

Topic 7: Multiple Continuous Random Variables

Lecture Outline

- Conditioning on an Event
- Joint Probability Density Functions
- Conditioning on a Random Variable
 - 條件機率 Conditional probability conditioned on a continuous RV
 - 條件期望值 Conditional expectation
 - 條件機率密度函數 Conditional PDF
- Continuous Bayes Rule
- Variants of Total Probability Theorem (**total probability theorem 的四個變形**)

Reading : Textbook 3.4- 3.6

Multiple Random Variables

Quite naturally, we definitely will have to deal with ***multiple RVs*** in a real-world problem:

- The total amount of time we need to wait in the line at the supermarket counter
- What information can the multiple antennas of a WiFi router provide to the receiver?

Again, we are particularly interested in the following:

- How different random variables are ***related*** to each other? 不同隨機變數間之關聯性
(正相關、負相關、零相關、獨立、條件機率、條件期望值、條件機率密度函數)
- How to learn the ***behavior*** of the sum of all RVs?
Ex: 分析網路系統中資料流量進出某網路節點所耗費的時間延遲 (排隊理論, queueing)
- How to learn the ***behavior*** or the ***true value*** of one RV when it is buried in the sum with other random variables?
Ex: 如何從 $Y = S + N$ 中擷取 S 的資訊。其中雜訊一般是 continuous RV，在數位通訊的技術下，傳送訊號 S 是 discrete。接收訊號 Y 為連續。

Conditioning on an Event

Definition

Similar to the definition of unconditional PDF, the *conditional PDF* of a continuous random variable X , conditioned on a particular event E with $P(E) > 0$, is a function $f_{X|E}$ that satisfies

$$P(X \in B|E) = \int_B f_{X|E}(x)dx \quad (1)$$

Special Case

Conditioning on the event E that X belongs to a subset A of the real line, we have

$$P(X \in B|X \in A) = \frac{P(X \in A, X \in B)}{P(X \in A)} = \frac{\int_{A \cap B} f_X(x)dx}{P(X \in A)} \quad (2)$$

Therefore, comparing (2) with (1), we have

$$f_{X|\{X \in A\}} = \begin{cases} \frac{f_X(x)}{P(X \in A)}, & \text{if } x \in A \\ 0, & \text{otherwise.} \end{cases}$$

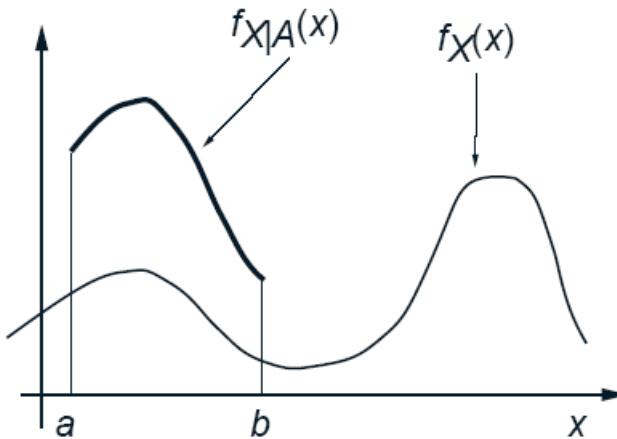
Illustration of Conditioning on an Event

Special Case

Conditioning on the set when X belongs to a subset A of the real line, we have

$$f_{X| \{X \in A\}} = \begin{cases} \frac{f_X(x)}{P(X \in A)}, & \text{if } x \in A \\ 0, & \text{otherwise.} \end{cases}$$

We see it's just a *scaling* of $f_x(x)$. So the PDF within the conditioning set has the same shape as the unconditional PDF



The unconditional PDF f_X and the conditional PDF $f_{X|A}$, where A is the interval $[a, b]$. Note that within the conditioning event A , $f_{X|A}$ retains the same shape as f_X , except that it is scaled along the vertical axis.

Example: The Exponential Random Variable Is Memoryless

Example 3.13 (Memoryless Property of Exponential)

The **time** T until a new light bulb burns out can be modeled by an **exponential** random variable with parameter λ . Alice turns the light on, leaves the room, and when she returns, t time units later, finds that the light bulb is still on, which corresponds to the event $A=\{T>t\}$.

Let X be the **additional time** until the light bulb burns out. What is the conditional CDF of X , given the event A ?

Total Probability Theorem Using Conditional PDF**

If A_1, A_2, \dots, A_n are disjoint events with $P(A_i) > 0$ for each i , that form a partition of the sample space, then

$$f_X(x) = \sum_{i=1}^n P(A_i) f_{X|A_i}(x)$$

- 這是常用的技巧，可協助計算出複雜 X 的PDF。
How to justify the equation?
- See Example 3.14 in the textbook
- Another example:

Let X be a standard normal random variable. A new random variable Y is defined as follows: We flip a coin. If the outcome is a head, then $Y = X$. And if the outcome of the coin flip is a tail, then $Y = -X$. Assume the coin flip is independent with X .
Please find the PDF of Y .

Total Probability Theorem using Conditional PDF

(Example 3.14)

The metro train arrives at the station near your home every quarter hour starting at 6:00 a.m. You walk into the station every morning between 7:10 and 7:30 a.m. and your arrival time is a uniform random variable over this interval. What is the **PDF** of the **time** you have to wait for the first train to arrive?

Total Expectation Theorem using Conditional Expectation

Definition

The conditional expectation is defined by

$$E[X|A] = \int_{-\infty}^{\infty} xf_{X|A}(x)dx$$

Total Expectation Theorem

If A_1, A_2, \dots, A_n are disjoint events with $P(A_i) > 0$ for each i , that form a partition of the sample space, then

$$E[X] = \sum_{i=1}^n P(A_i)E[X|A_i]$$

See [Example 3.17](#) in textbook.

Total Expectation Theorem using Conditional Expectation

(Example 3.17)

Suppose that the random variable X has the piecewise constant PDF

$$f_X(x) = \begin{cases} 1/3, & \text{if } 0 \leq x \leq 1, \\ 2/3, & \text{if } 1 < x \leq 2, \\ 0, & \text{otherwise.} \end{cases}$$

Find $E[X]$ and $\text{var}(X)$.

Joint Probability Density Function

We say that two continuous random variables associated with a common experiment are **jointly continuous** and can be described in terms of a **joint PDF** $f_{X,Y}$, if $f_{X,Y}$ is a nonnegative function that satisfies

$$P((X, Y) \in B) = \iint_{(x,y) \in B} f_{X,Y}(x, y) dx dy$$

for every subset B of the 2-dimensional plane.

- The probability that (X, Y) falls within a rectangle $B = \{(x, y) | a \leq x \leq b, c \leq y \leq d\}$ is

$$P(a \leq X \leq b, c \leq Y \leq d) = \int_c^d \int_a^b f_{X,Y}(x, y) dx dy$$

Joint Probability Density Function

- PDF $f_{X,Y}(x,y)$ must satisfy *normalization* equation, i.e.

$$\int_{-\infty}^{\infty} \int_{-\infty}^{\infty} f_{X,Y}(x, y) dx dy = 1$$

- To interpret the PDF, we let δ be very small and consider the probability of a small rectangle. We have

$$P(a \leq X \leq a + \delta, c \leq Y \leq c + \delta) \approx f_{X,Y}(a, c) \cdot \delta^2.$$

we can view $f_{X,Y}(a, c)$ as the “*probability per unit area*” in the vicinity of (a, c)

Marginal PDF from Joint PDF

The marginal PDFs $f_X(x)$ and $f_Y(y)$ of continuous X and Y are given by

$$f_X(x) = \int_{-\infty}^{\infty} f_{X,Y}(x, y) dy$$

$$f_Y(y) = \int_{-\infty}^{\infty} f_{X,Y}(x, y) dx$$

- Recall that for discrete random RVs, we have

$$p_X(x) = \sum_y p_{X,Y}(x, y)$$

$$p_Y(y) = \sum_x p_{X,Y}(x, y)$$

- 本質上，上述諸式均源自於 total probability theorem

Expectation

- If X and Y are jointly continuous random variables, and g is some function, then $Z = g(X, Y)$ is also a random variable. the **expected value rule** is

$$E[g(X, Y)] = \int_{-\infty}^{\infty} g(x, y) f_{X,Y}(x, y) dx dy$$

- Using the **expected value rule**, for any scalars a, b , we have

$$E[aX + bY] = aE[X] + bE[Y]$$

Conditioning on a Random Variable

Definition

Let X and Y be continuous random variables with joint PDF $f_{X,Y}$. For any fixed y with $f_Y(y) > 0$, the **conditional PDF** of X given that $Y = y$, is defined by

$$f_{X|Y}(x|y) \triangleq \frac{f_{X,Y}(x,y)}{f_Y(y)}$$

1. Analogous to the formula in the discrete PMF case
2. $f_{X|Y}(x|y)$ is a legitimate PDF (check normalization)

See example 3.19 in the textbook.

Interpretation of Conditioning on a Random Variable

Interpretation

Fix some small positive numbers δ_1 and δ_2 , and condition on the event

$B = \{y \leq Y \leq y + \delta_2\}$. We have

$$\begin{aligned} P(x \leq X \leq x + \delta_1 | y \leq Y \leq y + \delta_2) &\approx \frac{f_{X,Y}(x, y)\delta_1\delta_2}{f_Y(y)\delta_2} \\ &= f_{X|Y}(x|y)\delta_1 \end{aligned}$$

We can think of the limiting case where δ_2 decreases to zero and write

$$P(x \leq X \leq x + \delta_1 | Y = y) \approx f_{X|Y}(x|y)\delta_1$$

More generally,

$$P(X \in A | Y = y) = \int_A f_{X|Y}(x|y)dx$$

Conditional Expectation

Definition

The conditional expectation is defined by

$$E[X|Y = y] \triangleq \int_{-\infty}^{\infty} x f_{X|Y}(x|y) dx$$

- The expected value rule applies to conditional expectation

$$E[g(X)|Y = y] = \int_{-\infty}^{\infty} g(x) f_{X|Y}(x|y) dx$$

- Total expectation theorem

$$E[X] = \int_{-\infty}^{\infty} E[X|Y = y] f_Y(y) dy$$

Conditional Expectation

Example

The joint PDF of X and Y is $f_{X,Y}(x,y) = \frac{e^{-x/y}e^{-y}}{y}$, $0 < x < \infty, 0 < y < \infty$
Find the conditional expectation $E[X|Y=y]$.

Total Expectation Theorem

Total expectation theorem for continuous random variable

$$E[X] = \int_{-\infty}^{\infty} E[X|Y = y]f_Y(y)dy$$

- Consider any event A . Let X be the random variable that takes the value 1 if event A occurs and the value 0 otherwise (Such RV is called an **indicator**). In this case, we have $E[X]=P(A)$.

$$\begin{cases} X = 1, & \text{if } A \text{ occurs,} \\ X = 0, & \text{otherwise.} \end{cases}$$

Then, we have the following version of **total probability theorem**, when conditioned on continuous random variable Y (See page 30 of this topic)

$$P(A) = \int_{-\infty}^{\infty} P(A|Y = y)f_Y(y)dy$$

Independence

Two continuous random variables X and Y are **independent** if their joint PDF for all x and y is the product of the marginal PDFs:

$$f_{X,Y}(x,y) = f_X(x)f_Y(y)$$

Suppose X and Y are independent.

- Similar to the discrete case, the random variables $g(X)$ and $h(Y)$ are independent, for any functions g and h
- We have $E[XY] = E[X]E[Y]$

or more generally,

$$E[g(X)h(Y)] = E[g(X)]E[h(Y)]$$

- IF X and Y are **independent**, we have

$$\text{var}(X + Y) = \text{var}(X) + \text{var}(Y)$$

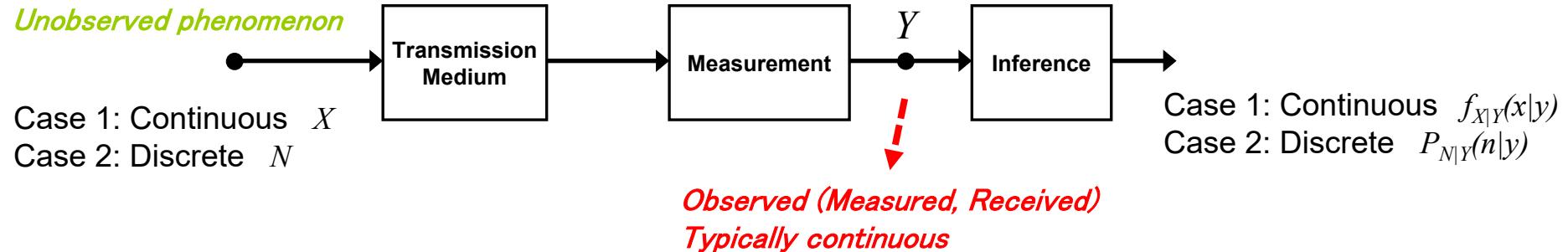
Continuous Bayes' Rule

The conditional PDF $f_{X|Y}(x|y)$ can be obtained via

$$f_{X|Y}(x|y) = \frac{f_X(x)f_{Y|X}(y|x)}{\int f_X(t)f_{Y|X}(y|t)dt}$$

Statistical Inference using Bayes' Rule

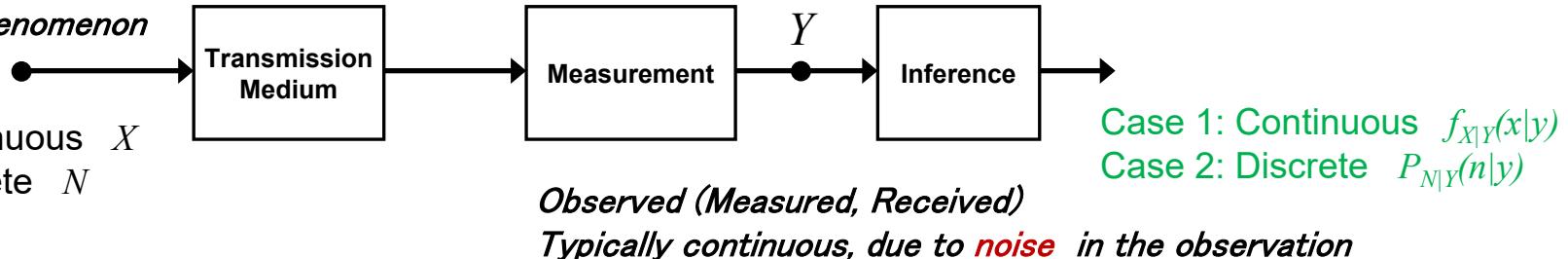
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- **Unobserved phenomenon (cause)** describes something we want to know about. But we can only observe a certain result (**effect**) at the receiving end.
For example,
 - Unknown transmitted bit (0 or 1) in a **communication system**
 - Unknown presence of an airplane in a **radar system**
 - Unknown disease of a patient
- **Objective:** (藉由觀察 Y 的值來推論 X 或 N 的值。此處 Y 是蒐集獲得之資料、訊號)
Determine which cause among many candidates is the most likely, by checking the **posterior probability** $P_{N|Y}(n|y)$ for all possible n (or posterior density $f_{X|Y}(x|y)$)

Statistical Inference using Bayes' Rule

Unobserved phenomenon



1. Continuous case: The *unobserved phenomenon* is a **continuous random variable** X , *the posterior density of the unknown X given observed $Y=y$ is*

$$f_{X|Y}(x|y) = \frac{f_X(x)f_{Y|X}(y|x)}{\int f_X(t)f_{Y|X}(y|t)dt}$$

2. Discrete case: The *unobserved phenomenon* is a **discrete random variable** N , *the posterior probability of N given the observed $Y=y$ is*

$$P(N = n|Y = y) = \frac{P_N(n)f_{Y|N}(y|N = n)}{\sum_i P_N(i)f_{Y|N}(y|N = i)}$$

Example – Signal Detection in Communication Systems**

A binary signal S is transmitted, and we are given that $P(S = 1) = p$ and $P(S = -1) = 1-p$. The received signal is $Y = N+S$, where N is normal noise, with zero mean and variance σ^2 , independent of S . What is the probability that $S = 1$, given that we have observed $Y=y$?

Joint CDF

If X and Y are two random variables associated with the same experiment, we define their joint CDF by

$$F_{X,Y}(x, y) \triangleq P(X \leq x, Y \leq y)$$

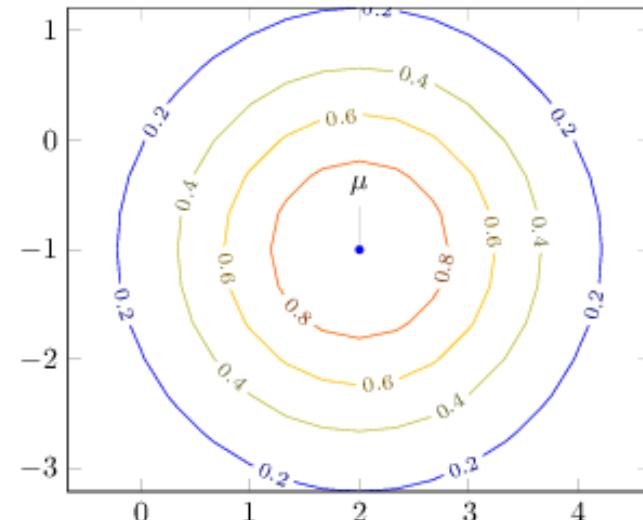
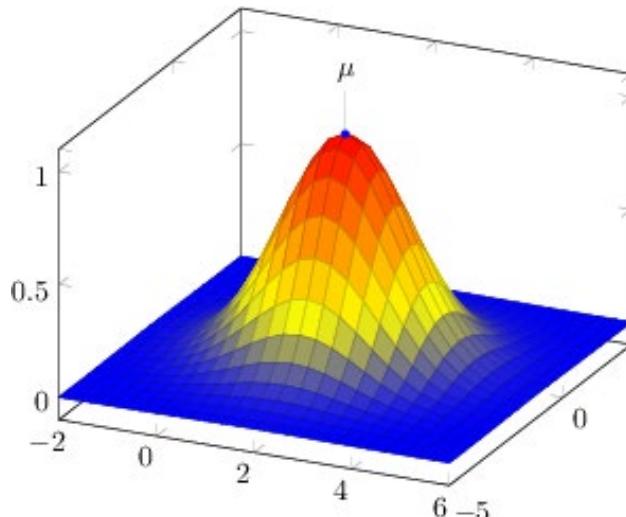
- The joint CDF $F_{X,Y}(x, y) \triangleq P(X \leq x, Y \leq y) = \int_{-\infty}^y \int_{-\infty}^x f_{X,Y}(s, t) ds dt$
- The PDF can be recovered from the PDF by differentiating:

$$f_{X,Y}(x, y) = \frac{\partial^2 F_{X,Y}(x, y)}{\partial x \partial y}$$

Joint PDF of Independent Gaussian RVs

Let X and Y be two independent Gaussian random variables with means μ_x, μ_y , and variances σ_x^2, σ_y^2 respectively. The joint PDF is given by

- The joint PDF has a bell shape centered at (μ_x, μ_y) . Whether the PDF is a *tall-thin* bell or *wide-fat* bell is determined by the variances σ_x^2, σ_y^2
- Very often we are interested in knowing the **contours** (等高線圖、切面圖) of the joint PDF. The contours are sets of points at which the PDF takes a constant value.



Summary of Total Probability Theorem and Its Variants

- We have already learned several different versions of *total probability theorem*. The original version is given in the form of probability of events (Section 1.4, page 28)

$$P(G) = \sum_{i=1}^K P(G | F_i)P(F_i)$$

- According to the types of underlying RVs in forming G and F_i , the TPT can be further generalized to the following variants:
 - Variant 1: Discrete RV in G – Discrete RV in F_i (DD Type)
 - Variant 2: Continuous RV in G – Continuous RV in F_i (CC Type)
 - Variant 3: Continuous RV in G – Discrete RV in F_i (CD Type)
 - Variant 4: Discrete RV in G – Continuous RV in F_i (DC Type)
- 通則: When continuous RV is involved, the sum and PMF in original TPT need to be modified to integral and PDF

第一類變形: Discrete RV in G – Discrete RV in F_i

➤ Variant 1: Discrete RV in G – Discrete RV in F_i (DD Type)

$$\begin{aligned} p_X(x) &= \sum_y p_{X,Y}(x,y) \\ &= \sum_y p_{X|Y}(x|y)p_Y(y) \end{aligned}$$

- ✓ This follows directly from original version by setting $G = \{X=x\}$ and $F_i = \{Y=y_i\}$
- ✓ See Example 2.14 and 2.15.
- ✓ Another possible forms:

$$p_X(x) = \sum_{i=1}^n p_{X|A_i}(x)P(A_i)$$

$$P(B) = \sum_y P(B|Y=y)p_Y(y)$$

- ✓ Example:

Let X, Y be discrete RVs. Compute the probability $P(X + Y = w)$.

第二類: Continuous RV in G – Continuous RV in F_i

- Variant 2: Continuous RV in G – Continuous RV in F_i (CCType)

$$\begin{aligned} f_X(x) &= \int_{-\infty}^{\infty} f_{X,Y}(x,y)dy \\ &= \int_{-\infty}^{\infty} f_{X|Y}(x|y)f_Y(y)dy \end{aligned}$$

- ✓ See page 172
- ✓ Example 3.19, on page 179 (continuous Bayes rule)

第三類: Continuous RV in G – Discrete RV in F_i

- Variant 3: Continuous RV in G – Discrete RV in F_i (**CDType**)

$$f_Y(y) = \sum_i f_{Y|N}(y|i)p_N(i)$$

- ✓ See page 180 and [Example 3.20](#).
- ✓ A more general form is (see page 167)

$$f_X(x) = \sum_{i=1}^n P(A_i)f_{X|A_i}(x)$$

- ✓ [Example 3.14](#), on page 168.
- ✓ See the example on page 6 of this topic.

第四類: Discrete RV in G – Continuous RV in F_i

➤ Variant 4: Discrete RV in G – Continuous RV in F_i (CDType)

$$p_N(n) = \int_{-\infty}^{\infty} P(N = n | Y = y) f_Y(y) dy$$

- ✓ See page 181 and page 182.
- ✓ A more general form is ([page 181 of textbook](#), [page 18 of this topic](#))

$$P(A) = \int_{-\infty}^{\infty} P(A | Y = y) f_Y(y) dy$$

- ✓ You can justify this relation by taking the integral over y to the 1st equation on page 181 of the textbook

- ✓ Example:

Let X, Y be independent **continuous** RVs. Compute the PDF of $X+Y$.

