

Subject:

Year. Month. Date.

باسمہ تعالیٰ

تمرین برائے امتحان

۹۵۲۱۰۲۱ - ۱۰/۱۰/۲۰۲۱

$$E[m] = \sum_{n=0}^{\infty} n p(n) = \sum_{n=a}^{\infty} n p(n) = \sum_{n=a}^{\infty} a p(n) \quad (1) \text{ الف}$$

$$E[m] > a p(m, a) \Rightarrow \frac{E[m]}{a}, P(m, a)$$

$$\frac{E[m]}{a}, P(m, a) \quad n \sim (m - \mu)^r \quad (ب)$$

$$\frac{E[(m - \mu)^r]}{a^r} > P((m - \mu)^r > a^r) \rightarrow \frac{\text{var}(m)}{a^r}, P(m, \mu > a)$$

$$\sqrt{\text{var}(m)} = \sigma \Rightarrow \text{var}(m) = \sigma^2 \quad (ج)$$

$$P \rightarrow P(a = 0) < \frac{P}{a} \rightarrow a = 0 \Rightarrow P(0) < \frac{P}{0} \Rightarrow P(0) < \infty, \infty$$

$$E[ax+b] = \sum_{n=0}^{\infty} (ax+b) h_n = a \sum_{n=0}^{\infty} n h_n + b \sum_{n=0}^{\infty} h_n \quad (2) \text{ الف}$$

$$= a E[m] + b = a E[m] + b$$

$$E[xy] = E[x] E[y]$$

$$E[xy] = \sum_{n=0}^{\infty} \sum_{y=0}^{\infty} xy h_{n,y} = \sum_{n,y=0}^{\infty} n y g_n h_{n,y} = \sum_{n=0}^{\infty} n g_n \sum_{y=0}^{\infty} y h_{n,y} = E[x] E[y] \quad (ب)$$

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$$\sigma^T[(aX + bY + c)^T] = \text{var}[(aX + bY + c)^T], \quad (C)$$

$$E[(aX + bY + c)^T] - (E[aX + bY + c])^T =$$

$$= E[a^T X^T + b^T Y^T + c^T + r_{ab}XY + r_{ac}X + r_{bc}Y] - (E[aX + bY + c])^T$$

$$= a^T E[X^T] + b^T E[Y^T] + c^T + r_{ab}E[XY] + r_{ac}E[X] + r_{bc}E[Y]$$

$$= (aE[X] + bE[Y] + c)^T = \text{var}[(aX + bY + c)^T]$$

$$a^T (E[X^T] - E[X]^T) + b^T (E[Y^T] - E[Y]^T) + r_{ab}(E[XY] - E[X]E[Y])$$

$$+ r_{ac}E[X] + r_{bc}E[Y] = a^T \sigma^T(X) + b^T \sigma^T(Y) + r_{ab}\sigma(XY)$$

$$\sigma^T_{aX+bY} = a^T \sigma^T_X + b^T \sigma^T_Y \quad (\rightarrow)$$

$$\text{var}[aX - bY] = E[(aX - bY)^T] - E[aX - bY]^T =$$

$$E[a^T X^T + b^T Y^T - r_{ab}XY] - (aE[X] - bE[Y])^T =$$

$$= a^T \sigma^T(X) + b^T \sigma^T(Y) - r_{ab}(E[X]E[Y] - E[X]E[Y]) = a^T \sigma^T(X) + b^T \sigma^T(Y)$$

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$$\text{Var}\left(\frac{X}{\sigma_X} + \frac{Y}{\sigma_Y}\right) = 1 + 2\rho(X, Y)$$

(Q1) (5)

$$\sigma^2\left[\frac{X}{\sigma_X} + \frac{Y}{\sigma_Y}\right] \left\{ \begin{array}{l} \rightarrow \sigma^2[aX + bY] = a^2\sigma_X^2 + b^2\sigma_Y^2 + 2ab\sigma_{XY} \\ \frac{1}{\sigma_X^2} = a \quad \frac{1}{\sigma_Y^2} = b \end{array} \right.$$

$$= \frac{1}{\sigma_X^2} \times \sigma_X^2 + \frac{1}{\sigma_Y^2} \times \sigma_Y^2 + \frac{2}{\sigma_X \sigma_Y}$$

$$= 1 + 2\rho(X, Y)$$

$$\text{Var}\left(\frac{X}{\sigma_X} - \frac{Y}{\sigma_Y}\right) = 1 - 2\rho(X, Y)$$

$$\sigma^2\left[\frac{X}{\sigma_X} - \frac{Y}{\sigma_Y}\right] \left\{ \begin{array}{l} \Rightarrow \sigma^2[aX - bY] = a^2\sigma_X^2 + b^2\sigma_Y^2 - 2ab\sigma_{XY} \\ \frac{1}{\sigma_X^2} = a \quad \frac{1}{\sigma_Y^2} = b \end{array} \right. \quad 1 - 2\rho(X, Y)$$

$$\text{Var}\left[\frac{X}{\sigma_X} + \frac{Y}{\sigma_Y}\right] = 1 + 2\rho(X, Y) \quad \left\{ \begin{array}{l} \Rightarrow 1 + 2\rho(X, Y) \geq 0 \Rightarrow \rho(X, Y) \geq -1 \\ \text{Var}\left[\frac{X}{\sigma_X} + \frac{Y}{\sigma_Y}\right] \geq 0 \end{array} \right.$$

$$\text{Var}\left[\frac{X}{\sigma_X} - \frac{Y}{\sigma_Y}\right] = 1 - 2\rho(X, Y) \quad \left\{ \begin{array}{l} \Rightarrow 1 - 2\rho(X, Y) \geq 0 \end{array} \right.$$

$$\text{Var}\left[\frac{X}{\sigma_X} - \frac{Y}{\sigma_Y}\right] \geq 0 \quad \left\{ \begin{array}{l} -1 \leq \rho(X, Y) \leq 1 \end{array} \right.$$

$$Y = aX + b$$

(Q2)

$$\text{Cov}(X, Y) = E[XY] - E[X]E[Y] = E[X(aX + b)] - E[X]E[aX + b]$$

$$= E[aX^2 + bX] - E[X](aE[X] + b) = aE[X^2] + bE[X] - aE[X]^2 - bE[X]$$

$$= a\sigma_X^2$$

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$$X = \frac{Y}{a} - \frac{b}{a} \quad \text{cov}(m, y) = E\left[Y\left(\frac{Y}{a} - \frac{b}{a}\right)\right] - E[Y]E\left[\frac{Y}{a} - \frac{b}{a}\right]$$

$$= E\left[\frac{Y^2}{a} - \frac{Yb}{a}\right] - E[Y]\left(\frac{E[Y]}{a} - \frac{b}{a}\right) = \frac{E[Y^2]}{a} - \frac{b}{a}E[Y] - \frac{E[Y]^2}{a} + \frac{b}{a}E[Y]$$

$$= \frac{\sigma_Y^2}{a} \Rightarrow a\sigma_X^2 = \frac{\sigma_Y^2}{a} \Rightarrow \sigma_Y^2 = a^2\sigma_X^2 \Rightarrow a\sigma_X = \sigma_Y$$

$$\text{var}\left[\frac{-b}{a\sigma_X}\right] = \text{var}[\text{const}] = 0 \Rightarrow 1 - \rho(X, Y) = 0 \Rightarrow \rho(X, Y) = 1$$

$$Y = aX + b \quad \text{demonstration} \quad (2)$$

$$\text{var}\left(\frac{X}{\sigma_X} - \frac{Y}{\sigma_Y}\right) = 1 - \rho(X, Y) = 0 \Rightarrow \frac{X}{\sigma_X} - \frac{Y}{\sigma_Y} = c \Rightarrow 1 = \frac{\sigma_Y}{\sigma_X} X + c \Rightarrow Y = aX + b$$

$$a = \frac{\sigma_Y}{\sigma_X} = \frac{\sqrt{\text{Var}(Y)}}{\sqrt{\text{Var}(X)}} > 0$$

$$P(\underbrace{m > y}_A) = 0.4 \quad P(\underbrace{y > z}_B) = 0.4 \quad P(\underbrace{z > m}_C) = 0.4 \quad (3)$$

$$P(A \cap B) = P(m > y \cap y > z) = P(m > z) = 0.4 \times 0.4 = 0.16$$

$$\underbrace{P(m > z)}_{0.16} + \underbrace{P(z > m)}_{0.4} = P(U) = 1$$

1.16 ≠ 1 - X.

$$P(k: P) = \begin{cases} P & \text{if } k=1 \\ 1-P & \text{if } k=0 \end{cases} \quad (4) \quad (5)$$

$$E[X] = 1 \times P + 0 \times (1-P) = P$$

$$\text{var}[X] = E[X^2] - E[X]^2 = 1^2 P + 0 \times (1-P) - (1 \times P + 0 \times (1-P))^2 = P(1-P)$$

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$$P(k; n, p) = \binom{n}{k} p^k (1-p)^{n-k} \quad E[X] = \sum_{k=0}^n \binom{n}{k} p^k (1-p)^{n-k} \cdot k = np$$
$$\text{var}[X] = E[X^2] - E[X]^2 = \sum_{k=0}^n k^2 \binom{n}{k} p^k (1-p)^{n-k} - (np)^2 = npq$$

$$P(k; \lambda) = \frac{\lambda^k e^{-\lambda}}{k!}, \quad k = 0, 1, 2, \dots \quad (C)$$

$$E[X] = \sum_{k=0}^{\infty} k e^{-\lambda} \frac{\lambda^k}{k!} = e^{-\lambda} \sum_{k=1}^{\infty} k \frac{\lambda^k}{k!} = e^{-\lambda} \lambda \sum_{k=1}^{\infty} \frac{\lambda^{k-1}}{(k-1)!} = e^{-\lambda} \lambda e^{\lambda} = \lambda$$

$$E[X^2] = \sum_{k=0}^{\infty} k^2 e^{-\lambda} \frac{\lambda^k}{k!} = e^{-\lambda} \sum_{k=1}^{\infty} k \frac{\lambda^k}{(k-1)!} = \lambda e^{-\lambda} \frac{d}{d\lambda} \sum_{k=1}^{\infty} \frac{\lambda^k}{(k-1)!} = \lambda e^{-\lambda} \frac{d}{d\lambda} (\lambda e^{\lambda})$$
$$= \lambda e^{-\lambda} (e^{\lambda} + \lambda e^{\lambda}) = \lambda + \lambda^2$$

$$\text{var}[X] = E[X^2] - E[X]^2 = \lambda + \lambda^2 - \lambda^2 = \lambda$$

$$f(n; \lambda) = \begin{cases} \lambda e^{-\lambda} & n=0 \\ n e^{-\lambda} & n > 0 \end{cases} \quad (D)$$

$$E[X] = \int_0^{\infty} n e^{-\lambda n} d\lambda \rightarrow -n e^{-\lambda n} \Big|_0^{\infty} + \int_0^{\infty} e^{-\lambda n} d\lambda = \frac{1}{\lambda}$$

$$E[X^2] = \int_0^{\infty} n^2 e^{-\lambda n} d\lambda = -n e^{-\lambda n} \Big|_0^{\infty} + \int_0^{\infty} 2n e^{-\lambda n} d\lambda = \frac{2}{\lambda^2}$$

$$\text{var}[X] = \frac{2}{\lambda^2} - \left(\frac{1}{\lambda}\right)^2 = \frac{1}{\lambda^2}$$

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$$P(m|y) = P(y) P(m|y)$$

(4) (ب)

$$E[m|y=y] = g(y) \Rightarrow E[X|Y] = g(Y) = g \circ Y$$

$$E[E[X|Y]] = \int_{y=-\infty}^{\infty} g(y) d_Y(y) dy$$

$$g(y) = \int_{m=-\infty}^{\infty} m d_{X|Y}(m|y) dm = \int_{m=-\infty}^{\infty} m \frac{d(m,y)}{d_Y(y)} dm$$

$$E[E[X|Y]] = \int_{y=-\infty}^{\infty} \int_{m=-\infty}^{\infty} m \frac{d(m,y)}{d_Y(y)} d_Y(y) dy = \int_{m=-\infty}^{\infty} \int_{y=-\infty}^{\infty} m d(m,y) d_Y(y) dy = E[X]$$

ب) $A: \frac{1}{4}$: اول
 $B: \frac{1}{4}$: 2
 $C: \frac{1}{4}$: 3
 $D: \frac{1}{4}$: 4

$$E[E[A|B]] + E[E[B|A]] = E[A] + E[B] = \frac{1}{4} + \frac{1}{4} = \frac{1}{2}$$

$$\frac{1}{2} = \Delta = E[X]$$

(5)

$$E[T] = E[t_1] + E[t_2] = 1$$

$$\begin{aligned} \text{بالفرض: } -\frac{1}{4} + \frac{1}{4} = -\frac{1}{4} & \left\{ \begin{array}{l} \frac{-1}{4} > -\frac{1}{4} \\ \frac{-1}{4} < -\frac{1}{4} \end{array} \right. \end{aligned}$$

$$E(T_R | \text{skub} = A) = 1 + E(T_R | B) \quad (I)$$

(6) (ب)

$$E(T_R | B) = \frac{1}{2} E(T_R | C) + \frac{1}{2} E(T_R | A) + \frac{1}{2} E(T_R | D) + 1 \quad (II)$$

$$E(T_R | C) = \frac{1}{2} E(T_R | B) + \frac{1}{2} E(T_R | E) + 1 = E(T_R | D) \quad (III)$$

$$E(T_R | E) = \frac{1}{2} E(T_R | C) + \frac{1}{2} E(T_R | D) + \frac{1}{2} E(T_R | F) + 1 \quad (IV)$$

$$(I) \text{ و } (II) \Rightarrow E(T_R | A) - 1 = \frac{1}{2} E(T_R | C) + \frac{1}{2} E(T_R | A) + \frac{1}{2} E(T_R | D) + 1$$

$$(II) \Rightarrow \frac{1}{2} E(T_R | A) = 1 + \frac{1}{2} E(T_R | C) \Rightarrow E(T_R | C) = E(T_R | A) - 2 \quad (V)$$

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⑨ الف اگر ما به شرط مفصل در نظر بگیریم که طالع / کسبه تا آزاد شود طالع:

$$E[X] = \frac{9}{10} (E[X] + 1) + \frac{1}{10} (E[X] + 2) = \frac{1}{10} E[X] + 1 + \frac{1}{10} \Rightarrow E[X] = 9.0$$

$$E[X] = \frac{1}{2} (1 + E(n|y)) + \frac{1}{2} (2 + E(n|z))$$

$$\frac{1}{2} (1 + \frac{1}{2} \cdot 1 + \frac{1}{2} \cdot 1) + \frac{1}{2} (2 + \frac{1}{2} \cdot 1 + \frac{1}{2} \cdot 1) = \frac{10}{4} = \frac{5}{2}$$

$$E[X^2] = \frac{9}{10} (E[X] + 1)^2 + \frac{1}{10} (E[X] + 2)^2 = \frac{9}{10} \cdot 10.0 + \frac{1}{10} \cdot 11.0 = 11.0$$

$$\text{Var}[X] = E[X^2] - E[X]^2 = 11.0 - 9.0 = 2.0, \text{ V.D}$$

$$E[X^2] = \frac{1}{2} (1 + \frac{1}{2} \cdot 1)^2 + \frac{1}{2} (2 + \frac{1}{2} \cdot 1)^2 = \frac{249}{12}$$

$$\text{Var}[X] = E[X^2] - E[X]^2 = \frac{249}{12} - \frac{25}{4} = \frac{198}{12} = \frac{99}{6}$$

⑩ الف (تابع چگالی) $f_{X_1}(x_1)$

$$f_{X_1}(x_1) = \int_{-\infty}^{\infty} f_{X_1, X_2}(x_1, x_2) dx_2 = \int_{-\infty}^{\infty} \frac{1}{2} e^{-x_1 - x_2} dx_2 = \frac{1}{2} e^{-x_1}$$

و $f_{X_1}(x_1) < 1$

$\int_{-\infty}^{\infty} f_{X_1}(x_1) dx_1 = \int_{-\infty}^{\infty} \frac{1}{2} e^{-x_1} dx_1 = 1$

تواند توزیع چگالی

$$P(X_1 < 1/2 | X_1 < 1, X_2 < 1/2) = \int_0^{1/2} \int_0^{1/2} \frac{1}{2} e^{-x_1 - x_2} dx_1 dx_2 = \frac{1}{4} e^{-1/2}$$

$$P(X_1 < 1/2 | X_1 < 1, X_2 < 1/2) = \frac{P(X_1 < 1/2, X_2 < 1/2)}{P(X_1 < 1, X_2 < 1/2)} = \frac{\frac{1}{4} e^{-1/2}}{\frac{1}{2} e^{-1/2}} = \frac{1}{2} = 0.5$$

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