'MATH+ECON+CODE' MASTERCLASS ON MATCHING MODELS, OPTIMAL TRANSPORT AND APPLICATIONS

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Block 9. Quantile methods

LEARNING OBJECTIVES: BLOCK 9

- ▶ Univariate quantiles: properties and uses; quantile regression
- ► Rosenblatt's quantiles
- ► Vector quantiles; vector quantile regression

REFERENCES FOR BLOCK 9

- ► [OTME], Ch. 6.3 and 9.5
- Hedonic models: Ekeland, Heckman and Nesheim, Heckman, Nesheim and Matzkin
- ▶ Quantile regression: Koenker and Bassett (1978), Koenker (2005)
- ► Vector quantiles: Ekeland, G and Henry (2012), Carlier, G and Santambrogio (2010), Carlier, G. and Chernozhukov (2016) Chernozhukov, G, Hallin and Henry (2016)

QUANTILES: DEFINITIONS

- ▶ In dimension one, there are several equivalent ways to define a quantile map $Q_Y : [0,1] \to \mathbb{R}$, among which:
 - ▶ the inverse of a cdf: $Q_Y(t) = F_Y^{-1}(t) = \inf\{y : F_Y(y) > t\};$
 - ▶ the minimizer of $\mathbb{E}\left[\rho_t\left(Y-q\right)\right]$ with $\rho_t\left(w\right)=tw^++(1-t)w^-$;
 - ▶ the map T that maximizes $\mathbb{E}\left[UT\left(U\right)\right]$ subject to $T\#\mathcal{U}\left(\left[0,1\right]\right)=P$.
- ▶ While equivalent in dimension one, the first two definitions cannot be generalized to the case when *Y* is multivariate, while the last one can, thanks to Brenier's theorem.
- ▶ The *vector quantile* associated with definition P is the (unique) gradient of a convex function $Q = \nabla \varphi$ such that if $U \sim \mu = \mathcal{U}\left(\left[0,1\right]^d\right)$, then $\nabla \varphi\left(U\right) \sim P$.

ROSENBLATT'S QUANTILE

Let P be a continuous distributions over \mathbb{R}^2 with finite second moments, and let $\mu = \mathcal{U}\left(\left[0,1\right]^2\right)$. The *Rosenblatt quantile* of distribution P is defined by

$$ar{\mathcal{T}}\left(u_{1},x_{2}
ight)=\left(ar{\mathcal{T}}_{1}\left(u_{1}
ight),ar{\mathcal{T}}_{2}\left(u_{1},u_{2}
ight)
ight)$$

where \bar{T} is given by

$$ar{T}_1\left(u_1,u_2
ight)=F_{Y_1}^{-1}\left(u_1
ight)$$
 , and $ar{T}_2\left(u_1,u_2
ight)=F_{Y_2|Y_1}^{-1}\left(u_2|Y_1=F_{Y_1}^{-1}\left(u_1
ight)
ight)$

▶ The fundamental property of this map is that $\bar{T}\#\mu=P$, and the Jacobian $D\bar{T}$ is lower triangular. The construction extends: the Rosenblatt quantile is the map \bar{T} such that $T\#\mu=P$, and such that

$$\begin{cases} Y_{1} = T_{1} (U_{1}) \\ Y_{2} = T_{2} (U_{1}, U_{2}) \\ ... \\ Y_{M} = T_{M} (U_{1}, U_{2}, ..., U_{M}) \end{cases}$$

has $Y \sim P$ where $U \sim \mu = U([0,1]^d)$, and $T_i(u)$ depends only on

FROM ROSENBLATT TO VECTOR QUANTILE

► For $\lambda > 0$, let $T^{\lambda}(u)$ be the optimal transport map between μ and P relative to surplus $\Phi^{\lambda}(u,y) = u_1y_1 + \lambda u_2y_2$. One has

$$\bar{T}(u) = \lim_{\lambda \to 0^+} T^{\lambda}(u).$$

- ▶ Intuition: because $\lambda \to 0$, the solution will tend to maximize $\mathbb{E}\left[U_1Y_1\right]$ which yields $Y_1 = F_{Y_1}^{-1}\left(U_1\right)$, and over set of couplings (U,Y) that verify this relation, will pick those maximizing $\mathbb{E}\left[U_2Y_2\right]$. Thus the $Y_2 = F_{Y_2|Y_1}^{-1}\left(u_2|F_{Y_1}^{-1}\left(u_1\right)\right)$.
- ► See a rigorous proof in https://arxiv.org/abs/0810.4153.

MOTIVATION: HEDONIC MODELS

- ▶ Hedonic model: A producer of observed characteristics $z \in \mathbb{R}^k$ and latent characteristics $u \in \mathbb{R}$ must choose to produce a good whose quality is a scalar $y \in \mathbb{R}$.
- ▶ The price of a unit of quality y is p(y) (observed) and the cost is C(z,y) (unobserved) so that the profit of choosing quality y is given by

$$p(y) - C(z, y) + uy = -\psi(z, y) + uy$$

where $\psi(z, y) = C(z, y) - p(y)$ is the observed part of minus the profit, which is assumed to be convex in y, and uy is a technology shock (high u's produce high quality at less cost).

► The indirect utility is given by

$$\varphi(z, u) = \max_{v} \left\{ -\psi(z, y) + uy \right\}$$

so by first order conditions, $\partial S\left(z,y\right)/\partial y+u=0$, thus, letting $\psi\left(z,y\right)=-S\left(z,y\right)$, quality y is chosen by consumer $\left(z,u\left(z,y\right)\right)$ such that

$$u(z,y) := \frac{\partial \psi(z,y)}{\partial y}$$

which is nondecreasing in y.

IDENTIFICATION BY QUANTILE (MATZKIN)

- ► The econometrician:
 - assumes U is independent from Z and postulates the distribution μ of U (say, $\mathcal{U}([0,1])$)
 - observes the distribution of choices Y given observable characteristics Z=z.
- ▶ Then (Matzkin), by monotonicity of y(z, u) in u, one has

$$\frac{\partial \psi(z, y)}{\partial y} = F_{Y|Z}(y|z)$$

which identifies $\partial_{\nu}\psi$, and hence the marginal cost $\partial_{\nu}C(z,y)$.

▶ By the same token,

$$\frac{\partial \varphi\left(z,u\right)}{\partial u} = F_{Y|Z}^{-1}\left(u|z\right)$$

identifies $\partial_u \varphi(z, u)$ to $F_{Y|Z}^{-1}$.

PARAMETERIZATION OF THE CONDITIONAL QUANTILE

- ▶ However, the conditional cdf $F_{Y|Z}(y|z)$ or the conditional quantile $F_{Y|Z}^{-1}(u|z)$ are not very easy to estimate nonparametrically. Indeed, the observations are given under the form (Z_i, Y_i) and if Z is continuous, there is not two units i and i' such that $Z_i = Z_{i'}$.
- Quantile regression therefore adopts a parameterization of the conditional quantile which is linear in Z. That is

$$Q_{Y|Z}(u|z) = z^{\mathsf{T}}\beta_u$$

(note that one can always augment z with nonlinear functions of z, so this parameterization is quite general).

► Note that this amounts to taking a linear parameterization of the indirect utility

$$\varphi(z, u) = z^{\mathsf{T}} b_u$$
 with $b_u = \int_0^u \beta_t dt$.

QUANTILE REGRESSION

▶ In order to estimate β_u , first note that

$$Q_{Y|Z}\left(u|z\right) = \arg\min_{q} \mathbb{E}\left[\rho_{u}\left(Y-q\right)|Z=z\right]$$

where $\rho_{u}(w) = tw^{+} + (1 - t)w^{-}$.

► Therefore, if the conditional quantile has the specified form, β_u is the solution to

$$\min_{\beta \in \mathbb{R}^{k}} \mathbb{E}\left[\rho_{u}\left(Y - Z^{\mathsf{T}}\beta\right) \middle| Z = z\right]$$

for each z, and therefore it is the solution to the quantile regression problem introduced by Koenker and Bassett (1978)

$$\min_{\beta \in \mathbb{R}^k} \mathbb{E}\left[\rho_u\left(Y - Z^{\mathsf{T}}\beta\right)\right].$$

QUANTILE REGRESSION AS LINEAR PROGRAMMING

 Koenker and Bassett showed that this problem has a linear programming formulation. Indeed, consider its sample version

$$\min_{\beta \in \mathbb{R}^k} \sum_{i=1}^n \rho_u \left(Y_i - Z_i^\mathsf{T} \beta \right)$$

▶ Introducing $Y_i - Z_i^T \beta = P_i - N_i$ with P_i , $N_i \ge 0$, we have

$$\begin{aligned} \min_{\substack{\beta \in \mathbb{R}^k \\ P_i \geq 0, N_i \geq 0}} \sum_{i=1}^n u P_i + (1-u) N_i \\ s.t. \ P_i - N_i = Y_i - Z_i^\mathsf{T} \beta \end{aligned}$$

therefore β can be obtained by simple linear programming.

MULTIVARIATE EXTENSION OF MATZKIN'S STRATEGY

Now assume quality is a vector $y \in \mathbb{R}^d$, and latent characteristics is $u \in \mathbb{R}^d$ (say, size+amenities). Assume utility of consumer choosing y is given by

$$-\psi(z,y)+u'y$$

where $\psi(z,y) = C(z,y) - P(y)$ is still assumed to be convex in y.

▶ By first order conditions, quality y is chosen by consumer (z, u(z, y)) such that

$$u(z,y) := \nabla_{y} \psi(z,y)$$

which, conditional on z, is "vector nondecreasing" in y in a generalized sense, where vector nondecreasing=gradient of a convex function.

- ► As before, assume:
 - ► The distribution of U given Z = z is μ (say $\mathcal{U}\left([0,1]^d\right)$)
 - ▶ The distribution $F_{Y|Z}$ of Y given Z is observed.

IDENTIFICATION VIA OPTIMAL TRANSPORT

▶ By Brenier's theorem, for each z, $\psi(z,y)$ and $\varphi(z,u)$ are solution to

$$\min_{\psi,\varphi} \mathbb{E}\left[\psi\left(Z,Y\right)\right] + \mathbb{E}\left[\varphi\left(Z,U\right)\right]$$

s.t. $\psi\left(z,y\right) + \varphi\left(z,u\right) > y^{\mathsf{T}}u$

▶ The solution potential $\psi(z,y)$ is convex in y and is such that for $(Z,Y) \sim F_{ZY}$,

$$U =
abla_{y} \psi \left(Z, Y
ight) \sim \mu$$
 and is independent from Z ,

or equivalently, for $U \sim \mu$ independent from Z, one has

$$(Z, Y = \nabla_{u} \varphi(Z, U)) \sim F_{ZY}.$$

▶ This is the "mass transportation approach" (MTA) to identification, applied to a number of contexts by G and Salanié (2012), Chiong, G, and Shum (2014), Bonnet, G, and Shum (2015), Chernozhukov, G, Henry and Pass (2015).

VECTOR QUANTILE REGRESSION

- ▶ In vector quantile regression, one would like to get a more parametric way to write down the dependence of Y in Z; more precisely, linear in Z as in classical quantile regression.
- ▶ One way to do this is to set $\varphi(Z, U) = Z^{\mathsf{T}}b(U)$. The M-K problem becomes

$$\min_{\psi,\varphi} \mathbb{E}\left[\psi\left(Z,Y\right)\right] + \mathbb{E}\left[Z^{\mathsf{T}}b\left(U\right)\right]$$
s.t. $\psi\left(z,y\right) + z^{\mathsf{T}}b\left(u\right) \geq y^{\mathsf{T}}u$

whose primal is

$$\max_{U,Z,Y} \mathbb{E} [U^{\mathsf{T}}Y]$$
s.t. $U \sim \mu$

$$(Z,Y) \sim F_{ZY}$$

$$\mathbb{E} [Z|U] = \mathbb{E} [Z]$$

Section 1

CODING

QUANTILE REGRESSION

- ► Following Koenker, we use Engel's dataset. The package 'quantreg' performs classical quantile regression.
- ► For vector quantile regression, we solve the problem using

```
A1 = kronecker(sparseMatrix(1:n,1:n),matrix(1,1,m))
A2 = kronecker(t(X),sparseMatrix(1:m,1:m))
f1 = matrix(t(nu),nrow=n)
f2 = matrix(mu %*% xbar,nrow=m*r)
e = matrix(1,m*n,1)
A = rbind2(A1,A2)
f = rbind2(f1,f2)
result = gurobi
(list(A=A,obj=c,modelsense="min",rhs=f,ub=e,sense="="),
params=NULL )
```