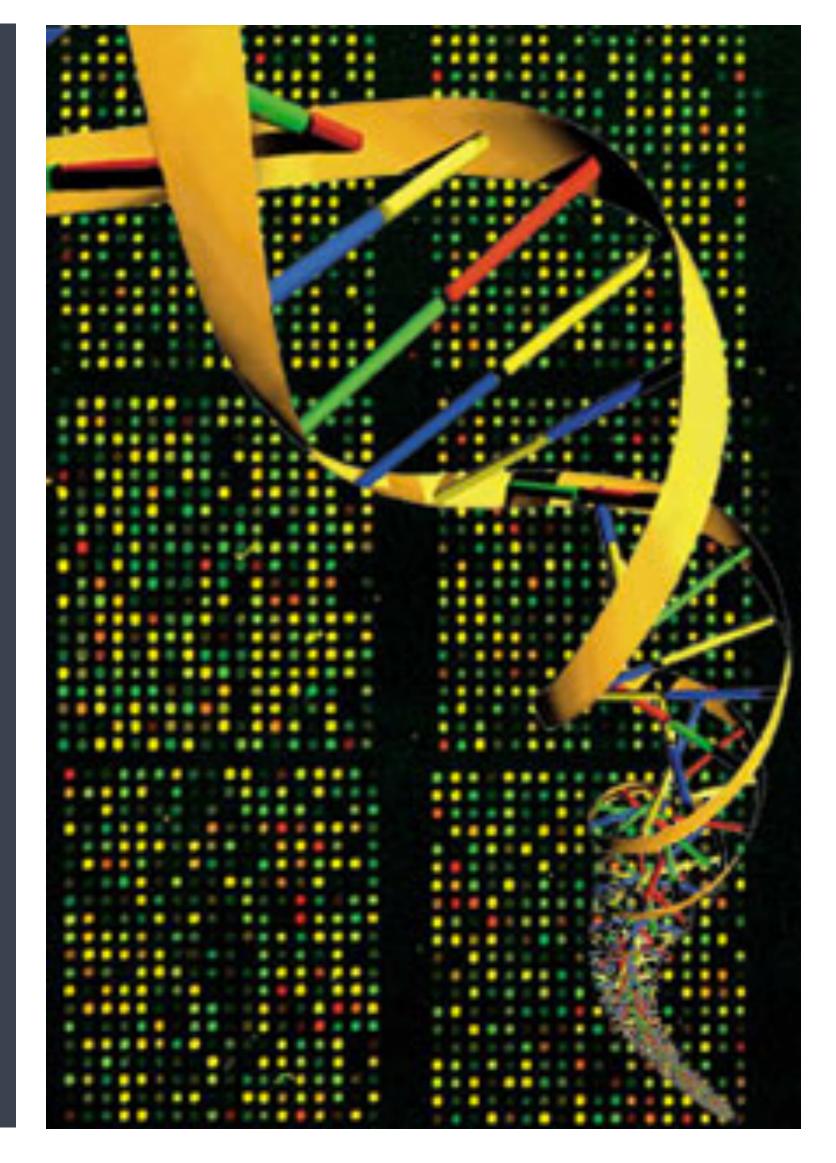
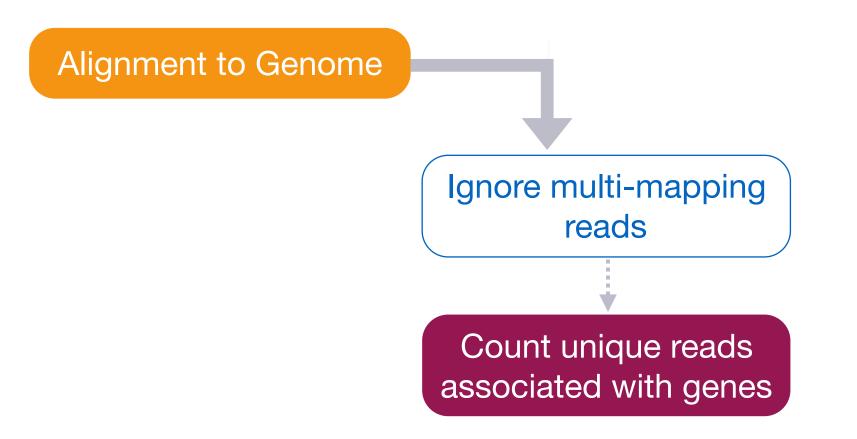
RNA-seq statistical analysis and gene-level differential expression

## Transcriptomics

b to discover functional patterns of biological response to conditions of interest (treatments, environmental influences, mutations etc.)



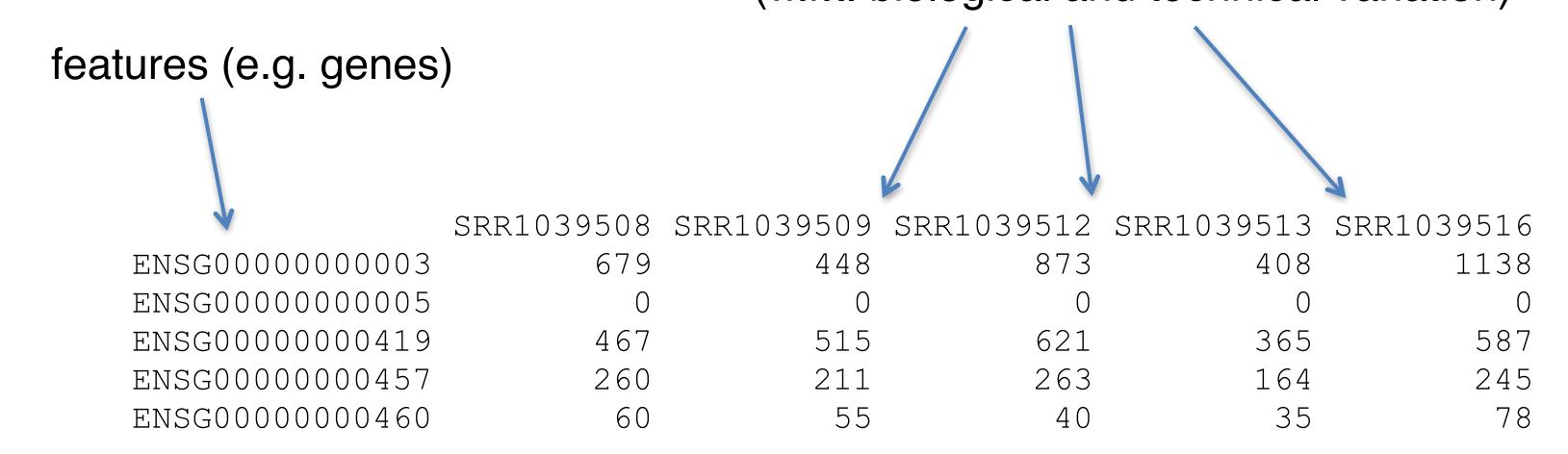


# High-throughput sequencing data

	SRR1039508	SRR1039509	SRR1039512	SRR1039513	SRR1039516
ENSG0000000003	679	448	873	408	1138
ENSG0000000005	0	0	0	0	0
ENSG00000000419	467	515	621	365	587
ENSG0000000457	260	211	263	164	245
ENSG0000000460	60	55	40	35	78

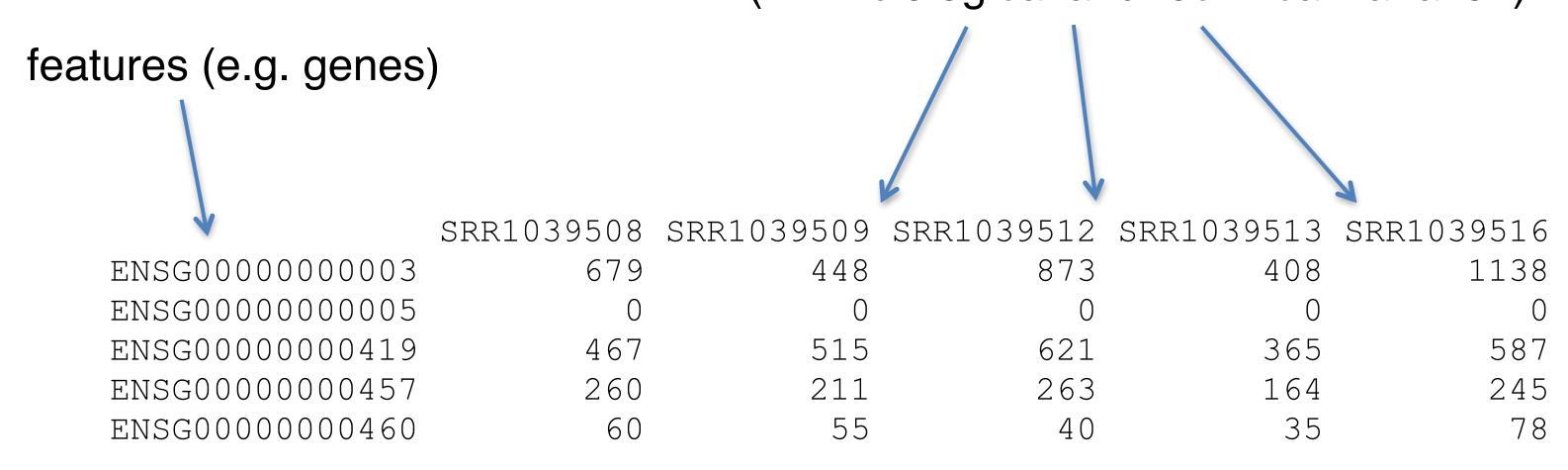
# High-throughput sequencing data

samples: want to see if differences across condition are significant (w.r.t. biological and technical variation)

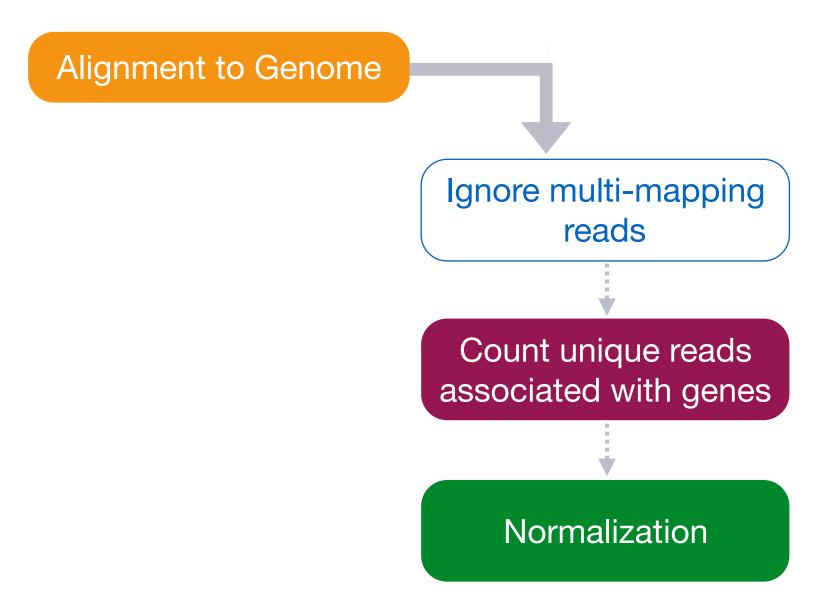


# High-throughput sequencing data

samples: want to see if differences across condition are significant (w.r.t. biological and technical variation)

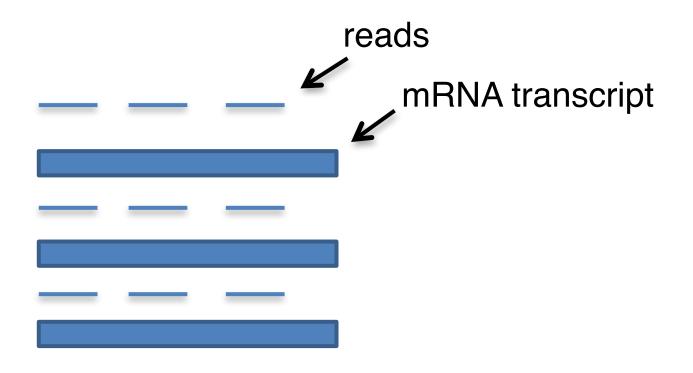


- often observed data consists of counts of reads across features (rows) and samples (columns)
- counts need an appropriate statistical model (normalization and variance modeling)

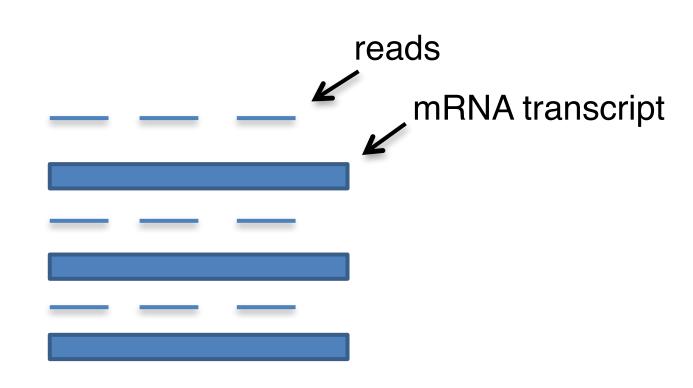


DE workflow :: normalization

#### mRNAs to reads



#### mRNAs to reads



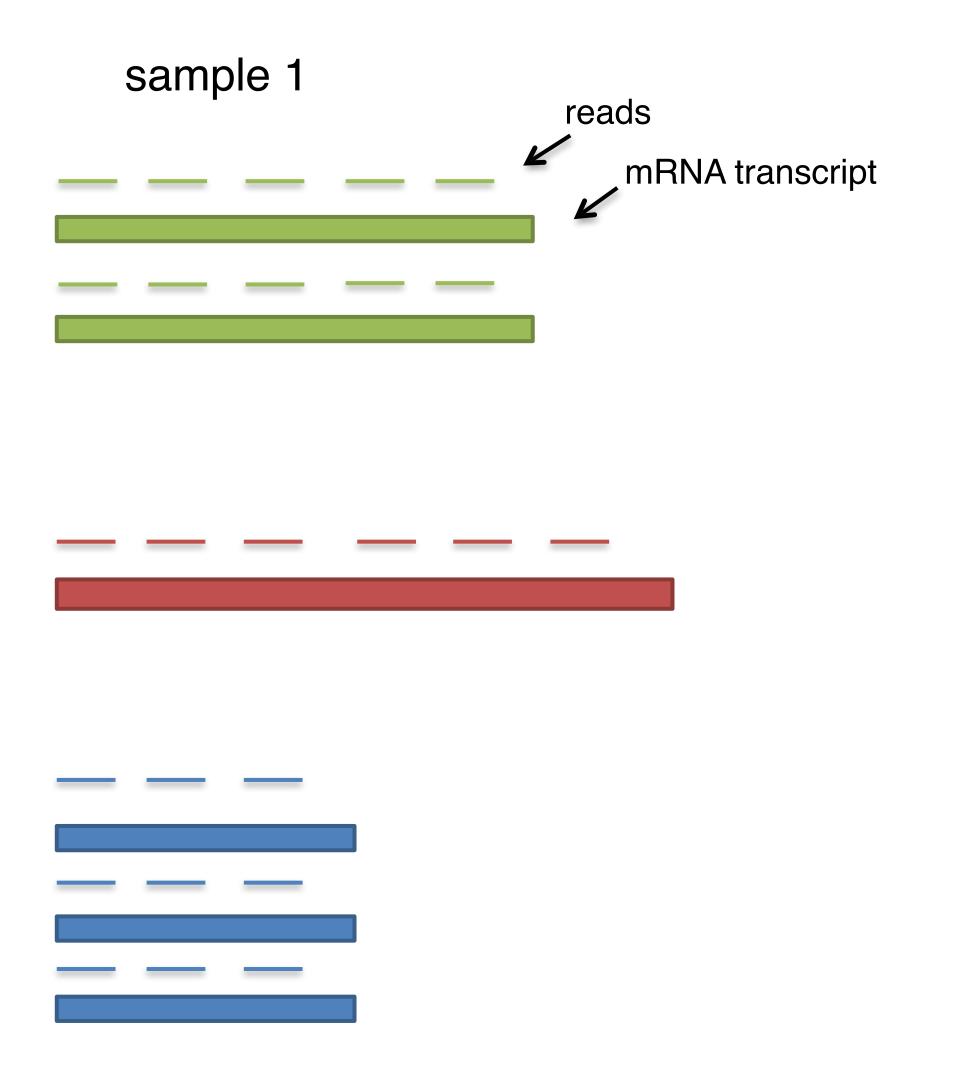
count of mapped reads proportional to:

- expression of RNA
- length of gene
- sequencing depth
- library prep. factors (PCR)
- etc...

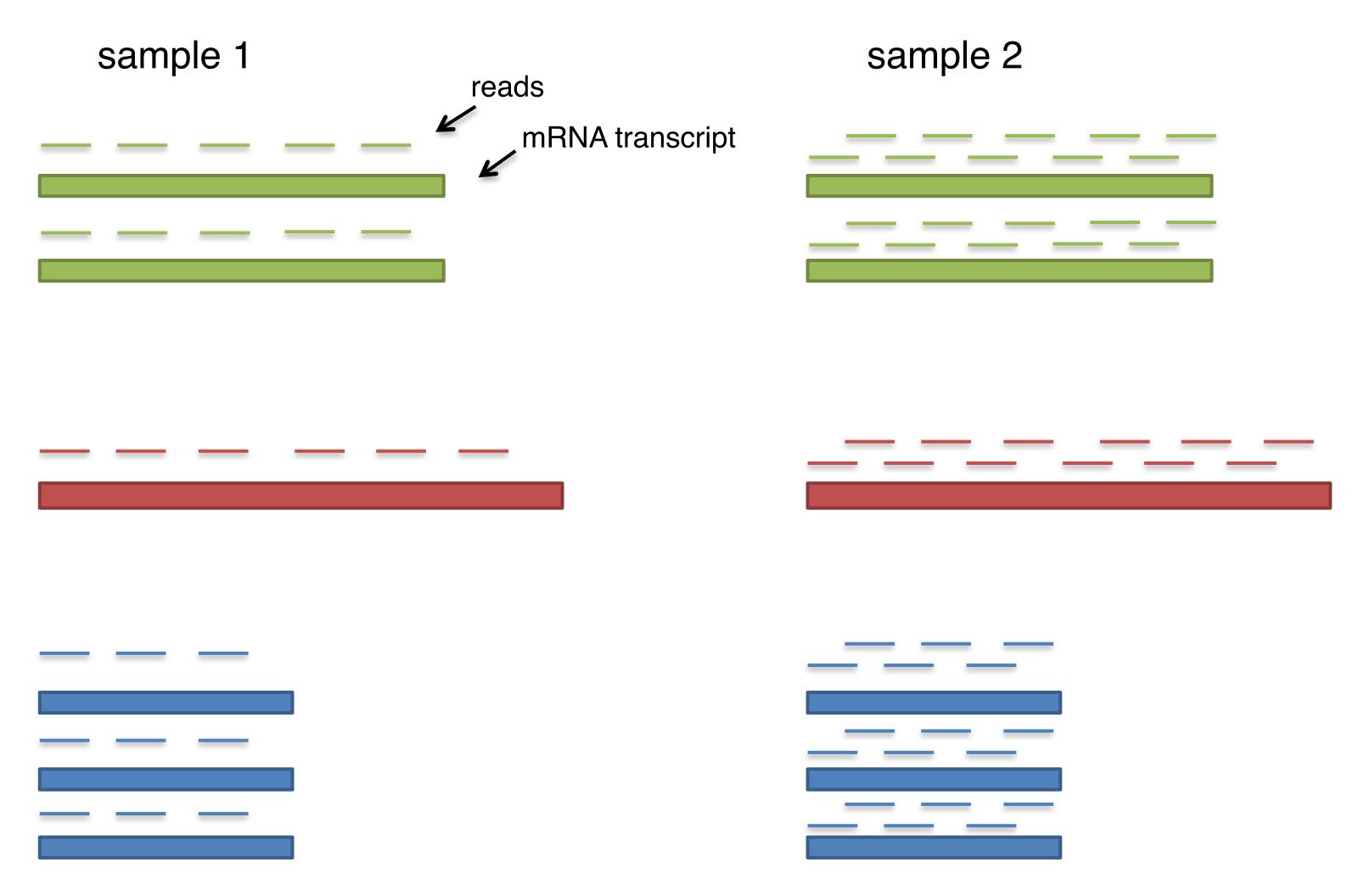
# Length of gene



# Sequencing depth



# Sequencing depth

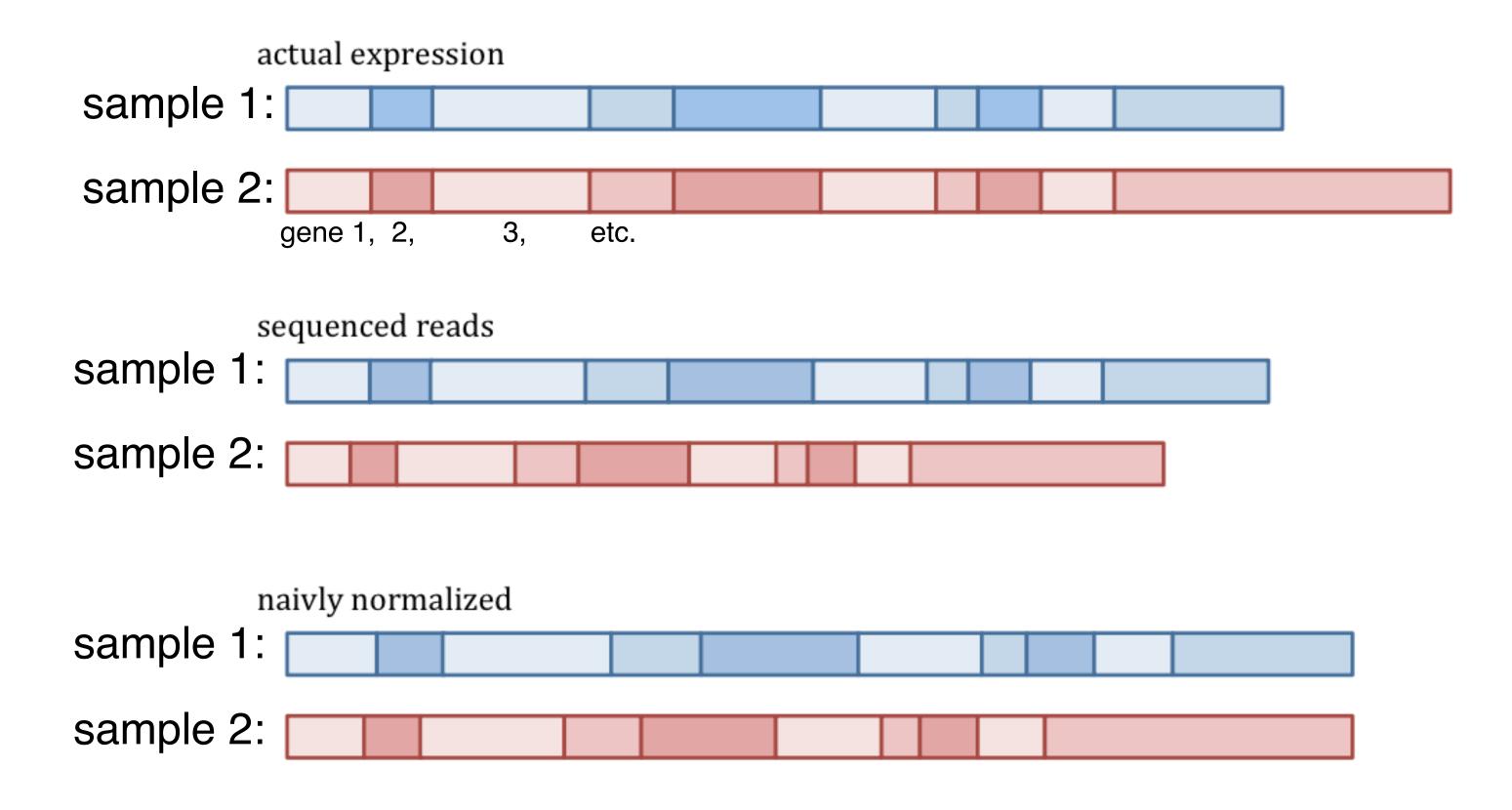


#### (Pre)-scaled measures of expression

- CPM (counts per million) counts scaled by the total number of reads
- > TPM (transcripts per million) the proportion of transcripts in your RNA
- PRESENTATION PROPERTY PROPERTY
- FPKM (fragments per kilobase of exon per million fragments mapped). Same idea for paired end sequencing

Actual counts contain information useful for statistical modeling, and should be used as input for differential expression tools

# Need to have a robust estimator for sequencing depth

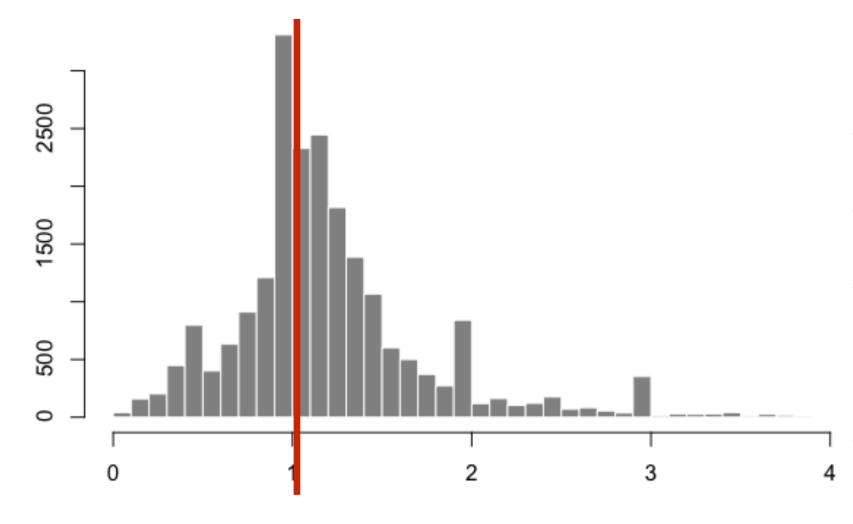


(slide from Simon Anders)

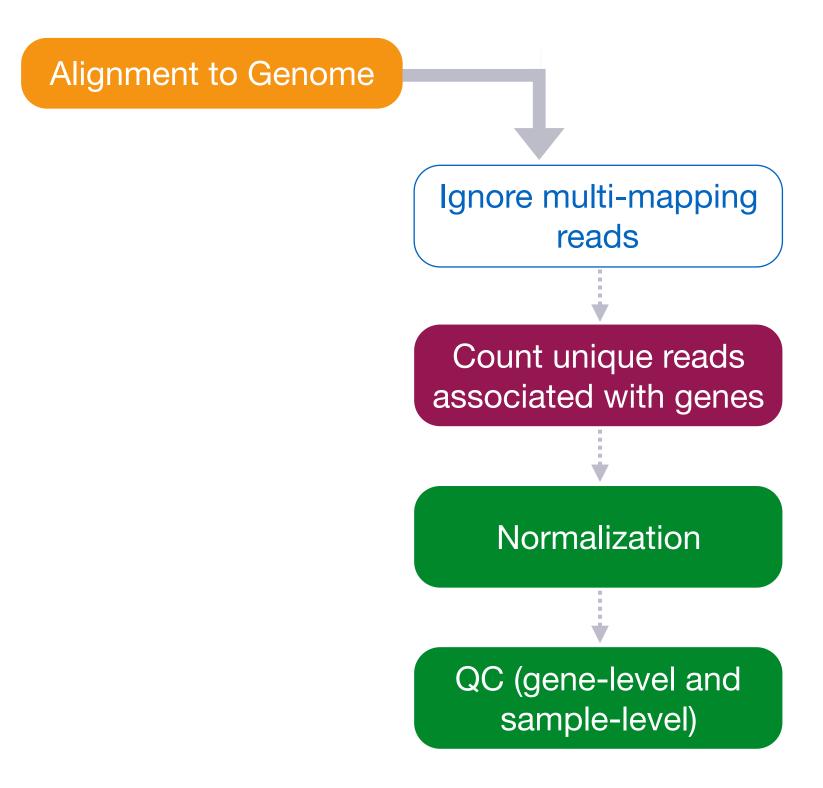
#### Median of ratios method

simple approach & works well for each gene look at the count ratios:

sample 1 / pseudo-reference sample



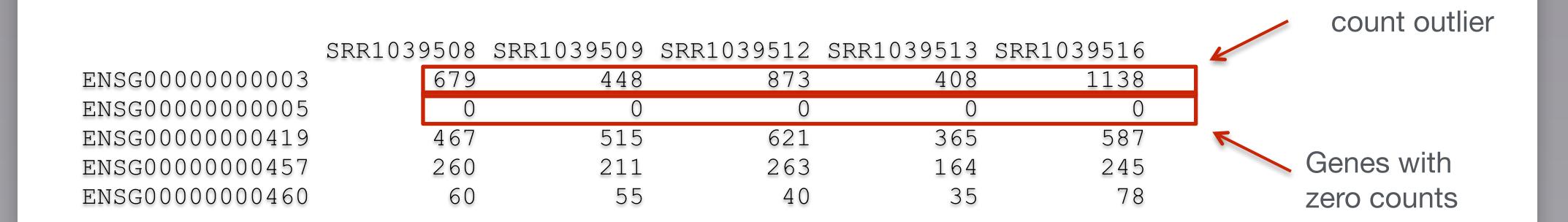
- in general: create a pseudo-reference-sample (rowwise geometric mean)
- calculate ratio of each sample to the reference
- take the median value as the normalization factor
- assumes that not ALL genes are DE (differentially expressed)
- robust to imbalance in up-/down- regulation and large numbers of DE genes



DE workflow :: quality control

	SRR1039508	SRR1039509	SRR1039512	SRR1039513	SRR1039516
ENSG0000000003	679	448	873	408	1138
ENSG0000000005	0	0	0	0	0
ENSG00000000419	467	515	621	365	587
ENSG00000000457	260	211	263	164	245
ENSG00000000460	60	55	40	35	78

	SRR1039508	SRR1039509	SRR1039512	SRR1039513	SRR1039516	
ENSG0000000003	679	448	873	408	1138	
ENSG0000000005	0	0	0	0	0	
ENSG0000000419	467	515	621	365	587	
ENSG0000000457	260	211	263	164	245	Genes with
ENSG0000000460	60	55	40	35	78	zero counts



Genes with extreme

SRR1039508 SRR1039509 SRR1039512 SRR1039513 SRR1039516

ENSG00000000003 ENSG000000000419 ENSG00000000457 ENSG00000000460

9500	SRR1039309	SRR1039312	SRR1039313	SRR1039316	
679	448	873	408	1138	
0	0	0	0	0	
467	515	621	365	587	
260	211	263	164	245	
60	55	40	35	78	

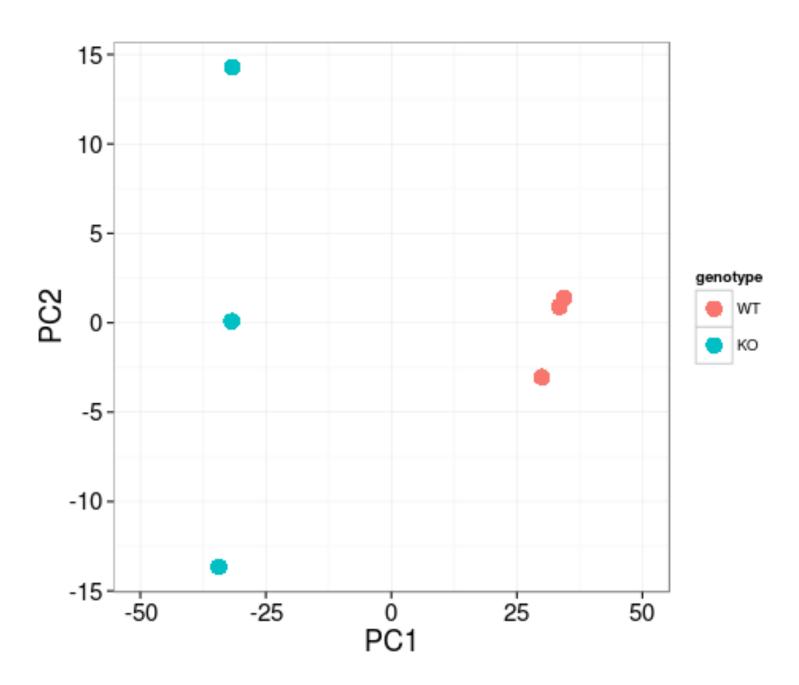
Genes with extreme count outlier



Genes with low mean normalized counts ('Independent filtering')

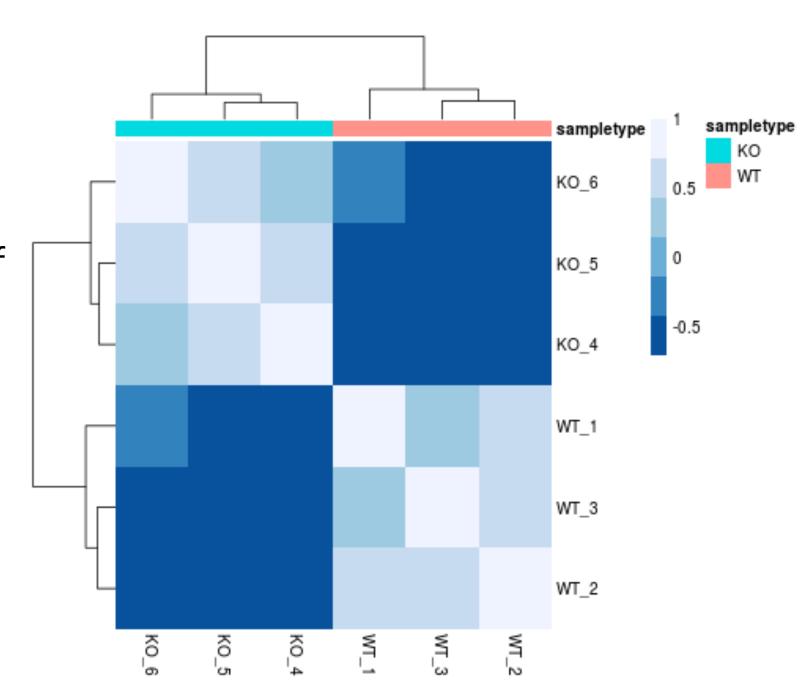
#### QC: Sample-level

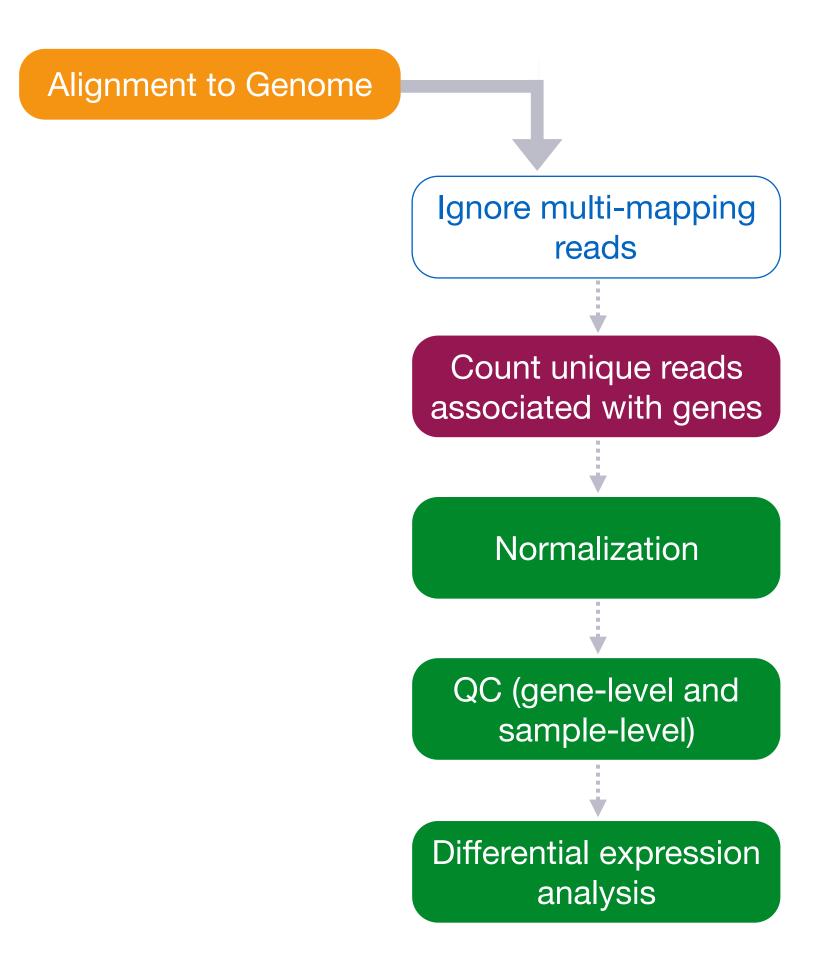
- Principal Component Analysis (PCA): A technique used to emphasize variation and bring out strong patterns in a dataset (dimensionality reduction)
- Project a line through the data points in n dimensional space (n = genes)
- Measure how much variance there is from that line (the distance from each point to the line).
- PC1 explains highest variance, PC2 next highest etc.



#### QC: Sample-level

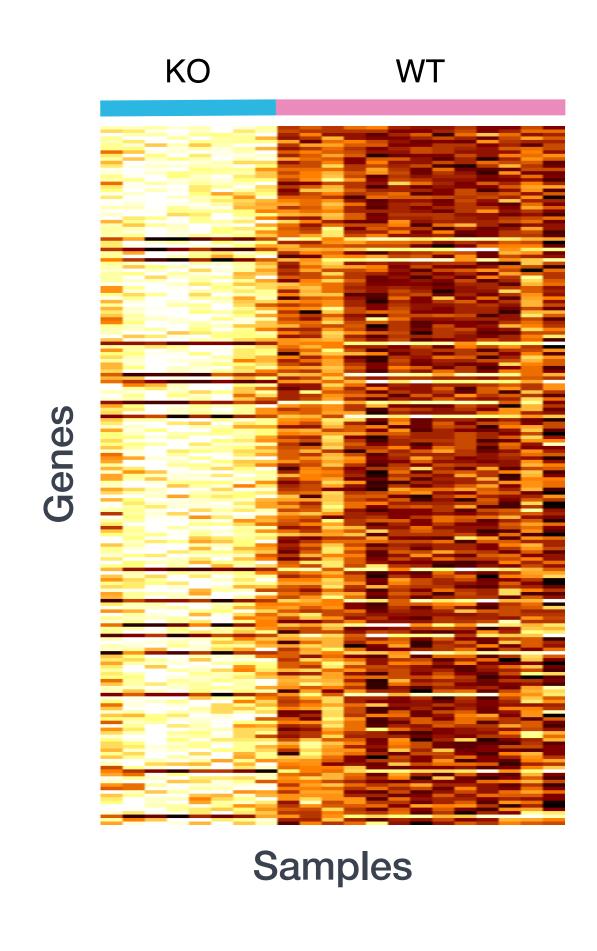
- Identify strong patterns in a dataset and potential outliers
- Correlation or distances for all pairwise combinations of samples.
- Generally high correlations with each other (values higher than 0.80)
- 'Blocks' indicate substructure in the data





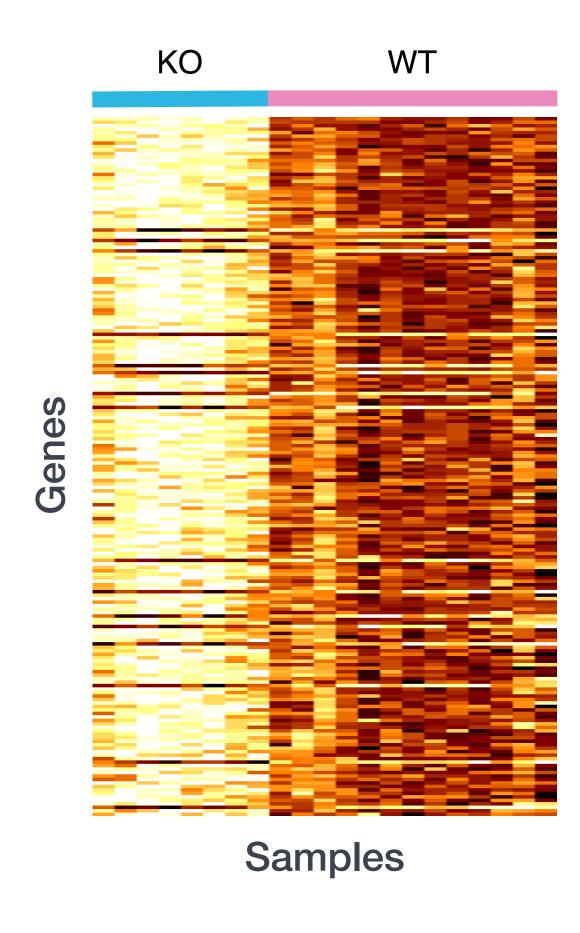
DE workflow:: differential expression

## Identifying differences in gene expression



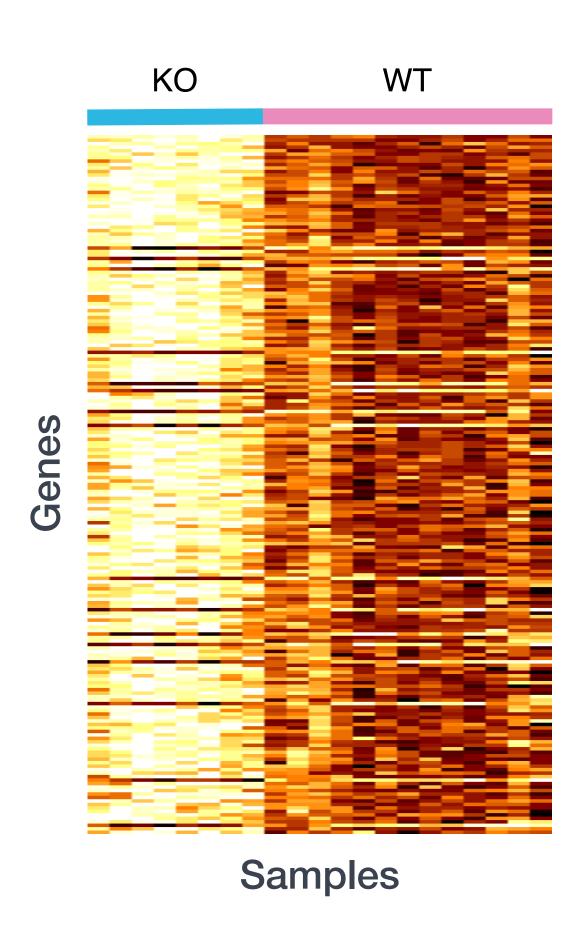
### Identifying differences in gene expression

- Looking for genes that change in expression between two or more groups
  - case vs. control
  - correlation of expression with some variable or clinical outcome

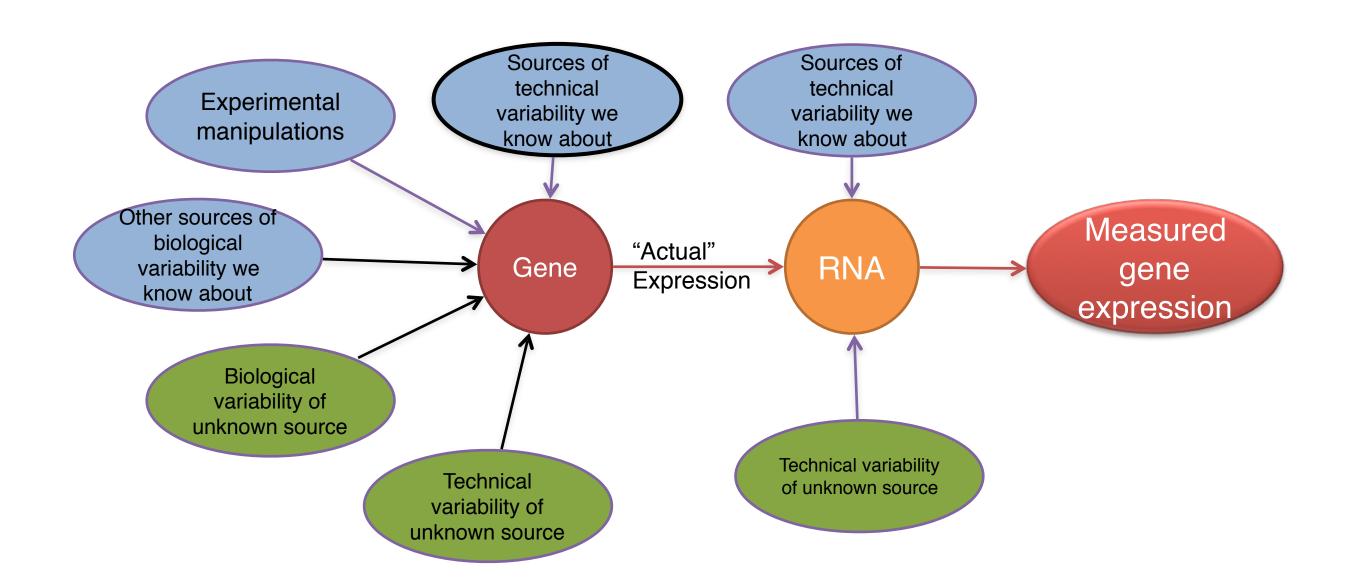


### Identifying differences in gene expression

- Looking for genes that change in expression between two or more groups
  - case vs. control
  - correlation of expression with some variable or clinical outcome
- ▶ Rank the genes by how different they are between the two groups (based on fold change values. Why does this not work?



#### The measurement is the "sum" of many effects

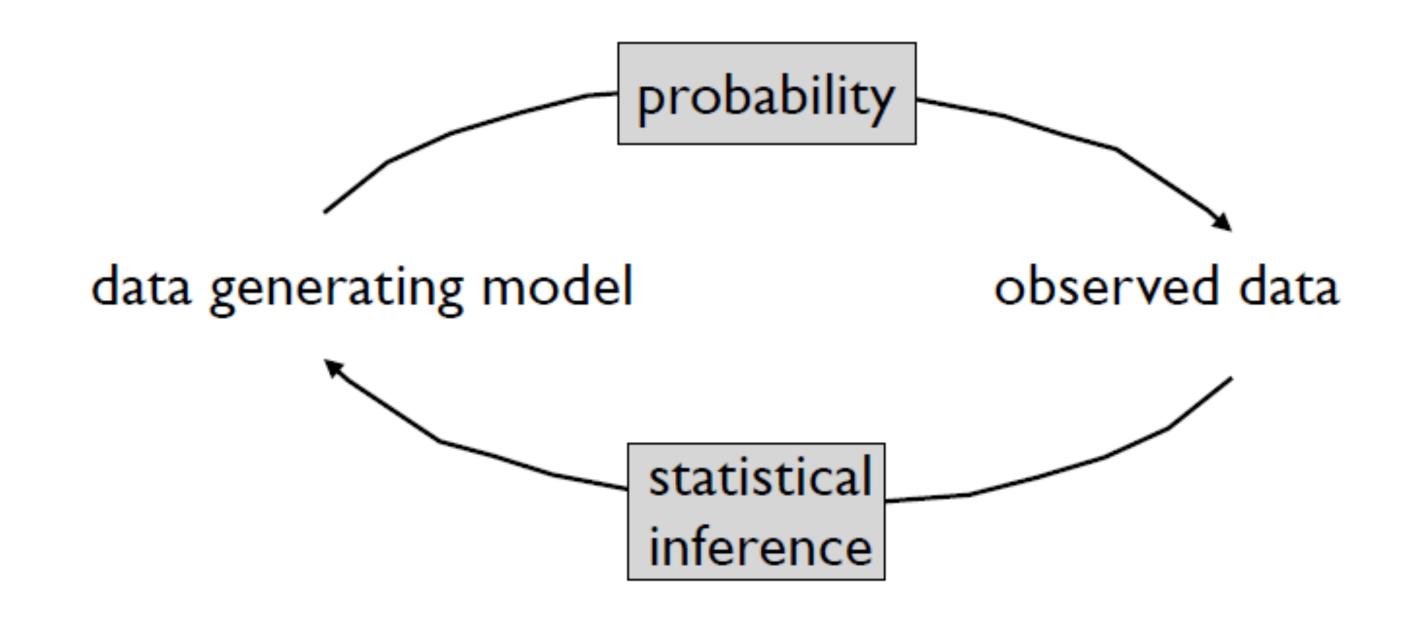


## Modeling gene-level data

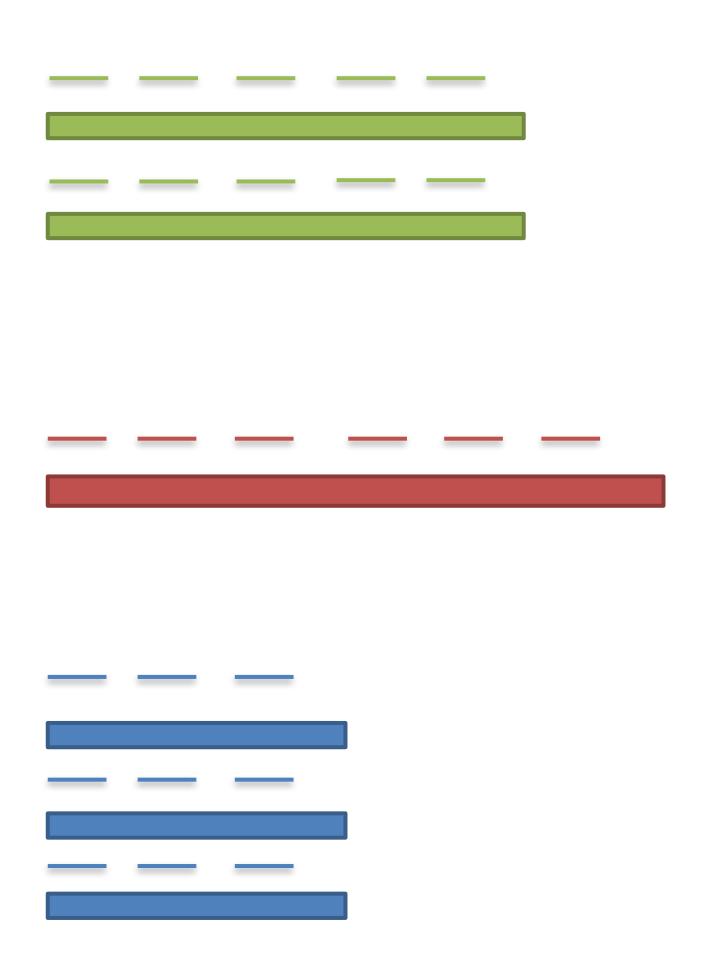
Courtesy of Paul Pavlidis, UBC

# Making sense of data

Data is what we observe. We want to infer something about "where it came from"

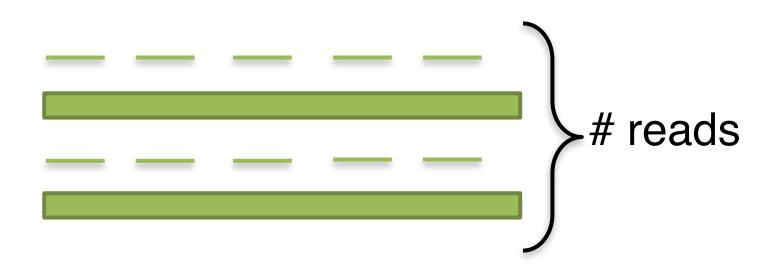


#### Variance of counts

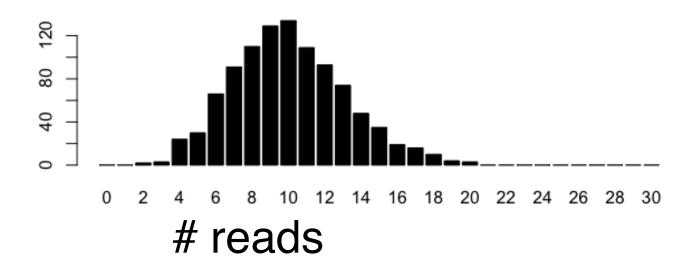


#### Variance of counts

#### Consider one gene:

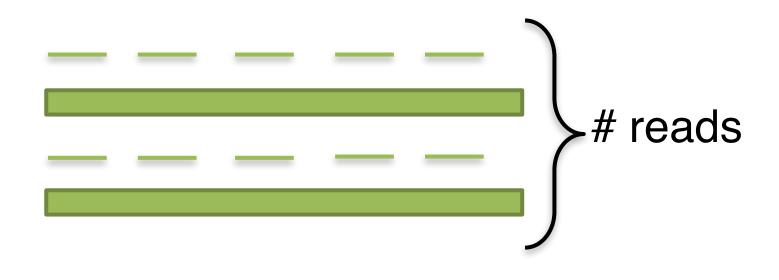


Binomial sampling distribution

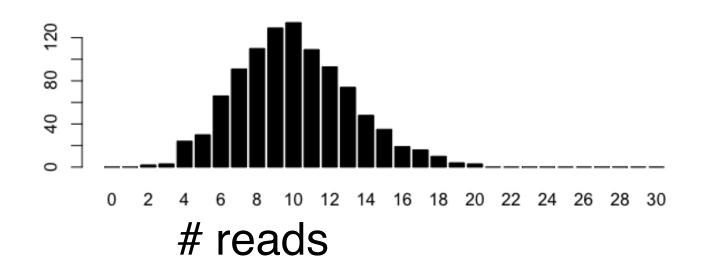


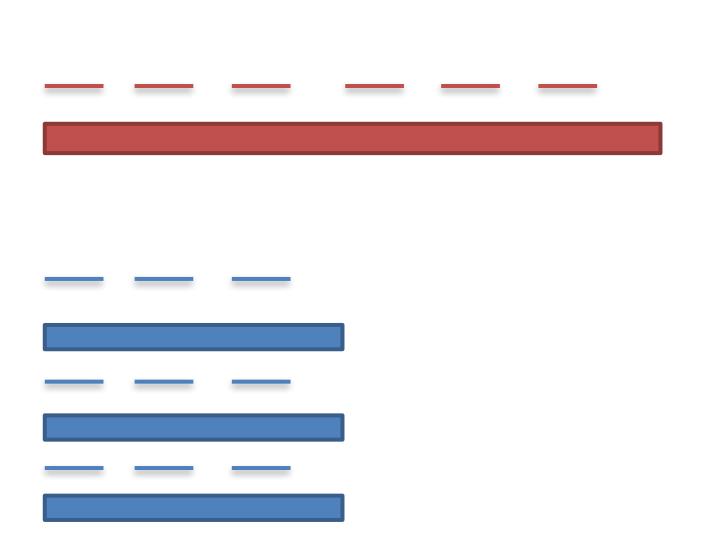
#### Variance of counts

#### Consider one gene:

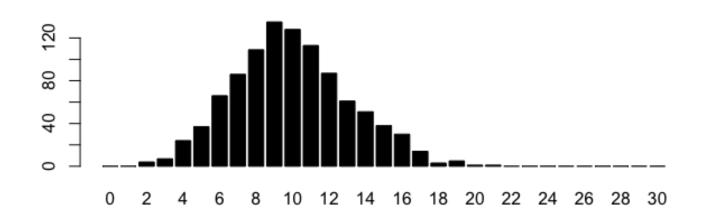


Binomial sampling distribution





 With millions of reads & small proportion for each gene => Poisson sampling distribution



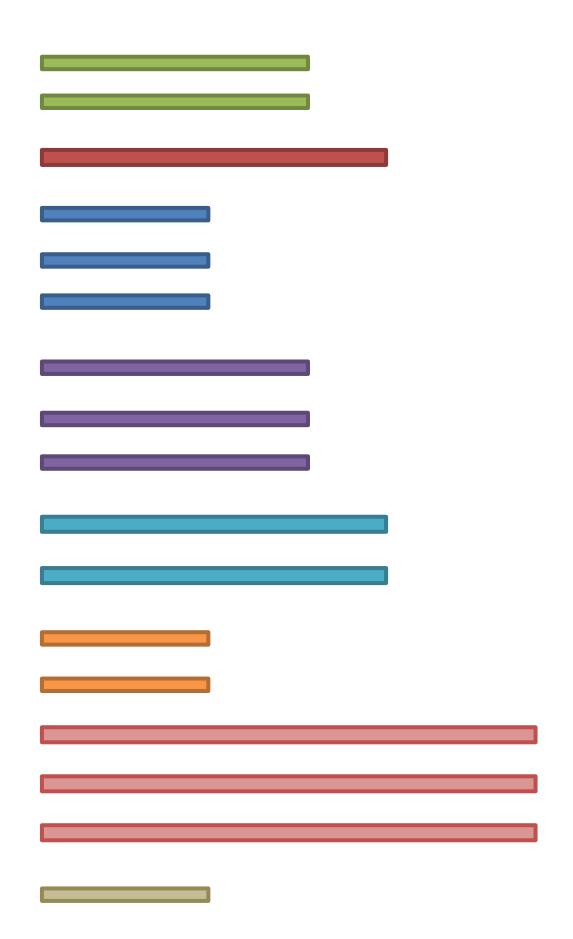
# Technical vs biological replicates: modeling dispersion

#### Biological vs. Technical Replicates

- Biological replicates contain multiple indviduals
  - > Technical replicates contain one individual with some technical steps replicated
  - Usually **biological variance > technical variance**, thus biological replicates are more useful. They also allow us to make inferences about treatment groups

# Biological replicates

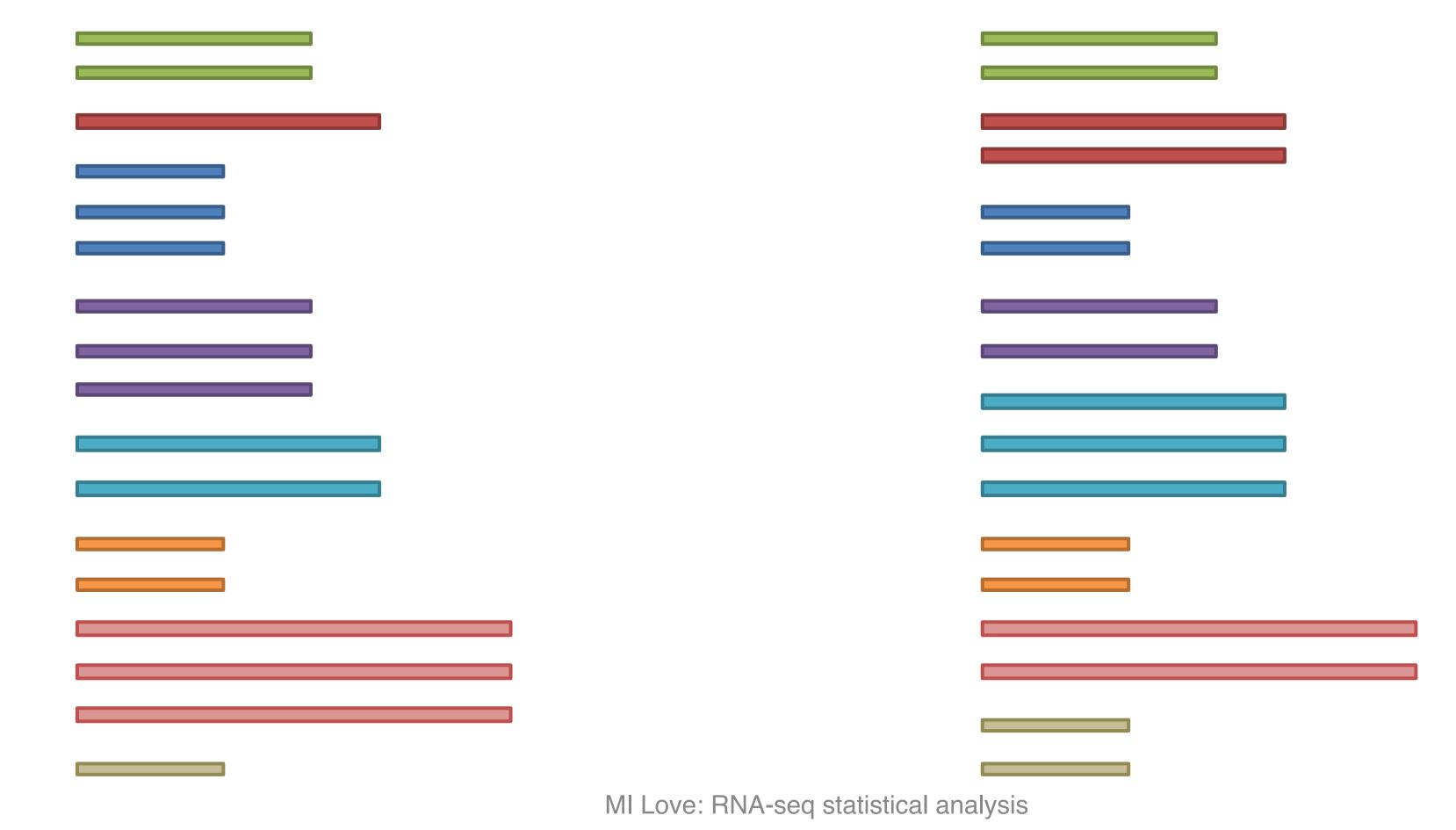
If the proportions of mRNA stays exactly constant ("technical replicate") we can expect Poisson dist.



# Biological replicates

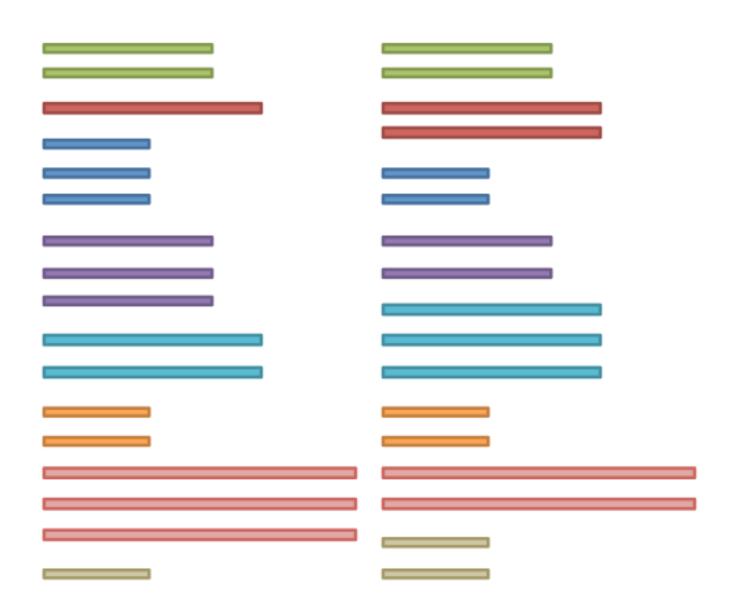
If the proportions of mRNA stays exactly constant ("technical replicate") we can expect Poisson dist.

But realistically, biological variation across sample units is expected



# Biological replicates

Biological variation for the abundance of a given gene produces "over-dispersion" relative to the Poisson dist.



### Biological replicates

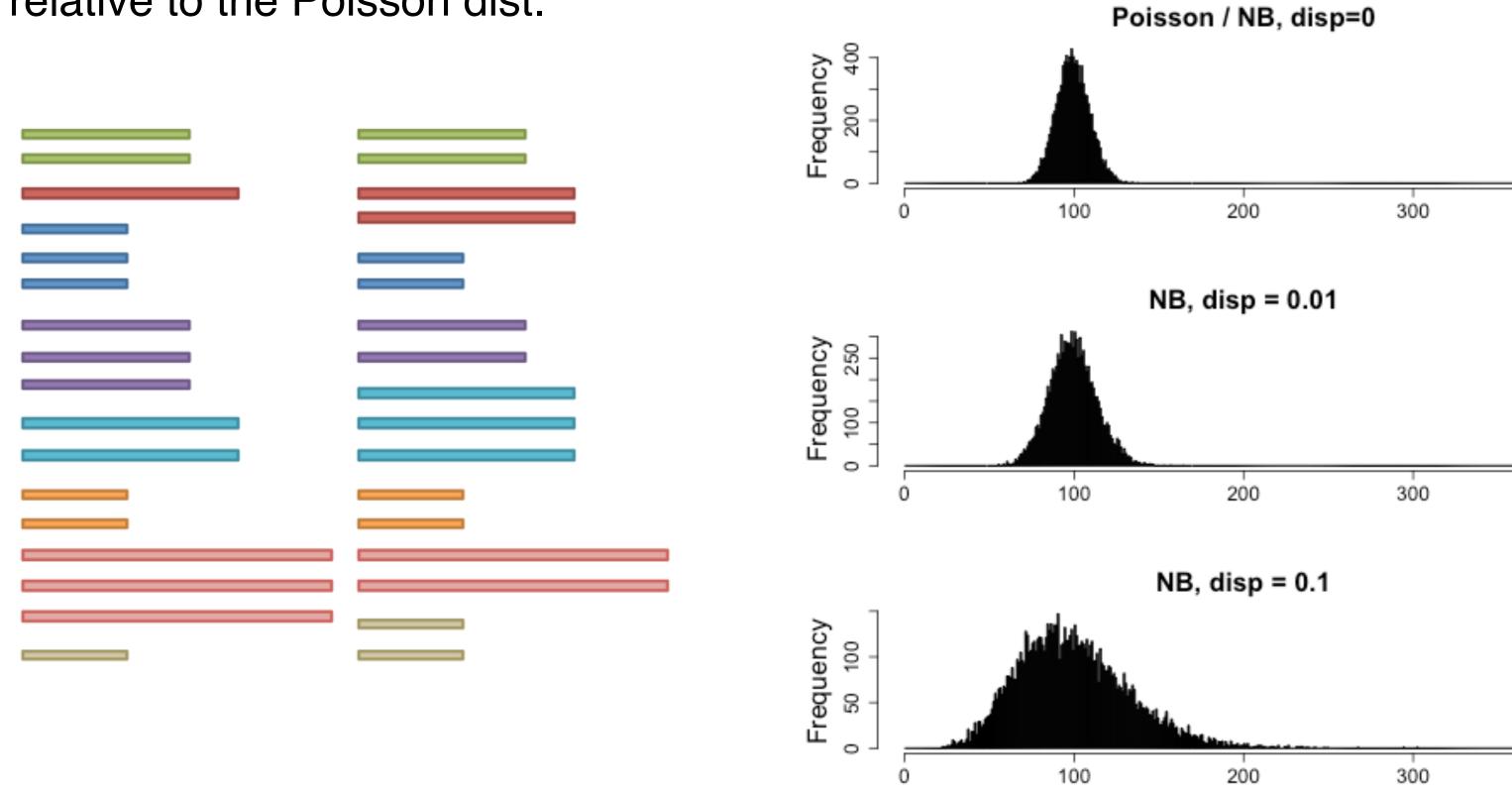
Biological variation for the abundance of a given gene produces "over-dispersion" relative to the Poisson dist.

Negative Binomial = Poisson with a varying mean

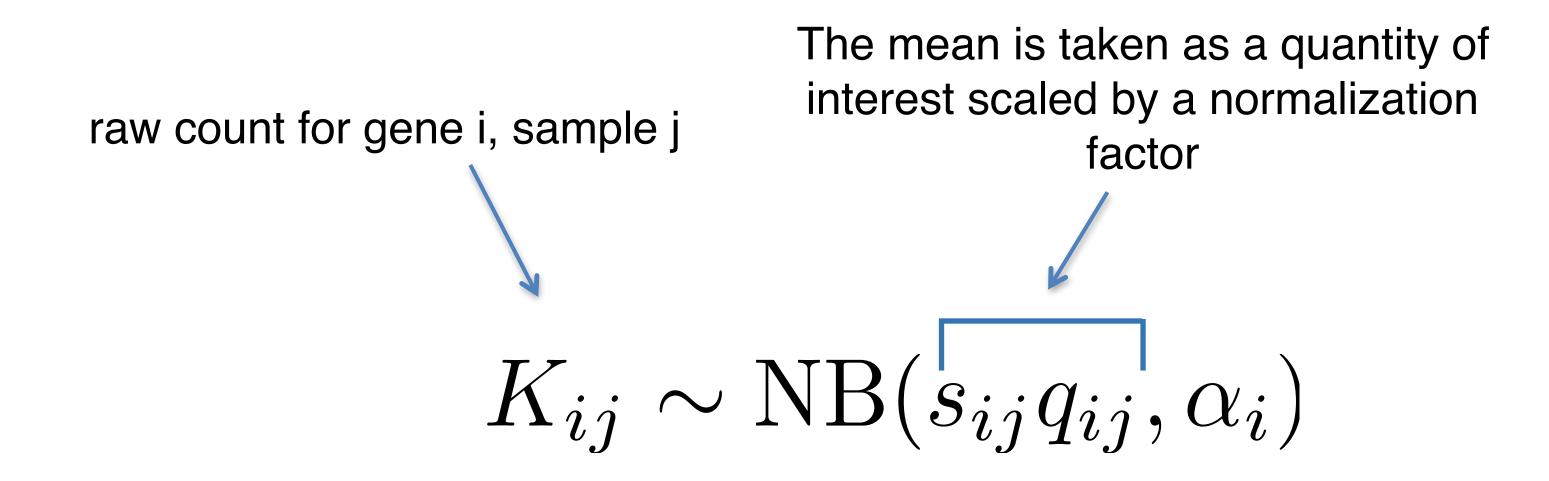
400

400

400

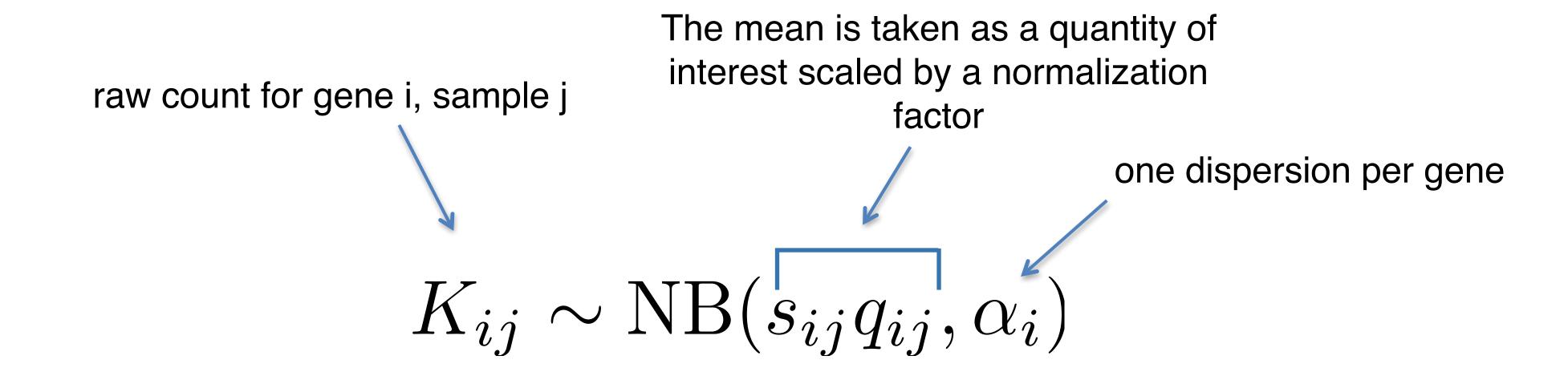


# Counts are modeled using the Negative Binomial (NB) distribution



We model read counts  $K_{ij}$ , as following the negative binomial distribution with mean  $\mu_i$  and **dispersion**  $\alpha_i$ 

# Counts are modeled using the Negative Binomial (NB) distribution



We model read counts  $K_{ij}$ , as following the negative binomial distribution with mean  $\mu_i$  and **dispersion**  $\alpha_i$ 

# Within group variability is accounted for using the dispersion parameter

The dispersion parameter  $\alpha_{i,}$  describes the variance of counts via:

Extra variation

due to biological variance

$$Var(K_{ij}) = \mu_{ij} + \alpha_i \mu_{ij}^2$$

sampling fragments

Poisson part:

MI Love: RNA-seq statistical analysis

## Dispersion is an approximation of the coefficient of variation

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$$\nabla^{2}$$

$$Var(K_{ij}) = \mu_{ij} + \alpha_{i}\mu_{ij}^{2}$$

for large counts: 
$$\sqrt{\alpha_i} pprox \frac{\sigma}{\mu} \equiv CV$$
 (coefficient of variation)

### Dispersion is an approximation of the coefficient of variation

$$Var(K_{ij}) = \mu_{ij} + \alpha_i \mu_{ij}^2$$

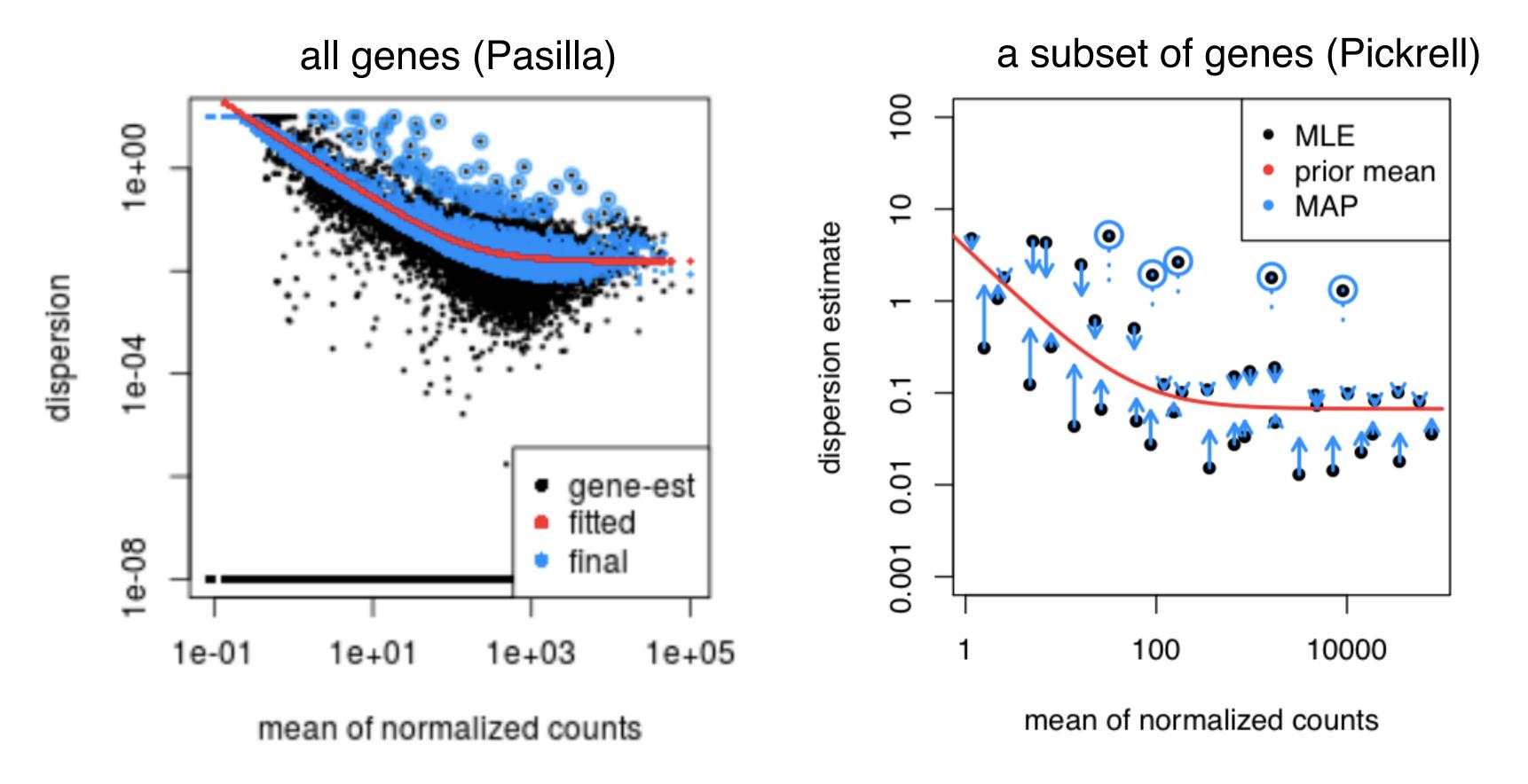
for large counts: 
$$\sqrt{\alpha_i} pprox \frac{\sigma}{\mu} \equiv CV$$
 (coefficient of variation)

alpha = 
$$0.01 => CV 10\%$$
  
alpha =  $0.25 => CV 50\%$ 

### Shrinkage and dispersion

- Different genes naturally have different scale of biological variability
- Over all genes, there will be a distribution of reasonable estimates of dispersion
- With small sample size (n=3-5 replicates per group), we will make very bad estimates of gene-wise dispersion unless we share information across genes

#### Shrinkage of dispersion



- 1. Gene-wise estimate = just look at one gene (MLE)
- 2. Fitted dispersion trend = the middle for the prior
- 3. Final estimate = posterior, uses shared information (MAP)

#### Differences across conditions

#### DESeq2

- Available through Bioconductor since 2013
- Publication: Genome Biology, Dec 2014.
   main text written with non-statisticians in mind
- Builds on good ideas for dispersion estimation and use of GLM from the DSS and edgeR methods
- See <u>bioconductor.org/install</u> for installation
- Note that the latest Bioconductor packages are only available with latest R version. Bioconductor and R versions are linked

#### DESeq2 steps

we tried to streamline the default pipeline.

count matrix (from featureCounts, htseq, tximport, etc.)

- 1. size factors (sequencing depth)
- 2. dispersion (biological variance)
- 3. Wald test or likelihood ratio test
- 4. build results table

#### Differences across two conditions

- Describe experiment with formula, e.g.: ~ condition
- Per gene the design matrix looks like:

```
\log_2 q1 1 0

\log_2 q2 1 0

\log_2 q3 = 1 1

\log_2 q4 1 1
```

#### Differences across two conditions

- Describe experiment with formula, e.g.: ~ condition
- Per gene the design matrix looks like:

$$log_{2} q1$$
 $log_{2} q2$ 
 $log_{2} q3$ 
 $log_{2} q4$ 
 $log_{2} q4$ 

All samples get an Intercept term

#### Differences across two conditions

- Describe experiment with formula, e.g.: ~ condition
- Per gene the design matrix looks like:

All samples get an Intercept term

The B condition samples also get a term that accounts for the difference between B and A

```
\log_2 q1 1 1 0 0 \log_2 q2 1 1 0 0 Intercept \log_2 q3 = 1 0 1 0 conditionA \log_2 q4 1 0 1 0 conditionB \log_2 q5 1 0 0 1 conditionC \log_2 q6 1 0 0 1
```

log <sub>2</sub> q1		1	1	0	0	
log <sub>2</sub> q2		1	1	0	0	Intercept
log <sub>2</sub> q3	=	1	0	1		conditionA
log <sub>2</sub> q4		1	0	1	0	conditionB
log <sub>2</sub> q5		1	0	0	1	conditionC
log <sub>2</sub> q6		1	0	0	1	

log <sub>2</sub> q1		1	1	0	0	
log <sub>2</sub> q2		1	1	0	0	Intercept
log <sub>2</sub> q3	=	1	0	1	0	conditionA
log <sub>2</sub> q4		1	0	1	0	conditionB
log <sub>2</sub> q5		1	0	0	1	conditionC
log <sub>2</sub> q6		1	0	0	1	

log <sub>2</sub> q1		1	1	0	0	
log <sub>2</sub> q2		1	1	0	0	Intercept
log <sub>2</sub> q3	=	1	0	1	0	conditionA
log <sub>2</sub> q4		1	0	1	0	conditionB
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log <sub>2</sub> q6		1	0	0	1	

```
\log_2 q1 1 1 0 0 \log_2 q2 1 1 0 0 Intercept \log_2 q3 = 1 0 1 0 conditionA \log_2 q4 1 0 1 0 conditionB \log_2 q5 1 0 0 1 conditionC \log_2 q6 1 0 0 1
```

#### GLM: Generalized linear models

- Extension of linear models to non-normally distributed response data (in our case, negative binomial)
- Helps address the different mean-variance relationships
- GLM fit for a gene will return coefficients indicating the overall expression strength of the gene for each design matrix element

log <sub>2</sub> q1		1	1	0	0	
log <sub>2</sub> q2		1	1	0	0	Intercept
log <sub>2</sub> q3	=	1	0	1	0	conditionA
log <sub>2</sub> q4		1	0	1	0	conditionB
log <sub>2</sub> q5		1	0	0	1	conditionC
log <sub>2</sub> q6		1	0	0	1	

log <sub>2</sub> q1		1	1	0	0	
log <sub>2</sub> q2		1	1	0	0	Intercept
$log_2 q3$	=	1	0	1	0	conditionA
$log_2 q4$		1	0	1	0	conditionB
$log_2 q5$		1	0	0	1	conditionC
log <sub>2</sub> q6		1	0	0	1	

log <sub>2</sub> q1		1	1	0	0	
log <sub>2</sub> q2		1	1	0	0	Intercept
log <sub>2</sub> q3	=	1	0	1	0	conditionA
log <sub>2</sub> q4		1	0	1	0	conditionB
log <sub>2</sub> q5		1	0	0	1	conditionC
log <sub>2</sub> q6		1	0	0	1	

#### Hypothesis testing

#### Wald test

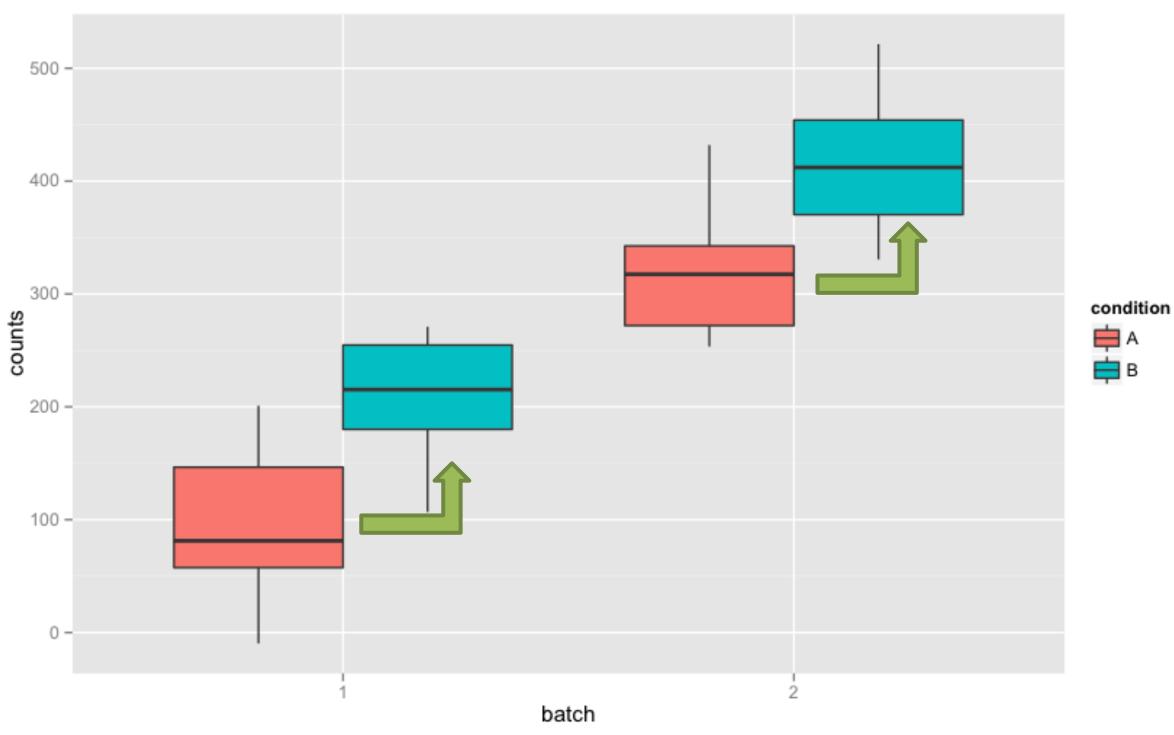
- use the shrunken estimate of the log2 fold-change divided by the SE
- this gives a Z-statistic which is compared to a standard normal distribution
- allows testing of individual coefficients, or contrasts of coefficients (i.e. two-level comparison)

#### Likelihood ratio test (LRT)

- examines two different models: full and reduced model (with some terms removed)
- determines if the increased likelihood of the data using the extra terms in the full model is more than expected if those extra terms are truly zero
- useful for identifying any gene that is changing in expression with respect to the biological factor of interest (useful for 3 or more levels)

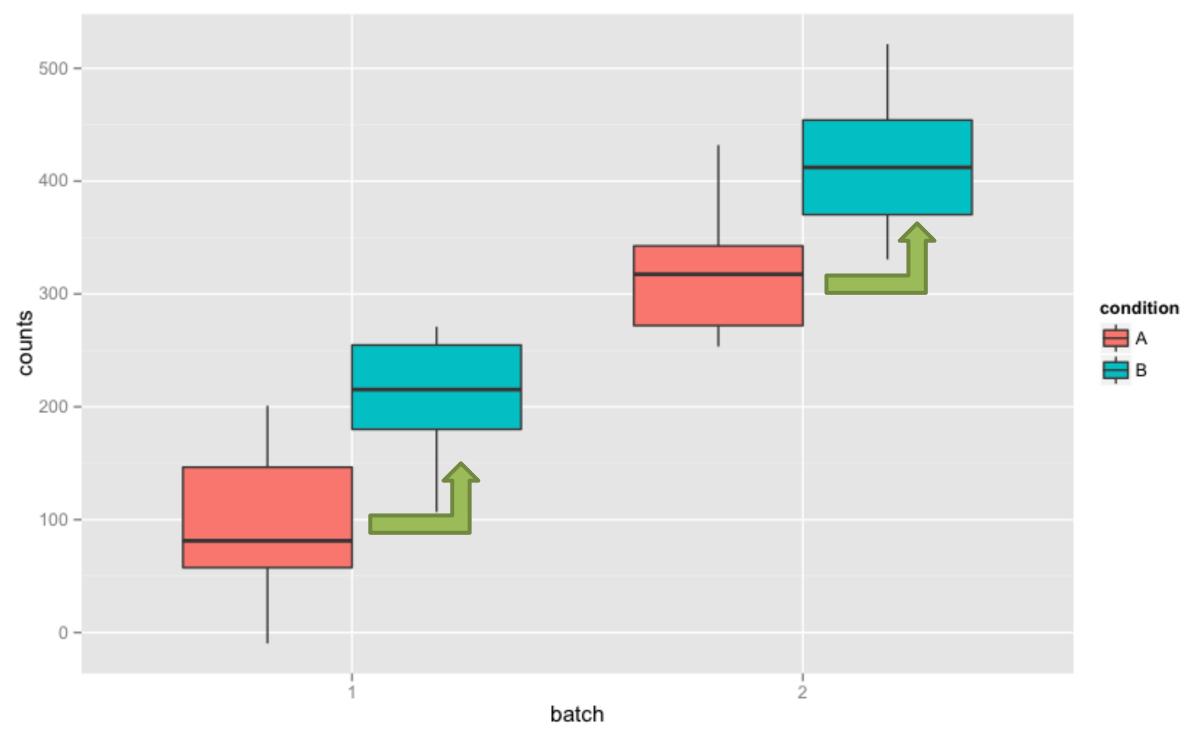
 Using a design formula: ~ batch + condition, adds terms that control for batch differences

Using a design formula: ~ batch + condition,
 adds terms that control for batch differences



MI Love: RNA-seq statistical analysis

- Using a design formula: ~ batch + condition, adds terms that control for batch differences
- If batches are unknown, possible to detect these with other methods: svaseq, RUVSeq



MI Love: RNA-seq statistical analysis

### Complex designs

Want to test: treatment changes for enriched samples over baseline, controlling for individual effects

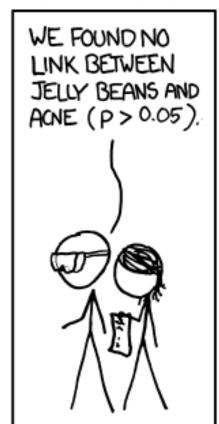
~individual + enrichment + treatment + enrichment:treatment

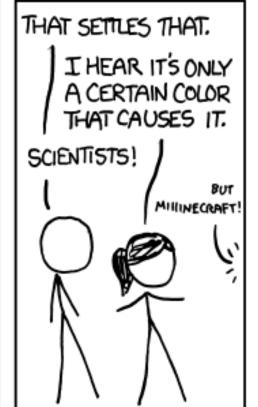
indiv	. en	rich. treat.
1	input	control
1	IP	control
1	input	treat
1	IP	treat
2	input	control
2	IP	control
2	input	treat
2	IP	treat

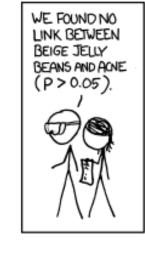
. . .

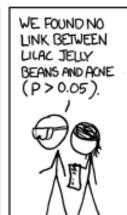
### Multiple test correction

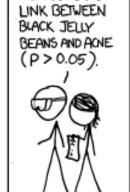








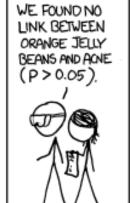


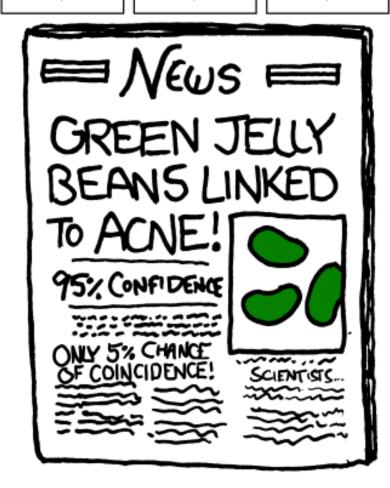


WE FOUND NO



WE FOUND NO LINK BETWEEN





#### Multiple test correction

#### Multiple test correction

▶ In our example using condition A vs. condition B we found 2,898 genes to be differentially expressed at p < 0.05

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- This is 12% of the genes we tested

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- ▶ In our example using condition A vs. condition B we found 2,898 genes to be differentially expressed at p < 0.05
- This is 12% of the genes we tested
- Expect 5% by chance (this is what p < 0.05)
- Probably ~1/3 genes we found are false positives

This is the multiple testing problem. The more tests we perform, the more we inflate the number of false positives observed.

#### Controlling the FWER

- Control  $\alpha$ , the probability of making an error (false positive)
- **Bonferroni:** Reject any hypothesis with p-value  $\leq \alpha/m$ 
  - Conservative; high probability of false negatives

#### Controlling the FDR

- FDR: false discovery rate: the expected percent of false predictions in the set of predictions
- ► Benjamini-Hochberg: Rank j / m multiplied by the FDR level
  - designed to control the FDR
- Q-value: The minimum FDR that can be attained when calling that feature significant

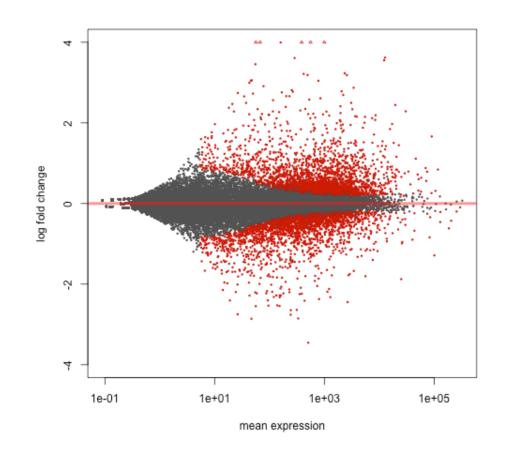
## DE vs EDA

# Two paths in DESeq2



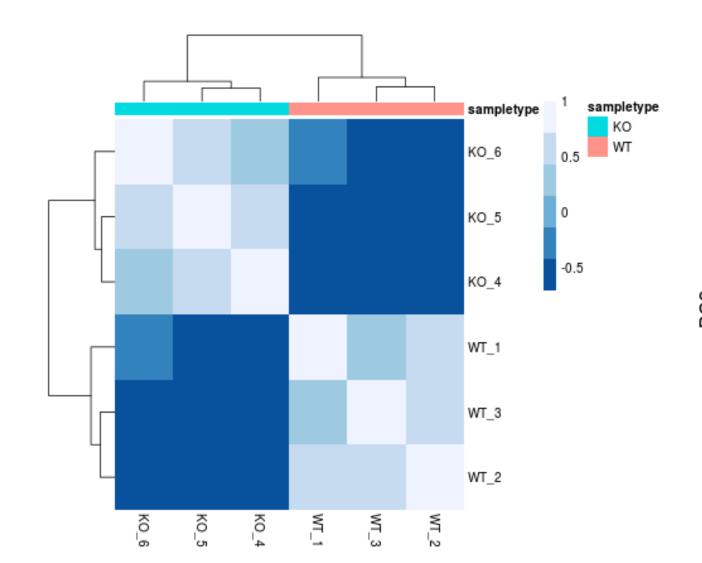
**Differential expression** 

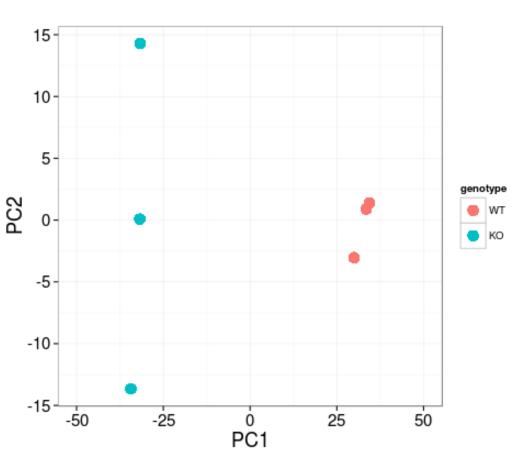
testing, p-values, FDR



<u>Transformations and</u> <u>Exploratory Data Analysis (EDA)</u>

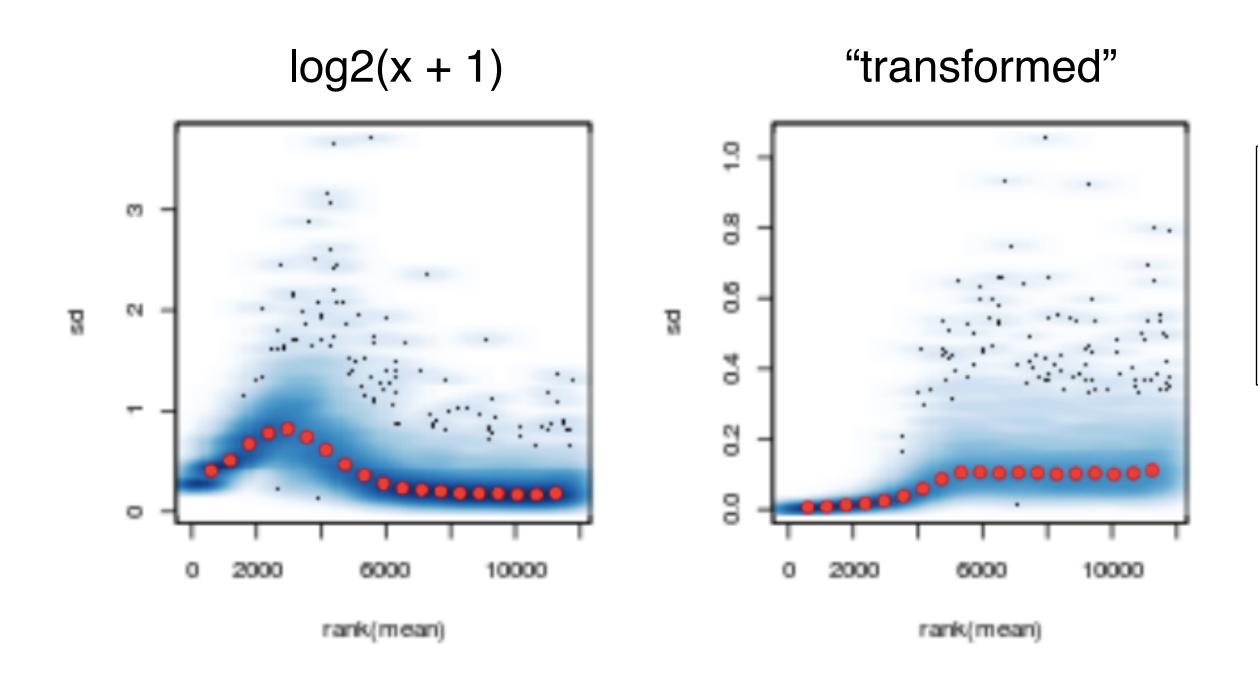
clustering, heatmaps, sample-sample distances





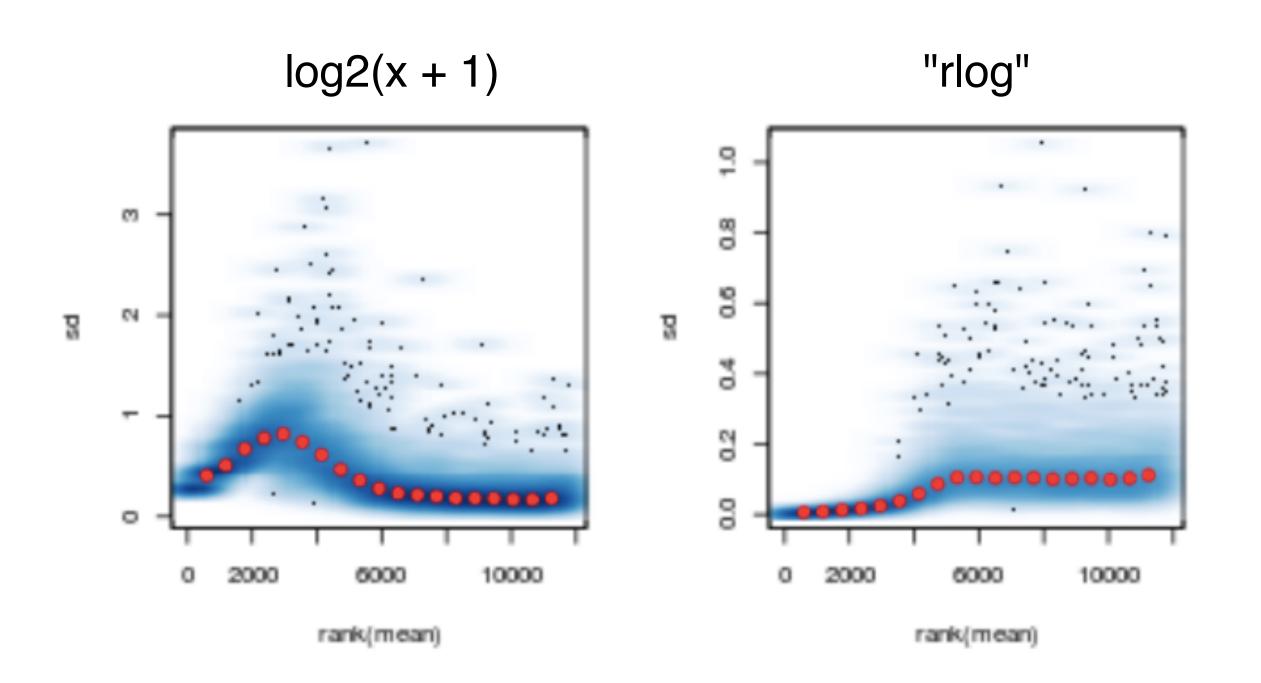
#### Transformations

For comparison analyses when using unsupervised techniques, it can be useful to *transform* data. These techniques (VST and log) perform better when values have a similar dynamic range

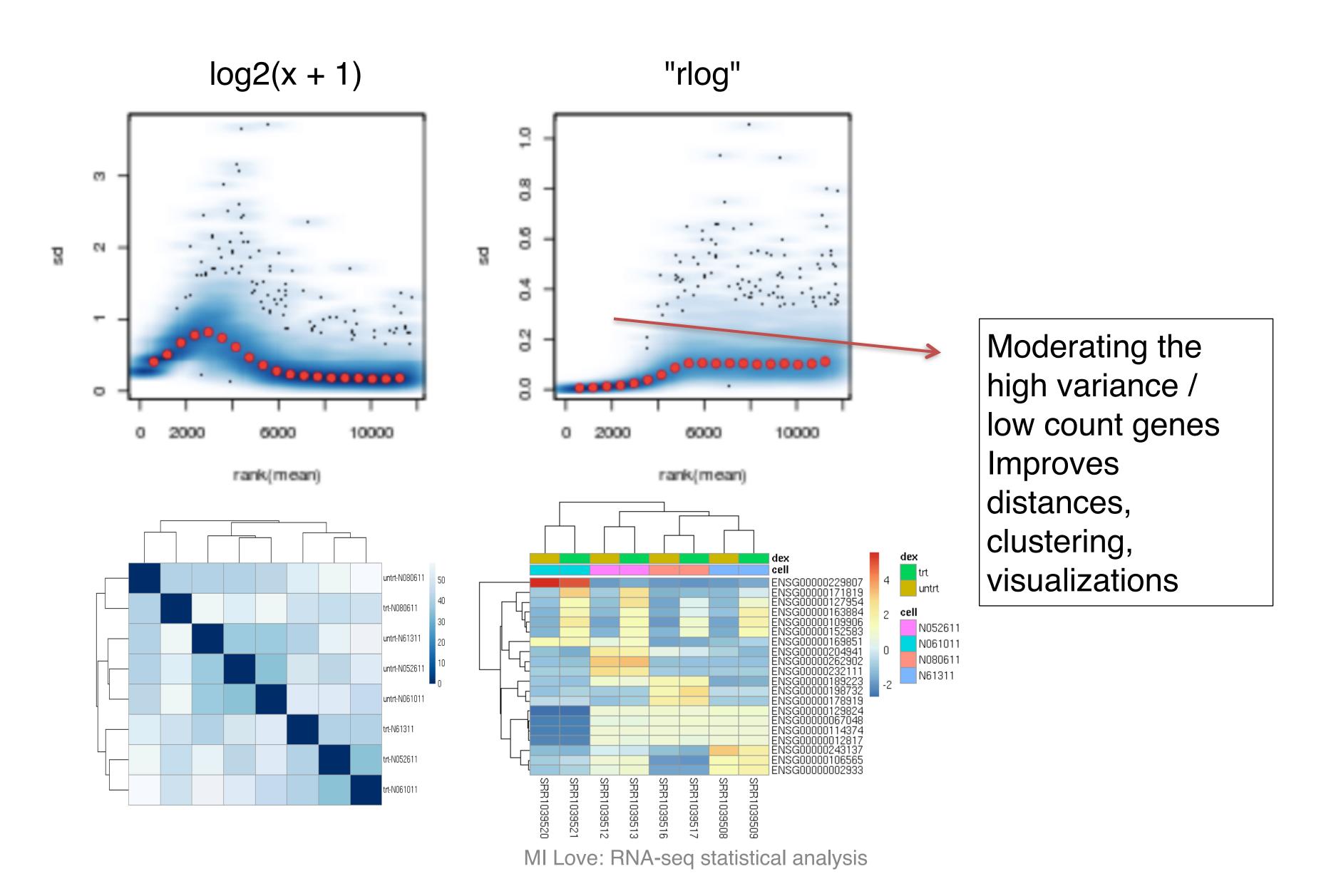


renders data *homoskedastic* (variance of the gene is stabilized across expression levels

#### rlog stabilizes variances along the mean



#### rlog stabilizes variances along the mean



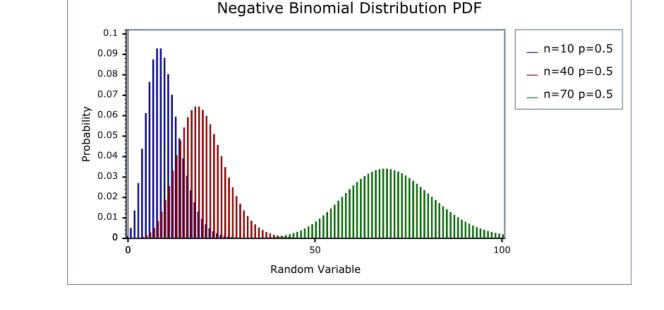
# Variance stabilizing transform (VST)

- The variance stabilizing transformation is an earlier approach (from original DESeq) for transforming counts.
- Uses a function which is  $\log$ -like but doesn't go to  $-\ln t$  at x=0.
- VST doesn't use size factors so is better for data that has consistent sequence depth across samples
- VST is closed form, so can be better for large datasets (n > 50) due to speed.

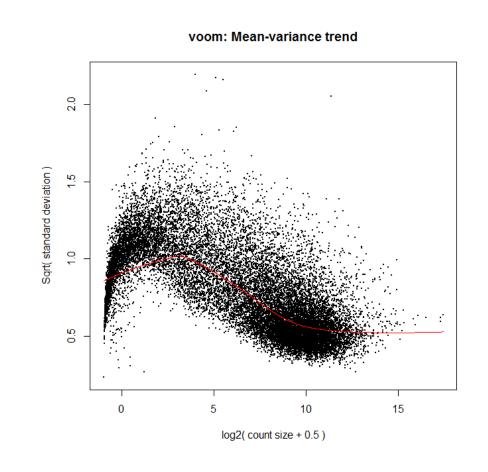
# Comparison of methods

## Count model vs linear model

- DESeq2 and edgeR similar approach, similar results
  - very sensitive, may sometimes underestimate FDR



- limma+voom uses a linear model,
   weights determined by variance over mean
  - strong control of FDR, may be less sensitive for small sample size
  - recommended when number of biological replicates
     per group grows large (e.g. > 20), because CLT kicks in



#### Credits

#### RNA-seq statistical analysis and gene-level differential expression

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Acknowledgements:

DESeq2/DEXSeq:

Work supported by:

Wolfgang Huber

Simon Anders

Alejandro Reyes

Rafael Irizarry

Dana-Farber Cancer

Institute

Harvard TH Chan School

of Public Health

VE RI TASI

NIH Training Grant