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QMB 6911 Capstone Project in Business Analytics

Solutions: Problem Set #9

1 Data Description

This analysis follows the script <code>Trucks_Reg_Model.R</code> to produce a more accurate model for used trucks cost with the data from <code>UsedTrucks.dat</code> in the <code>Data</code> folder. The dataset includes the following variables. From

 $type_i$ = sale type

pauc = price when sold at auction
pret = price when sold retail

mileage = odometer

make = make of vehicule
year = model year of vehicle

damage = an index of damage to vehicle, 1 little damage, 10 a lot

dealer = dealer id
ror = rate-of-return

cost = net amount given to trade in

problem set 7, we build a model that can predict the cost of the vehicule depending on the damage of the trucks. we will add quadratic specification for damage which will increase the relationship between cost and damage.

	Model 1			
(Intercept)	8.76488***			
	(0.08135)			
squared_damage	0.00459^*			
	(0.00214)			
make	-0.00619			
	(0.00537)			
damage	-0.04677			
	(0.02543)			
dealer	0.00196			
	(0.00417)			
mileage	-0.00001^{***}			
	(0.00000)			
age	-0.07163***			
	(0.01114)			
type	0.04438^*			
	(0.02182)			
\mathbb{R}^2	0.19087			
Adj. R ²	0.19029			
Num. obs.	9861			
*** $p < 0.001$; ** $p < 0.01$; * $p < 0.05$				

Tab. 1: Quadratic Model for Trucks cost

2 Linear Regression Model

A natural staring point is the recommended linear model from Problem Set #7.

2.1 Quadratic Specification for Damage

we can say that the more a Truck is damaged the less likely it will sell. to impliment this in the model we squared the damage variable to have more impact in he realtionship.

The results of this regression specification are shown in Table 1.

according to our model milage and age are more significant predictor than damge with -5.053e-06 and -7.163e-02 coefficient respectively. the P value for damage 0.06 and for squared damage 0.03 which is slightly significant when squared

3 Nonlinear Specifications

3.1 Nonparametric Specification for Damage

The specification in Table 1 assumes a quadratic functional form for the relationship between cost and damage. To consider the damage variable alone, while accounting for the effects of other variables, one can fit a non-parametric model to the residuals from a model of Trucks cost, after regressing Trucks cost on the other variables. This leaves only the variation in trucks cost that is not explained by the other variables. Going one step further, perform the same transformation to the damage variable: take the residuals from a model of damage, after regressing damage on the other variables. This allows a model that would fit exactly the same as if it were estimated within a full model with all variables included.

To illustrate the fit of the model, Figure 1 shows a scatter plot of the residual log cost on damage. The observations are shown in blue and the fitted values are shown in red. The variation in the fitted values results from the fact that it is not plotted against the transformed excess damage variable used in the regressions. Still, the quadratic pattern is apparent and appears to match the data.

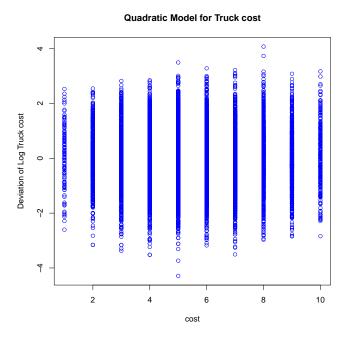


Fig. 1: Linear-Quadratic Model for Truck cost

As a comparison, Figure ?? augments the above by showing the plot against the residuals from the regression for damage: the "excess damage" compared to what would be expected given the other characteristics of a truck.

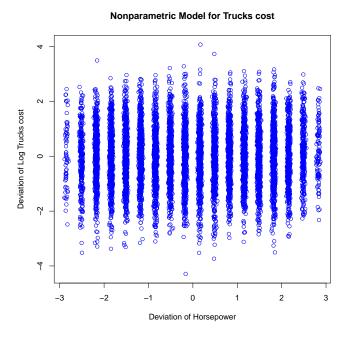


Fig. 2: Linear-Quadratic Model for Tractor Prices: damage

Now consider a nonparametric specification for the relationship between cost and damage. Figure 3 overlays the nonparametric estimate (shown in green) with the above in Figure ??. The pattern has more variation in slope but closely follows the prediction from the quadratic model. So far, it appears that the quadratic form is close enough.

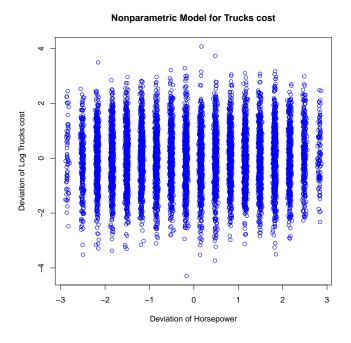


Fig. 3: Nonparametric Model for Tractor Prices: Damage

3.2 Nonparametric Specification for Age

As above, first conduct FWL regressions to reduce the problem to two dimensions. To illustrate the fit of the model, Figure 4 shows a scatter plot of the residual log cost on the residuals from the regression for age: the "excess age" of a truck compared to what would be expected given the other characteristics of the Truck. The observations are shown in blue and the fitted values are shown in red.

Next we considered a nonparametric specification for the relationship between cost and age. Figure 4 overlays the nonparametric estimate (shown in green) with the linear model. The pattern has more variation in slope but closely follows the prediction from the linear model. Although the nonparametric estimate varies around the linear estimate, it appears that the linear form is a close enough approximation without the added complexity. Next, I will revisit the remaining continuous variable.

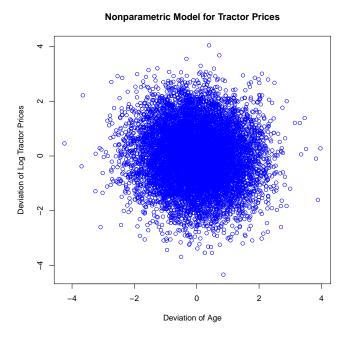


Fig. 4: Nonparametric Model for Tractor Prices: Excess Age

3.3 Nonparametric Specification for mileage

As above, first conduct FWL regressions to reduce the problem to two dimensions. To illustrate the fit of the model, Figure 5 shows a scatter plot of the residual log cost on residuals from the regression for mileage: the "excess mileage" of a Truck compared to what would be expected given the other characteristics of the Truck. The observations are shown in blue and the fitted values are shown in red. As with age, the linear fit follows a straight line, since we have a single variable with no quadratic transformation. I moved directly to the nonparametric specification for the relationship between prices and engine hours. Figure 5 overlays the nonparametric estimate, shown in green. The pattern has more variation in slope but closely follows the prediction from the linear model. Although the nonparametric estimate varies around the linear estimate, it appears that the linear form is also a close enough approximation, just as was found for the age variable.

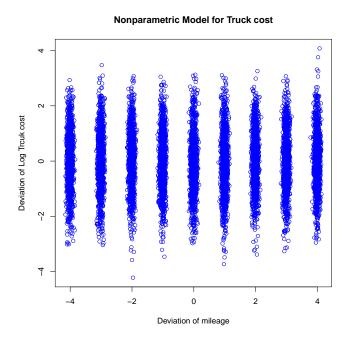


Fig. 5: Nonparametric Model for Tractor Prices: Excess Engine Hours

4 Semiparametric Estimates

As I was building the above nonparametric models, I stored the predictions and will now use them as variables in linear models. Table 2 shows the estimates from a set of models. Model 1 is the benchmark linear model in Table ??. Model 2 is a semi-parametric model with a nonparametric fit on horsepower substituted in for the horsepower variables. Models 3 and 4 are semi-parametric models with nonparametric fits on age and engine hours, respectively. Model 5 is a maximally semiparametric model, with nonparametric fits for all continuous variables. For each of the single-variable semiparametric models, the coefficients are near one and the fits are similar to the linear model. Even with maximal flexibility, the fit of Model 5 is not much better than the benchmark linear model. Across all models, the adjusted \bar{R}^2 values are all hovering around 0.80. All things considered, these are excellent models and the linear model is sufficient.

	Model 1	Model 2	Model 3	Model 4	Model 5
(Intercept)	8.76488***	8.65711***	8.95824***	8.76488***	8.74458***
•	(0.08135)	(0.05515)	(0.07530)	(0.08135)	(0.11028)
squared_damage	0.00459*		0.00454*	0.00459*	
	(0.00214)		(0.00213)	(0.00214)	
make	-0.00619	-0.00619	-0.00580	-0.00619	-0.00587
	(0.00537)	(0.00537)	(0.00536)	(0.00537)	(0.00536)
damage	-0.04677	-0.00394	-0.05654^*	-0.04677	-0.01388
	(0.02543)	(0.00808)	(0.02536)	(0.02543)	(0.00873)
dealer	0.00196	0.00232	0.00245	0.00196	0.00270
	(0.00417)	(0.00417)	(0.00416)	(0.00417)	(0.00417)
mileage	-0.00001***	-0.00001***	-0.00001***	-0.00001***	-0.00001^*
	(0.00000)	(0.00000)	(0.00000)	(0.00000)	(0.00000)
age	-0.07163^{***}	-0.06884^{***}		-0.07163^{***}	-0.03239
	(0.01114)	(0.01118)		(0.01114)	(0.03528)
type	0.04438^*	0.05465^*	0.04512^*	0.04438^*	0.05433^*
	(0.02182)	(0.02212)	(0.02181)	(0.02182)	(0.02211)
damage_np		1.28304^{**}			1.24683**
		(0.43922)			(0.43894)
age_np			1.01862^{***}		0.96938**
			(0.13621)		(0.31910)
mileage_np					0.69829
					(0.36384)
\mathbb{R}^2	0.19087	0.19119	0.19206	0.19087	0.19269
Adj. R^2	0.19029	0.19061	0.19149	0.19029	0.19195
Num. obs.	9861	9861	9861	9861	9861

^{***}p < 0.001; **p < 0.01; *p < 0.05

Tab. 2: Semiparametric Models for Truck cost

5 Generalized Additive Model

5.1 Linear Model

As an example of the output from the GAM specification, I first estimated the model with no nonlinear terms, which is essentially a linear regression.

```
Family: gaussian
Link function: identity
Formula:
log_cost ~ damage + squared_damage + age + make + damage + dealer +
   mileage + age + type
Parametric coefficients:
                Estimate Std. Error t value Pr(>|t|)
(Intercept)
               8.765e+00 8.135e-02 107.740 < 2e-16 ***
              -4.677e-02 2.543e-02 -1.839
                                            0.0659 .
damage
squared_damage 4.595e-03 2.135e-03 2.152
                                           0.0314 *
age
              -7.163e-02 1.114e-02 -6.430 1.34e-10 ***
              -6.194e-03 5.368e-03 -1.154 0.2486
make
dealer
               1.964e-03 4.168e-03 0.471
                                             0.6376
             -5.053e-06 1.062e-06 -4.756 2.01e-06 ***
mileage
               4.438e-02 2.182e-02 2.034 0.0420 *
type
Signif. codes: 0 *** 0.001 ** 0.01 * 0.05 . 0.1
R-sq.(adj) =
              0.19
                     Deviance explained = 19.1%
GCV = 1.1444 Scale est. = 1.1435 n = 9861
```

5.2 Semiparametric Model

Further investigating the results of the full semiparametric specification in Model 5 of Table 2, I estimated the model with all three continuous variables specified as nonparametric functions. The result was that almost all the variables—both linear and nonlinear—were statistically significant. The only exception was a loss in significance of the diesel indicator.

```
Family: gaussian
Link function: identity
Formula:
log_cost ~ s(damage) + s(age)
Parametric coefficients:
           Estimate Std. Error t value Pr(>|t|)
(Intercept) 7.53777 0.01077
                               700.2 <2e-16 ***
Signif. codes: 0 *** 0.001 ** 0.01 * 0.05 . 0.1
Approximate significance of smooth terms:
           edf Ref.df
                            F p-value
s(damage) 1.000 1.000
                        0.908 0.341
        4.149 5.111 243.022 <2e-16 ***
s(age)
Signif. codes: 0 *** 0.001 ** 0.01 * 0.05 . 0.1
R-sq.(adj) = 0.191
                     Deviance explained = 19.1%
GCV = 1.1436 Scale est. = 1.1429
                                   n = 9861
```

On the other hand, the adjusted R-squared has not increased very much, from 0.799 to 0.819 under this specification, which may not justify the added complexity of the model. Perhaps more importantly, the coefficients on the linear terms are very similar across models, indicating that the models support similar conclusions relating to any business decision involving the John Deere premium. With this second model, we have even more support for those conclusions and are certain that the conclusions are not coincidental results of the functional form decisions for previous models.

Perhaps as a middle ground, we can estimate a model with a nonparametric specification for the horsepower variable alone, since it seems to have a nonlinear relationship with value in either case. This retains most of the predictive value of the maximally semiparametric model and accommodates the nonlinear relationship with value of horsepower.

```
Family: gaussian
Link function: identity
Formula:
log_cost ~ s(damage) + age + make + damage + dealer + mileage +
   age + type
Parametric coefficients:
             Estimate Std. Error t value Pr(>|t|)
(Intercept) 1.209e+00 3.007e-02 40.192 < 2e-16 ***
age
           -7.148e-02 1.114e-02 -6.417 1.46e-10 ***
          -6.227e-03 5.368e-03 -1.160 0.2461
make
damage
           1.310e+00 9.977e-03 131.253 < 2e-16 ***
dealer
           1.870e-03 4.169e-03 0.448
                                         0.6538
          -5.072e-06 1.062e-06 -4.774 1.84e-06 ***
mileage
           4.453e-02 2.182e-02 2.041
                                         0.0413 *
type
Signif. codes: 0 *** 0.001 ** 0.01 * 0.05 . 0.1
Approximate significance of smooth terms:
          edf Ref.df F p-value
s(damage) 4.75 5.914 3432 <2e-16 ***
Signif. codes: 0 *** 0.001 ** 0.01 * 0.05 . 0.1 1
Rank: 15/16
R-sq.(adj) = 0.191 Deviance explained = 19.1%
GCV = 1.1444 Scale est. = 1.1431 n = 9861
```

6 The Box-Tidwell Transformation

The Box–Tidwell function tests for non-linear relationships to the mean of the dependent variable. The nonlinearity is in the form of an exponential transformation in the form of the Box-Cox transformation, except that the transformation is taken on the explanatory variables.

6.1 Transformation of Damage

Performing the transformation on the Damage variable produces a modified form of the linear model. This specification allows a single exponential transformation on Damage, rather than a quadratic form.

```
MLE of lambda Score Statistic (z) Pr(>|z|)
4.4456
2.132 0.03301 *
---
Signif. codes: 0 *** 0.001 ** 0.01 * 0.05 . 0.1 1
iterations = 10
```

The R output is the statistics for a test of nonlinearity: that the exponent λ in the Box–Tidwell transformation is zero. The "MLE of lambda" statistic is the optimal exponent on Damage. Similar to the Box-Cox transformation, with Box-Tidwell, the exponents are on the explanatory variables and are all called lambda, in contrast to the parameter τ in our class notes. The exponent is significantly different from 0, although it is a small positive value, which suggests an increasing relationship for the value of horsepower with a slope that is sharply declining. Next I consider the possibility of a changing relationship for the next continuous variable.

6.2 Transformation of Age

```
MLE of lambda Score Statistic (z) Pr(>|z|)
-0.084324 5.4043 6.507e-08 ***
---
Signif. codes: 0 *** 0.001 ** 0.01 * 0.05 . 0.1 1
iterations = 3
```

This coefficient is effectively 1, which is more evidence of a purely linear relationship between log_saleprice and age: the percentage depre-

ciation rate is constant. Next, I will consider the possibility of nonlinearity in depreciation from hours of use.

6.3 Transformation of Mileage

Although $\hat{\lambda}$ is not statistically significant, this suggests a moderately increasing relationship between the log of tractor prices and engine hours, which means that tractors with high hours of use depreciate more quickly with each additional hour of use.

Since a nonlinear relationship was detected with horsepower, I will next estimate a model with nonlinearity in all three continuous variables.

6.4 Transformation of All Three Continuous Variables

```
MLE of lambda Score Statistic (z) Pr(>|z|)
damage 8.24044 0.0417 0.9667305
age 0.24381 3.7234 0.0001965 ***
mileage 1.42031 -1.0083 0.3133037
---
Signif. codes: 0 *** 0.001 ** 0.01 * 0.05 . 0.1 1
iterations = 9
```

The performance is similar to the other models with forms of nonlinearity for the value of horsepower. Now consider the full set of such models in a table for a final comparison.

	Model 1	Model 2	Model 3
(Intercept)	8.76488***	8.65711***	8.66231***
•	(0.08135)	(0.05515)	(0.05080)
squared_damage	0.00459^{*}	,	,
	(0.00214)		
make	-0.00619	-0.00619	-0.00618
	(0.00537)	(0.00537)	(0.00537)
damage	-0.04677	-0.00394	
	(0.02543)	(0.00808)	
dealer	0.00196	0.00232	0.00196
	(0.00417)	(0.00417)	(0.00417)
mileage	-0.00001^{***}	-0.00001^{***}	-0.00001^{***}
	(0.00000)	(0.00000)	(0.00000)
age	-0.07163***	-0.06884***	-0.07347^{***}
	(0.01114)	(0.01118)	(0.01103)
type	0.04438^*	0.05465^*	0.04237
	(0.02182)	(0.02212)	(0.02175)
damage_np		1.28304^{**}	
		(0.43922)	
damage_bt			0.00000^*
			(0.00000)
\mathbb{R}^2	0.19087	0.19119	0.19077
Adj. \mathbb{R}^2	0.19029	0.19061	0.19027
Num. obs.	9861	9861	9861
***	* ^ ~		

^{***}p < 0.001; **p < 0.01; *p < 0.05

Tab. 3: Alternate Models for Tractor Prices

7 Final Comparison of Candidate Models

I created one more variable mileage_bt by raising horsepower to the optimal exponent $\hat{\lambda}=4.4456.$ Then, I included this variable in the place of the damage variables a the linear regression model.