

Topic 11: Limit Theorems

Outline

- Markov Inequality
- Chebyshev Inequality
- Weak Law of Large Number (LLN) 大數法則
- The Central Limit Theorem (CLT) 中央極限定理
- Convergence of a Sequence of RVs 隨機序列的收斂性

Reading:

Textbook Section 5.1~5.5

Introduction

- Why study limit theorems?

- Study the *asymptotic behavior* of sequences of RVs. We'd like to know whether a sequence of RVs converges

(如鐵穹防禦系統要能鎖定並持續追蹤直到擊中目標)

- 大數法則(Law of Large Numbers) 陳述sample mean $\frac{1}{n} \sum_{i=1}^n X_i$ 在樣本

數很大時會收斂到真實的期望值。

應用：可估計群體中某現象存在個數（某事件發生次數）。在群體樣本非常大時，此估計愈準確。

- 我們經常對一連串i.i.d.隨機變數之總和感興趣，如計算個數、多次賽局輸贏總合、長隊伍等待時間等。中央極限定理(Central Limit Theorem)宣告我們可借用常態分佈來近似此總和的PDF

Review of Sample Mean

- Law of large numbers and Central Limit Theorem are related to the sample mean. Let's have a quick review.
- X_1, X_2, \dots, X_n are **i.i.d.** RVs with $E[X_i] = \mu$, $\text{var}(X_i) = \sigma^2$

Sample mean: $M_n = (X_1 + X_2 + \dots + X_n) / n$

Then,

$$E[M_n] = \mu; \quad \text{var}(M_n) = \frac{\sigma^2}{n}$$

Markov Inequality

If a random variable X can only take **nonnegative** values, then we must have the following Markov inequality

$$P(X \geq a) \leq \frac{E[X]}{a} \quad \text{for all } a > 0.$$

Remark: 僅需知道期望值即可!

Markov inequality provides a loose bound, but is **easy to calculate!**

Example 5.1: Uniform in $[0,4]$ (p.266)

Markov: $\Pr(X \geq 3) \leq \frac{2}{3} = 0.67$; Exact: $\Pr(X \geq 3) = 0.25$

Chebyshev Inequality

If X is a random variable with mean μ and variance σ^2 , then

$$P(|X - \mu| \geq c) \leq \frac{E[(X - \mu)^2]}{c^2} \quad \text{for all } c > 0.$$

Remarks:

- Markov and Chebyshev inequalities allow us to derive bounds on probabilities when only the **mean**, or both the **mean** and the **variance** are available
- In general, Markov and Chebyshev inequality are NOT very tight upper bounds, as the required information is only the mean and the variance. (Examples 5.1 and 5.2)

Chebyshev Inequality

$$P(|X - \mu| \geq c) \leq \frac{E[(X - \mu)^2]}{c^2} \quad \text{for all } c > 0.$$

In general, Markov and Chebyshev inequality are NOT very tight upper bound.

Example 5.2:

Let X be uniform in $[0,4]$. Use the Chebyshev inequality to bound the probability of $|X - 2| \geq 1$.

Chebyshev Inequality Example 1: Weak Law of Large Numbers (LLN)

Let X_1, X_2, \dots, X_n be a sequence of independent and identically distributed (i.i.d.) RVs with mean μ and finite variance σ^2 . For every $\epsilon > 0$ (no matter how small it is), we have the following (weak) law of large numbers

$$P\left(\left|\frac{X_1 + \dots + X_n}{n} - \mu\right| > \epsilon\right) \rightarrow 0, \quad \text{as } n \rightarrow \infty$$

where $M_n = (X_1 + X_2 + \dots + X_n)/n$ is the sample mean.

The weak LLN (大數法則) says that the sample mean will be very close to the true mean, when the sample size is sufficiently large.

(Proof) Use Chebyshev inequality!

$$P(|M_n - \mu| \geq \epsilon) \leq \frac{\text{Var}(M_n)}{\epsilon^2} = \frac{\sigma^2}{n\epsilon^2}$$

Chebyshev Inequality Example: The Pollster's Problem

The Pollster's Problem (民意調查問題)

You may often see something like this, particularly during elections:

「這次選前民意調查，XX的支持率約為52%。此結果有95% 信心水準，抽樣誤差於正負 3% 以內。」

Questions:

- 1) 直覺上，抽樣樣本數越多，則估計出來的支持率越接近真實。但，最少需抽樣幾份，才能保證有上述之抽樣誤差與信心水準？
- 2) 另外一個問法是，民調公司預算僅能允許抽樣1000份。若欲保證預估的支持率可有95%信心水準，那抽樣誤差要如何修正？

The Pollster's Problem

Example 5.5: Accuracy of Polling (p.270) (民調數據的準確率)

- The i th person' polled: $X_i = \begin{cases} 1 & \text{If "Yes"} \\ 0 & \text{If "No"} \end{cases}$
- Fraction of "Yes" : $M_n = \frac{X_1 + X_2 + \cdots + X_n}{n}$
- ◆ p : (unknown) fraction of population voting "Yes": $\mu_x = p$
- ◆ The question here is:
95% 信心水準，抽樣誤差於正負 1% 以內，則需要抽樣幾份？

The Pollster's Problem

- Accuracy of Polling (p.270) (民調數據的準確率)

95% 信心水準，抽樣誤差於正負 1% 以內

$$\Pr(|M_n - \mu_X| \geq 0.01) \leq 0.05$$

- ◆ 使用 Chebyshev: $\Pr(|M_n - \mu_x| \geq \varepsilon) \leq \frac{\sigma_x^2}{n\varepsilon^2} \quad \mu_x = p$

- In fact, X_i is i.i.d. Bernoulli RV: $\mu_x = p$; $\sigma_x^2 = p(1-p) \leq 1/4$

$$\Pr(|M_n - p| \geq 0.01) \leq \frac{1}{0.0004n} \Rightarrow n > 50,000$$

- *Remarks:*

1. Chebyshev 算出的 n 可保證不等式成立，but...
2. CLT can also do the tricks, and give a much smaller n (i.e., fewer samples. See page 14.)

The Central Limit Theorem (CLT)

Central Limit Theorem (中央極限定理)

Let X_1, X_2, \dots, X_n be a sequence of i.i.d. RVs with mean μ and variance σ^2 , and define

$$Z_n \triangleq \frac{X_1 + \dots + X_n - n\mu}{\sigma\sqrt{n}}$$

Then, the CDF of Z_n converges to the standard normal CDF $\Phi(z)$ in the sense that

$$\lim_{n \rightarrow \infty} P(Z_n \leq z) = \Phi(z) \quad \text{for every } z.$$

Remarks:

- Note that X_1, X_2, \dots, X_n can be of **any distributions** (this is why CLT is so **powerful**), and the CLT is also true.

不論 X_1, X_2, \dots, X_n 原本的統計行為如何， Z_n 在 n 很大時會趨近於常態分佈

The Central Limit Theorem (CLT)

Central Limit Theorem (中央極限定理)

Let X_1, X_2, \dots, X_n be a sequence of i.i.d. RVs with mean μ and variance σ^2 , and define

$$Z_n \triangleq \frac{X_1 + \dots + X_n - n\mu}{\sigma\sqrt{n}}$$

Then, the CDF of Z_n converges to the standard normal

- 該如何去理解上式 Z_n 的結構？減去 $n\mu$ 的意義，分母的物理意義各為何？
- CLT well models many practical real-world problems. **Sum** of i.i.d. RVs with arbitrary distributions having finite mean and variance can be approximated by Gaussian
 - ✓ The number of occurrence of a certain event, the sample mean
 - ✓ Noise, channel effects in the received signal of your smartphone

Proof of CLT (p.290)

- Sum of indep. RVs → use moment generating function

- We have:

$$M_{Z_n}(s) = \mathbf{E}[e^{sZ_n}] = \mathbf{E} \left[e^{(s/\sqrt{n})(X_1 + \dots + X_n)} \right]$$

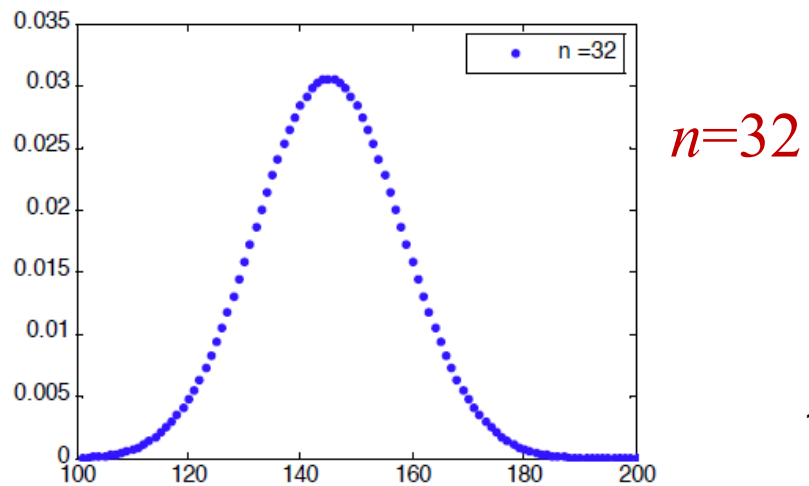
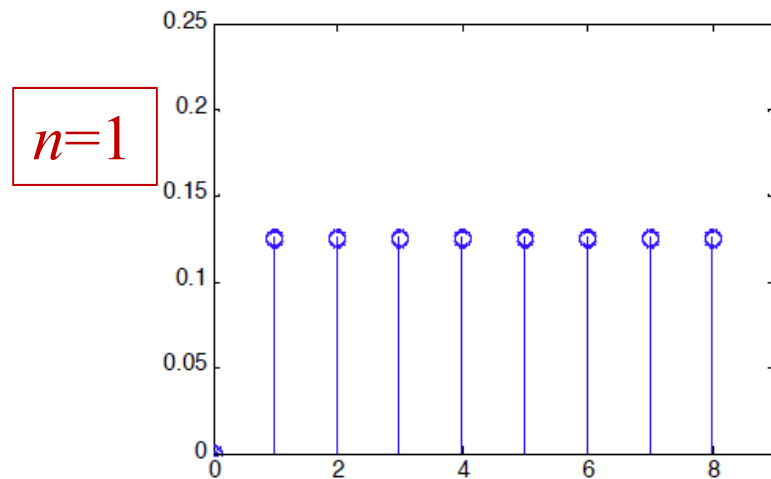
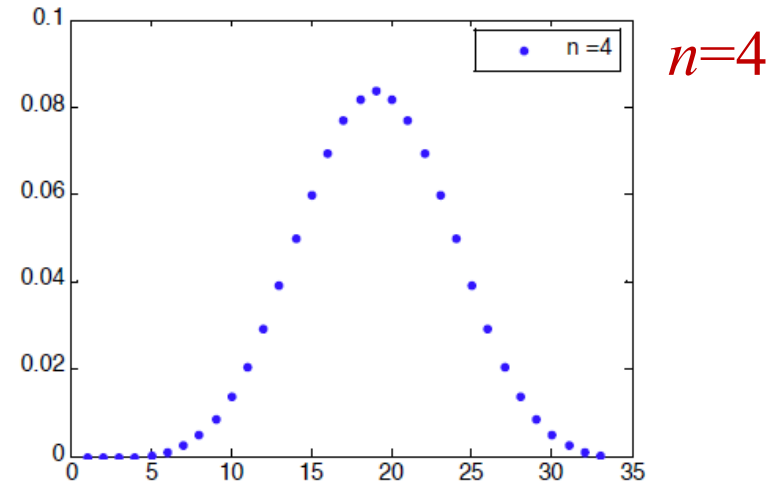
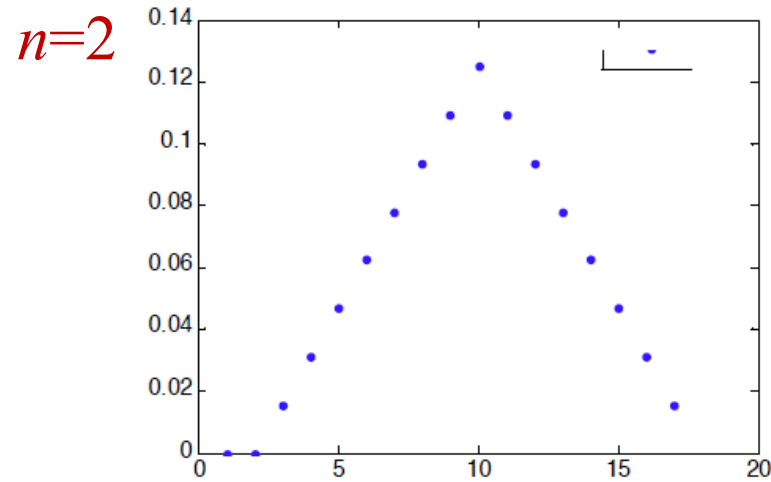
$$\mathbf{E}[e^{sX/\sqrt{n}}] \approx 1 + \frac{s}{\sqrt{n}}\mathbf{E}[X] + \frac{s^2}{2n}\mathbf{E}[X^2]$$

$$M_{Z_n}(s) = (\mathbf{E}[e^{sX/\sqrt{n}}])^n \approx \left(1 + \frac{s^2}{2n} \right)^n \longrightarrow e^{s^2/2}$$

which is the transform of the standard normal.

“N” in CLT

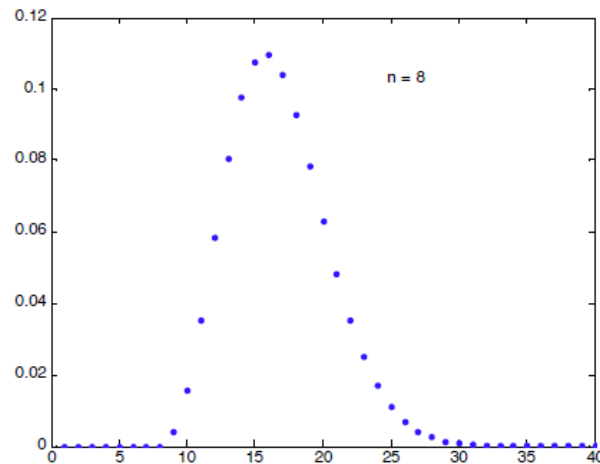
- To make a “good” Gaussian, how large is n ?



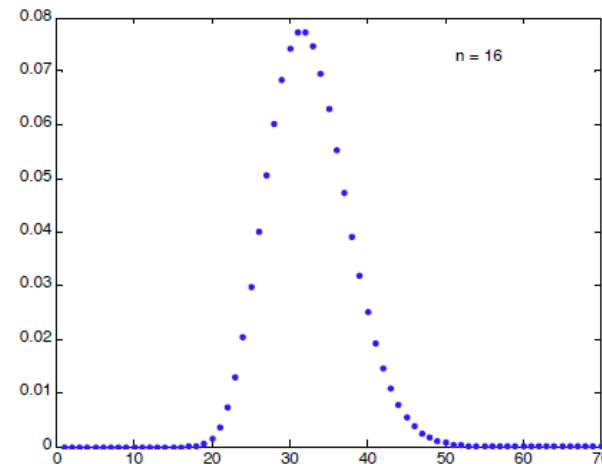
“N” in CLT (2)

- The shape of PDF of X has “some” impact on n .

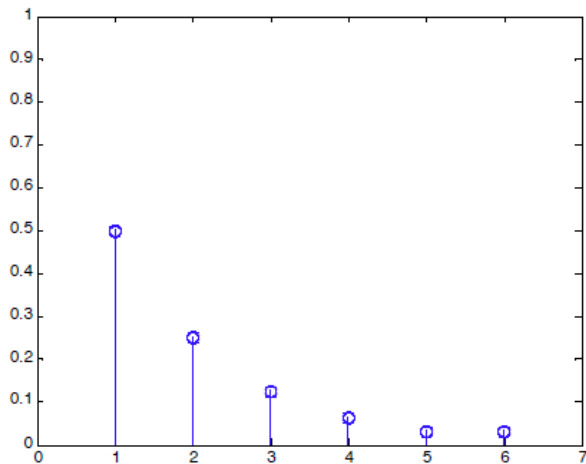
$n=8$



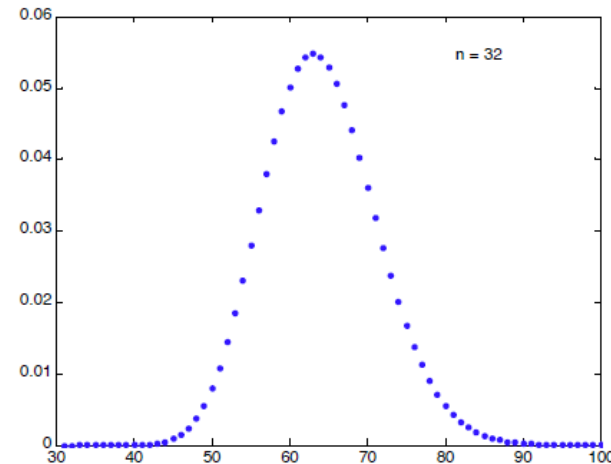
$n=18$



$n=1$



$n=32$



The Pollster's Problem (p.277)

- For the i th person: $X_i = \begin{cases} 1 & \text{If "Yes"} \\ 0 & \text{If "No"} \end{cases}$
- Fraction of "Yes": $M_n = \frac{X_1 + X_2 + \dots + X_n}{n}$ $\sigma_x^2 = p(1-p) \leq \frac{1}{4}$
- ◆ p : fraction of population voting "Yes" = μ_X
- ◆ Suppose we want: **95% 信心水準，抽樣誤差於正負 1% 以內**

$$\Pr(|M_n - \mu_X| \geq 0.01) \leq 0.05$$

- ◆ CLT facilitates calculations of the probability:

$$\left| \frac{X_1 + X_2 + \dots + X_n - np}{n} \right| \geq 0.01 \rightarrow \left| \frac{X_1 + X_2 + \dots + X_n - np}{\sqrt{n}\sigma} \right| \geq 0.01 \frac{\sqrt{n}}{\sigma}$$

$$\Pr(|M_n - \mu_X| \geq 0.01) \approx \Pr(|Z| \geq 0.02\sqrt{n}) \leq 0.05$$

$$2 \cdot 0.01 \cdot \sqrt{n} \geq 1.96 \Rightarrow n > 9604$$

CLT Applied to Binomial

- Bernoulli: $\mu_x = p; \quad \sigma_x^2 = p(1-p)$

- Binomial: $S_n = X_1 + X_2 + \cdots + X_n$
 $\mu_S = np; \quad \sigma_S^2 = np(1-p)$

- Approximation:

CDF of $\frac{S - np}{\sqrt{np(1-p)}} \rightarrow$ standard normal

- *Example:* $n=36, p=0.5$; find $\Pr(S_n=19)$

- Exact answer:

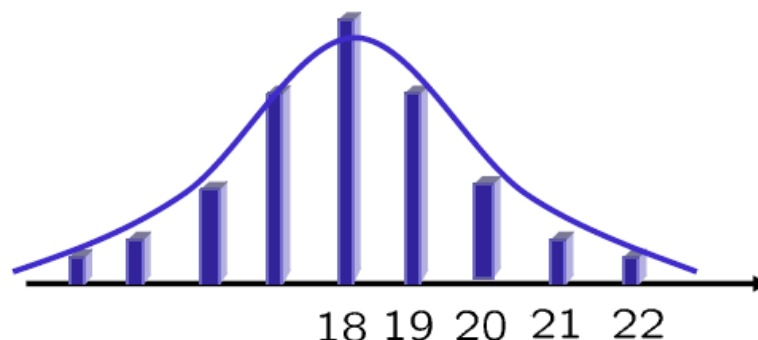
$$\binom{36}{19} \left(\frac{1}{2}\right)^{36} = 0.1251$$

De Moivre-Laplace Approximation to Binomial

- The $\frac{1}{2}$ Correction for Binomial Approximation

Let S_n be normalized to zero mean and unit variance:

$$\Pr(k \leq S_n \leq l) \approx \Phi(l+0.5) - \Phi(k-0.5)$$



■ Ex: $n=36, p=0.5; \Pr(S_n=19) = \Pr(18.5 \leq S_n \leq 19.5)$



$$\begin{aligned} 18.5 \leq S_n \leq 19.5 & \iff \\ \frac{18.5 - 18}{3} \leq \frac{S_n - 18}{3} \leq \frac{19.5 - 18}{3} & \iff \\ 0.17 \leq Z_n \leq 0.5 & \end{aligned}$$

$$\begin{aligned} \Pr(S_n = 19) & \approx \Pr(0.17 \leq Z \leq 0.5) \\ & = \Pr(Z \leq 0.5) - \Pr(Z \leq 0.17) \\ & = 0.6915 - 0.5675 \\ & = 0.124 \end{aligned}$$

Convergence

- **Deterministic limit:** a sequence \rightarrow a number

- a_n converges to a : $\lim_{n \rightarrow \infty} a_n = a$

- a_n eventually gets and stays (arbitrarily) close to a

- Def: For every $\varepsilon > 0$ there exists n_0 , such that for all

- $$n \geq n_0, \quad |a_n - a| \leq \varepsilon.$$

- **Stochastic limits:** there exists several definitions

- Convergence “in probability”

- Convergence “almost surely” or “with probability 1”

Convergence “in Probability”

- Y_1, Y_2, \dots is a sequence of RVs, and let a be a real number. We say Y_n converges to a **in probability**, if **for every** $\varepsilon > 0$, we have

$$\lim_{n \rightarrow \infty} \Pr(|Y_n - a| \geq \varepsilon) = 0$$

- (Almost) all of the PMF/PDF of Y_n eventually gets concentrated (arbitrarily) close to a
- 例子: (大數法則) The weak law of large numbers \rightarrow the **sample mean** converges in prob to the **true mean**.
- **Accuracy level:** ε , **Confidence level:** δ ;

$$\Pr(|Y_n - a| \geq \varepsilon) \leq \delta, \quad \text{for all } n \geq n_0$$

Example (p.272)

Example 5.6:

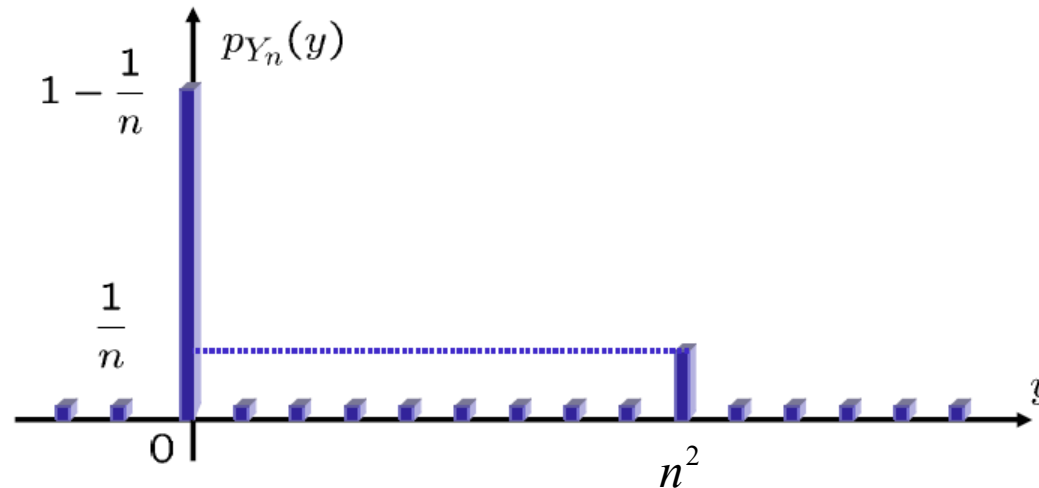
Consider a sequence of independent random variables X_n that are uniform in $[0,1]$. Let

$$Y_n = \min\{X_1, \dots, X_n\}$$

Does Y_n converge to in probability?

Example (p.272)

- Consider a sequence of RVs Y_n with the following PMFs.



- Does Y_n converge? Yes.

$$\lim_{n \rightarrow \infty} \Pr(|Y_n - 0| \geq \varepsilon) = \lim_{n \rightarrow \infty} \frac{1}{n} = 0$$

- What is $E[Y_n]$? “ n ”

$$0 \cdot \left(1 - \frac{1}{n}\right) + n^2 \frac{1}{n} = n$$

Convergence “with Probability 1”

- Y_1, Y_2, \dots is a sequence of RVs, and let c be a real number. We say Y_n converges to c **with probability 1**, if

$$\Pr(\lim_{n \rightarrow \infty} Y_n = c) = 1$$

- Remarks:

➤ $\left\{ \lim_{n \rightarrow \infty} Y_n = c \right\}$ 寫得精確一些應為 $\left\{ \omega: \lim_{n \rightarrow \infty} Y_n(\omega) = c \right\}$

大括號表示一個事件，此事件發生機率為 1 時稱 Y_n converges to c **with probability 1**

- 隨機變數所構成的序列，其收斂的不同定義會在研究所“隨機過程”課程中有更詳細解說

Example -- Convergence with Probability 1

Example

Consider a sample space $\Omega = [0,1]$, the closed interval between 0 and 1. Assume that each sample point $\epsilon \in \Omega$ follows continuous uniform distribution. Then, the random sequence $X_n(\epsilon) = \exp(-n^2(\epsilon - 1/n))$ converges to 0 with probability 1.

Example

Example 5.14

Consider a sequence of independent random variables X_n that are uniform in $[0,1]$. Let

$$Y_n = \min\{X_1, \dots, X_n\}$$

Show that Y_n converge to 0 with probability 1.

Convergence “with Probability 1”

- Y_1, Y_2, \dots is a sequence of RVs, and let c be a real number. We say Y_n converges to c **with probability 1**, if

$$\Pr(\lim_{n \rightarrow \infty} Y_n = c) = 1$$

- Remarks

- Convergence with probability 1 定義中有 傳統數列收斂 的概念(see p.19)。
樣本空間中若 幾乎 每一個 sample point 對應到的數列都會收斂，僅部分的 sample points 對應到的數列不收斂，若那些對應到不收斂數列的 sample points 所成集合發生機率是零，即為此處所謂收斂 **with probability 1** 或稱為 **convergence almost surely**
- Convergence with probability 1 implies convergence in probability. The converse is not necessarily true. (Problem 15*, p.292)
Example 5.15 (p.282) says a random sequence may converge **in probability**, but not **with probability 1**

Example

Example 5.15

Consider a discrete time arrival process. The set of times is $I_k = \{2^k, 2^k + 1, \dots, 2^{k+1} - 1\}$ for $k=1,2,\dots$. During each I_k , there is exactly one arrival, and all times within an interval are equally likely. The arrival times in different intervals are assumed to be independent. Define $Y_n=1$ if there is an arrival at time n , and $Y_n=0$ if there is no arrival.

Show that Y_n converge to 0 in probability, but NOT with probability 1

The Strong Law of Large Numbers

- Let X_1, X_2, \dots be a sequence of i.i.d. RVs with mean μ .

$$\text{Let } M_n = (X_1 + X_2 + \dots + X_n) / n,$$

$$\Pr(\lim_{n \rightarrow \infty} M_n = \mu) = 1$$

- Sample space: set A = collections of (x_1, x_2, \dots)

$$\omega = (x_1, x_2, \dots) \text{ such that } \left[\lim_{n \rightarrow \infty} \frac{x_1 + x_2 + \dots + x_n}{n} = \mu \right]$$

→ Prob of $A = 1$.

➤ Weak Law: $\lim_{n \rightarrow \infty} \{ \Pr(|M_n - \mu| \geq \varepsilon) \} = 0$

這是 **convergence in probability** 的概念

➤ Strong Law: For every $\varepsilon > 0$ there exists n_0 ,

$$\Pr(\{ \omega_n : |M_n - \mu| < \varepsilon, n \geq n_0 \}) = 1$$

這是 **convergence with probability 1** 的概念