Quiz I Review Probabilistic Systems Analysis 6.041/6.431

Massachusetts Institute of Technology

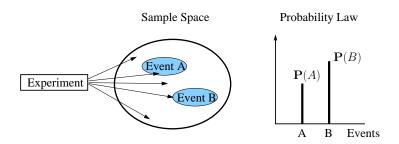
October 7, 2010

Quiz Information

- Closed-book with one double-sided 8.5 x 11 formula sheet allowed
- Date: Tuesday, October 12, 2010
- Time: 7:30 9:00 PM
- Location: 54-100
- Content: Chapters 1-2, Lecture 1-7, Recitations 1-7, Psets 1-4, Tutorials 1-3
- Show your reasoning when possible!

A Probabilistic Model:

- Sample Space: The set of all possible outcomes of an experiment.
- Probability Law: An assignment of a nonnegative number P(E) to each event E.



Probability Axioms

Given a sample space Ω :

- 1. **Nonnegativity:** $P(A) \ge 0$ for each event A
- 2. **Additivity:** If A and B are disjoint events, then

$$\mathbf{P}(A \cup B) = P(A) + P(B)$$

If A_1, A_2, \ldots , is a sequence of disjoint events,

$$\mathbf{P}(A_1 \cup A_2 \cup \cdots) = P(A_1) + P(A_2) + \cdots$$

3. Normalization $P(\Omega) = 1$

Properties of Probability Laws

Given events A, B and C:

- 1. If $A \subset B$, then $\mathbf{P}(A) \leq \mathbf{P}(B)$
- 2. $P(A \cup B) = P(A) + P(B) P(A \cap B)$
- 3. $P(A \cup B) \le P(A) + P(B)$
- 4. $\mathbf{P}(A \cup B \cup C) = \mathbf{P}(A) + \mathbf{P}(A^c \cap B) + \mathbf{P}(A^c \cap B^c \cap C)$

Discrete Models

• **Discrete Probability Law:** If Ω is finite, then each event $A \subseteq \Omega$ can be expressed as

$$A = \{s_1, s_2, \dots, s_n\}$$
 $s_i \in \Omega$

Therefore the probability of the event A is given as

$$\mathbf{P}(A) = \mathbf{P}(s_1) + \mathbf{P}(s_2) + \cdots + \mathbf{P}(s_n)$$

 Discrete Uniform Probability Law: If all outcomes are equally likely,

$$\mathbf{P}(A) = \frac{|A|}{|\Omega|}$$

Conditional Probability

• Given an event B with $\mathbf{P}(B) > 0$, the conditional probability of an event $A \subseteq \Omega$ is given as

$$\mathbf{P}(A|B) = \frac{\mathbf{P}(A \cap B)}{\mathbf{P}(B)}$$

- P(A|B) is a valid probability law on Ω , satisfying
 - 1. $P(A|B) \ge 0$
 - **2**. **P**($\Omega | B$) = 1
 - 3. $\mathbf{P}(A_1 \cup A_2 \cup \cdots | B) = \mathbf{P}(A_1 | B) + \mathbf{P}(A_2 | B) + \cdots$, where $\{A_i\}_i$ is a set of disjoint events
- P(A|B) can also be viewed as a probability law on the restricted universe B

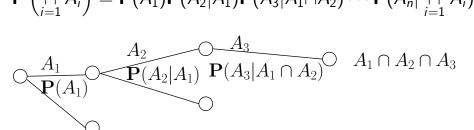
Multiplication Rule

• Let A_1, \ldots, A_n be a set of events such that

$$\mathbf{P}\left(\bigcap_{i=1}^{n-1}A_i\right)>0.$$

Then the joint probability of all events is

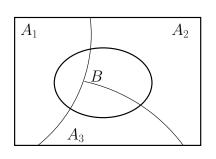
$$\mathbf{P}\left(\bigcap_{i=1}^{n}A_{i}\right)=\mathbf{P}(A_{1})\mathbf{P}(A_{2}|A_{1})\mathbf{P}(A_{3}|A_{1}\cap A_{2})\cdots\mathbf{P}(A_{n}|\bigcap_{i=1}^{n-1}A_{i})$$

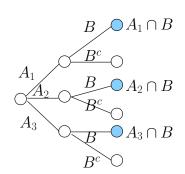


Total Probability Theorem

Let A_1, \ldots, A_n be disjoint events that partition Ω . If $\mathbf{P}(A_i) > 0$ for each i, then for any event B,

$$\mathbf{P}(B) = \sum_{i=1}^{n} \mathbf{P}(B \cap A_i) = \sum_{i=1}^{n} \mathbf{P}(B|A_i)\mathbf{P}(A_i)$$

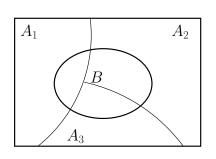


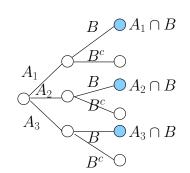


Bayes Rule

Given a finite partition A_1, \ldots, A_n of Ω with $\mathbf{P}(A_i) > 0$, then for each event B with $\mathbf{P}(B) > 0$

$$\mathbf{P}(A_i|B) = \frac{\mathbf{P}(B|A_i)\mathbf{P}(A_i)}{\mathbf{P}(B)} = \frac{\mathbf{P}(B|A_i)\mathbf{P}(A_i)}{\sum_{i=1}^{n}\mathbf{P}(B|A_i)\mathbf{P}(A_i)}$$





Independence of Events

Events A and B are independent if and only if

$$\mathbf{P}(A\cap B)=\mathbf{P}(A)\mathbf{P}(B)$$

or

$$P(A|B) = P(A)$$
 if $P(B) > 0$

• Events A and B are **conditionally independent** given an event C if and only if

$$P(A \cap B|C) = P(A|C)P(B|C)$$

or

$$P(A|B \cap C) = P(A|C)$$
 if $P(B \cap C) > 0$

Independence
 ⇔ Conditional Independence.

Independence of a Set of Events

• The events A_1, \ldots, A_n are **pairwise independent** if for each $i \neq j$

$$\mathbf{P}(A_i \cap A_j) = \mathbf{P}(A_i)\mathbf{P}(A_j)$$

• The events A_1, \ldots, A_n are **independent** if

$$\mathbf{P}\left(\bigcap_{i\in S}A_i\right)=\prod_{i\in S}\mathbf{P}(A_i)\quad\forall\ S\subseteq\{1,2,\ldots,n\}$$

 Pairwise independence ⇒ independence, but independence ⇒ pairwise independence.

Counting Techniques

• Basic Counting Principle: For an m-stage process with n_i choices at stage i,

$$\#$$
 Choices = $n_1 n_2 \cdots n_m$

Permutations: k-length sequences drawn from n distinct items without replacement (order is important):

$$\#$$
 Sequences $= n(n-1)\cdots(n-k+1) = \frac{n!}{(n-k)!}$

• **Combinations:** Sets with *k* elements drawn from *n* distinct items (order within sets is not important):

Sets =
$$\binom{n}{k}$$
 = $\frac{n!}{k!(n-k)!}$

Counting Techniques-contd

 Partitions: The number of ways to partition an n-element set into r disjoint subsets, with n_k elements in the kth subset:

$$\binom{n}{n_1, n_2, \dots, n_r} = \binom{n}{n_1} \binom{n - n_1}{n_2} \cdots \binom{n - n_1 - \dots - n_r - 1}{n_r}$$
$$= \frac{n!}{n_1! n_2! \cdots n_r!}$$

where

$$\binom{n}{k} = \frac{n!}{k!(n-k)!}$$
$$\sum_{i=1}^{r} n_i = n$$

Discrete Random Variables

 A random variable is a real-valued function defined on the sample space:

$$X:\Omega\to R$$

• The notation $\{X = x\}$ denotes an event:

$${X = x} = {\omega \in \Omega | X(\omega) = x} \subseteq \Omega$$

• The **probability mass function (PMF)** for the random variable X assigns a probability to each event $\{X = x\}$:

$$p_X(x) = \mathbf{P}(\{X = x\}) = \mathbf{P}(\{\omega \in \Omega | X(\omega) = x\})$$

PMF Properties

- Let X be a random variable and S a countable subset of the real line
- The axioms of probability hold:
 - 1. $p_X(x) \ge 0$
 - 2. $P(X \in S) = \sum_{x \in S} p_X(x)$
 - $3. \sum_{x} p_X(x) = \overline{1}$
- If g is a real-valued function, then Y = g(X) is a random variable:

$$\omega \xrightarrow{X} X(\omega) \xrightarrow{g} g(X(\omega)) = Y(\omega)$$

with PMF

$$p_Y(y) = \sum_{x \mid g(x) = y} P_X(x)$$

Expectation

Given a random variable X with PMF $p_X(x)$:

- $\mathbf{E}[X] = \sum_{x} x p_X(x)$
- Given a derived random variable Y = g(X):

$$\mathbf{E}[g(X)] = \sum_{x} g(x)p_X(x) = \sum_{y} yp_Y(y) = E[Y]$$
$$\mathbf{E}[X^n] = \sum_{x} x^n p_X(x)$$

• Linearity of Expectation: $\mathbf{E}[aX + b] = a\mathbf{E}[X] + b$.

Variance

The expected value of a derived random variable g(X) is

$$\mathbf{E}[g(X)] = \sum_{x} g(x) p_X(x)$$

The variance of X is calculated as

- $var(X) = E[(X E[X])^2] = \sum_{x} (x E[X])^2 p_X(x)$
- $var(X) = \mathbf{E}[X^2] \mathbf{E}[X]^2$
- $var(aX + b) = a^2 var(X)$

Note that $var(x) \ge 0$

Multiple Random Variables

Let X and Y denote random variables defined on a sample space Ω .

The joint PMF of X and Y is denoted by

$$p_{X,Y}(x,y) = \mathbf{P}(\{X = x\} \cap \{Y = y\})$$

 The marginal PMFs of X and Y are given respectively as

$$p_X(x) = \sum_{y} p_{X,Y}(x,y)$$
$$p_Y(y) = \sum_{y} p_{X,Y}(x,y)$$

Functions of Multiple Random Variables

Let Z = g(X, Y) be a function of two random variables

• PMF:

$$p_Z(z) = \sum_{(x,y)|g(x,y)=z} p_{X,Y}(x,y)$$

Expectation:

$$\mathbf{E}[Z] = \sum_{x,y} g(x,y) p_{X,Y}(x,y)$$

• Linearity: Suppose g(X, Y) = aX + bY + c.

$$\mathbf{E}[g(X,Y)] = a\mathbf{E}[X] + b\mathbf{E}[Y] + c$$

Conditioned Random Variables

 Conditioning X on an event A with P(A) > 0 results in the PMF:

$$p_{X|A}(x) = \mathbf{P}(\lbrace X = x \rbrace | A) = \frac{\mathbf{P}(\lbrace X = x \rbrace \cap A)}{\mathbf{P}(A)}$$

• Conditioning X on the event Y = y with $\mathbf{P}_Y(y) > 0$ results in the PMF:

$$p_{X|Y}(x|y) = \frac{\mathbf{P}(\{X = x\} \cap \{Y = y\})}{\mathbf{P}(\{Y = y\})} = \frac{p_{X,Y}(x,y)}{p_Y(y)}$$

Conditioned RV - contd

Multiplication Rule:

$$p_{X,Y}(x,y) = p_{X|Y}(x|y)p_Y(y)$$

• Total Probability Theorem:

$$egin{aligned}
ho_X(x) &= \sum_{i=1}^n \mathbf{P}(A_i)
ho_{X|A_i}(x) \
ho_X(x) &= \sum_{y}
ho_{X|Y}(x|y)
ho_Y(y) \end{aligned}$$

Conditional Expectation

Let X and Y be random variables on a sample space Ω .

• Given an event A with P(A) > 0

$$\mathbf{E}[X|A] = \sum_{x} x p_{X|A}(x)$$

• If $P_Y(y) > 0$, then

$$\mathbf{E}[X|\{Y=y\}] = \sum_{x} x p_{X|Y}(x|y)$$

• Total Expectation Theorem: Let A_1, \ldots, A_n be a partition of Ω . If $\mathbf{P}(A_i) > 0 \ \forall i$, then

$$\mathbf{E}[X] = \sum_{i=1}^{n} \mathbf{P}(A_i) \mathbf{E}[X|A_i]$$

Independence

Let X and Y be random variables defined on Ω and let A be an event with $\mathbf{P}(A) > 0$.

• *X* is independent of *A* if either of the following hold:

$$p_{X|A}(x) = p_X(x) \ \forall x$$

 $p_{X,A}(x) = p_X(x)\mathbf{P}(A) \ \forall x$

 X and Y are independent if either of the following hold:

$$p_{X|Y}(x|y) = p_X(x) \ \forall x \forall y$$
$$p_{X,Y}(x,y) = p_X(x)p_Y(y) \ \forall x \forall y$$

Independence

If X and Y are independent, then the following hold:

- If g and h are real-valued functions, then g(X) and h(Y) are independent.
- $\mathbf{E}[XY] = \mathbf{E}[X]\mathbf{E}[Y]$ (inverse is not true)
- var(X + Y) = var(X) + var(Y)

Given independent random variables X_1, \ldots, X_n ,

$$var(X_1+X_2+\cdots+X_n) = var(X_1)+var(X_2)+\cdots+var(X_n)$$

Some Discrete Distributions

	X	$p_X(k)$	E [X]	var(X)
Bernoulli	$\left\{ egin{array}{ll} 1 & {\sf success} \ 0 & {\sf failure} \end{array} ight.$	$\begin{cases} p & k=1\\ 1-p & k=0 \end{cases}$	р	p(1-p)
Binomial	Number of successes in n Bernoulli trials	$\binom{\binom{n}{k}p^k(1-p)^{n-k}}{k=0,1,\ldots,n}$	np	np(1-p)
Geometric	Number of trials until first success	$(1-p)^{k-1}p$ $k=1,2,\dots$	$\frac{1}{p}$	$\frac{1-p}{p^2}$
Uniform	An integer in the interval [a,b]	$\begin{cases} \frac{1}{b-a+1} & k = a, \dots, b \\ 0 & \text{otherwise} \end{cases}$	<u>a+b</u> 2	(b-a)(b-a+2) 12