

WEIGHTED LEAST-SQUARES

1

Consider the model
general linear

$$Y_i = X_i \beta + \epsilon_i$$

$\epsilon_i \sim \text{normal}$

(*)

where $E[\epsilon_i] = 0$ for all i and $\text{Var}(\epsilon_i) = \sigma_i^2$,
and the σ_i^2 can be different. Really, what
we mean here is

$$E[\epsilon_i | X_i] = 0 \text{ and } \text{Var}(\epsilon_i | X_i) = \sigma_i^2,$$

so the variances of the errors can depend on
values of the predictor variables. For now
we assume the covariances are 0:

$$\text{Cov}(\epsilon_i, \epsilon_j | X_i, X_j) = 0 \text{ for } i \neq j.$$

Further, let W denote the inverse of the
covariance matrix. Also define

Note

$$W^{-1} = \begin{bmatrix} \frac{1}{w_1} & 0 & 0 & \dots & 0 \\ 0 & \frac{1}{w_2} & 0 & \dots & 0 \\ 0 & 0 & \frac{1}{w_3} & \dots & 0 \\ \vdots & \vdots & \vdots & \ddots & \vdots \\ 0 & 0 & 0 & \dots & \frac{1}{w_n} \end{bmatrix}$$

$$W^{1/2} := \begin{bmatrix} \sqrt{w_1} & 0 & 0 & \dots & 0 \\ 0 & \sqrt{w_2} & 0 & \dots & 0 \\ 0 & 0 & \sqrt{w_3} & \dots & 0 \\ \vdots & \vdots & \vdots & \ddots & \vdots \\ 0 & 0 & 0 & \dots & \sqrt{w_n} \end{bmatrix}$$

Here

$$W = \begin{bmatrix} w_1 & 0 & 0 & \dots & 0 \\ 0 & w_2 & 0 & \dots & 0 \\ 0 & 0 & w_3 & \dots & 0 \\ \vdots & \vdots & \vdots & \ddots & \vdots \\ 0 & 0 & 0 & \dots & w_n \end{bmatrix}$$

is KNOWN

Note $W^{1/2}$ is symmetric and $W^{1/2} W^{1/2} = W$. Also,

~~$(W^{1/2})^{-1} (W^{1/2})^{-1} = W^{-1}$~~

$(W^{-1})^{1/2} (W^{-1})^{1/2} = W^{-1}$, and which also equals $(W^{1/2})^{-1} (W^{1/2})^{-1} = W^{-1}$

So we just write $W^{-1/2} = (W^{1/2})^{-1}$

Now vectorize (*) and left-multiply by $W^{1/2}$:

2

(**)

$$\underbrace{W^{1/2} Y}_{Y^*} = \underbrace{W^{1/2} X}_{X^*} \beta + \underbrace{W^{1/2} \epsilon}_{\epsilon^*}$$

So that $Y^* = X^* \beta + \epsilon^*$. Now

$$E[\epsilon^*] = E[W^{1/2} \epsilon] = W^{1/2} E[\epsilon] = 0$$

↑ MATRIX OF KNOWN CONSTANTS

$$\underbrace{\sigma^2(\epsilon^*)}_{\text{Covariance Matrix of the } n \times 1 \text{ errors}} = W^{1/2} \underbrace{\sigma^2(\epsilon)}_{\text{Covariance Matrix of original errors}} W^{1/2} = W^{1/2} W^{-1} W^{1/2}$$

$$= \underbrace{W^{1/2} W^{-1/2}}_{= I_{n \times n}} \underbrace{W^{-1/2} W^{1/2}}_{= I_{n \times n}}$$

Remember $\sigma^2(\epsilon^*)$ and $\sigma^2(\epsilon)$ are $n \times n$, $n = \#$ of observations

Thus, the model in (**) satisfies the mean = 0 and const. error variance assumptions for the generalized linear model (also, normality of the ϵ^* is inherited from the normality of the ϵ). Thus

$$b^* = (X^{*T} X^*)^{-1} X^{*T} Y^*$$

contain the least-squares estimators for the β vector.

We can express this as

$$b^* = \left[\underbrace{(W^{1/2} X)^T}_{= X^T W} \underbrace{(W^{1/2} X)}_{= X^T W} \right]^{-1} \underbrace{(W^{1/2} X)^T W^{1/2} Y}_{= Y^*}$$

③

$$= [X^T W X]^{-1} X^T W Y.$$

What if (and this is usually the case) the w -values are unknown? How should we estimate them?? Either estimate the standard deviation

function with \hat{s}_i or the variance function with \hat{v}_i , and set

$$w_i = \frac{1}{\hat{s}_i^2}$$

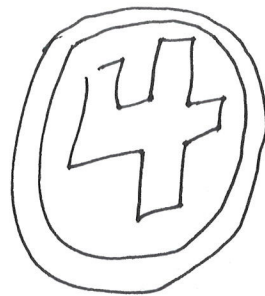
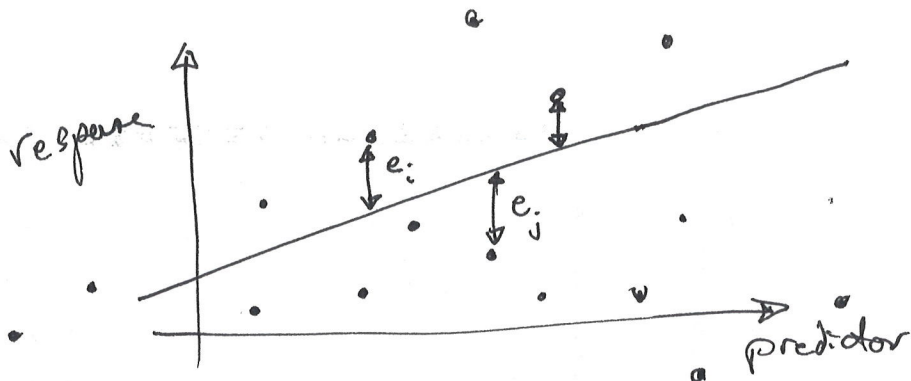
or

$$w_i = \frac{1}{\hat{v}_i}$$

(these methods are NOT the same...)

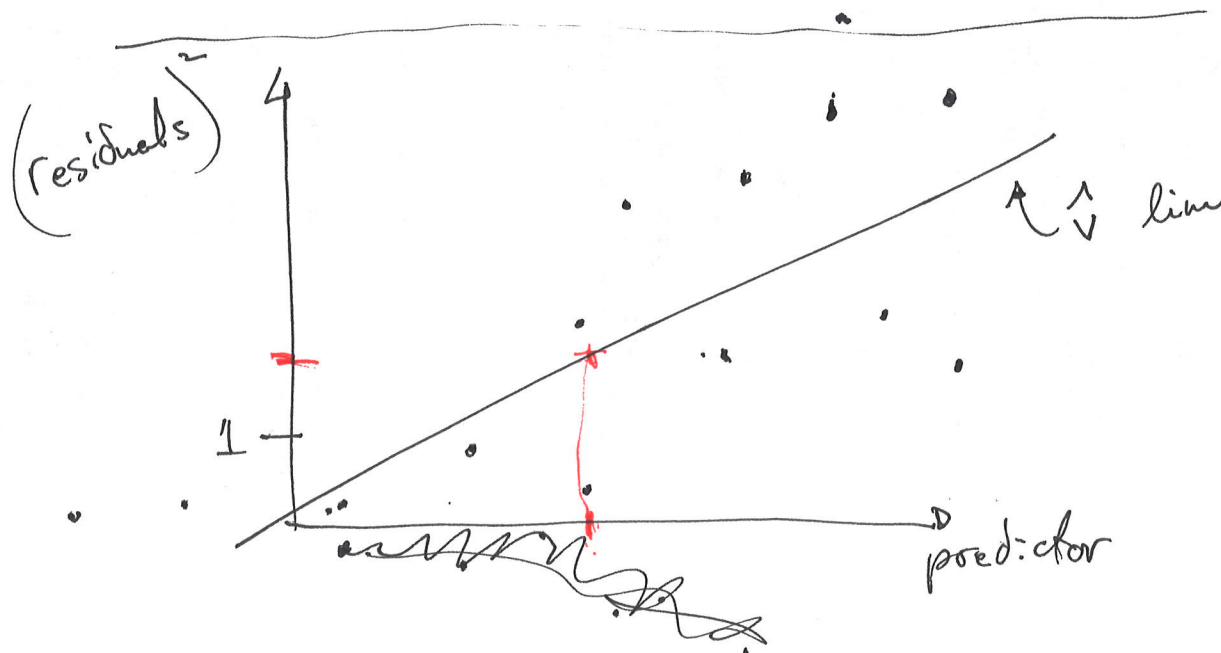
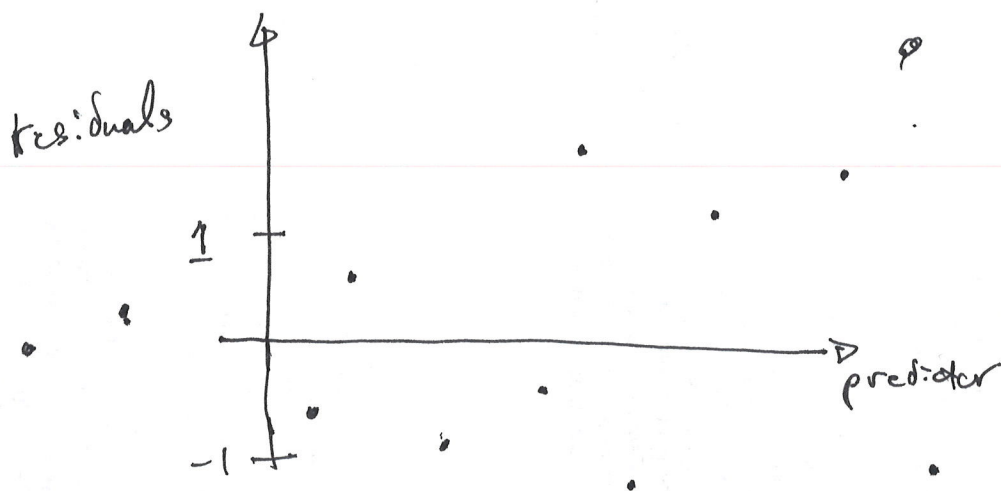
How to do this?

- ① Fit regression model with unweighted least-squares (the usual way: $\hat{\beta} = (X^T X)^{-1} X^T Y$).
- ② Estimate the variance function w/ \hat{v} or the std. dev. function \hat{s} with the fits of a least squares regression of the squared errors or errors against the predictor variables of interest.
- ③ Use the fits from ② to get weights.
- ④ Use *** and weights from ③ to get b^* .
- ⑤ Can iterate... start back @ ① with $\hat{\beta}_{new} = b^*$, get new residuals for ②, etc.



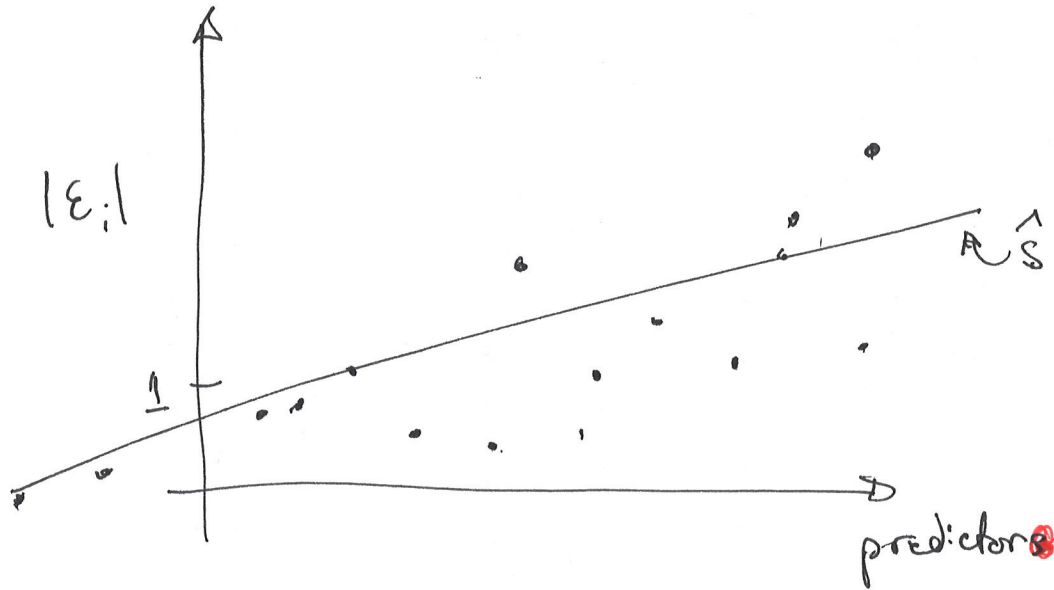
Recall $E[e_i] = 0$, and $\text{Var}(e_i) = E[e_i^2] - \underbrace{(E[e_i])^2}_{=0} = E[e_i^2]$.

This is why we can ~~keep estimate~~ think to model the variance of the errors from the squared residuals vs. predictors of interest.



or, can model (w/ least-squares regression) the
std. dev. ~~var~~ of least-squares reg. line, regressing
 $|e_i|$ vs. predictor(s):

(5)



Use $|e_i|$ vs. preds or e_i^2 vs. preds as appropriate.

Other than $w_i = \frac{1}{\hat{s}_i^2}$ or $w_i = \frac{1}{v_i}$ you can try
other things: If the response ~~is~~ ^{y_i is an} average of n_i observations,
consider $w_i = n_i$. An alternative to these ...
could try $w_i = \frac{1}{\hat{s}_i}$. Try something that ~~looks~~
comes from a good linear fit like this.

If two or more predictors (say x_1, x_2) contribute to the
heteroskedasticity, try regressing e_i^2 or $|e_i|$ onto $x_1, x_2, x_1 x_2$,
and maybe x_1^2 and x_2^2 . Then build a regression model, get the fits, ...