2010 Spring Quiz2 Study Notes by Joongwoo Brian Park Send error report to jwbpark@mit.edu

 $\frac{ \text{Probability Density Functions (PDF)}}{ \text{For a continuous RV X with PDF $f_X(x)$}} (\geq 0),$

$$\begin{array}{lcl} P(a \leq X \leq b) & = & \int_a^b f_X(x) dx \\ P(x \leq X \leq x + \delta) & \approx & f_X(x) . \delta \\ \\ P(X \in A) & = & \int_A f_X(x) dx \end{array}$$

Remarks: - if X is continuous, $P(X=x)=0 \quad \forall x!!$ - $f_X(x)$ may take values larger than 1. Normalization property:

$$\int_{-\infty}^{\infty} f_X(x) dx = 1$$

2. Mean and variance of a continuous RV

$$\begin{split} E[X] &= \int_{-\infty}^{\infty} x f_X(x) dx \\ E[g(X)] &= \int_{-\infty}^{\infty} g(x) f_X(x) dx \\ \operatorname{Var}(X) &= \int_{-\infty}^{\infty} (x - E[X])^2 f_X(x) dx \\ &= E[X^2] - (E[X])^2 \ (\geq 0) \\ E[aX + b] &= aE[X] + b \\ \operatorname{Var}(aX + b) &= a^2 \operatorname{Var}(X) \end{split}$$

3. Cumulative Distribution Functions

Definition:

$$F_X(x) = P(X \le x)$$

monotonically increasing from 0 (at $-\infty$) to 1 (at $+\infty$).

$$F_X(x) = P(X \le x) = \int_{-\infty}^x f_X(t) dt \ \ ({\rm continuous})$$

$$f_X(x) = \frac{dF_X}{dx}(x)$$

• Discrete RV:

$$F_X(x) = P(X \le x) = \sum_{k \le x} p_X(k) \ \ (\text{piecewise constant})$$

$$p_X(k) = F_X(k) - F_X(k-1)$$

4. Normal/Gaussian Random Variables Standard Normal RV: N(0, 1):

$$\begin{split} f_X(x) &=& \frac{1}{\sqrt{2\pi}}e^{-x^2/2} \\ E[X] &=& 0, \quad \operatorname{Var}(X) = 1 \end{split}$$

General normal RV: $N(\mu, \sigma^2)$:

$$f_X(x) = \frac{1}{\sigma\sqrt{2\pi}}e^{-(x-\mu)^2/2\sigma^2}$$

- if Y = aX + b, then $Y \sim N(a\mu + b, a^2\sigma^2)$.
- CDF for standard normal $\phi(.)$ can be read in a table.
- To evaluate CDF of a general standard normal, express it as a function of a standard normal:

$$X \sim N(\mu, \sigma^2) \Leftrightarrow \frac{X - \mu}{\sigma} \sim N(0, 1)$$

$$P(X \leq x) = P\Big(\frac{X - \mu}{\sigma} \leq \frac{x - \mu}{\sigma}\Big) = \phi\Big(\frac{x - \mu}{\sigma}\Big)$$

where $\phi(.)$ denotes the CDF of a standard normal

5. Joint PDF

Joint PDF of two continuous RV X and Y: $f_{X,Y}(x,y)$.

$$\begin{split} &P(x \leq X \leq x + \delta, \ y \leq Y \leq y + \delta) \approx f_{X,Y}(x,y).\delta^2 \\ &P(A) = \int \int_A f_{X,Y}(x,y) dx dy \\ &E[g(X,Y)] = \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} g(x,y) f_{X,Y}(x,y) dx dy \\ &f_X(x) = \int_{-\infty}^{\infty} f_{X,Y}(x,y) dy \end{split}$$

By definition.

$$X,\ Y\ \text{independent} \Leftrightarrow f_{X,Y}(x,y) = f_X(x)f_Y(y)$$

6. Conditioning on an event

X a continuous RV, A a subset of the real line

$$\begin{array}{rcl} f_{X\mid A}(x) & = & \left\{ \begin{array}{l} \frac{f_X(x)}{P(X\in A)} & \text{if } x\in A \\ 0 & \text{otherwise} \end{array} \right. \\ P(X\in B\mid X\in A) & = & \int_B f_{X\mid A}(x) dx \\ & E[X\mid A] & = & \int_{-\infty}^\infty x f_{X\mid A}(x) dx \\ & E[g(X)\mid A] & = & \int_{-\infty}^\infty g(x) f_{X\mid A}(x) dx \end{array}$$

If A_1, \ldots, A_n are disjoint events that form a partition of the sample space,

$$\begin{array}{lcl} f_X(x) & = & \displaystyle \sum_{i=1}^n P(A_i) f_{X \mid A_i}(x) \ (\approx \ {\rm total \ probability \ theorem}) \\ \\ E[X] & = & \displaystyle \sum_{i=1}^n P(A_i) E[X \mid A_i] \ ({\rm total \ expectation \ theorem}) \\ \\ E[g(X)] & = & \displaystyle \sum_{i=1}^n P(A_i) E[g(X) \mid A_i] \end{array}$$

7. Conditioning on a RV

$$\begin{split} &P(x \leq X \leq x + \delta | Y \approx y) \approx f_{X|Y}(x|y).\delta \\ &f_{X|Y}(x|y) = \frac{f_{X,Y}(x,y)}{f_{Y}(y)} \\ &f_{X}(x) = \int_{-\infty}^{\infty} f_{Y}(y) f_{X|Y}(x|y) dy \\ &P(A) = \int_{-\infty}^{\infty} P(A|X = x) f_{X}(x) dx \end{split}$$

$$\begin{split} E[Y|X=x] &= \int_{-\infty}^{\infty} y f_{Y|X}(y|x) dy \\ E[g(Y)|X=x] &= \int_{-\infty}^{\infty} g(y) f_{Y|X}(y|x) dy \\ E[g(X,Y)|X=x] &= \int_{-\infty}^{\infty} g(x,y) f_{Y|X}(y|x) dy \\ E[Y] &= \int_{-\infty}^{\infty} E[Y|X=x] f_X(x) dx \\ E[g(Y)] &= \int_{-\infty}^{\infty} E[g(Y)|X=x] f_X(x) dx \\ E[g(X,Y)] &= \int_{-\infty}^{\infty} E[g(X,Y)|X=x] f_X(x) dx \end{split}$$

8. Continuous Bayes' Rule X, Y continuous RV, N discrete RV, A an event.

$$\begin{split} f_{X|Y}(x|y) &= \frac{f_{Y|X}(y|x)f_{X}(x)}{f_{Y}(y)} = \frac{f_{Y|X}(y|x)f_{X}(x)}{\int_{-\infty}^{\infty} f_{Y|X}(y|t)f_{X}(t)dt} \\ P(A|Y=y) &= \frac{P(A)f_{Y|A}(y)}{f_{Y}(y)} = \frac{P(A)f_{Y|A}(y)}{f_{Y|A}(y)P(A) + f_{Y|A}c(y)P(A^{c})} \\ P(N=n|Y=y) &= \frac{p_{N}(n)f_{Y|N}(y|n)}{f_{Y}(y)} = \frac{p_{N}(n)f_{Y|N}(y|n)}{\sum_{i} p_{N}(i)f_{Y|N}(y|i)} \end{split}$$

9. Independence of continuous RV

$$\begin{array}{lll} X, \ Y \ \mathrm{independent} & \Leftrightarrow & f_{X \mid Y}(x \mid y) = f_X(x) \\ \\ \Rightarrow & g(X), \ h(Y) \ \mathrm{independent} \\ \\ \Rightarrow & E[XY] = E[X]E[Y] \\ \\ \Rightarrow & E[g(X)h(Y)] = E[g(X)]E[h(Y)] \\ \\ \Rightarrow & \mathrm{Var}(X + Y) = \mathrm{Var}(X) + \mathrm{Var}(Y) \end{array}$$

- 10. Derived distributions Def: PDF of a function of a RV X with known PDF: Y = g(X). Method:

$$F_{Y}(y) = P(Y \leq y) = P(g(X) \leq y) = \int_{x \mid g\left(x\right) \leq y} f_{X}(x) dx$$

• Differentiate: $f_Y(y) = \frac{dF_Y}{dx}(y)$

11. Convolution

W = X + Y, with X, Y independent.

Discrete case:

$$p_{W}(w) = \sum_{x} p_{X}(x) p_{Y}(w - x)$$

• Continuous case

$$f_W(w) = \int_{-\infty}^{\infty} f_X(x) f_Y(w-x) dx$$

Mechanics

- put the PMFs (or PDFs) on top of each other
- flip the PMF (or PDF) of Y
- shift the flipped PMF (or PDF) of Y by w
- cross-multiply and add (or evaluate the integral)

In particular, if X, Y are independent and normal, then

• W = X + Y is normal

• $f_{X|W}(x|w)$ is a normal PDF for any given w.

12. Law of iterated expectations

E[X|Y] is a random variable that is a function of Y (the expectation is taken with respect to X). To compute E[X|Y], first express E[X|Y=y] as a function of y. Law of iterated expectations:

$$E[X] = E[E[X|Y]]$$

(equality between two real numbers)

13. Law of conditional variances Var(X|Y) is a random variable that is a

(the variance is taken with respect to X). To compute $\operatorname{Var}(X|Y)$, first express

$$\mathrm{Var}(X|Y=y) = E[(X-E[X|Y=y])^2|Y=y]$$

as a function of y. Law of conditional variances:

$$Var(X) = E[Var(X|Y)] + Var(E[X|Y])$$

(equality between two real numbers)

14. Sum of a random number of iid RVs

N discrete RV, \boldsymbol{X}_i i.i.d and independent of N. $\boldsymbol{Y} = \boldsymbol{X}_1 + \ldots + \boldsymbol{X}_N.$ Then:

$$\begin{split} E[Y] &= E[X]E[N] \\ \mathrm{Var}(Y) &= E[N]\mathrm{Var}(X) + (E[X])^2\mathrm{Var}(N) \end{split}$$

15. Least square prediction

Goal: estimate RV X with a real number c. By definition of least square prediction, the best estimator minimizes the mean square error $E[(X-c)^2]$. Best estimator in the absence of information: c=E[X].

Corresponding mean square error: Var(X).

16. Covariance and Correlation

$$\begin{array}{lcl} \mathrm{Cov}(X,Y) & = & E[(X-E[X])(Y-E[Y])] \\ \\ & = & E[XY]-E[X]E[Y] \end{array}$$

Correlation: (has no dimension)

$$\rho = \frac{\mathrm{Cov}(X,Y)}{\sigma_X \sigma_Y} \quad \in [-1,1]$$

By definition, X, Y are uncorrelated if and only if Cov(X, Y) = 0.

Remark: X, Y independent $\Rightarrow \operatorname{Cov}(X, Y) = 0$ (the converse is not true)

17. Uniform continuous RV over [a, b]

$$f_X(x) = \left\{ \begin{array}{ll} \frac{1}{b-a} & \text{ if } a \leq x \leq b \\ 0 & \text{ otherwise} \end{array} \right.$$

$$F_{X}(x) = \left\{ \begin{array}{ll} 0 & \text{if } x \leq a \\ \frac{x-a}{b-a} & \text{if } a \leq x \leq b \\ 1 & \text{otherwise } (x > b) \end{array} \right.$$

$$E[X] = \frac{a+b}{2}$$
 $Var(X) = \frac{(b-a)^2}{12}$

18. Exponential RV with parameter λ

$$f_{X}\left(x\right)=\left\{ \begin{array}{cc} \lambda e^{-\lambda x} & \text{ if } x\geq 0 \\ 0 & \text{ otherwise} \end{array} \right.$$

$$F_{X}(x) = \left\{ \begin{array}{ll} 1 - e^{-\lambda x} & \text{ if } x \geq 0 \\ 0 & \text{ otherwise} \end{array} \right.$$

$$E[X] = \frac{1}{\lambda} \quad Var(X) = \frac{1}{\lambda^2}$$

19. Normal RV with parameters (μ, σ^2)

$$f_X(x) = \frac{1}{\sigma\sqrt{2\pi}}e^{-(x-\mu)^2/2\sigma^2}$$

$$E[X] = \mu$$

 $Var(X) = \sigma^2$

20. Bernoulli Process Bernoulli process is a sequence $X_1,X_2\dots$ of independent Bernoulli random variables with

$$\begin{array}{lcl} P(X_i=1) & = & p \\ P(X_i=0) & = & 1-p \end{array}$$

<u>Memoryless property</u> For any given time n, the sequence $X_{n+1}, X_{n+2} \cdots$ is also a Bernoulli process, and is **independent** from $X_1, X_2 \cdots X_n$.

22. $\underline{\mathbf{Fresh\text{-}Start}}$ Every arrival restarts the process

23. Important RV associated with Bernoulli Processes

• First arrival : The time to first arrival (T) is a geometric RV

$$p_T(t) = (1-p)^{t-1}p, t = 1, 2...$$

• Number of arrivals: The number of arrivals (K) in n trials is a binomial RV

$$p_K(k) = \binom{n}{k} p^k (1-p)^{n-k}, k = 0, 1 \dots n$$

Note: (n-fixed,k-random)

ullet K^{th} arrival: The time to the K^{th} arrival Y_K is a Pascal RV

$$p_{Y_K}(t) = {t-1 \choose k-1} p^k (1-p)^{(t-k)}$$

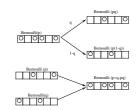
Note: (k-fixed,t-random)

24. Alternate description of the Bernoulli Process

- $\bullet~$ Start with a sequence of independent geometric RVs T_1 , T_2 . . . , with common parameter p.
- Record success(arrival) at times, T_1 , $T_1 + T_2$, $T_1 + T_2 + T_3$
- ullet K^{th} arrival time Y_k is the sum of the first k inter-arrival times

$$\begin{array}{rcl} Y_k & = & T_1 + T_2 \dots T_k \\ \\ [Y_k] & = & [T_1 + T_2 \dots T_k] = \frac{k}{p} \\ \\ (Y_k) & = & (T_1 + T_2 \dots T_k) = \frac{k(1-p)}{p^2} \end{array}$$

25. Splitting and Merging of Bernoulli Processes



- If arrivals from a Bernoulli process are split into two processes with probability q and (1-q), each process is an **independent** Bernoulli process with parameters pq and p(1-q)
- Conversely, if we merge two Bernoulli processes with parameters p and q, we get an **independent** Bernoulli process with parameter