#### L13 p. 2

### **LECTURE 13**

• Readings: Section 6.1

# Lecture outline

- Random processes
- Bernoulli process
- Definition and basic properties
- Interarrival times
- Distribution of kth arrival
- Merging and splitting

L13 p. 3

# The Bernoulli process

- ullet A sequence of independent Bernoulli trials  $(X_1,\,X_2,\,X_3,\,\ldots)$
- At each trial (each  $i \in \{1, 2, \ldots\}$ ):
- $P(success) = P(X_i = 1) = p$
- $P(failure) = P(X_i = 0) = 1 p$
- Examples:
- Sequence of coin tosses
- Sequence of lottery wins/losses
- Sequence of ups and downs of the Dow Jones
- Arrivals of tasks to computer (in time slots)

- Ch. 7: Markov processes, discrete time

- rarely before had infinite collections

interpret the index as time

Third quarter of course:

• 
$$E[X_i] = var(X_i) =$$

 $\bullet$  Let S be the number of successes/arrivals up to and including time n

$$p_S(k) = E[S] = var(S) =$$

Basic properties

Random processes

 A discrete-time random process is a collection of random variables (defined in the same probabilistic model), e.g.,

- focus is often on dependencies and long-term behavior

- Ch. 6: memoryless processes, discrete and continuous time

 $(X_1, X_2, X_3, \ldots), (X_0, X_1, X_2, \ldots), \text{ or } (\ldots, X_{-1}, X_0, X_1, \ldots)$ 

 Let T be the number of trials up to and including the first success/arrival

$$p_T(t)$$
 =  $E[T]$  =  $var(T)$  =

•  $P(X_i = 1 \text{ for all } i) =$ 

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# Independence and memorylessness

- Trials at disjoint sets of times are independent
- Example:  $X_4, X_5, X_6, \ldots$  is independent of  $X_1, X_2, X_3$
- From any time, the future is independent of the past
- Any fixed reindexing gives a Bernoulli process, e.g.,

$$(X_4, X_5, X_6, \ldots)$$
 or  $(X_1, X_3, X_5, \ldots)$ 

• Memoryless property

$$\mathbf{P}(T - n = t \mid T > n) =$$

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L13 p. 4

# Arrival and interarrival times

- $\bullet \ \ \mbox{Let} \ Y_k$  be the time of the  $k{\rm th}$  arrival
- Increments

$$T_1 = Y_1,$$
  $T_k = Y_k - Y_{k-1},$   $k = 2, 3, ...$ 

are called interarrival times

- Distribution of  $T_1$
- Distribution of other  $T_k \mathbf{s}$
- After a rainy day, the number of days until it rains again is geometrically distributed with parameter p, independent of the past. What is the probability that it rains on April 3 and April 7?

#### L13 p. 8

# Time of the kth arrival

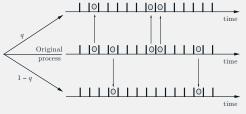
•  $Y_k$ : number of trials to kth success

$$- E[Y_k] =$$

$$- var(Y_k) =$$

$$- P(Y_k = t) =$$

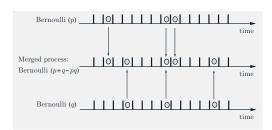
# Splitting of a Bernoulli process (using independent coin flips)



yields Bernoulli processes

# L13 p. 9

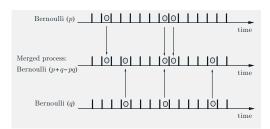
# Merging of independent Bernoulli processes



yields a Bernoulli process

L13 p. 10

# Refining the time scale



- "Collisions" are a limitation in the model:
  - cannot tell the difference between 1 and 2 arrivals
- Shorter time slots decreases p, can make  $p^2 \ll p$

L13 p. 11

# Poisson approximation to binomial

• Number of arrivals in 
$$n$$
 slots is binomial 
$$p_S(k) \ = \ \frac{n!}{(n-k)!k!} \cdot p^k (1-p)^{n-k}, \qquad \text{for } k \geq 0$$

• Interesting to think of  $n\to\infty$  with  $\lambda=np$  constant

$$p_S(k) = \frac{n!}{(n-k)!k!} \cdot p^k (1-p)^{n-k}$$

$$= \frac{n(n-1)\cdots(n-k+1)}{k!} \cdot \frac{\lambda^k}{n^k} \cdot \left(1-\frac{\lambda}{n}\right)^{n-k}$$

$$= \frac{n}{n} \cdot \frac{n-1}{n} \cdot \frac{n-k+1}{n} \cdot \frac{\lambda^k}{k!} \cdot \left(1-\frac{\lambda}{n}\right)^{n-k}$$
The proof fixed  $k > 0$ ,  $\lim_{n \to \infty} (1-x)^{n-k} = e^{-\lambda}$  for

• For any fixed  $k \ge 0$ ,  $\lim_{n \to \infty} (1 - \lambda/n)^{n-k} = e^{-\lambda}$  so

$$\lim_{n \to \infty} p_S(k) = \frac{\lambda^k}{k!} e^{-\lambda}$$