

$$X = \sum_{i=1}^n X_i$$

(1)

$$X_i = \begin{cases} 1 & \text{اندر شرایط: } \varepsilon \text{ یا } \mu \\ 0 & \text{وگرنه} \end{cases}$$

$$\mu \cdot E[X_i] = 1/4$$

$$\sigma^2 = \text{Var}(X_i) = p_h = 3/14$$

طبق قضیه حد مرکزی

$$\frac{\sum_{i=1}^n X_i - n\mu}{\sigma\sqrt{n}} \sim \text{Normal}(0,1)$$

اندر  $\alpha$  درصد با اطمینان

$$1 - \alpha = 0.90$$

$$P(-z_{\alpha/2} < \frac{X - n\mu}{\sigma\sqrt{n}} < z_{\alpha/2}) = 1 - \alpha$$

$$P(n\mu - z_{\alpha/2}\sigma\sqrt{n} < X < n\mu + z_{\alpha/2}\sigma\sqrt{n}) = 0.90$$

برای  $n=100$

$$P\left(\frac{n\mu}{4} - 1.96 \times \frac{\sqrt{3}}{4} \times 10 < X < \frac{n\mu}{4} + 1.96 \times \frac{\sqrt{3}}{4} \times 10\right) = 0.90$$

$$\rightarrow [97, 99.6] \quad X \sim \text{احتمال } 90\%$$

$$X_i \sim \text{unif}(0, \theta) \rightarrow f_{X_i}(x) = \begin{cases} 1/\theta & 0 < x < \theta \\ 0 & \text{o.w} \end{cases}$$

الف

$$L(x_1, x_2, x_3, \dots, x_n; \theta) = p(x_1, x_2, \dots, x_n; \theta) = p(x_1 | \theta) \dots p(x_n | \theta) \\ = \left(\frac{1}{\theta}\right)^n$$

برای اینکه  $\left(\frac{1}{\theta}\right)^n$  بیشینه شود باید  $\theta$  تا حد امکان کوچک شود  
اما از آنجایی که  $x_i$  نمی تواند از  $\theta$  بزرگ تر باشد

$$\max(x_1, \dots, x_n) \leq \theta$$

$$\max(x_1, \dots, x_n) = \text{MLE}_{\theta}$$

$$Y = \max(x_1, x_2, \dots, x_n)$$

$$F_Y(a) = p(Y \leq a) = p(x_1 \leq a) p(x_2 \leq a) \dots p(x_n \leq a)$$

$$= F_{X_1}(a) F_{X_2}(a) \dots F_{X_n}(a)$$

$$= F_X^n(a) = \left(\frac{a}{\theta}\right)^n \rightarrow \frac{d}{dy} F_Y(a) = \frac{n a^{n-1}}{\theta^n}$$

$$E[Y] = E[\text{MLE}_{\theta}] = \int_0^{\theta} y \frac{n y^{n-1}}{\theta^n} dy \rightarrow f_Y(y) = \frac{n y^{n-1}}{\theta^n}$$

$$= \int_0^{\theta} n \frac{y^n}{\theta^n} dy = \frac{y^{n+1}}{n+1} \times \frac{n}{\theta^n} \Big|_0^{\theta} = \frac{n}{n+1} \theta$$

$$E[\text{MLE}_{\theta}] - \theta = \frac{-1}{n+1} \theta \rightarrow \text{biased}$$

ب

$$\bar{X} = \frac{\left( \sum_{i=1}^n x_i \right)^2}{n^2}$$

$$E[\bar{X}] = \frac{1}{n^2} \left( \sum_{i=1}^n E[x_i^2] + \sum_{\substack{i \neq j \\ 1 \leq i, j \leq n}} E[x_i x_j] \right)$$

جس میں  $x_i, x_j$  مستقل ہیں  
 $E(x_i)E(x_j)$

$$= \frac{1}{n^2} \left( \sum_{i=1}^n E[x_i^2] + \sum_{i \neq j} E[x_i]E[x_j] \right)$$

$$= \frac{1}{n^2} \left( \sum_{i=1}^n \text{Var}(x_i) + \mu^2 + \sum_{i \neq j} \mu^2 \right)$$

$$= \frac{1}{n^2} \left( n\sigma^2 + n\mu^2 + 1\left(\frac{n}{2}\right)\mu^2 \right) = \frac{\sigma^2}{n} + \frac{\mu^2}{n} + \frac{n-1}{n} \mu^2$$

$$= \frac{\sigma^2}{n} + \mu^2$$

$\bar{X}$  unbiased estimator of  $\mu$  is not unbiased.  $\bar{X}$  is biased estimator of  $\mu$ .

$$\rightarrow Y = \bar{X} - \frac{\sigma^2}{n} = \left( \frac{\sum_{i=1}^n x_i^2}{n} \right) - \frac{\sigma^2}{n}$$

$$E[Y] - \mu^2 = 0$$

$\leftarrow$   $Y$  is unbiased estimator of  $\mu^2$ .

(۴) انت  $X$  از روش کار، آمار است و داریم:

$$L(X|n) = p(X|n) = \frac{\binom{1^v}{2} \binom{n-3}{4}}{\binom{n}{6}} = \frac{1^v \times \frac{(n-4)}{2!}}{\frac{n \times (n-1) \times (n-2)}{6!}}$$

$$= \frac{q_0}{2!} \times \frac{(n-4)}{n(n-1)(n-2)} = \frac{q_0 (n-4)}{n(n-1)(n-2)}$$

استیمن این باید صورت گیرد

$$\frac{d}{dn} \left( \frac{q_0 (n-4)}{n(n-1)(n-2)} \right) = 0$$

$$\frac{dA}{dn} = \frac{n(n-1)(n-2) - (n-4)((n-1)(n-2) + n(n-2) + n(n-1))}{n^2(n-1)^2(n-2)^2} = 0$$

$$\rightarrow n^3 - 3n^2 + 2n = 3n^3 - 7n^2 + 5n - 4$$

$$\rightarrow 2n^3 - 10n^2 + 3n - 4 = 0$$

$$n_1 = 0.45, n_2 = 1.45, n_3 = 1.45$$

و بی در نظر می‌اندازیم، می‌دانیم که ۲ و ۳ از روش کار، آمار است و اینها - بزرگی بوده اند

$$\rightarrow n \approx 1.45$$

$$\rightarrow n = 1.45$$

$$L(X|1) = \frac{q_0 \times 2}{1 \times 1 \times 1} = 10\%$$

$$L(X|9) = \frac{q_0 \times 2}{9 \times 8 \times 7} = 1\%$$

$$n=1 \leq n=9 \leftarrow \text{بسیار کم}$$

$$n=9 \leq n=1$$

که همین خوب نیست - چون آمار بسیار کم است

$$E[\hat{\eta}] = \eta$$

مطلوب: unbiased  $\hat{\eta}$   $\hat{\eta} = \eta$  (د)

$$\begin{aligned} \rightarrow E[\hat{\eta}] = \eta &= \alpha_1 E[x_1] + \dots + \alpha_n E[x_n] \\ &= \alpha_1 (E[\eta] + E[\epsilon_1]) + \dots + \alpha_n (E[\eta] + E[\epsilon_n]) \\ &= \alpha_1 \eta + \alpha_1 \epsilon_1 + \dots + \alpha_n \eta \\ &= \eta (\alpha_1 + \alpha_2 + \dots + \alpha_n) \end{aligned}$$

$$\begin{aligned} \rightarrow \alpha_1 + \alpha_2 + \dots + \alpha_n &= 1 \\ \text{لدينا الحد الأدنى لـ } \min \text{Var}(\hat{\eta}) & \\ \text{نريد الحد الأدنى لـ } \text{Var}(\hat{\eta}) &= \sum_{i=1}^n \alpha_i^2 \text{Var}(x_i) \\ &= \sum_{i=1}^n \alpha_i^2 (\text{Var}(\eta) + \text{Var}(\epsilon_i)) \\ &= \sum_{i=1}^n \alpha_i^2 \sigma_i^2 \end{aligned}$$

$$\rightarrow F(\alpha) = \sum_{i=1}^n \alpha_i^2 \sigma_i^2 + \lambda \left( \sum_{i=1}^n \alpha_i - 1 \right)$$

$$\rightarrow \frac{\partial F}{\partial \lambda} = \sum_{i=1}^n \alpha_i - 1 = 0$$

$$\frac{\partial F}{\partial \alpha_i} = 2\alpha_i \sigma_i^2 + \lambda = 0 \rightarrow \alpha_i = \frac{-\lambda}{2\sigma_i^2}$$

$$\rightarrow \left[ \frac{-\lambda}{2} \sum_{i=1}^n \frac{1}{\sigma_i^2} = 1 \rightarrow \lambda = \frac{-2}{\sum_{i=1}^n \frac{1}{\sigma_i^2}} \right]$$

$$\alpha_i = \frac{1}{\sigma_i^2 \left( \sum_{i=1}^n \frac{1}{\sigma_i^2} \right)}$$