) = \delta(x)E_{\theta}[Y(1)|X = x] + (1 - \delta(x))E_{\theta}[Y(0)|X = x];$$

$$W(\theta, \delta) = \int W(\theta, \delta, x) dF_X(x).$$

The ideal policy is

$$\delta^*(x) = \mathbf{1} \{E_{\theta}[Y(1)|X = x] \geq E_{\theta}[Y(0)|X = x]\}.$$

Statistical Treatment Rule

This is not feasible in general b/c we do not know θ .

Suppose we have some data that is informative about θ . How to use the past data to inform the future choice of treatment rule?

Statistical Treatment Rule:

Before making our treatment assignment, we observe some data
 $Z \sim P_\theta$

(Assume Z independent of the future individual to be treated.)

We then choose δ based on the data z .

Note the timing:

1. observe Z (e.g. from a randomized controlled trial);
2. take a new individual, and observe their X ;
3. assign this individual to treatment based on her own X as well as the data of others collected in Z .

Notation:

$$\delta(x, z)$$

indicating that the policy depends on data and on any information we have about the new individual.

Ex ante probability of assigning individuals with $X = x$ to treatment:

$$\beta(\delta, x, \theta) = E_{\theta}[\delta(x, Z)] = \int \delta(x, z) dP_{\theta}(z).$$

Ex ante expected social welfare for a given rule δ :

$$\begin{aligned} E_{\theta}[W(\theta, \delta(\cdot, Z))] = \\ \int \int \left\{ \delta(x, z) \cdot E_{\theta}[Y(1)|X = x] \right. \\ \left. + (1 - \delta(x, z)) \cdot E_{\theta}[Y(0)|X = x] \right\} dF_X(x) dP_{\theta}(z). \end{aligned}$$

Example: Dehejia (2005)

GAIN experiment, a randomized evaluation of a job training program in California. (Data from Alameda County.)

Tobit model for earnings of individual i in quarter t :

$$Y_{it}^* = x'_{it1}\beta_1 + T_i \cdot x'_{it1}\beta_2 + x'_{it2}\beta_3 + \epsilon_{it}, \quad \epsilon_{it} \stackrel{iid}{\sim} N(0, \sigma^2),$$

$$Y_{it} = 1(Y_{it}^* > 0)Y_{it}^*.$$

Parameter vector: $\theta = (\beta, \sigma^2)$. The data are:

$$Z = \{(x_{it1}, x_{it2}, T_i, Y_{it}) : i = 1, \dots, n, t = 1, \dots, T\}.$$

Use Bayesian methods to simulate posterior distribution $p(\theta|Z)$.

Hypothetical decision problem: counselor is dealing with a new individual (person $n + 1$), whose covariates $x_{n+1,t}$ are observed and whose earnings will follow the same Tobit model.

Can simulate outcomes for person $n + 1$:

- ▶ Draw θ from posterior $p(\theta|Z)$.
- ▶ Simulate $Y_{n+1}(0)$ given $x_{n+1,t}$ and setting $T_{n+1} = 0$.
- ▶ Simulate $Y_{n+1}(1)$ given $x_{n+1,t}$ and setting $T_{n+1} = 1$.

Then choose treatment that has higher expected outcome.

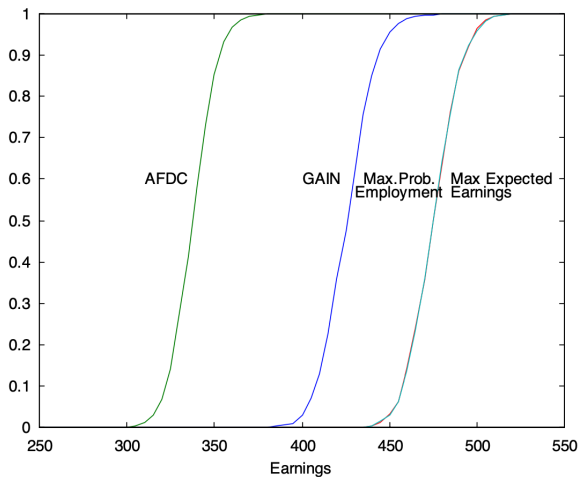


Fig. 7. Predictive distributions for average earnings.

From: Dehejia (2005)

Some other economic applications

- ▶ Job Training Programs: JTPA (Kitagawa and Tetenov 2018)
- ▶ Environmental Policy (Assuncao et al, 2019)
- ▶ Energy Incentives (Knittel and Stolper, 2019)
- ▶ Marketing (Rossi et al 1996, Dube et al 2017)

Manski (2004)

Suppose the covariate X is discrete, taking on possible values $\{x_j : j = 1, \dots, k\}$.

Suppose data Z are obtained from a randomized experiment:

N_j units with $X = x_j$, of which
 N_j^1 treated and N_j^0 controls.

Conditional Empirical Success (CES) Rule:

$$\hat{\beta}_j := \frac{1}{N_j^1} \sum_{i=1}^{N_j} T_{ji} Y_{ji} - \frac{1}{N_j^0} \sum_{i=1}^{N_j} (1 - T_{ji}) Y_{ji}.$$

Then define

$$\hat{\delta}(x_j, Z) = 1(\hat{\beta}_j > 0).$$

Manski's CES rule is nonparametric and intuitive, but not immediately clear whether it is in some sense optimal.

Related question: if covariate X takes on many values (or is continuous), will CES work well or are there alternatives?

To analyze further, we need some measure of a statistical treatment rule's performance.

Consider expected welfare regret:

$$R(\theta, \delta) = E_{\theta} [W(\theta, \delta^*) - W(\theta, \delta)],$$

where δ^* is the ideal rule.

Maximum expected welfare regret can be used as a measure of performance for a rule δ ,

$$\max_{\theta} R(\theta, \delta)$$

For the class of CES rules, Manski provides bounds on maximum expected welfare regret when the randomized experiment is conducted by:

- ▶ Simple randomization
- ▶ Stratified block randomization

(see Class 6 for definitions/discussion of these random assignment mechanisms)

Stoye (2009)

Provides the minmax expected welfare regret optimal rule for some important cases.

First case

- ▶ No covariates
- ▶ Y_0 and Y_1 are binary variables
- ▶ Let \bar{Y}_1 and \bar{Y}_0 be the averages in the two groups.

Consider two different random assignment mechanisms:

(i) Suppose random assignment is by matched pairs, where n is even and we observe exactly $n/2$ treated and $n/2$ controls in our data Z .

Let

$$\hat{\delta}_{MP}(Z) = \begin{cases} 0 & \text{if } \bar{Y}_1 < \bar{Y}_0 \\ \frac{1}{2} & \text{if } \bar{Y}_1 = \bar{Y}_0 \\ 1 & \text{if } \bar{Y}_1 > \bar{Y}_0 \end{cases}$$

(Essentially the CES rule.)

(ii) Simple random assignment, where n^0 and n^1 denote the number of control and treated observations in the sample.

Let

$$\hat{\delta}_S(Z) = \begin{cases} 0 & \text{if } n_1 \left(\bar{Y}_1 - \frac{1}{2} \right) < n_0 \left(\bar{Y}_0 - \frac{1}{2} \right) \\ \frac{1}{2} & \text{if } n_1 \left(\bar{Y}_1 - \frac{1}{2} \right) = n_0 \left(\bar{Y}_0 - \frac{1}{2} \right) \\ 1 & \text{if } n_1 \left(\bar{Y}_1 - \frac{1}{2} \right) > n_0 \left(\bar{Y}_0 - \frac{1}{2} \right) \end{cases}$$

Result: Stoye shows that $\hat{\delta}_{MP}$ and $\hat{\delta}_S$ are minmax rules with respect to expected welfare regret.

Further, this minmax regret result:

- ▶ Extends to allow for bounded outcomes.
- ▶ Extends to hold with covariates: then the minmax regret rule conditions *fully* on X , even if this means very few observations per cell, or even some empty cells.

Results as sharp as Stoye's are difficult to obtain in more complex situations with

- ▶ More complex outcome distributions
- ▶ Structured/parametrized outcome distributions
- ▶ Nonexperimental (observational) data, or data from adaptive experiments
- ▶ Restrictions on the class of rules (e.g. constraints on complexity of rule)
- ▶ etc.

Then it can be useful to turn to large-sample approximations to study alternative rules.

Local Asymptotics for Treatment Assignment

Consider a setting without covariates, but where data are not necessarily from a RCT: there is just some general statistical model

$$Z^n \sim P_\theta, \quad \theta \in \Theta$$

(where n indicates sample size).

The model parameter θ is informative about average treatment effect through:

$$\text{ATE} = W(\theta, 1) - W(\theta, 0) = g(\theta)$$

for some known function g .

As $n \rightarrow \infty$, we will often be able to estimate θ consistently:

$$\hat{\theta} \xrightarrow{P} \theta \quad \Rightarrow \quad g(\hat{\theta}) \xrightarrow{P} g(\theta),$$

and therefore we can learn the optimal rule in the limit.

However, this type of analysis does not capture the finite-sample risk arising from estimation error in $\hat{\theta}$.

One useful way to better reflect finite-sample properties is to consider the behavior of rules when $g(\theta) \approx 0$: let θ_0 satisfy

$$g(\theta_0) = 0,$$

and consider parameters local to θ_0 :

$$\theta = \theta_0 + \frac{h}{\sqrt{n}}.$$

Under this local parametrization:

- ▶ uncertainly about whether $g(\theta) \leq 0$ does not vanish;
- ▶ but classic asymptotic normality theory for parametric and semiparametric statistical models largely carries through.

Hirano and Porter (2009): if P_θ is a regular parametric model, and if $\hat{\theta}$ is an asymptotically efficient estimator (e.g. MLE), then the “plug-in” rule

$$\hat{\delta} = 1(g(\hat{\theta}) > 0)$$

is locally asymptotically minmax regret.

In semiparametric settings, if \hat{g} is a semiparametrically efficient estimator of the ATE, then $\hat{\delta} = 1(\hat{g} > 0)$ is locally asymptotically minmax regret.

Empirical Welfare Maximization

This suggests to replace unknown welfare with a “good” estimate.

Next consider the problem with covariates: $\delta(\cdot)$ is a function of X .

Let

$$W(\delta) = E_X [\delta(X)E[Y(1)|X] + (1 - \delta(X))E[Y(0)|X]].$$

Suppose

$$\widehat{W}(\delta) = \text{estimate of } W(\delta),$$

and we set

$$\hat{\delta} = \arg \max_{\delta} \widehat{W}(\delta).$$

This is the general empirical welfare maximization principle.

Suppose we have (conditionally) randomized experimental data and X has finite support.

Then we can estimate $E[Y|T = 1, X]$ and $E[Y|T = 0, X]$ by empirical conditional averages $\hat{\mu}_1(X)$ and $\hat{\mu}_0(X)$.

Then set

$$\widehat{W}(\delta) = \frac{1}{n} \sum_{i=1}^n [\delta(X_i) \hat{\mu}_1(X_i) + (1 - \delta(X_i)) \hat{\mu}_0(X_i)].$$

This leads to Manski's CES rule.

Next suppose X is continuous.

Then the space of possible functions $\delta(X)$ is very large, and it is not generally possible to estimate $W(\delta)$ uniformly well.

Kitagawa and Tetenov (2018) propose to restrict the class of possible rules δ .

For example, consider only rules of the form

$$\delta(X) = 1(\alpha + \beta X > 0).$$

In applications, it may be more practical to consider simple classes of rules such as this.

For $\delta \in \mathcal{A}$ where \mathcal{A} is sufficiently “small,” it may be possible to construct welfare estimators $\widehat{W}(\delta)$ s.t.

$$\widehat{W}(\delta) \xrightarrow{P} W(\delta),$$

and

$$\sqrt{n} \left(\widehat{W}(\delta) - W(\delta) \right) \xrightarrow{d} N(0, V_\delta),$$

uniformly in $\delta \in \mathcal{A}$.

Then

$$\hat{\delta}(X) = \arg \max_{\delta \in \mathcal{A}} \widehat{W}(\delta)$$

will typically have good properties.