

Topic 8: Derived Distributions

Lecture Outline

- Functions of One Continuous Random Variable
Given the PDF of X , find the PDF of $Y=g(X)$
- Function of Two Continuous Random Variable $W=g(X,Y)$
Given the PDF of X and Y , find the PDF of $W=g(X,Y)$
- Joint PDF of **Multiple Functions** of Two Random Variables

$$V = g(X, Y)$$

$$W = h(X, Y)$$

References:

1. Textbook: Section 4.1 for functions of one RV
2. H. Stark and J. Woods, ***Probability and Random Process with Applications to Signal Processing***, 3rd ed., 2002, (pp. 152 ~ 161)
3. S. Ghahramani, ***Fundamentals of Probability with Stochastic Processes***, 4th ed., 2019 (Sec. 8.4)

Review

Discrete

$$p_X(x)$$

$$p_{X,Y}(x,y)$$

$$p_{X|Y}(x|y) = \frac{p_{X,Y}(x,y)}{p_Y(y)}$$

$$p_X(x) = \sum_y p_{X,Y}(x,y)$$

$$F_X(x) = \mathbf{P}(X \leq x)$$

$$\mathbf{E}[X], \text{ var}(X)$$

Continuous

$$f_X(x)$$

$$f_{X,Y}(x,y)$$

$$f_{X|Y}(x|y) = \frac{f_{X,Y}(x,y)}{f_Y(y)}$$

$$f_X(x) = \int_{-\infty}^{\infty} f_{X,Y}(x,y) dy$$

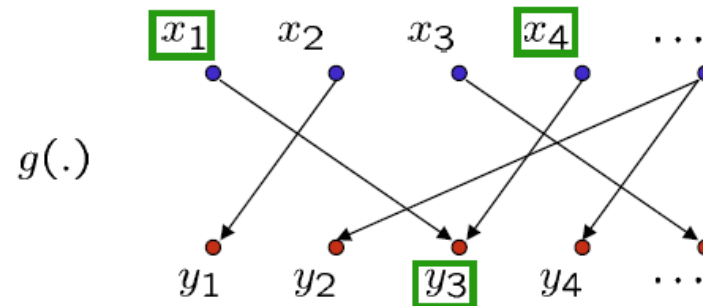
Review: Derived **PMF**

Recall derived **PMF**

Given $p_X(x)$ and a new random variable $Y = g(X)$, derive new PMF $p_Y(y)$ by

$$p_Y(y) = \sum_{x: g(x)=y} p_X(x)$$

Pick up a y value, find all x 's that give $g(x)=y$. We sum over all x 's that give $g(x)=y$.



Note:

The above method cannot be immediately extended to deriving PDFs since X is continuous. PDFs are **not** probabilities, and we cannot add probabilities of $\{X=x\}$ for continuous X

PDF of $Y=g(X)$ for **Continuous** X

Two steps to derive

1. First, calculate the **CDF** $F_Y(y)$ of $Y=g(X)$ using the formula

$$F_Y(y) = P(Y \leq y) = \int_{x:g(x) \leq y} f_X(x) dx$$

2. **Differentiate** to obtain the PDF of Y :

$$f_Y(y) = \frac{dF_Y(y)}{dy}$$

Example 4.3

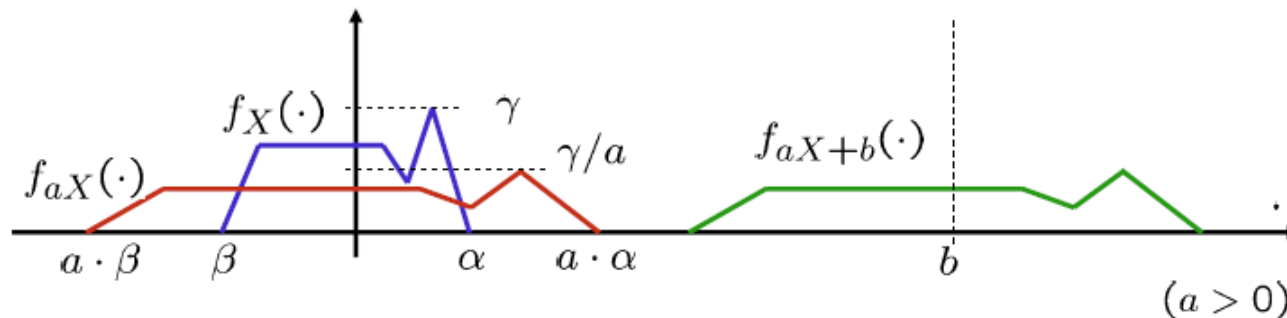
Let $Y = g(X) = X^2$, where X is a uniform random variable over $[-2, 2]$. Find the PDF of Y .

(Sol) $f_Y(y) = \frac{1}{4}y^{-1/2}$ for $0 \leq y \leq 4$

The Linear Case

Let X be a continuous random variable with PDF $f_X(x)$, and let $Y = aX + b$, for some scalars $a \neq 0$ and b . Then,

$$f_Y(y) = \frac{1}{|a|} f_X\left(\frac{y-b}{a}\right)$$



Two steps: **Scaling**: $z=ax$; **Shift**: $y=z+b$

Example:

1. A linear function of a normal random variable is normal
2. A linear function of an exponential random variable is **NOT** necessarily exponential

Example 4.4: A linear function of an exponential random variable

- Let X be an **exponential random variable** with parameter λ .

We have the PDF

$$f_X(x) = \begin{cases} \lambda e^{-\lambda x} & \text{if } x \geq 0, \\ 0 & \text{otherwise,} \end{cases}$$

where λ is a positive parameter. Let $Y = aX + b$. Then,

$$f_Y(y) = \begin{cases} \frac{\lambda}{|a|} e^{-\lambda(y-b)/a} & \text{if } (y-b)/a \geq 0, \\ 0 & \text{otherwise.} \end{cases}$$

Note that if $b = 0$ and $a > 0$, then Y is an exponential random variable with parameter λ/a . In general, however, Y need not be exponential. For example, if $a < 0$ and $b = 0$, then the range of Y is the negative real axis.

Monotonic Function of a Continuous Random Variable

Suppose that g is **strictly monotonic** and that for some function h and all x in the range of X we have

$$y = g(x) \quad \text{if and only if} \quad x = h(y)$$

Assume that h is differentiable. Then, **the PDF of $Y=g(X)$** in the region is given by

$$f_Y(y) = f_X(h(y)) \left| \frac{dh}{dy}(y) \right|$$

This can be justified by first **finding the CDF of Y** , then **taking the derivative**.

- Monotonically increasing: $g(x) < g(x')$ for all x, x' satisfying $x < x'$
- Monotonically decreasing
- 上式絕對值的物理意義？

Illustration (page 210)

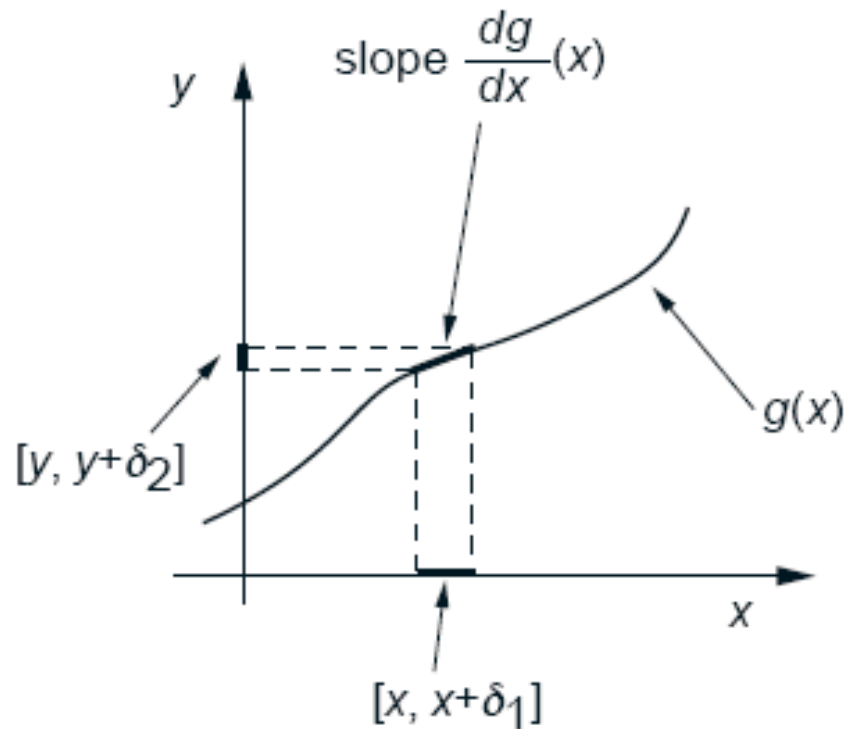


Fig. 4.4: Consider an interval, $[x, x+\delta_1]$, where δ_1 is a small number. Under the mapping g , the image of this interval is another interval $[y, y+\delta_2]$. Since (dg/dx) is the slope

of g , we have $\frac{\delta_2}{\delta_1} \approx \frac{dg}{dx}(x)$

Or in terms of the inverse function,

$$\frac{\delta_1}{\delta_2} \approx \frac{dh}{dy}(y), \quad h = g^{-1}$$

- We now note the event $\{x \leq X \leq x+\delta_1\}$ is the same as the event $\{y \leq Y \leq y+\delta_2\}$. Thus,
 $f_Y(y)\delta_2 \approx \text{Pr}(y \leq Y \leq y + \delta_2)$
 $= \text{Pr}(x \leq X \leq x + \delta_1) \approx f_X(x)\delta_1$

Examples

- Let $Y = g(X) = X^2$, where X is a uniform random variable on $[0,2]$. Find the PDF of Y .

- Sol:

$$f_Y(y) = \begin{cases} \frac{1}{4}y^{-\frac{1}{2}}, & 0 \leq y \leq 4, \\ 0, & \text{otherwise.} \end{cases}$$

- Let $Y = g(X) = X^2$, where X is a uniform random variable on $[-2,2]$. Find the PDF of Y . <Non-monotonic>

- Sol:

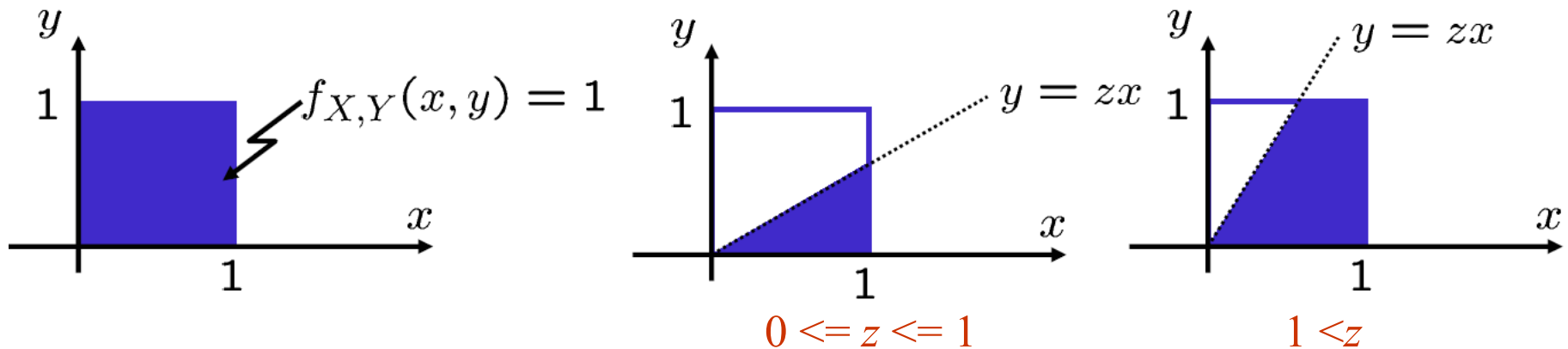
Functions of **Two** Random Variables

Two-step procedure

1. Calculate the CDF $F_Z(z)$ of $Z=g(X,Y)$
2. Differentiate to obtain the PDF of Z

Example: (4.8) (p.211)

Let X and Y be two independent uniform (continuous) random variables over $[0,1]$. Find the PDF of $Z=Y/X$.

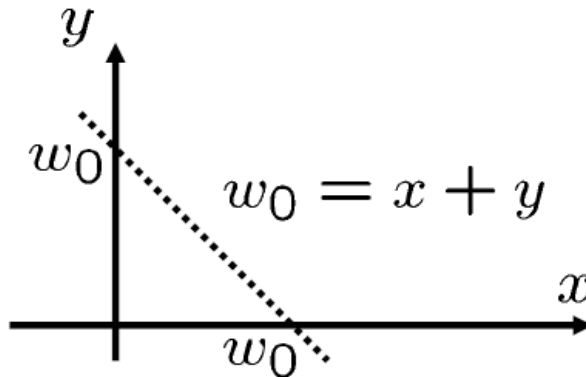


$$F_Z(z) = z/2 \quad 0 \leq z \leq 1$$

$$F_Z(z) = 1 - 1/2z \quad z \geq 1$$

The Distribution of $X+Y$

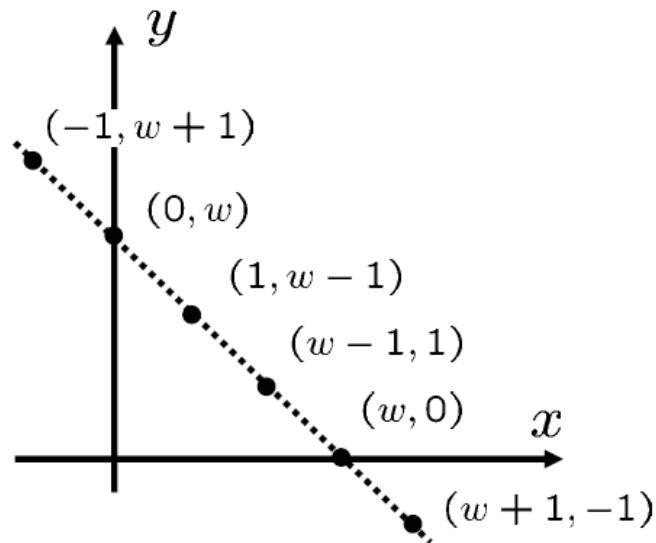
- Let X, Y be two RVs, and let $W = X + Y$. Find the PMF (if X, Y *discrete*) or PDF of W (if X, Y *continuous*).
- Points where the value $W = w_0$ is some constant lie on the following line



- Idea
 - Discrete case: add probabilities of all points on this line.
 - Continuous case: integrate the joint density on this line.

$X+Y$: Independent Discrete Integers

- Let X, Y be discrete, independent RVs.
- Then $W = X + Y$ is also integer-valued.



- Thus,
$$p_W(w) = \Pr(X + Y = w) = \sum_x \Pr(X = x) \Pr(Y = w - x)$$
$$= \sum_x p_X(x) p_Y(w - x)$$

This operation is called **discrete convolution** (摺積、迴旋積)

Obtaining $p_W(w)$ by Discrete Convolution

- Problem10 (p. 246): Let X, Y be two indep. RVs with PMFs

$$p_X(x) = \begin{cases} 1/3, & \text{if } x = 1, 2, 3, \\ 0, & \text{otherwise.} \end{cases} \quad p_Y(y) = \begin{cases} 1/2, & \text{if } y = 0, \\ 1/3, & \text{if } y = 1, \\ 1/6, & \text{if } y = 2, \\ 0, & \text{otherwise.} \end{cases}$$

Find the PMF of $W = X + Y$.

$$p_W(w) = \sum_x p_X(x) p_Y(w-x)$$

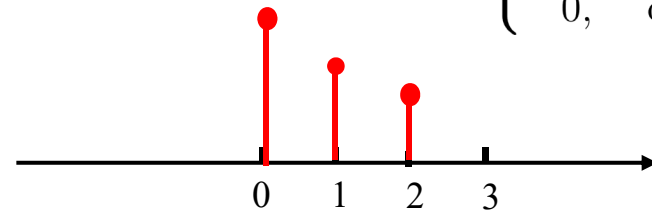
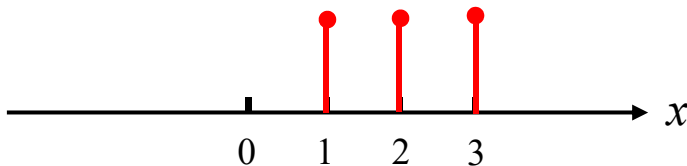
Discrete Convolution

$$p_W(w) = \sum_x p_X(x)p_Y(w-x)$$

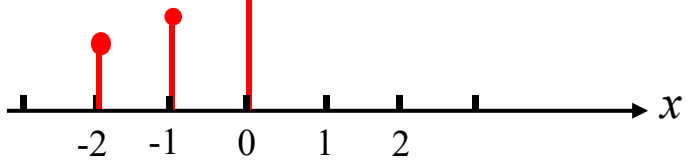
$$p_X(x) = \begin{cases} 1/3, & \text{if } x = 1, 2, 3, \\ 0, & \text{otherwise.} \end{cases}$$

$$p_Y(y) = \begin{cases} 1/2, & \text{if } y = 0, \\ 1/3, & \text{if } y = 1, \\ 1/6, & \text{if } y = 2, \\ 0, & \text{otherwise.} \end{cases}$$

$p_X(x)$

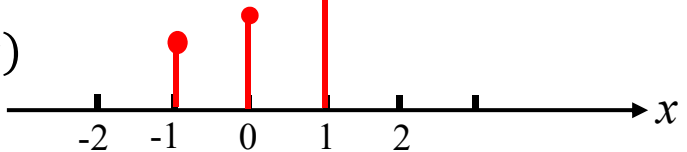


$p_Y(-x)$



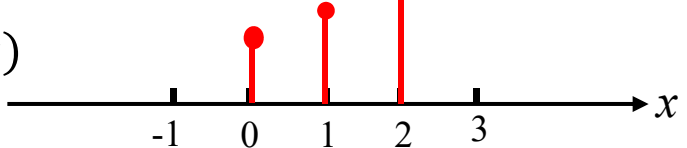
$$w = 0: \sum_x p_X(x)p_Y(-x) = 0$$

$p_Y(1-x)$



$$w = 1: \sum_x p_X(x)p_Y(1-x) = 1/6$$

$p_Y(2-x)$



$$w = 2: \sum_x p_X(x)p_Y(2-x) = \frac{1}{3} \times \frac{1}{3} + \frac{1}{3} \times \frac{1}{2} = \frac{5}{18}$$

X+Y: Convolution Integral

- Let X, Y be independent, continuous RVs.
- Then the PDF of $W = X + Y$ is given by

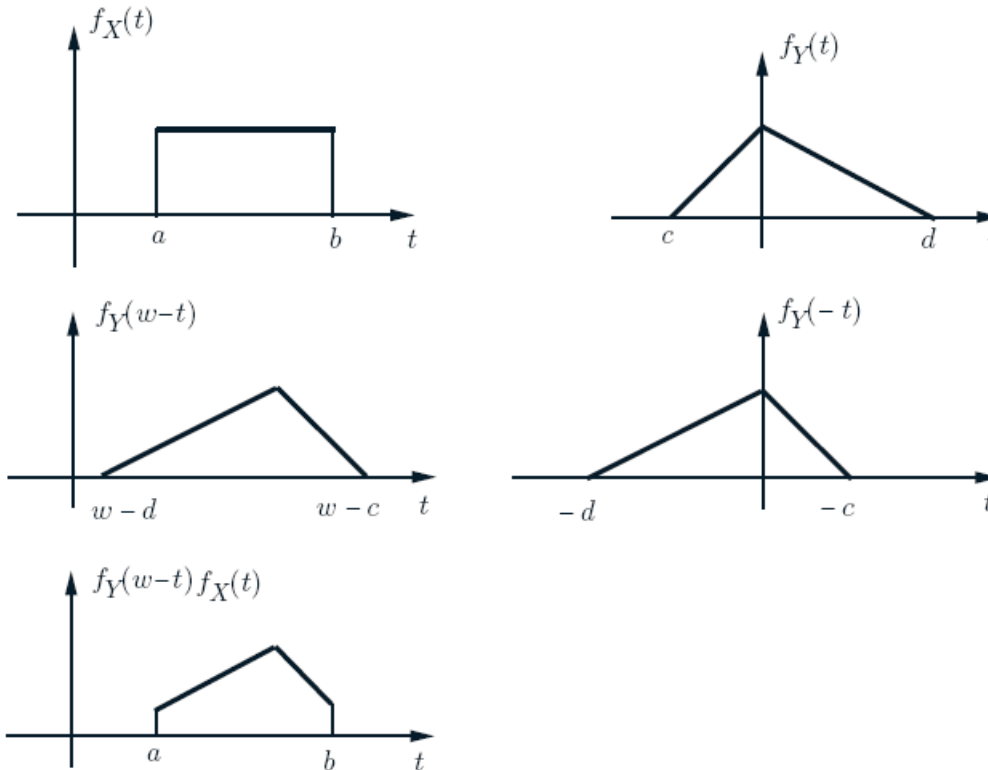
$$f_W(w) = \int_{-\infty}^{\infty} f_X(x) f_Y(w-x) dx$$

➤ This integral is called (continuous) **convolution** (摺疊積分)

Graphical Understanding of Continuous Convolutions

Convolution can be perceived with the help of [graphical illustration](#).

- Key step: **how to draw $f_Y(w-t)$ from $f_Y(t)$ for a fixed w and variable t**



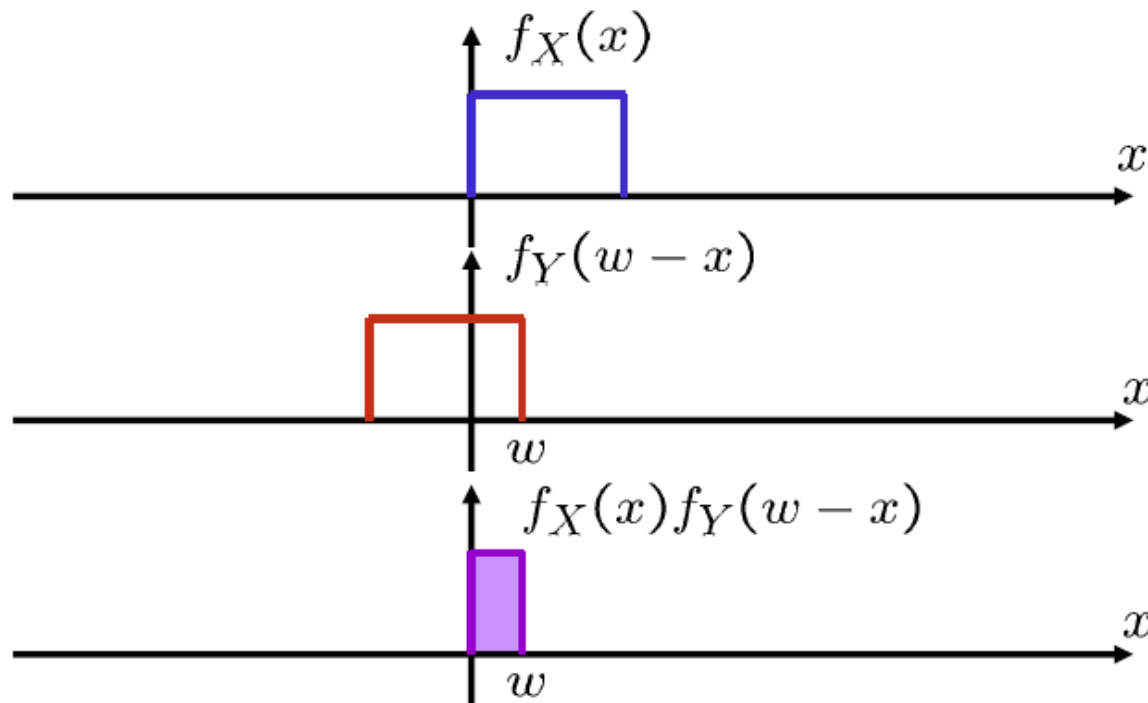
$X+Y$ Example: Independent Continuous

- Let X, Y be independent, uniform on $[0, 1]$.

Find the PDF of $W = X + Y$.

- Convolution idea applies:

$$f_W(w) = \int_{-\infty}^{\infty} f_X(x) f_Y(w-x) dx$$



Sum of Independent Normals

- Let X, Y be independent, **normal** RVs.

$$X \sim N(0, \sigma_x^2) \quad Y \sim N(0, \sigma_y^2)$$

- What is the **PDF of $W = X + Y$** ?

$$\begin{aligned} f_W(w) &= \int_{-\infty}^{\infty} f_X(x) f_Y(w - x) dx \\ &= \frac{1}{2\pi\sigma_x\sigma_y} \int_{-\infty}^{\infty} \exp(-x^2/2\sigma_x^2) \exp[-(w - x)^2/2\sigma_y^2] dx \\ &= C \cdot \exp(-w^2/2(\sigma_x^2 + \sigma_y^2)), \quad C = 1/(2\pi(\sigma_x^2 + \sigma_y^2))^{1/2} \end{aligned}$$

- 上式積分的結果很 neat，但推導有些繁瑣。推導關鍵有兩點：1) 配方法；2) $\int_{-\infty}^{\infty} f_X(x) dx = 1$ for any PDF $f_X(x)$.
- Hence, **W is normal** with mean = 0, and $\sigma_w^2 = \sigma_x^2 + \sigma_y^2$
- Note: More generally, **$aX + bY$** is also normal for any two constants a and b
- 此結果若將 Independent 的條件移除則 **未必成立**

Review

Review: Monotonic Function of a Continuous RV

Suppose that g is **strictly monotonic** and that for some function h and all x in the range of X we have

$$y = g(x) \quad \text{if and only if} \quad x = h(y)$$

Assume that h is differentiable. Then, the **PDF of $Y=g(X)$** in the region is given by

$$f_Y(y) = f_X(h(y)) \left| \frac{dh}{dy}(y) \right|$$

這個式子不需要背，其實**本質上就是先找 Y 的 CDF，再取微分即可！**

Joint Density of Two Functions of Two RVs

Problem:

Given joint PDF $f_{X,Y}(x,y)$ of two continuous random variables X and Y .

We now have two newly defined random variables V and W

$$V = g(X, Y)$$

$$W = h(X, Y)$$

where $g(\cdot)$ and $h(\cdot)$ are two real-valued functions. How do we compute the joint PDF $f_{V,W}(v,w)$ from $f_{X,Y}(x,y)$?

Important Remarks:

1. In general, $g(\cdot)$ and $h(\cdot)$ can be **any** 2 real-valued functions
2. In this topic, we will consider a special case of $g(\cdot)$ and $h(\cdot)$. We assume that the system of two equations in two unknowns has a **unique** solution $x=p(v,w)$ and $y=q(v,w)$ for x and y in terms of v and w , similar to the monotonic case mentioned on the previous page.

Basic Concept

- By properly selecting the values of dv , dw , dx , and dy , we have
[probability before mapping] = [probability after mapping].

$$P(v \leq V \leq v + dv, w \leq W \leq w + dw) = P(x \leq X \leq x + dx, y \leq Y \leq y + dy)$$

When dv , dw , dx , and dy are very small,

$$P(v \leq V \leq v + dv, w \leq W \leq w + dw) \approx f_{VW}(v, w)(dv dw)$$

$$P(x \leq X \leq x + dx, y \leq Y \leq y + dy) \approx f_{XY}(x, y)(dx dy)$$

- The above amounts to

$$f_{VW}(v, w) = f_{XY}(x, y) \cdot \underbrace{\left(\frac{dx dy}{dv dw} \right)}_{\text{ratio of two areas}}$$

So, the question now is how do we find the ratio?

⇒ Use change of variable theorem in calculus

Formula to the Joint Density

$$f_{VW}(v, w) = f_{XY}(x, y) \cdot \underbrace{\left(\frac{dxdy}{dvdw} \right)}_{\text{ratio of two areas}}$$

Using **change of variable theorem** in calculus,

$$\boxed{f_{VW}(v, w) = f_{XY}(x, y) \cdot |J|^{-1}} \quad (\text{evaluated at } x=p(v, w) \text{ and } y=q(v, w))$$

where J is the **Jacobian** representing the (inverse) **ratio of the areas** and is the determinant

$$J = \begin{vmatrix} \frac{\partial g(x, y)}{\partial x} & \frac{\partial g(x, y)}{\partial y} \\ \frac{\partial h(x, y)}{\partial x} & \frac{\partial h(x, y)}{\partial y} \end{vmatrix} = \frac{\partial g(x, y)}{\partial x} \cdot \frac{\partial h(x, y)}{\partial y} - \frac{\partial g(x, y)}{\partial y} \cdot \frac{\partial h(x, y)}{\partial x}$$

- H. Stark and J. Woods, ***Probability and Random Process with Applications to Signal Processing***, 3rd ed., 2002, (pp. 152 ~ 161)

Example

- Let $V=X+Y$, and $W=X-Y$. Find the joint PDF of V and W in terms of the joint PDF of X and Y . (S&W, p.153)

- **Solutions:**

We use the formula $f_{VW}(v, w) = f_{XY}(x, y) \cdot |J|^{-1}$

First, we solve for x and y in terms of v and w . It is clear that

$$\begin{aligned}x &= \frac{1}{2}(v + w) \\ y &= \frac{1}{2}(v - w)\end{aligned}$$

The Jacobian is

$$J = \begin{vmatrix} 1 & 1 \\ 1 & -1 \end{vmatrix} = -2$$

Thus, $f_{VW}(v, w) = \frac{1}{2}f_{XY}\left(\frac{v+w}{2}, \frac{v-w}{2}\right)$ ■