

$$\zeta_t = \sigma v_t = \exp\left\{\frac{w_t}{2}\right\} v_t \quad \begin{pmatrix} v_t \\ \zeta_t \end{pmatrix} \sim N\left(\begin{pmatrix} 0 \\ 0 \end{pmatrix}; \begin{pmatrix} I & 0 \\ 0 & \kappa^2 \end{pmatrix}\right)$$

$$w_t = w + \rho w_{t-1} + \gamma_t$$

Ex 1a

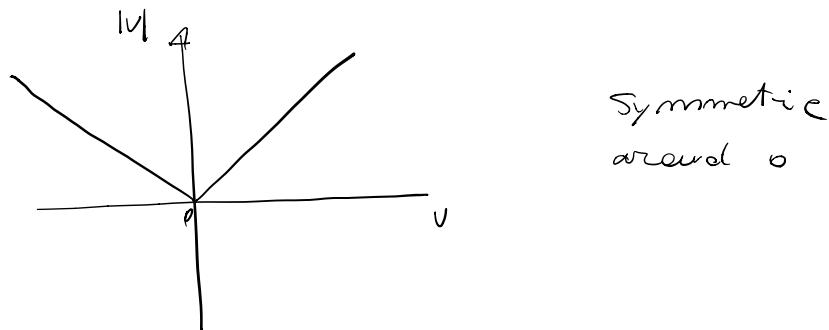
$$(i) E[|\zeta_t|] = E[|\zeta_t v_t|] = E[|\zeta_t| |v_t|] = E[\zeta_t] E[|v_t|]$$

Note that  $|v_t|$  is distributed as a folded Gaussian distribution such that  $E[|v_t|] = \sqrt{\frac{2}{\pi}}$

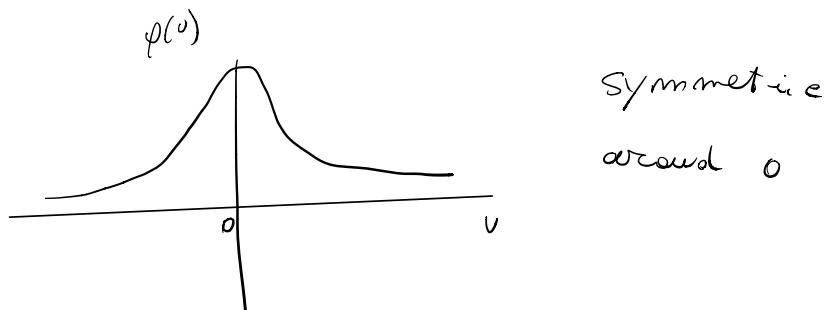
Formal derivation: Not required, but welcome

$$E[|v_t|] = \int_{-\infty}^{\infty} |v| \varphi(v) dv, \text{ where } \varphi(v) = \frac{1}{\sqrt{2\pi}} \exp\left\{-\frac{v^2}{2}\right\}$$

Note that



normal



such that  $|v| \varphi(v)$  is also going to be symmetric around 0.

Using this argument we can write

$$\begin{aligned} E[|v_t|] &= \int_{-\infty}^{\infty} |v| \varphi(v) dv = \int_{-\infty}^0 |v| \varphi(v) dv + \int_0^{\infty} |v| \varphi(v) dv \\ &= 2 \int_0^{\infty} |v| \varphi(v) dv = 2 \int_0^{\infty} v \varphi(v) dv \end{aligned}$$

Note that

$$\begin{aligned} \int_0^{\infty} v \varphi(v) dv &= \int_0^{\infty} \frac{v}{\sqrt{2\pi}} \exp\left\{-\frac{v^2}{2}\right\} dv = \frac{1}{\sqrt{2\pi}} \int_0^{\infty} v \exp\left\{-\frac{v^2}{2}\right\} dv \\ &= -\frac{1}{\sqrt{2\pi}} \int_0^{\infty} -v \exp\left\{-\frac{v^2}{2}\right\} dv = -\frac{1}{\sqrt{2\pi}} \left. \exp\left\{-\frac{v^2}{2}\right\} \right|_0^{\infty} \\ &= -\frac{1}{\sqrt{2\pi}} \left[ \frac{\partial}{\partial v} \exp\left\{-\frac{v^2}{2}\right\} \right]_0^{\infty} = -\frac{1}{\sqrt{2\pi}} \left[ -v \exp\left\{-\frac{v^2}{2}\right\} \right]_0^{\infty} \end{aligned}$$

such that

$$2 \int_0^{\infty} v \varphi(v) dv = \frac{2}{\sqrt{2\pi}} = \sqrt{\frac{2}{\pi}}$$


$$\text{ii) } E\left[z_e^2\right] = E\left[v_e^2 \sigma_e^2\right] = E\left[v_e^2\right] E\left[\sigma_e^2\right] = E\left[\sigma_e^2\right]$$

because  $E\left[v_e^2\right] = 1$  by assumption

Formal derivation: Not required, but welcome

$$E\left[v_e^2\right] = \int_{-\infty}^{\infty} v_e^2 \varphi(v) dv = 2 \int_0^{\infty} v_e^2 \varphi(v) dv$$

$$= \frac{2}{\sqrt{2\pi}} \int_0^{\infty} v^2 \exp\left\{-\frac{v^2}{2}\right\} dv =$$

$$= \frac{2}{\sqrt{2\pi}} \int_0^{\infty} v \cdot v \exp\left\{-\frac{v^2}{2}\right\} dv$$

remember integration by parts

$$\int_a^b f'(v) g(v) dv = f(v)g(v) \Big|_a^b - \int_a^b f(v)g'(v) dx$$

Set

$$g(v) = v \quad g'(v) = 1$$

$$f'(v) = v \exp\left\{-\frac{v^2}{2}\right\} \quad f(v) = -\exp\left\{-\frac{v^2}{2}\right\}$$

such that

$$\frac{2}{\sqrt{2\pi}} \int_0^\infty u \cdot v \exp\left\{-\frac{v^2}{2}\right\} du =$$

$$= \frac{2}{\sqrt{2\pi}} \left[ -u \exp\left\{-\frac{v^2}{2}\right\} \Big|_0^\infty + \int_0^\infty \exp\left\{-\frac{v^2}{2}\right\} du \right]$$

Note that  $\left. -u \exp\left\{-\frac{v^2}{2}\right\} \right|_0^\infty = 0$  can you show this?

$$= \frac{2}{\sqrt{2\pi}} \int_0^\infty \exp\left\{-\frac{v^2}{2}\right\} du = 2 \int_0^\infty \frac{1}{\sqrt{2\pi}} \exp\left\{-\frac{v^2}{2}\right\} du$$

$$= 2 \int_0^\infty \varphi(v) du = 2(1 - \Phi(0)) = 2\left(1 - \frac{1}{2}\right) = 1$$

where  $\Phi(x) = \int_{-\infty}^x \varphi(v) dv$  is the

cumulative density function of a Gaussian distribution.

The same result is obtained by noting that

$$\int_{-\infty}^\infty \varphi(v) du = 1 \text{ oral}$$

$$\int_{-\infty}^{\infty} \varphi(v) dv = 2 \int_0^{\infty} \varphi(v) dv$$

such that  $\int_0^{\infty} \varphi(v) dv = \frac{1}{2}$

$$(iii) E[|Z_\epsilon|^3] = E\left[\epsilon^3 E[|V_\epsilon|^3]\right]$$

where  $E[|V_\epsilon|^3]$  is the third moment of a folded normal distribution and is equal to  $2\sqrt{\frac{2}{\pi}}$

Formal derivation: Not required, but welcome

$$E(|V_\epsilon|^3) = \int_{-\infty}^{\infty} |v|^3 \varphi(v) dv = 2 \int_0^{\infty} v^3 \varphi(v) dv$$

$$= 2 \int_0^{\infty} \frac{v^3}{\sqrt{2\pi}} \exp\left\{-\frac{v^2}{2}\right\} dv =$$

$$= \frac{2}{\sqrt{2\pi}} \int_0^{\infty} v^2 v \exp\left\{-\frac{v^2}{2}\right\} dv =$$

by parts  $f(v) = v^2$   $f'(v) = 2v$

$$g'(v) = v \exp\left\{-\frac{v^2}{2}\right\} \quad g(v) = -\exp\left\{-\frac{v^2}{2}\right\}$$

$$= \frac{2}{\sqrt{2\pi}} \left[ -v^2 \exp\left\{-\frac{v^2}{2}\right\} \Big|_0^\infty + 2 \int_0^\infty v \exp\left\{-\frac{v^2}{2}\right\} dv \right]$$

Note that  $\left. -v^2 \exp\left\{-\frac{v^2}{2}\right\} \right|_0^\infty = 0 \rightarrow \underline{\text{show this}}$

$$\begin{aligned} &= \frac{2}{\sqrt{2\pi}} \cdot 2 \int_0^\infty v \exp\left\{-\frac{v^2}{2}\right\} dv = 2 \cdot 2 \int_0^\infty \frac{v}{\sqrt{2\pi}} \exp\left\{-\frac{v^2}{2}\right\} dv \\ &= 2 \cdot 2 \int_0^\infty v \varphi(v) dv = 2 E[v] = 2 \sqrt{\frac{2}{\pi}} \end{aligned}$$



$$\text{iv) } E[z_e^3] = E[\zeta_e^3] + E[v_e^3] = 3 E[\zeta_e^3]$$

because the third moment of a standard Gaussian distribution is 3

Formal derivation: Not required, but welcome

$$\begin{aligned} 4E[v_e^3] &= \int_{-\infty}^{\infty} v^3 \varphi(v) dv = 2 \int_0^{\infty} v^3 \varphi(v) dv \\ 2 \int_0^{\infty} \frac{v^3}{\sqrt{2\pi}} \exp\left\{-\frac{v^2}{2}\right\} dv &= \frac{2}{\sqrt{2\pi}} \int_0^{\infty} v^3 \exp\left\{-\frac{v^2}{2}\right\} dv \end{aligned}$$

by parts

$$f(v) = v^3 \quad f'(v) = 3v^2$$

$$g(v) = v \exp\left\{-\frac{v^2}{2}\right\} \quad g(0) = -\exp\left\{-\frac{v^2}{2}\right\}$$

so

$$\frac{2}{\sqrt{2\pi}} \left[ -v^3 \exp\left\{-\frac{v^2}{2}\right\} \right]_0^\infty + 3 \int_0^\infty v^2 \exp\left\{-\frac{v^2}{2}\right\} dv =$$

$$-v^3 \exp\left\{-\frac{v^2}{2}\right\} \Big|_0^\infty = 0 \quad \text{show this}$$

$$= \frac{2}{\sqrt{2\pi}} \cdot 3 \int_0^\infty v^2 \exp\left\{-\frac{v^2}{2}\right\} dv = 3 \cdot 2 \int_0^\infty v^2 \varphi(v) dv$$

$$= 3 \mathbb{E}[v^2] = 3 \quad \blacksquare$$

$$\begin{aligned} \text{v) } \mathbb{E}\left[\sum_{t=1}^T \sum_{t'=1}^{T-1} \right] &= \mathbb{E}\left[\left(\sum_t v_t \sum_{t'=1}^{T-1} \mathbb{E}_{t'} \mathbb{E}_{t'-1} \right)\right] = \\ &= \mathbb{E}\left[\left(\sum_t v_t \sum_{t'=1}^{T-1} \right)\right] \mathbb{E}\left[\sum_t v_t \sum_{t'=1}^{T-1} \right] = \mathbb{E}\left[\left(\sum_t v_t \right) \mathbb{E}\left[\sum_{t'=1}^{T-1} \right] \right] \mathbb{E}\left[\sum_t v_t \right] \\ &= \mathbb{E}\left[v_t\right]^2 \mathbb{E}\left[\sum_t v_t \right] = \frac{2}{\pi} \mathbb{E}\left[\sum_t v_t \right] \end{aligned}$$

Since identically distributed

$$ri) \mathbb{E} \left[ \zeta_t^2 \zeta_{t-j}^2 \right] = \mathbb{E} \left[ \sigma_t^2 \sigma_{t-j}^2 \right]$$

Exploiting that  $\mathbb{E} \left[ v_t^2 \right] = \mathbb{E} \left[ v_{t-1}^2 \right] = 1$

and  $v_t \perp v_{t-j}$  if  $j > 0$

$$rii) \mathbb{E} \left[ \sigma_t^2 \right]$$

$$\zeta_t = \exp \left\{ \frac{w_t}{2} \right\} \Rightarrow \zeta_t^P = \exp \left\{ \frac{P w_t}{2} \right\}$$

$$w_t = w + \phi w_{t-1} + \eta_t = \frac{w}{1-\phi} + \sum_{s=0}^{\infty} \phi^s \eta_{t-s}$$

$$w_t \sim N \left( \frac{w}{1-\phi}; \frac{\sigma^2}{1-\phi^2} \right)$$

$$\frac{P}{2} w_t = x_t \sim N \left( \frac{Pw}{2(1-\phi)}, \frac{P^2 \sigma^2}{4(1-\phi^2)} \right)$$

so

$$\mathbb{E} \left[ \sigma_t^2 \right] = \mathbb{E} \left[ \exp \left\{ x_t \right\} \right]$$

We know that if  $Z \sim N(a, b^2)$

then  $\exp\{Z\}$  is Log-Normal with mean

$$\mathbb{E}[\exp\{Z\}] = \exp\left\{a + \frac{b^2}{2}\right\}$$

In our case  $a = \frac{\rho w}{2(1-\phi)}$ ;  $b^2 = \frac{\rho^2 k^2}{2(1-\phi^2)}$

such that

$$\mathbb{E}[\exp\{X_c\}] = \mathbb{E}[e^P] = \exp\left\{\frac{\rho w}{2(1-\phi)} + \frac{\rho^2 k^2}{2(1-\phi^2)}\right\}$$

Call  $\alpha_w = \frac{w}{1-\phi}$ ;  $\beta_w^2 = \frac{k^2}{1-\phi^2}$

such that  $\mathbb{E}[e^P] = \exp\left\{\frac{\rho \alpha_w}{2} + \frac{\rho^2 \beta_w^2}{8}\right\}$



$$\text{viii) } E[\zeta_t^2 \zeta_{t-j}^2]$$

$$\text{Note that } E[\zeta_t^2 \zeta_{t-j}^2] = E[\zeta_t^2 \zeta_{t+j}^2]$$

write  $\zeta_{t+j}^2 = \exp\{w_{t+j}\}$  and note that

$$E[\zeta_t^2 \zeta_{t+j}^2] = E[\exp\{w_t + w_{t+j}\}]$$

by the law of iterated expectation

$$\begin{aligned} E[\exp\{w_t + w_{t+j}\}] &= E[E[\exp\{w_t + w_{t+j}\} | w_t]] = \\ &= E[\exp\{w_t\} E[\exp\{w_{t+j}\} | w_t]] \end{aligned}$$

To compute  $E[\exp\{w_{t+j}\} | w_t]$  note that

$$w_{t+j} = w + \phi w_{t+j-1} + \eta_{t+j}$$

$$w_{t+1} = w + \phi w_t + \eta_{t+1}$$

$$w_{t+2} = w + \phi w_{t+1} + \eta_{t+2} = w + \phi(w + \phi w_t + \eta_{t+1}) + \eta_{t+2}$$

$$= w(1 + \phi) + \phi^2 w_t + \phi \eta_{t+1} + \eta_{t+2}$$

$$w_{t+3} = w(1 + \phi + \phi^2) + \phi^3 w_t + \phi^2 \eta_{t+1} + \phi \eta_{t+2} + \eta_{t+3}$$

In general

$$w_{t+j} = w \sum_{s=0}^{j-1} \phi^s + \phi^j w_t + \sum_{s=0}^{j-1} \phi^s \gamma_{t+j-s}$$

and

$$\exp\{w_{t+j}\} = \exp\left\{w \sum_{s=0}^{j-1} \phi^s\right\} \exp\{\phi^j w_t\} \exp\left\{\sum_{s=0}^{j-1} \phi^s \gamma_{t+j-s}\right\}$$

so

$$\mathbb{E}\left[\exp\{w_{t+j}\} \mid w_t\right] = \exp\left\{w \sum_{s=0}^{j-1} \phi^s\right\} \exp\{\phi^j w_t\} \mathbb{E}\left[\exp\left\{\sum_{s=0}^{j-1} \phi^s \gamma_{t+j-s}\right\} \mid w_t\right]$$

Note that

$$\begin{aligned} \mathbb{E}\left[\exp\left\{\sum_{s=0}^{j-1} \phi^s \gamma_{t+j-s}\right\} \mid w_t\right] &= \mathbb{E}\left[\exp\left\{\sum_{s=0}^{j-1} \phi^s \gamma_{t+j-s}\right\}\right] = \mathbb{E}\left[\prod_{s=0}^{j-1} \exp\left\{\phi^s \gamma_{t+j-s}\right\}\right] \\ &= \prod_{s=0}^{j-1} \mathbb{E}\left[\exp\left\{\phi^s \gamma_{t+j-s}\right\}\right] = \prod_{s=0}^{j-1} \exp\left\{\frac{\phi^s k^2}{2}\right\} = \exp\left\{\frac{k^2}{2} \sum_{s=0}^{j-1} \phi^{2s}\right\} \end{aligned}$$

so

$$\begin{aligned} &= \mathbb{E}\left[\exp\{w_t\} \mathbb{E}\left[\exp\{w_{t+j}\} \mid w_t\right]\right] = \\ &= \mathbb{E}\left[\exp\{w_t\} \left[ \exp\{\phi^j w_t\} \exp\left\{w \sum_{s=0}^{j-1} \phi^s\right\} \exp\left\{\frac{k^2}{2} \sum_{s=0}^{j-1} \phi^{2s}\right\} \right]\right] \\ &= \mathbb{E}\left[\exp\{w_t\} \exp\{\phi^j w_t\}\right] - \mathbb{E}\left[\exp\{w_t(1+\phi^j)\}\right] \end{aligned}$$

$$= \exp\left\{\frac{w(1+\phi^j)}{1-\phi} + \frac{\frac{k^2}{2}(1+\phi^j)^2}{1-\phi^2}\right\}$$

Such that

$$\begin{aligned} \mathbb{E} \left[ \exp \left\{ w_t + w_{t+5} \right\} \right] &= \\ = \exp \left\{ \frac{w(1+\phi^5)}{1-\phi} + \frac{\kappa^2 (1+\phi^5)^2}{2(1-\phi^2)} + w \sum_{s=0}^{j-1} \phi^s + \frac{\kappa^2 \sum_{s=0}^{j-1} \phi^{2s}}{2} \right\} \end{aligned}$$

Remember that

$$\sum_{s=0}^{m-1} a^s = \frac{1-a^m}{1-a} \quad \text{for } a \neq 1$$

Since  $\phi \neq 1$  we have

$$\sum_{s=0}^{j-1} \phi^s = \frac{1-\phi^j}{1-\phi} ; \quad \sum_{s=0}^{j-1} \phi^{2s} = \frac{1-\phi^{2j}}{1-\phi^2}$$

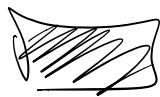
so

$$= \exp \left\{ \frac{w(1+\phi^5)}{1-\phi} + \frac{\kappa^2 (1+\phi^5)^2}{2(1-\phi^2)} + \frac{w(1-\phi^5)}{1-\phi} + \frac{\kappa^2 (1-\phi^{2j})}{2(1-\phi^2)} \right\}$$

$$= \exp \left\{ \frac{2w}{1-\phi} + \frac{\kappa^2 (1+\phi^5)}{1-\phi^2} \right\} \quad \text{call } \alpha_w = \frac{w}{1-\phi} ; \quad \beta_w^2 = \frac{\kappa^2}{1-\phi^2}$$

$$= \exp \left\{ 2\alpha_w + \beta_w^2 (1+\phi^5) \right\} = \exp \left\{ 2\alpha_w + \beta_w^2 \right\} \exp \left\{ \beta_w^2 \phi^5 \right\}$$

$$= E \left[ \sigma_t^2 \right]^2 \exp \left\{ \beta_w^2 \epsilon_t^2 \right\}$$



Ex 1b

independence  
↓

$$E \left[ \sigma_t^2 \sigma_{t-1} \right] = E \left[ \sigma_t^2 \sigma_{t-1} v_{t-1} \right] = E \left[ \sigma_t^2 \sigma_{t-1} \right] E \left[ v_{t-1} \right]$$

$$= 0 \quad \text{because } E \left[ v_{t-1} \right] = 0$$

Absence of correlation between past returns and current volatility levels implies no leverage effect.

Ex 1c

$$\text{i) } \sigma_t^2 = \exp \{ u_t \}; \quad u_t \sim N(\alpha_u, \beta_u^2)$$

such that  $\sigma_t^2$  is log Normal

the density of a log Normal distribution is

$$P(s) = \frac{\exp \left\{ -\frac{(\log s - \alpha_u)^2}{2 \beta_u^2} \right\}}{s \sqrt{2\pi} \beta_u}$$

Formal derivation: Not required, but welcome

$$P(\zeta_t^2 \leq s) = P(\exp\{w_t\} \leq s) = P(w_t \leq \log(s))$$

$$= P(\alpha_w + \beta_w z \leq \log(s)) \quad \text{where } z \sim N(0, 1)$$

$$= P(z \leq \frac{\log(s) - \alpha_w}{\beta_w}) = \Phi\left(\frac{\log(s) - \alpha_w}{\beta_w}\right)$$

so

$$P(s) = \frac{\partial P(\zeta_t^2 \leq s)}{\partial s} = \frac{\partial \Phi\left(\frac{\log(s) - \alpha_w}{\beta_w}\right)}{\partial s}$$

$$= \varphi\left(\frac{\log(s) - \alpha_w}{\beta_w}\right) \frac{1}{s \beta_w} = \frac{1}{s \beta_w \sqrt{2\pi}} \exp\left\{-\frac{(\log(s) - \alpha_w)^2}{2 \beta_w^2}\right\}$$

(i)

$$P(\zeta_t) = \int_0^\infty P(\zeta_t | \zeta_t^2) P(\zeta_t^2) d\zeta_t^2$$

$$= \int_0^\infty \frac{1}{\zeta_t^2} \varphi\left(\frac{\zeta_t}{\beta_w}\right) \frac{1}{\zeta_t^2 \beta_w} \varphi\left(\frac{\log \zeta_t^2 - \alpha_w}{\beta_w}\right) d\zeta_t^2$$

iii)

$$P(\zeta_e^2, \zeta_e) = P(\zeta_e | \zeta_e^2) P(\zeta_e^2) =$$
$$= \frac{1}{\zeta_e} \varphi\left(\frac{\zeta_e}{\zeta_e^2}\right) \frac{1}{\zeta_e^2 \beta_w} \varphi\left(\frac{\log \zeta_e^2 - \alpha w}{\beta_w}\right)$$

civ)

$$P(\zeta_e | \zeta_e^2) = \frac{1}{\zeta_e} \varphi\left(\frac{\zeta_e}{\zeta_e^2}\right)$$

