



Contents lists available at ScienceDirect

International Journal of Forecasting

journal homepage: www.elsevier.com/locate/ijforecast

Nowcasting GDP and its components in a data-rich environment: The merits of the indirect approach

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ARTICLE INFO

Keywords:

Mixed-frequency data
 Dynamic factor models
 Growth accounting
 Model averaging
 Ledoit–Wolf Shrinkage

ABSTRACT

The national accounts provide a coherent and exhaustive description of the current state of the economy, but are available at the quarterly frequency and are released with a nonignorable publication lag. The paper illustrates a method for nowcasting and forecasting the sixteen main components of Gross Domestic Product (GDP) by output and expenditure type at the monthly frequency, using a high-dimensional set of monthly economic indicators spanning the space of the common macroeconomic and financial factors. The projection on the common space is carried out by combining the individual nowcasts and forecasts arising from all possible bivariate models of the unobserved monthly GDP component and the observed monthly indicator. We discuss several pooling strategies and we select the one showing the best predictive performance according to a pseudo-real-time forecasting experiment. Monthly GDP can be indirectly estimated by the contemporaneous aggregation of the value added of the different industries and of the expenditure components. This enables the comparative assessment of the indirect nowcasts and forecasts vis-à-vis the direct approach and a growth accounting exercise. Our approach meets the challenges posed by the dimensionality, since it can handle a large number of time series with a complexity that increases linearly with the cross-sectional dimension, while retaining the essential heterogeneity of the information about the macroeconomy. An application to the Italian case leads to several interesting discoveries concerning the time-varying predictive content of the information carried by the monthly indicators.

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1. Introduction

The national accounts offer a coherent and consistent set of macroeconomic aggregates for the analysis of economic structure and change. They are compiled quarterly according to the methods and definitions outlined in Eurostat (2013a, 2013b).

The most authoritative measure of aggregate economic activity is gross domestic product (GDP). From the output, or production, side of the economy it is defined as the sum of the value added of the productive sectors; when it is evaluated at market prices, taxes less subsidies on products need to be added as well. If imports are added to the domestic supply of goods and services, then the total resources need to equal the total final uses for consumption, investment, and exports; hence, according to the expenditure approach, GDP results from the sum of final consumption expenditure, gross capital formation,

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and net exports. As explained below, these definitions hold strictly when the aggregates are evaluated at current prices (and at the average prices of the previous year).

A key observation is that the compilation of the GDP accounts requires processing and combining a variety of sources of information on the supply and demand sides of the economy, both primary and secondary, based on quantitative surveys of firms and households. The amount of data and the complexity of the process of reconciling the information do not make it currently possible for statistical agencies to compile the national accounts in a more timely way and at an observation frequency faster than quarterly.

The objective of this paper is to apply the methodology proposed in Proietti and Giovannelli (2020) to the problem of nowcasting and forecasting the production and expenditure national accounts, using a large-dimensional set of economic indicators available at the monthly frequency. Though we focus on the Italian case, the methods and main issues can be extended to other countries of the European Union, and to the euro area itself, which produces a harmonized system of national accounts (the European System of National and Regional Accounts, known as ESA 2010) and monthly economic indicators, with the same timing and nomenclature.

There is ample literature on nowcasting GDP with mixed-frequency data: one approach directly targets GDP at the quarterly frequency of observation, as in bridge models, see Baffigi et al. (2004), the MlXed frequency DAta Sampling approach (Ghysels et al., 2004, 2006), as in Kuzin et al. (2011), and the density combination approach by Aastveit et al. (2014). Another strand of the literature focuses on monthly GDP as the target variable; see Aruoba et al. (2012), Camacho and Pérez-Quirós (2010), Foroni and Marcellino (2014), Mitchell et al. (2005), and Proietti et al. (2017), among others. consider the problem of nowcasting the quarterly growth of the output and expenditure components of the euro area GDP and find that aggregating the components forecasts can result in an improvement of the quality of the nowcasts of total GDP growth. Bańbura and Rünstler (2011) deal with nowcasting the euro area aggregate GDP quarterly growth using a large-dimensional mixed-frequency factor model and propose a method to assess the contribution of each individual indicator to the nowcasts. Koop et al. (2020) estimate quarterly output growth for 12 UK regions with the support of a mixed-frequency vector autoregressive model subject to temporal and cross-sectional aggregation restrictions. Arencibia et al. (2017) nowcast Spanish monthly GDP and its demand components via a mixed-frequency parametric dynamic factor model.

We contribute to the literature by applying and illustrating a method for nowcasting and forecasting the 16 main components of GDP by output and expenditure type at the monthly frequency, using a high-dimensional set of monthly economic indicators, spanning the space of the common economic factors. As an application of the model averaging methodology proposed by Proietti and Giovannelli (2020), the projection of the quarterly components on the monthly common space is carried out by combining the individual nowcasts and forecasts

arising from all possible mixed-frequency bivariate models that jointly consider the quarterly national accounts component and the observed monthly indicator.

The components of GDP at market prices by output type that we consider are the value added of the 10 main branches of economic activity and taxes less subsidies:

| Label | Value added of branch | |
|-------|--|---|
| A–B | Agriculture, forestry, and fishing | + |
| C–D–E | Industry (except construction) | + |
| F | Construction | + |
| G–H–I | Wholesale and retail trade, transport, accommodation, and food service activities | + |
| J | Information and communication | + |
| K | Financial and insurance activities | + |
| L | Real estate activities | + |
| M–N | Professional, scientific, and technical activities; administrative and support service activities | + |
| O–Q | Public administration, defense, education, human health, and social work activities | + |
| R–U | Arts, entertainment, and recreation; other service activities; activities of household and extra-territorial organizations | = |
| | Total Gross Value Added (GDP at basic+ prices) | |
| TIS | Taxes less subsidies on products | = |
| | GDP at market prices | |

As for the breakdown of GDP by expenditure type, we consider the following components:

| Label | Component | |
|-------------------|---|---|
| FCE | Final consumption expenditure | |
| FCE _{gg} | Final consumption expenditure of general government | + |
| FCE _h | Household and NPISH final consumption expenditure | + |
| GCF | Gross capital formation | + |
| EXP | Exports of goods and services | – |
| IMP | Imports of goods and services | = |
| GDP | GDP at market prices | |

The quarterly series are taken at chained volumes with the reference year 2015 from the Italian quarterly national accounts compiled by Istat (<https://www.istat.it/en/national-accounts>). The information set also includes a set of $N = 433$ monthly indicators that are selected on the basis of their coverage, representativeness, and timeliness, and are currently monitored by the Italian Ministry of Economy and Finance.

For each of the above 16 components and the GDP we estimate N bivariate dynamic factor models that provide monthly nowcasts, forecasts, and historical estimates of the components. These are then pooled according to different aggregation strategies. The optimal model averaging method requires estimating the nowcast (forecast) error covariance matrix. We argue that the Ledoit–Wolf shrinkage estimator (Ledoit & Wolf, 2004a, 2004b), with

a compound symmetry shrinkage target, offers a solution that is both effective and feasible. Other weighting schemes, based on the goodness of fit of the indicator models within the training sample, and simple averaging are considered and compared.

Once consensus nowcasts and forecasts are obtained, they can be contemporaneously aggregated to form indirect GDP nowcasts and forecasts from the output and the expenditure side. This enables the comparison of the direct and indirect approaches to nowcasting and forecasting GDP.

The theoretical literature on the comparative merits of forecasting an aggregate by combining the components' forecasts is thoroughly reviewed in Lütkepohl (1986, 2011), for the case when the true model is vector ARMA. Poncela and García-Ferrer (2014) extended the results to the class of unobserved components models with shared and common trends that is used in this paper. Our contribution focuses on the comparison of direct and indirect predictions that use the same model averaging methodology, and thus it is in the same vein as Espasa et al. (2002), Heinisch and Scheufele (2018), Hendry and Hubrich (2011), and Hubrich (2005).

A recursive nowcasting and forecasting experiment provides firm evidence that the indirect approach provides more accurate nowcasts and forecasts than the direct approach. The experiment generates a wealth of information leading to several interesting discoveries, enabling us, among other things, to identify the contribution of the monthly indicators to the nowcasts and forecasts of monthly GDP and its components, highlighting the role of business and consumer surveys when information on hard indicators has not yet accrued. This also allows us to carry out a growth accounting exercise, which is informative about the contribution to aggregate growth of sectorial value added, final consumption, and investment during the phases of the business cycle.

The plan of the paper is the following. Section 2 presents the bivariate model for the GDP components and the individual indicators at the basis of our methodology. In Section 3 we outline how consensus nowcasts and forecasts are obtained for the GDP components. Section 4 deals with the contemporaneous aggregation of the GDP components according to the output and the expenditure approaches, so as to form indirect GDP nowcasts. The empirical illustration starts with a description of the data available (Section 5). The results are presented in Section 6: the monthly indirect GDP estimates and growth accounting decomposition are discussed in 6.1; and 6.2 presents the results of a recursive pseudo-real-time nowcasting and forecasting experiment. Section 7 concludes the paper.

2. Bivariate dynamic factor model for the GDP component and the monthly indicator

Let y_{it} denote the i th monthly unobservable GDP component at time t , for $i = 0, 1, \dots, M$, $t = 0, \dots, n$, and let $Y_{it} = y_{it} + y_{i,t-1} + y_{i,t-2}$ be the observable quarterly total. In our application, $M = 16$ and $i = 0$ corresponds to GDP. The information set also consists of N monthly indicators,

$\{x_{jt}, j = 1, \dots, N, t = 0, \dots, n\}$, which are available with a specific publication schedule. We shall denote by $\mathcal{Y}_{it} = \{Y_{i,3\tau}; \tau = 1, \dots, \lfloor (t-\delta)/3 \rfloor\}$ the information available at time t on the i th component from the national statistical agency, where δ is the delay in the release of the quarterly national accounts. Also, the information available at time t for the j th monthly indicator will be denoted by $\mathcal{X}_{jt} = \{x_{j1}, \dots, x_{j,t-\delta_j}\}$, where δ_j is the delay in their release with respect to the reference month.

Proietti and Giovannelli (2020) consider the problem of estimating the nowcasting or forecasting target $E(y_{i,t+l}|\mathcal{Y}_{it}, \mathcal{X}_t)$, where $\mathcal{X}_t = \bigcup_{j=1}^N \mathcal{X}_{jt}$ is the complete information on the monthly indicators and $l \geq 0$ is the forecast horizon, by combining the estimates of the partial projections $E(y_{i,t+l}|\mathcal{Y}_{it}, \mathcal{X}_{jt})$, $j = 1, \dots, N$, by means of the estimator

$$\hat{E}(y_{i,t+l}|\mathcal{Y}_{it}, \mathcal{X}_t) = \sum_{j=1}^N w_{ij} \hat{y}_{i,t+l}^{(j)}, \quad (1)$$

where $\hat{y}_{i,t+l}^{(j)} = \hat{E}(y_{i,t+l}|\mathcal{Y}_{it}, \mathcal{X}_{jt})$. As the nowcasts or forecasts $\hat{y}_{i,t+l}^{(j)}$ can be easily obtained by a suitably specified bivariate model that can handle the mixed frequency and ragged-edge structure of the information set, the main issue lies with the choice of the weights w_{ij} , $j = 1, \dots, N$. If $\hat{y}_{i,t+l}^{(j)}$ had a stationary distribution, and the monthly GDP component were observable, then the optimal weights could be estimated from the covariance matrix of the partial predictors $\hat{y}_{i,t+l}^{(j)}$ and their covariance with the target variable. However, neither of these two circumstances occur, and the optimal weights will depend on the covariance matrix of the prediction errors, estimated from a training or a test sample. This point will be followed up in Section 3. We now discuss the specification of the model leading to the partial estimates $\hat{y}_{i,t+l}^{(j)}$.

Our model assumes that the indicators are generated according to a nonstationary model, such that $\Delta x_{jt} = x_{jt} - x_{j,t-1}$ has an approximate factor structure, in the sense specified by Forni et al. (2000) and Forni and Lippi (2001):

$$x_{jt} = x_{j,t-1} + m_j + \chi_{jt} + \xi_{jt}, \quad j = 1, \dots, N, \quad (2)$$

where $m_j = E(\Delta x_{jt})$ is the constant drift; χ_{jt} represents the zero mean common component, which is a linear combination of q common dynamic factors driving the comovements among the series, whereas ξ_{jt} is the idiosyncratic component. The latter is specific to the j th series, or weakly correlated across a finite number of time series. As in Forni et al. (2000) and Forni and Lippi (2001), we write $\chi_{jt} = \mathbf{b}_j(L)' \mathbf{u}_t$, where \mathbf{u}_t is a $q \times 1$ vector of common dynamic factors, $\mathbf{u}_t \sim WN(\mathbf{0}, \mathbf{I}_q)$ and, denoting $b_{jr}(L) = \sum_{k \geq 0} b_{jr,k} L^k$, where L is the lag operator, $\mathbf{b}_j(L) = [b_{j1}(L), \dots, b_{jr}(L), \dots, b_{jq}(L)]'$. We assume that the common factors are pervasive and that ξ_{jt} and χ_{jt} are orthogonal at all leads and lags and across the series. Representation (2) has been at the core of applied macroeconomic analysis and forecasting; see, e.g., Forni et al. (2009) and Forni et al. (2018) and the references therein.

As far as the underlying monthly GDP component is concerned, we assume that it follows a similar structure,

$\Delta y_{it} = m_i + \mathbf{b}_i(L)' \mathbf{u}_t$, but with no idiosyncratic component. As a matter of fact, y_{it} results from the contemporaneous aggregation of many elementary component series, whose idiosyncratic components are averaged out. The projection of Δy_{it} on χ_{jt} is $\Delta y_{it} = m_i + \theta_{ij} \chi_{jt} + \xi_{it}$, where $\theta_{ij} = \sum_r \sum_k b_{irk} b_{jrk} / \sum_r \sum_k b_{jrk}^2$. The component ξ_{it} is orthogonal to χ_{jt} and to ξ_{jt} , by construction, since it measures the contribution of the common factors unaccounted for by χ_{jt} .

Under the additional assumption that χ_{jt} is generated by a parametric model, namely an ARMA(1,1) process, and that ξ_{it} is white noise, we consider the following bivariate model for the i th GDP component and the j th monthly indicator:

$$\begin{aligned} \begin{bmatrix} \chi_{jt} \\ y_{it} \end{bmatrix} &= \begin{bmatrix} \chi_{j,t-1} \\ y_{i,t-1} \end{bmatrix} + \begin{bmatrix} m_j \\ m_i \end{bmatrix} + \begin{bmatrix} 1 \\ \theta_{ij} \end{bmatrix} \chi_{jt} + \begin{bmatrix} \xi_{jt} \\ \xi_{it} \end{bmatrix}, \\ t &= 1, \dots, n, \\ \chi_{jt} &= \phi_{ij} \chi_{j,t-1} + \eta_{jt} - \vartheta_{ij} \eta_{j,t-1}, \quad \eta_{jt} \sim \text{IID } N(0, \sigma_{\eta_j}^2), \\ \begin{bmatrix} \xi_{jt} \\ \xi_{it} \end{bmatrix} &\sim \text{IID } N\left(\begin{bmatrix} 0 \\ 0 \end{bmatrix}, \begin{bmatrix} \sigma_j^2 & 0 \\ 0 & \sigma_i^2 \end{bmatrix}\right). \end{aligned} \quad (3)$$

In (3) we restrict the autoregressive coefficients ϕ_j to lie in the open interval $(0, 1)$ and ϑ_j in $[0, 1]$. The independence of ξ_{it} and ξ_{jt} follows from the structure of our model. On the contrary, serial independence is imposed for the identifiability of the model in the presence of temporal aggregation.

Model (3) provides a parsimonious representation capable of capturing the comovements between the GDP component and the monthly indicator. The ARMA(1,1) specification for the common component is sufficiently rich to capture its persistence. Obviously, it holds exactly if for the j th indicator $\mathbf{b}_j(L) = [0, \dots, 0, (1 - \vartheta_j L)/(1 - \phi_j L), 0, \dots, 0]$, and only approximately if χ_{jt} loads on multiple common shocks.

Summarizing, our bivariate model assumes that χ_{jt} and y_{it} are difference stationary and conditionally independent, given the common component of the indicator variable, χ_{jt} , which is assumed to follow an ARMA(1,1) process. The common component of the j th indicator is a dynamic weighted combination of the common factors, whose number and nature, to our advantage, does not need to be specified.

3. Nowcasting and forecasting the monthly GDP components

The baseline model (3) is estimated for all pairs (i, j) , $i = 0, \dots, M$, and $j = 1, \dots, N$, under the observational constraints posed by temporal aggregation, as only the quarterly totals, $Y_{it} = y_{it} + y_{i,t-1} + y_{i,t-2}$, are observable for times $t = 3\tau$, $\tau = 1, \dots, \lfloor n/3 \rfloor$, where $\lfloor \cdot \rfloor$ denotes integer division, with a fixed publication delay with respect to the quarter τ and a revision schedule.

Maximum likelihood estimation of the unknown parameters $(m_i, m_j, \theta_{ij}, \phi_{ij}, \vartheta_{ij}, \sigma_{\eta_j}^2, \sigma_i^2, \sigma_j^2)$ is performed by casting the model in state-space form and evaluating the likelihood by the associated Kalman filter. The technical details are given in Proietti and Giovannelli (2020).

Nowcasting and forecasting of the unobserved monthly GDP components, conditional on the maximum likelihood estimates of the parameters, is carried out with the support of the Kalman filter and updating equations, which compute the conditional means

$$\hat{y}_{i,t+l}^{(j)} = \hat{E}(y_{i,t+l} | \mathcal{Y}_{it}, \chi_{jt}), \quad l = 0, 1, \dots, L.$$

The N nowcasts and forecasts of the i th GDP component arising for $j = 1, \dots, N$, can be pooled to produce the conditional mean averaging estimator

$$\hat{E}_t(y_{i,t+l}) = \sum_{j=1}^N w_{ij} \hat{y}_{i,t+l}^{(j)}, \quad \sum_{j=1}^N w_{ij} = 1. \quad (4)$$

The weights $\{w_{ij}, j = 1, \dots, N\}$ should be chosen so as to minimize the mean square nowcast/forecast error; the latter can be estimated by performing a pseudo-real-time nowcasting/forecasting exercise. For instance, let $v_{ij,\tau} = Y_{i,3\tau} - E(Y_{i,3\tau} | \mathcal{Y}_{i,3\tau-1}, \chi_{j,3\tau})$, $j = 1, \dots, N$, $\tau = 1, \dots, T$ denote the nowcasting error in predicting the quarterly value of the GDP component on the basis of the information available in real time, and let $\hat{\Sigma}_i$ be a positive definite estimator of the covariance matrix of the nowcast errors $(v_{i1,\tau}, \dots, v_{ij,\tau}, \dots, v_{iN,\tau})'$. The optimal combination weights (Bates & Granger, 1969) are given by the elements of the vector

$$\hat{\mathbf{w}} = \frac{\hat{\Sigma}_i^{-1} \mathbf{i}}{\mathbf{i}' \hat{\Sigma}_i^{-1} \mathbf{i}}, \quad (5)$$

where \mathbf{i} is an $N \times 1$ vector of ones.

The main difficulty with implementing (5) lies with the fact that the cross-sectional dimension N is large compared to the number of available quarterly forecast errors, T , so that the sample covariance matrix $\hat{\Sigma}_i = \{\hat{s}_{i,hk}, (h, k) = 1, \dots, N\}$, $\hat{s}_{i,hk} = \frac{1}{T} \sum_{\tau} (v_{ih,\tau} - \bar{v}_{ih})(v_{ik,\tau} - \bar{v}_{ik})$, is singular. Moreover, it is a highly inaccurate estimator of the true nowcasting error covariance matrix. The solution is offered by the optimal linear shrinkage estimator proposed by Ledoit and Wolf (2004a, 2004b), which achieves positive definiteness and estimation accuracy by setting the elements of the matrix $\hat{\Sigma}_i$ as follows:

$$\hat{\Sigma}_{i,hk} = \{(1 - \lambda_i) \hat{s}_{i,hk} + \lambda_i \tilde{\omega}_{i,hk}, (h, k) = 1, \dots, N\}. \quad (6)$$

This is a weighted linear combination of the sample covariance and a shrinkage target covariance, where $\lambda_i \in [0, 1]$ is the shrinkage intensity, and $\tilde{\omega}_{i,hk}$ is the covariance between the nowcast errors estimated by assuming a compound symmetry covariance structure. This assumes that the nowcast errors are equally correlated, due to the presence of a single common factor. The correlation can be estimated using the average sample correlation: $\bar{r}_i = \frac{2}{(N-1)N} \sum_{h=1}^{N-1} \sum_{k=h+1}^N \frac{\hat{s}_{i,hk}}{\sqrt{\hat{s}_{i,hh} \hat{s}_{i,kk}}}$. Then, we set

$$\tilde{\omega}_{i,hk} = \begin{cases} \hat{s}_{i,hh}, & h = k, \\ \bar{r}_i \sqrt{\hat{s}_{i,hh} \hat{s}_{i,kk}}, & h \neq k. \end{cases}$$

The shrinkage parameter λ_i is estimated by minimizing the mean square nowcast error, e.g., by performing a grid search over the unit interval or applying the closed-form estimator by Ledoit and Wolf (2004a).

An alternative weighting scheme is obtained from the estimation output. Since the series are processed sequentially, see [Anderson and Moore \(1979\)](#) and [Koopman and Durbin \(2000\)](#), we can obtain the deviance measure:

$$D_{ij} = -2 \sum_{\tau=1}^{\lfloor n/3 \rfloor} \ln f(Y_{i,3\tau} | \mathcal{Y}_{i,3\tau-1}, \mathcal{X}_{j,3\tau}), \quad j = 1, \dots, N, \quad (7)$$

where $f(Y_{i,3\tau} | \mathcal{Y}_{i,3\tau-1}, \mathcal{X}_{j,3\tau})$ is the Gaussian nowcast density of the quarterly i th component, conditional on its available past information, consisting of the quarterly values up to quarter $\tau - 1$, and the monthly values of the i th indicator available up to the current time, $t = 3\tau$.

Then we can set

$$w_{ij} = \frac{\exp(-D_{ij})}{\sum_k \exp(-D_{ik})}. \quad (8)$$

As the number of parameters of the bivariate models remains constant across j , this weighting scheme is related to post-selection model averaging according to exponential AIC or BIC weights, see [Claeskens and Hjort \(2008\)](#) and the references therein.

4. The direct and indirect approach to GDP nowcasting and the issues posed by aggregation

When applied to aggregate GDP, our model averaging methodology delivers a direct nowcast/forecast of monthly GDP. An indirect nowcast/forecast can be obtained by the aggregation of the estimated monthly GDP components' nowcast/forecast.

As it is well known ([Bloem et al., 2001](#)), chain-linked volume measures are not consistent in cross-sectional aggregation. For example, the sum of the value added of the 10 branches plus taxes less subsidies would not deliver GDP at market prices.

The non-additivity is a consequence of the adoption of chain-linked methods in the construction of the national accounts. In particular, the annual overlap method is used, which calculates the chain-linked data with the annual data of the previous year as a point of reference. While the adoption is motivated by the need to remedy the problem of price structure obsolescence present in the old fixed-base approach, it has as a consequence the loss of additivity of volumes in all years except for the reference year and the year following the reference year.

Hence, in order to perform the aggregation of the estimates we need to restore the additivity by expressing the estimates at the average prices of the previous year. Subsequently, the GDP estimates can be chain-linked and expressed in chained volumes with the same reference year.

With a change of notation, let $y_{i,ms}^{(r)}$ denote the value of the i th GDP component in month m of year s evaluated at the average prices of year r . Let us assume that the initial year, $r = 0$, is taken as the reference year and that $i = 1, \dots, 11$ indexes the value added of the 10 branches and taxes less subsidies, which are the GDP components in the output approach. The value of the chained volume component can be expressed recursively as

$$y_{i,ms}^{(0)} = \left(\sum_m y_{i,m,s-1}^{(0)} \right) \frac{y_{i,ms}^{(s-1)}}{\sum_m y_{i,m,s-1}^{(s-1)}}, \quad s \geq 1. \quad (9)$$

The denominator is the value of the component in year $s - 1$ at current prices, so that the multiplicative factor is a Laspyres quantity index.

Our methodology provides estimates of $y_{i,ms}^{(0)}$ on the left-hand side of (9), and by inverting the above expression, we can obtain the component evaluated at the previous year average prices:

$$y_{i,ms}^{(s-1)} = y_{i,ms}^{(0)} \frac{\sum_m y_{i,m,s-1}^{(s-1)}}{\sum_m y_{i,m,s-1}^{(0)}}. \quad (10)$$

This operation can be referred to as *dechaining*. Notice that $\sum_m y_{i,m,s-1}^{(s-1)}$ is the annual total at current prices and this is available from the official national accounts.

At this point, the GDP components can be aggregated. Summing with respect to i we obtain the GDP at the previous year's prices:

$$y_{ms}^{(s-1)} = \sum_i y_{i,ms}^{(s-1)}. \quad (11)$$

As a final step we chain-link the GDP estimates to express them in chained volumes with reference year 0. This is achieved by computing recursively:

$$y_{ms}^{(0)} = \left(\sum_m y_{m,s-1}^{(0)} \right) \frac{y_{ms}^{(s-1)}}{\sum_m y_{m,s-1}^{(s-1)}}. \quad (12)$$

Note that, using (10) and (11), (12) can be written as follows:

$$y_{ms}^{(0)} = \sum_i y_{i,ms}^{(0)} \frac{P_{i,s-1}}{P_{s-1}}, \quad (13)$$

with

$$P_{i,s-1} = \frac{\sum_m y_{i,m,s-1}^{(s-1)}}{\sum_m y_{i,m,s-1}^{(0)}}, \quad P_{s-1} = \frac{\sum_m y_{m,s-1}^{(s-1)}}{\sum_m y_{m,s-1}^{(0)}}.$$

Hence, we can express chained volumes GDP as a weighted linear combination of the components' chained volumes, where the weights are the ratio of the component Laspyres price index for the previous year with respect to the reference year and the price index of GDP for the same year.

The discussion in this section and expression (13) make it clear that GDP growth cannot be decomposed into a weighted average of the components' growth rates. Hence, the lack of additivity is an issue for any growth accounting exercise that is conducted with "real" aggregates, i.e., chained volumes.

With reference to the annual growth rates, after some manipulation of (13),

$$\frac{y_{ms}^{(0)} - y_{m,s-1}^{(0)}}{y_{m,s-1}^{(0)}} = \sum_i \left[\frac{y_{i,ms}^{(0)} - y_{i,m,s-1}^{(0)}}{y_{i,m,s-1}^{(0)}} W_{i,m,s-1}^{(0)} \frac{P_{i,s-1}}{P_{s-1}} + W_{i,m,s-1}^{(0)} \left(\frac{P_{i,s-1}}{P_{s-1}} - \frac{P_{i,s-2}}{P_{s-2}} \right) \right], \quad (14)$$

where

$$W_{i,m,s-1}^{(0)} = \frac{y_{i,m,s-1}^{(0)}}{y_{m,s-1}^{(0)}}.$$

As a result, the weighted sum of the yearly growth rates $\frac{y_{i,ms}^{(0)} - y_{i,m,s-1}^{(0)}}{y_{i,m,s-1}^{(0)}}$ does not add up to GDP growth and can only be seen as an approximation to the contribution of the i th component to aggregate growth.

5. Description of the dataset

Our dataset comprises the time series of GDP and its 16 components taken from the production and expenditure quarterly national accounts compiled by ISTAT (Istituto Nazionale di Statistica; see [Bisio and Moauro \(2018\)](#) for a thorough discussion of the methodology). The series are available for the period 1996.q1–2019.q3 at chained volumes. The corresponding series at current prices were used to compute the weights for contemporaneous aggregation according to expression (13).

A collection of $N = 433$ monthly time series currently in use at the Italian Ministry of Economics and Finance, collected from various sources and electronic databases, completes our dataset. The selection criteria were representativeness, coverage, timeliness, and availability: the set of monthly indicators need to provide ample coverage of economic and financial phenomena; indicators that are released with a publication delay longer than three months are unlikely to provide useful information for the objectives of this paper. The series cover the period 1996.m1–2019.m12 with a ragged-edge structure. A complete list is available in [Appendix B](#), which also reports the publication delay with respect to the reference month.

The indicators can be grouped according to their main theme and geographical reference: [Table 1](#) provides the distribution of the indicators according to the two characteristics. While one-fourth of the indicators refer to the Italian economy, the dataset also contains a number of indicators pertaining to the euro area, to account, e.g., for the developments of the German manufacturing sector, which are very important for the Italian one. More than one-third of the indicators relate to consumer and business surveys, which provide a timely and easily accessible assessment of the state of the economy.

Overall, the dataset is unbalanced towards the supply side of the economy: quantitative consumer surveys are much less timely and serve more the compilation of the national accounts (and microeconomic research). The qualitative surveys partially compensate for this. Reweighting the individual contribution of the indicators according to their ability to predict the national accounts target aims at readdressing the balance.

The preliminary transformations that are adopted for modeling the series according to (3) are the following: all the consumer and production price indices are transformed into monthly inflation rates; the survey series are cumulated: if z_{jt} denotes a survey series, then $x_{jt} = \sum_{r=0}^t z_{jr}$. All the other series are taken in levels.

6. Empirical results

6.1. Historical estimation and growth accounting

Historical estimation deals with the construction of monthly GDP estimates, by expenditure and output components, conditional on the full sample available. For this

Table 1

Breakdown of the 433 monthly indicators by theme and geographical reference.

| Theme | Geographical reference | | | Total |
|----------|------------------------|-------|-------|-------|
| | Euro area | Italy | Other | |
| Industry | 3.2 | 8.1 | 1.2 | 12.5 |
| Trade | – | 7.9 | 0.2 | 8.1 |
| Services | – | 6.2 | – | 6.2 |
| Finance | 3.9 | 4.6 | 3.7 | 12.2 |
| Labor | – | 5.5 | – | 5.5 |
| Prices | 3.9 | 9.9 | 2.1 | 15.9 |
| Survey | 6.0 | 30.7 | 0.9 | 37.6 |
| Other | – | 1.8 | – | 1.8 |
| Total | 17.1 | 75.1 | 8.1 | 100.0 |

purpose the weights can be validly obtained by considering the in-sample predictive performance, as measured by the conditional deviance of the quarterly GDP estimation error.

Using the batch of data described in the previous section, the N monthly estimates of the GDP components, $\tilde{y}_{it}^{(j)} = \hat{E}(y_{it} | \mathcal{Y}_{in}, \mathcal{X}_{jn})$, have been computed and are later combined to produce the historical estimates $\tilde{y}_{it} = \sum_{j=1}^N w_{ij} \tilde{y}_{it}^{(j)}$, $t = 1, \dots, n$. The weights used for aggregation are proportional to $\exp(-D_{ij})$, given in (7), and thus are an expression of the conditional likelihood of the observed quarterly GDP series.

The GDP components are then aggregated to form GDP monthly estimates according to the output and expenditure approach using the methodology outlined in [Section 4](#). If we replace the time index with two indices (m, s), respectively labeling the month, $m = 1, \dots, 12$, and the year, $s = 1, \dots, \lfloor n/12 \rfloor$, so that $t = m + 12(s - 1)$, and if i indexes the value added of the 10 branches and taxes less subsidies, GDP by output components is obtained as

$$\tilde{y}_{ms}^{(o)} = \sum_i \tilde{y}_{i,ms} \frac{P_{i,s-1}}{P_{s-1}},$$

where the cross-sectional aggregation weights are the ratio of the i th component deflator of year $s - 1$ with respect to the reference year ($P_{i,s-1}$) to the GDP deflator for the same year (P_{s-1}). Both are obtained by the ratio of the annual value added or GDP at current prices to that at chained volumes, available from the national accounts. For the expenditure approach we apply the same aggregation rule, but imports have to be subtracted:

$$\tilde{y}_{ms}^{(e)} = \sum_i \tilde{y}_{i,ms} \frac{P_{i,s-1}}{P_{s-1}} - \tilde{y}_{IMP,ms} \frac{P_{IMP,s-1}}{P_{s-1}},$$

where $i = \{FCE_{gg}, FCE_h, GCF, EXP\}$.

The monthly GDP estimates are presented in [Fig. 1](#) for the two approaches, along with the annual growth rate. The interval estimates refer to a nominal coverage probability set equal to 0.95. They are based on the assumptions that the estimates of the components are independent both within and between, i.e., across j and i , and that the conditional density of y_{it} is Gaussian. In particular, the variance of the monthly estimate for each component is computed as the average (with weights w_{ij}) of the

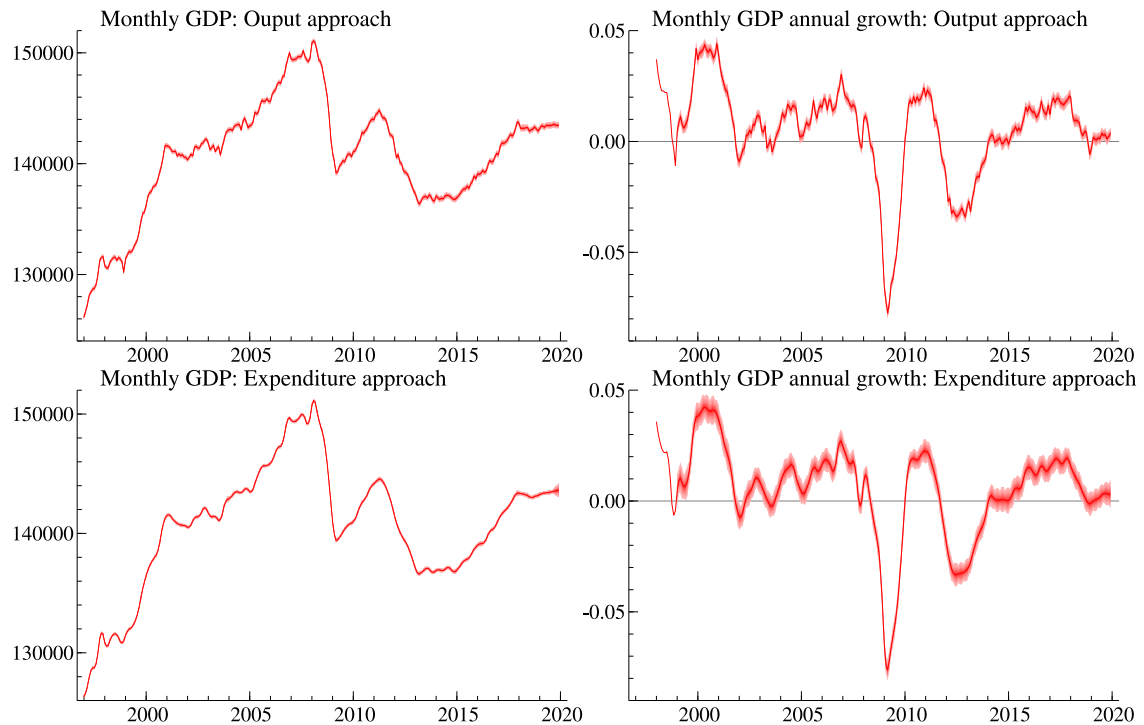


Fig. 1. Historical estimates of monthly GDP at market prices (chained volumes 2015) and its yearly relative changes, according to the output approach (top panels) and the expenditure approach (bottom panels).

variances $\widehat{\text{Var}}(y_{it}|\mathcal{Y}_{in}, \mathcal{X}_{jn})$, provided by the Kalman filter and smoother (KFS), plus the variance of the monthly estimates of the same component. This second source of variability typically plays a minor role, except at the end of the sample, where the variability of the predictions is higher and can vary relevantly depending on the indicator used. The underlying idea is that the conditional distribution of y_{it} can be characterized as a mixture of N Gaussian distributions with mean \hat{y}_{it} and variance $\widehat{\text{Var}}(y_{it}|\mathcal{Y}_{in}, \mathcal{X}_{jn})$, each one arising from selecting the j th indicator, with mixture probabilities $w_{ij}, j = 1, \dots, N$.

The conditional variance of the growth rates is computed similarly, where $\widehat{\text{Var}}(\Delta_{12}y_{it}/y_{i,t-12}|\mathcal{Y}_{in}, \mathcal{X}_{jn})$ is obtained by the delta method based on the output of the KFS. Hence, the confidence bands capture the so-called “filtering” uncertainty, which arises from the fact that only the quarterly sums are available for the components of GDP and the variability of the estimates that are obtained by conditioning on a different indicator.

In interpreting the plots, it must be brought in mind that neither parameter uncertainty nor model uncertainty is considered. A more complete assessment of these additional sources of uncertainty, along the lines of [Corona et al. \(2020\)](#) and [Proietti et al. \(2017\)](#), is left to future research.

Inspection of [Fig. 1](#) induces the following considerations:

1. The filtering and indicator uncertainty is not large in historical estimation. The pseudo-real-time estimation experiment, not conditioning on the past

and future history of the national accounts components, will bring to the surface the kind of uncertainty that is faced in real time. In [Fig. 1](#) only the last three estimates, referring to October–December 2019 can be considered as a nowcast; in fact, the interval estimates widen.

2. Higher uncertainty characterizes the monthly estimates of GDP using the expenditure approach. This stylized fact, previously documented by [Frale et al. \(2011\)](#) and [Proietti et al. \(2017\)](#), is related to the availability of representative and relevant monthly indicators. A tentative explanation is that the statistical system is highly unbalanced towards the supply side of the economy.
3. The monthly estimates arising from the expenditure approach are generally smoother. This is a consequence of the higher volatility and idiosyncrasy of the shocks affecting industry and construction, which account for a large share of total value added in Italy.
4. For brevity, the estimates of monthly GDP arising from the direct approach are not presented. They are qualitatively similar and are a compromise between the two estimates from the expenditure and the output side. Also, the estimation error uncertainty lies in between the two approaches.

The availability of the estimates of the individual components opens the way to performing a growth accounting exercise. [Figs. 2](#) and [3](#) present the contribution of the individual output components and expenditure components to the aggregate GDP annual growth rate.

The contribution is computed by multiplying the weight of the component as a fraction of total GDP, calculated on the annual value at current prices for the previous year, multiplied by the growth rate of the component, i.e., in the notation of (14), $\frac{y_{i,m,s}^{(0)} - y_{i,m,s-1}^{(0)}}{y_{i,m,s-1}^{(0)}} W_{i,m,s-1}^{(0)}$. As a result, the decomposition of GDP growth is additive only up to an approximation that depends on the time pattern of the deflators of the components, relative to that of GDP.

This provides valuable information on the sources of growth in any particular month. For instance, during the Great Recession the strongest negative contribution was given by the industrial value added, followed with a lag by Sector G-H-I. These sectors also provide the largest positive contribution to GDP growth. The contribution of the construction sector has been negative thereafter until 2014. Export and household consumption are the components making the largest contribution to growth in 2016 from the expenditure side. The service sectors make a contribution that is relatively small; the most procyclical one is made by the value added of self-employed professionals in Sectors M–Z. The contribution of TLS is obviously strongly correlated to that of the industrial and trading sectors.

From the expenditure side, investments made the largest negative contribution during the Great Recession; interestingly, during the sovereign debt crisis (2011–2013) final household consumption made a sizable contribution, much larger than during the Great Recession. Net exports made a negative contribution to GDP growth during the Great Recession, and a positive one during the sovereign debt crisis, which compensated for the strong and negative growth in investment and private consumption.

6.2. A recursive pseudo-real-time nowcasting and forecast-experiment

6.2.1. Design of the experiment

We consider an econometrician using our methodology in real time to nowcast and forecast GDP and compare the predictions with the realized values. Her objectives would be multifold: (i) assess the predictive ability of the model for the components and aggregate GDP, both per se and comparatively; (ii) decide whether the aggregation of the components offers advantages over predicting GDP directly using the N indicators; (iii) evaluate the importance of the monthly indicators and their variability during business-cycle phases.

To address these issues she implements a recursive pseudo-real-time experiment, which at the end of each month produces a nowcast and a forecast of quarterly GDP by aggregating the monthly estimates. The design follows the flow of economic information available in three typical situations, positioned at the end of the three months making up each quarter.

Starting from January 2008, for every month m , $m = 1, 2, 3$, of reference quarter τ , she produces the nowcasts $\hat{E}(y_{i,3\tau-k} | \mathcal{Y}_{it}, \mathcal{X}_{jt})$, for $k = 0, 1, 2$, and $t = m + 3(\tau - 1)$, for all the GDP components and GDP itself, for all $i = 1, \dots, M$ and $j = 1, \dots, N$. Furthermore, she produces the

forecasts for the three months making up quarter $\tau + 1$, $\hat{E}(y_{i,3(\tau+1)-k} | \mathcal{Y}_{it}, \mathcal{X}_{jt})$, $k = 0, 1, 2$.

Temporal aggregation across the index k will produce the nowcasts of the components and GDP for quarter τ , $\hat{Y}_{ij,\tau} = \sum_{k=0}^2 \hat{E}(y_{i,3\tau-k} | \mathcal{Y}_{it}, \mathcal{X}_{jt})$, that she can compare with the quarterly published value, $Y_{i\tau}$. Let $v_{ij,\tau}$ denote the nowcast error $Y_{i\tau} - \hat{Y}_{ij,\tau}$. As for the next quarter forecast, she computes the prediction $\hat{Y}_{ij,\tau+1}^{(f)} = \sum_{k=0}^2 \hat{E}(y_{i,3(\tau+1)-k} | \mathcal{Y}_{it}, \mathcal{X}_{jt})$ and the corresponding forecast error $v_{ij,\tau+1}^{(f)} = Y_{i,\tau+1} - \hat{Y}_{ij,\tau+1}^{(f)}$.

At the end of the m th month she will have estimated N models for the i th component and she will combine the individual nowcasts with weights determined according to Section 3—either from the deviance of the estimated model or from the Ledoit–Wolf optimal shrinkage estimator, using $v_{ij,\tau}$ or $v_{ij,\tau+1}^{(f)}$, respectively—to construct a regularized estimator of the nowcast and forecast error covariance matrix. The model averaging nowcasts and forecasts are then $\hat{Y}_{i,\tau} = \sum_{j=1}^N w_{ij,t} \hat{Y}_{ij,\tau}$ and $\hat{Y}_{i,\tau+1}^{(f)} = \sum_{j=1}^N w_{ij,t}^{(f)} \hat{Y}_{ij,\tau+1}^{(f)}$. The aggregation weights $w_{ij,t}$ and $w_{ij,t}^{(f)}$ are indexed by t since they are re-estimated in each month.

Along with the direct nowcast of GDP she produces the indirect ones by performing the cross-sectional aggregation according to the method outlined in Section 4. Hence, three GDP nowcasts and forecasts will be made available: the direct one and two indirect ones from the output and the expenditure approaches. The nowcast and one-quarter-ahead forecast errors constitute the basis for the aggregation weights and for the assessment of the methodology.

It is assumed that the exercise is conducted at the end of each month m of quarter τ for all the months from January 2008 up to December 2018. While the nowcast and forecast target remain the same, the information set available to the econometrician varies with m . In particular, it increases as we progress towards the end of the reference quarter.

The first nowcast and forecast are made at the end of the first month of the reference quarter ($m = 1$, i.e., January, April, July, and October). The information available on GDP of the reference quarter and its components includes the soft indicators (business and consumer surveys), which are available up to the current month (all the indicators of Table B.1 with publication delay equal to 0 in the penultimate column), whereas the hard indicators are available according to a ragged-edge structure, such that, e.g., industrial production is available up to the last month of the previous quarter (all the indicators with publication delay equal to 1 in Table B.1). As far as the target series is concerned, the information set also includes the GDP series up to the previous quarter¹ ($\tau - 1$).

¹ The flash estimate of the GDP of quarter $\tau - 1$ would have been just published; since this is not available for a large part of the test sample (it becomes available only from April 2016), we replace it by the quarterly GDP figure published by Istat.

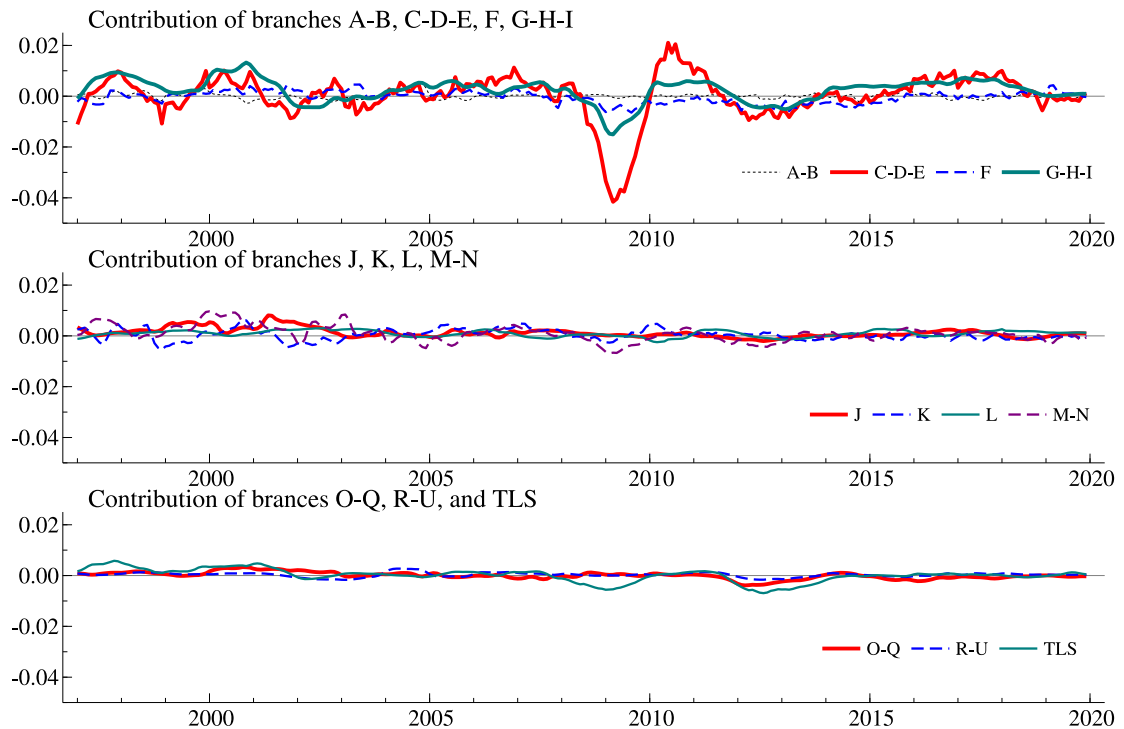


Fig. 2. Growth accounting: contribution of the 11 output components to yearly GDP growth.

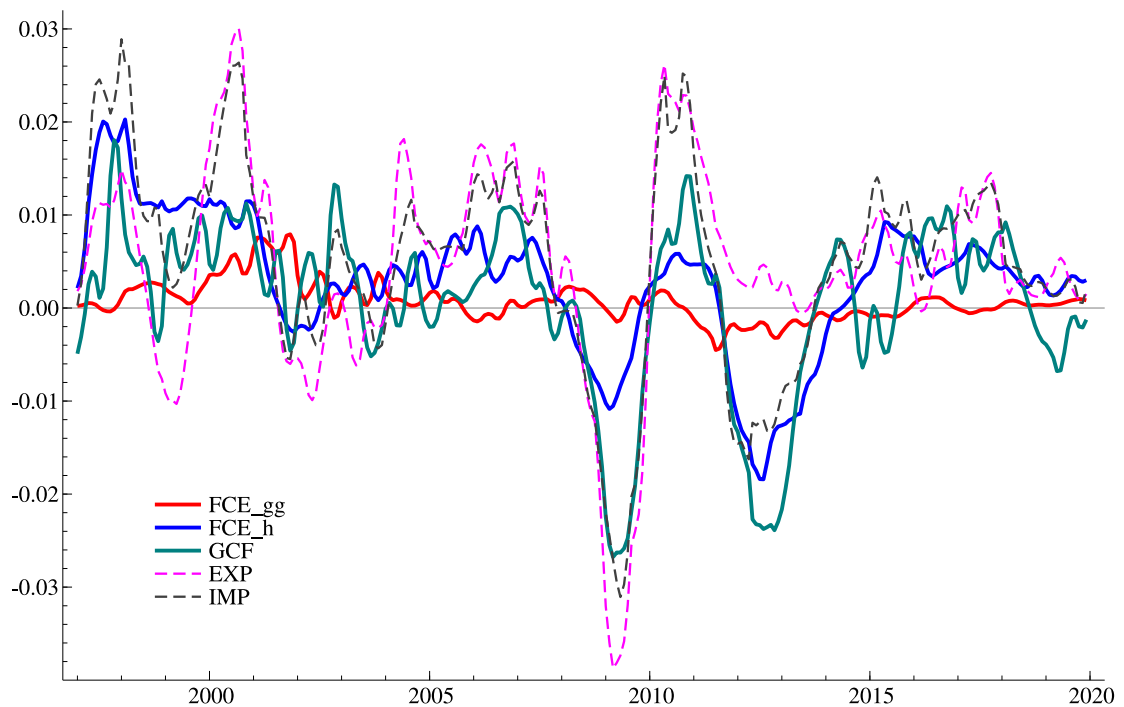


Fig. 3. Growth accounting: contribution of the five expenditure components (bottom panel) to yearly GDP growth. That of IMP has the sign reversed.

The second nowcast and forecast are made the last day of the second month of the reference quarter ($m = 2$, i.e., February, May, August, and November). The information set available includes the soft indicators up to

the second month of the reference quarter, producer and consumer prices, and monetary and financial aggregates for the first month of the quarter become. Moreover, the monthly information concerning quarter $\tau - 1$ is

complete, except for building permits and production in construction.

The third nowcast and forecast are made at the end of the reference quarter ($m = 3$, March, June, September, and December). The information set on the previous quarter is complete, but also that concerning the current quarter is relatively rich, featuring the index of industrial production and retail sales for the first month of the quarter.

6.2.2. Empirical assessment of predictive ability

The main results, referring to GDP, are presented in [Tables 2](#) and [4](#). Detailed results for the individual GDP components are available in the Supplement. The tables report the root mean square errors (RMSEs), computed over $T = 44$ prediction errors, for the test sample 2008.q1–2018.q4 (nowcasting) and 2008.q2–2019.q1 (forecasting) of the direct ([Table 2](#)) and indirect nowcast and forecast of the Italian GDP, by the output approach ([Table 3](#)), and the expenditure approach ([Table 4](#)).

Each table is divided into four panels: the first reports the RMSEs of four model averaging strategies using the deviance-based weights (8), with no selection (DEV_N), and selecting the 10, 30, and 50 best performing bivariate models in terms of their in-sample performance, as measured by the deviance within the training sample.

The second panel refers to the pooling scheme (5), using the Ledoit–Wolf (LW) estimator of the nowcast and forecast error covariance matrix. The latter is constructed according to (6), using the out-of-sample prediction errors for the last three years of data before time t . The shrinkage intensity is estimated according to [Appendix A](#).

We also consider an equally weighted model averaging scheme (Panel C) and compare the predictive accuracy of our pooled nowcasts and forecasts with a univariate benchmark and a multivariate one based on a dynamic factor model (DFM, Panel D). The univariate benchmark latter is an $ARIMA(p,1,0)$ model selected recursively according to the Bayesian information criterion (BIC) applied to quarterly GDP, for which we report the one-step-ahead RMSE under the nowcasting column and the two-step-ahead RMSE in the forecasting one. The benchmark should provide an upper bound to the nowcasting and forecasting RMSE in the absence of information concerning the monthly indicators.

A more informative benchmark arises from extracting from the monthly indicators r dynamic common factors, following [Stock and Watson \(2002a, 2002b\)](#). The forecasts of (the) quarterly GDP (components) are obtained from an autoregressive model augmented by the estimated common factors:

$$\hat{Y}_{i,\tau+h} = \hat{c} + \sum_{j=1}^r \hat{\gamma}_j \hat{F}_{j\tau} + \sum_{k=1}^p \hat{\phi}_k Y_{i,\tau-j+1},$$

where Y denotes the GDP component, $h \geq 1$ is the forecast lead, and $\hat{F}_{j\tau}$ is the j th estimated factor, obtained by averaging the monthly factor scores for the three months composing the quarter. The monthly factors are obtained by principal components analysis. The ragged-end structure of the dataset has been handled by applying

the expectation–maximization-type algorithm proposed by [Stock and Watson \(2002a, 2002b\)](#). The parameters are estimated by least squares. Regarding the model specification, the number of factors is estimated through the information criterion proposed by [Alessi et al. \(2010, ABC\)](#), whereas the lag order p is estimated according to the BIC. We also tested the inclusion of lags for $\hat{F}_{j\tau}$, and using the selection criteria by [Bai and Ng \(2002\)](#), but the results were not significantly better and are not reported here for brevity. Finally, the indirect predictions are obtained by means of the aggregating scheme according to [Section 4](#). The best results were obtained with a number of factors selected a priori ($r = 4$). For the indirect approach, the forecasts were aggregated using the annual overlap method described in [Section 6.1](#).

The main evidence arising from [Table 2](#) can be summarized as follows.

- The monthly indicators carry information that is useful for nowcasting and forecasting GDP directly. Their use leads to a systematic and sizable reduction of the RMSE with respect to the univariate and DFM benchmarks. The RMSE reduction can be as large as 28% (19%) in nowcasting (forecasting) GDP at the end of the reference quarter ($m = 3$) with the LW predictor. The DFM approach using four factors shows superior performance only when forecasting GDP with the expenditure approach (see [Table 4](#)).
- As the monthly information on the GDP of the reference quarter increases, the nowcast and forecast accuracy improves: this is evidenced by the monotonic decrease in the RMSE as we move from $m = 1$ to $m = 3$.
- The best performing method for aggregating the individual GDP predictions is provided by the optimal weighting scheme in (5) using the [Ledoit and Wolf \(2004a\)](#) shrinkage estimator of the nowcast and forecast error covariance matrix.
- The second best strategy is to adopt deviance-based averaging. In general, it is beneficial to perform a preliminary screening, and restricting attention to the best 50 bivariate nowcasts and forecasts leads to the best performance in the subset.
- Simple averaging does not compare well with the other model averaging schemes.

The superior performance of the LW aggregate nowcasts and forecasts can be ascribed to the fact that the weights optimize the local out-of-sample predictive performance, based on the last two years of recursive prediction errors, as well as to the effectiveness of the shrinkage estimator (6).

The above findings are confirmed for both indirect nowcasting and forecasting methods. For comparison, [Tables 3](#) and [4](#) report the RMSE of the indirect univariate and multivariate benchmarks. While there is no evidence that the indirect univariate predictor is superior to the direct one (the nowcast RMSE is smaller for the direct method, but the forecast RMSE is larger), there is decisive support for the indirect model averaging method, especially when the LW aggregation weights are adopted. In particular, the GDP nowcasts and forecasts by output,

Table 2

Nowcasting and forecasting results for Italy GDP (direct approach). The table reports the mean square prediction errors associated with the different methods. The predictand is quarterly GDP for the sample periods 2008.q1–2018.q4 (nowcasting) and 2008.q2–2019.q1 (forecasting).

| | PANEL A: Deviance (In sample) | | | | | |
|-----------------------|--|---------|---------|-----------------------------|---------|---------|
| | Nowcasting current quarter | | | Forecasting 1 quarter ahead | | |
| | $m = 1$ | $m = 2$ | $m = 3$ | $m = 1$ | $m = 2$ | $m = 3$ |
| | | | | | | |
| DEV_N | 2823.8 | 2753.0 | 2625.3 | 6098.9 | 5945.0 | 5699.3 |
| DEV_{10} | 2925.9 | 2803.9 | 2405.4 | 6154.7 | 6052.3 | 5554.6 |
| DEV_{30} | 2790.4 | 2659.0 | 2328.6 | 6087.3 | 5925.4 | 5489.2 |
| DEV_{50} | 2759.9 | 2609.6 | 2310.0 | 6002.0 | 5781.1 | 5369.9 |
| | PANEL B: Ledoit – Wolf (Out of sample) | | | | | |
| | Nowcasting current quarter | | | Forecasting 1 quarter ahead | | |
| | $m = 1$ | $m = 2$ | $m = 3$ | $m = 1$ | $m = 2$ | $m = 3$ |
| | | | | | | |
| $OPT_{LW}^{\lambda*}$ | 2659.6 | 2530.6 | 2294.7 | 5688.1 | 5376.9 | 4961.5 |
| | PANEL C: Simple average | | | | | |
| | Nowcasting current quarter | | | Forecasting 1 quarter ahead | | |
| | $m = 1$ | $m = 2$ | $m = 3$ | $m = 1$ | $m = 2$ | $m = 3$ |
| | | | | | | |
| | 2827.1 | 2765.6 | 2651.9 | 6106.9 | 5959.3 | 5732.6 |
| | PANEL D: Benchmarks | | | | | |
| | Nowcasting current quarter | | | Forecasting 1 quarter ahead | | |
| | $m = 1$ | $m = 2$ | $m = 3$ | $m = 1$ | $m = 2$ | $m = 3$ |
| | | | | | | |
| $AR(BIC)^{Dir}$ | 2930.5 | | | 5918.3 | | |
| $DFM_{ABC,IC1}$ | 5467.6 | 5209.2 | 3976.7 | 8377.2 | 8757.5 | 5872.5 |
| $DFM_{T=4}$ | 2788.1 | 3517.8 | 2686.2 | 6075.0 | 5919.2 | 6269.4 |

Table 3

Nowcasting and forecasting results for output-side GDP (chain-link method). The table reports the mean square prediction errors associated with the different methods. The predictand is quarterly GDP for the sample periods 2008.q1–2018.q4 (nowcasting) and 2008.q2–2019.q1 (forecasting).

| | PANEL A: Deviance (In sample) | | | | | |
|-----------------------|--|---------|---------|-----------------------------|---------|---------|
| | Nowcasting current quarter | | | Forecasting 1 quarter ahead | | |
| | $m = 1$ | $m = 2$ | $m = 3$ | $m = 1$ | $m = 2$ | $m = 3$ |
| | | | | | | |
| DEV_N | 2838.6 | 2756.5 | 2648.5 | 5946.3 | 5791.9 | 5614.9 |
| DEV_{10} | 2784.7 | 2628.0 | 2163.3 | 5817.4 | 5620.8 | 5133.6 |
| DEV_{30} | 2748.6 | 2574.1 | 2208.1 | 5873.9 | 5576.5 | 5117.8 |
| DEV_{50} | 2687.7 | 2521.6 | 2202.5 | 5741.0 | 5459.3 | 5038.0 |
| | PANEL B: Ledoit – Wolf (Out of sample) | | | | | |
| | Nowcasting current quarter | | | Forecasting 1 quarter ahead | | |
| | $m = 1$ | $m = 2$ | $m = 3$ | $m = 1$ | $m = 2$ | $m = 3$ |
| | | | | | | |
| $OPT_{LW}^{\lambda*}$ | 2219.9 | 2221.3 | 1829.2 | 4669.7 | 5149.3 | 4594.6 |
| | PANEL C: Simple average | | | | | |
| | Nowcasting current quarter | | | Forecasting 1 quarter ahead | | |
| | $m = 1$ | $m = 2$ | $m = 3$ | $m = 1$ | $m = 2$ | $m = 3$ |
| | | | | | | |
| | 2887.6 | 2780.4 | 2693.8 | 6006.3 | 5834.9 | 5680.1 |
| | PANEL D: Benchmarks | | | | | |
| | Nowcasting current quarter | | | Forecasting 1 quarter ahead | | |
| | $m = 1$ | $m = 2$ | $m = 3$ | $m = 1$ | $m = 2$ | $m = 3$ |
| | | | | | | |
| $AR(BIC)^{Ind}$ | 2982.8 | | | 5909.5 | | |
| $DFM_{ABC,IC1}$ | 5488.6 | 5198.4 | 3767.9 | 8400.1 | 8565.2 | 5629.1 |
| $DFM_{T=4}$ | 3155.5 | 3645.6 | 3076.7 | 5890.9 | 5797.0 | 6277.1 |

resulting from the contemporaneous aggregation of the value added of the branches, are the most accurate, despite a non-monotonic behavior with respect to m , with accuracy gains that reach 60% in nowcasting GDP at the end of the third month of the reference quarter. The expenditure approach offers a relative improvement only when starting from $m = 2$.

6.2.3. Inside the black box

When and how did the predictive accuracy gain of the LW pooling method accrue? What indicators contributed most to the outcome? What are the sources of the comparative advantages of the indirect approach?

To address these questions we set off by evaluating the local relative predictive performance of the LW model averaging method with respect to the univariate ARIMA(1,1,0) predictor by means of the fluctuation test

Table 4

Nowcasting and forecasting results for expenditure-side GDP (chain-link method). The table reports the mean square prediction errors associated with the different methods. The predictand is quarterly GDP for the sample periods 2008.q1–2018.q4 (nowcasting) and 2008.q2–2019.q1 (forecasting).

| | PANEL A: Deviance (In sample) | | | | | |
|-----------------------|--|---------|---------|-----------------------------|---------|---------|
| | Nowcasting current quarter | | | Forecasting 1 quarter ahead | | |
| | $m = 1$ | $m = 2$ | $m = 3$ | $m = 1$ | $m = 2$ | $m = 3$ |
| | | | | | | |
| DEV_N | 2941.4 | 2854.8 | 2748.6 | 6059.7 | 5912.1 | 5719.9 |
| DEV_{10} | 2943.3 | 2764.5 | 2205.7 | 5794.7 | 5658.0 | 5183.7 |
| DEV_{30} | 2779.8 | 2565.0 | 2181.2 | 5756.6 | 5478.2 | 5044.8 |
| DEV_{50} | 2764.6 | 2540.8 | 2245.1 | 5773.8 | 5489.4 | 5098.3 |
| | PANEL B: Ledoit – Wolf (Out of sample) | | | | | |
| | Nowcasting current quarter | | | Forecasting 1 quarter ahead | | |
| | $m = 1$ | $m = 2$ | $m = 3$ | $m = 1$ | $m = 2$ | $m = 3$ |
| | | | | | | |
| $OPT_{LW}^{\lambda*}$ | 3118.2 | 2230.0 | 2081.6 | 6203.3 | 4548.6 | 4667.1 |
| | PANEL C: Simple average | | | | | |
| | Nowcasting current quarter | | | Forecasting 1 quarter ahead | | |
| | $m = 1$ | $m = 2$ | $m = 3$ | $m = 1$ | $m = 2$ | $m = 3$ |
| | | | | | | |
| | 2948.9 | 2878.1 | 2788.7 | 6067.9 | 5950.7 | 5778.1 |
| | PANEL D: Benchmarks | | | | | |
| | Nowcasting current quarter | | | Forecasting 1 quarter ahead | | |
| | $m = 1$ | $m = 2$ | $m = 3$ | $m = 1$ | $m = 2$ | $m = 3$ |
| | | | | | | |
| $AR(BIC)^{Ind}$ | 3353.6 | | | 5588.9 | | |
| $DFM_{ABC,JC1}$ | 5054.5 | 4739.5 | 3678.7 | 7307.6 | 7810.8 | 5262.0 |
| $DFM_{\tau=4}$ | 2542.7 | 3069.4 | 2792.0 | 5010.0 | 4728.9 | 5472.2 |

statistic proposed by [Giacomini and Rossi \(2010\)](#). The statistic is the standardized average squared loss differential of the univariate and LW prediction errors calculated over rolling windows of two years.

[Fig. 4](#) displays the fluctuation tests for the direct LW case, according as to whether the nowcast (left panels) or nowcasts (right panels) are performed in months $m = 1, 2, 3$, of the reference quarter (nowcast) or of the quarter before (forecast). Values above zero signify that the LW method outperforms the benchmark, and vice versa. The dashed lines are drawn at the 5% critical values provided by [Giacomini and Rossi \(2010, Sec. 3.2\)](#) for the null hypothesis that the LW method and the univariate benchmark have equal out-of-sample performance at each point in time.

It should be noticed that the first point shown in each graph refers to the years 2008–2010, which include the Great Recession. Hence, it is evident that the direct LW predictor outperformed the benchmark significantly during the Great Recession, and that no significant accuracy gains are detectable thereafter. It was then that the monthly indicators carried relevant information in real time that was not contained in the GDP history.

It is remarkable that when it comes to comparing the indirect GDP predictors by output, see [Fig. 5](#), the LW predictor shows a sizable and sometimes significantly greater increase in predictive accuracy during the downturn associated with the sovereign debt crisis. This characteristic can also be observed for $m = 2$ and $m = 3$ when GDP is predicted according to the expenditure approach ([Fig. 6](#)).

We performed an indicator-importance analysis by counting how many times each one belonged to the set of 50 indicators with largest weights w_{ijt} across the 44 recursive prediction exercises taking place in month $m = 1, 2, 3$, of quarter $\tau = 1, \dots, 44$. In the *balloon plots 7–9*

the top 50 indicators are listed on the left in increasing order of importance from top to bottom. The size and color of the balloons are proportional to the square of the nowcast and forecast error that arises from the bivariate models featuring a particular indicator. The latter reduce as we move from $m = 1$ to $m = 3$; the reduction is quite dramatic in nowcasting GDP indirectly using the output approach; see the panels on the left in [Fig. 8](#).

The plots point out a very important fact: in nowcasting the current quarter GDP using the information accrued in the first month ($m = 1$), a prominent, if not exclusive, role is played by business-cycle indicators (BCIs), and economic sentiment indicators (ESIs) based on harmonized surveys of businesses and consumers. Thus, the main driver of the gains in nowcasting accuracy at the end of the first month of the quarter are the so-called soft indicators. If the output approach is used ([Fig. 8](#)), some hard indicators are also relevant, relating to the labor market (for the service sectors) or to economic activity (car registration and retail sales for Sector G-H-I, and motorway flow of trucks for industry). In other words, sector-specific indicators emerge now as drivers of the accuracy of the nowcast of a GDP component.

As we move to second and third month of the quarter ($m = 2, 3$) and more quantitative information on the current quarter accrues, hard indicators gain progressively more importance: in particular, industrial production indices, retail and industrial turnover, new orders, and import and export indices; see the second and third panels on the right in [Figs. 7, 8](#), and [9](#). The motorway flow of trucks stands out as one of the most important coincident indicators. As a matter of fact, this is one of the indicators that is monitored by the authors in their

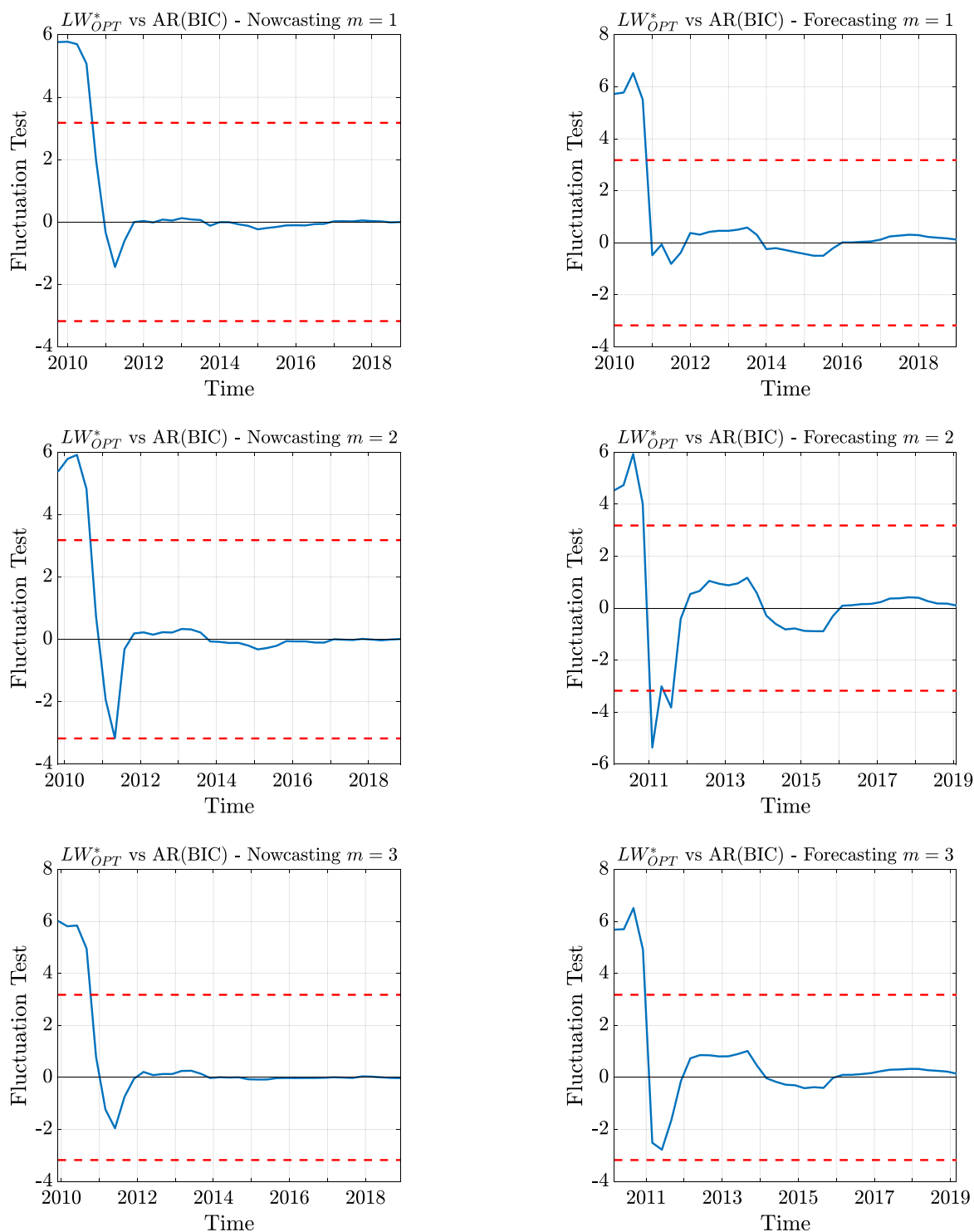
LW_{OPT}* vs AR(BIC) - Direct

Fig. 4. Fluctuation test for the LW direct GDP predictor. The solid line in each figure is the average difference between the squared nowcast (forecast) errors of two LW methods and the univariate benchmark, normalized by its estimated standard deviation, computed over a rolling window of eight quarters. The zero horizontal line indicates equal performance, and the dotted lines indicate the 5% critical values. LW_{OPT}* outperforms (underperforms) the AR(BIC) locally, at the 5% significance level, when the solid line is above (below) the upper (lower) dashed line.

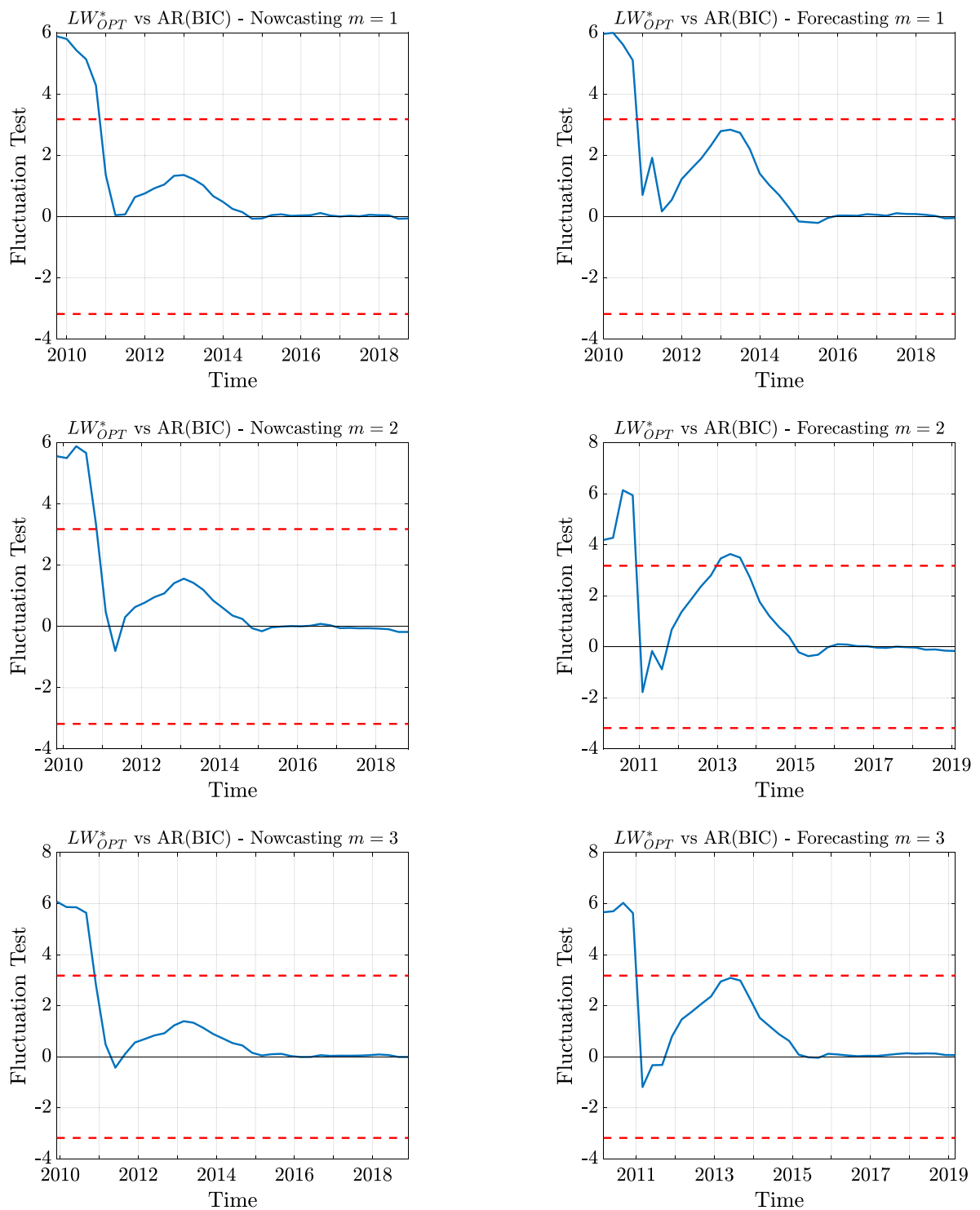
LW_{OPT}^* vs AR(BIC) - Output

Fig. 5. Fluctuation test for the LW indirect predictor by the output approach. The solid line in each figure is the average difference between the squared nowcast (forecast) errors of two LW methods and the univariate benchmark, normalized by its estimated standard deviation, computed over a rolling window of eight quarters. The zero horizontal line indicates equal performance, and the dotted lines indicate the 5% critical values. LW_{OPT}^* outperforms (underperforms) the AR(BIC) locally, at the 5% significance level, when the solid line is above (below) the upper (lower) dashed line.

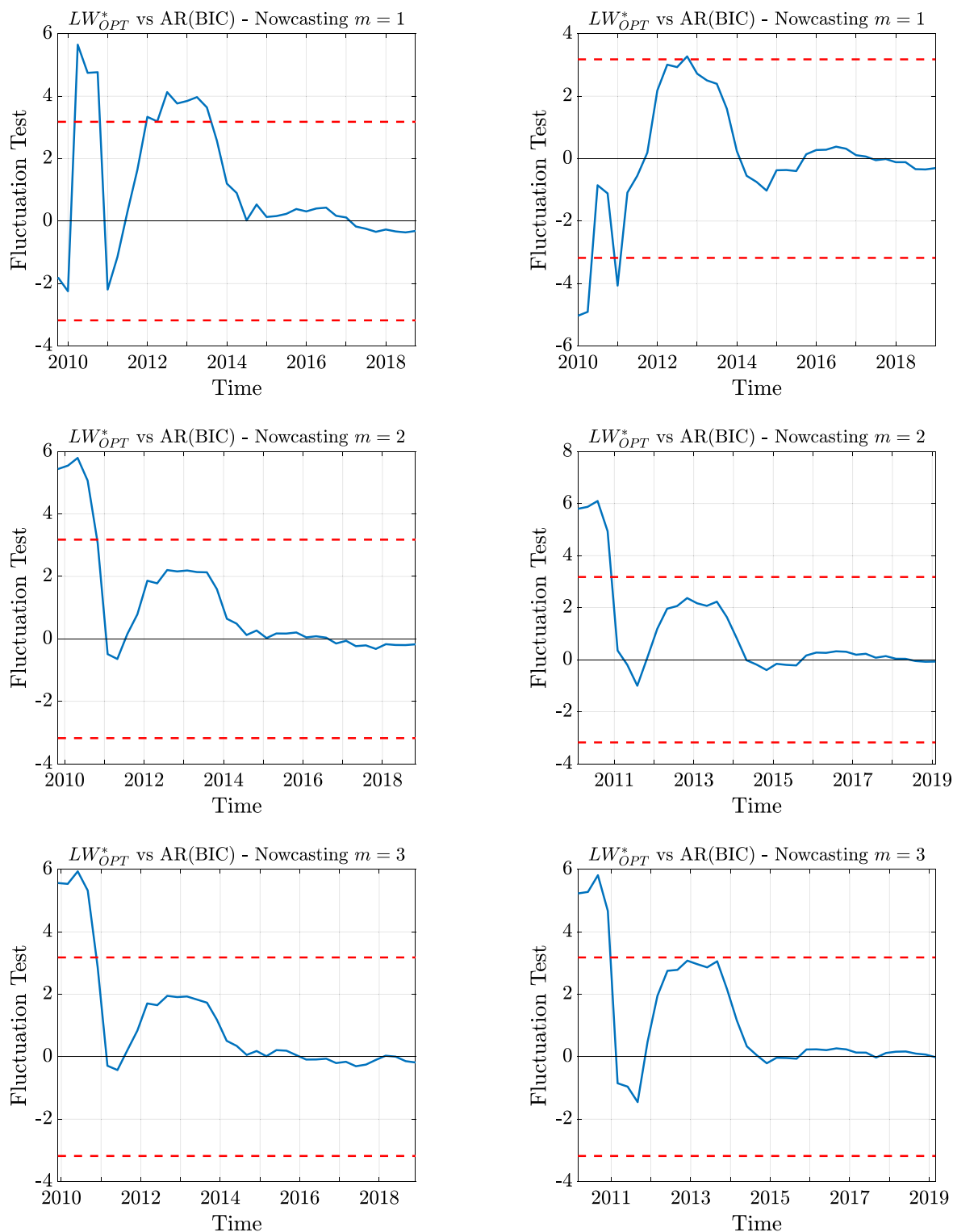
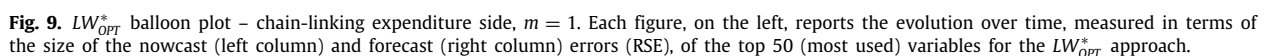
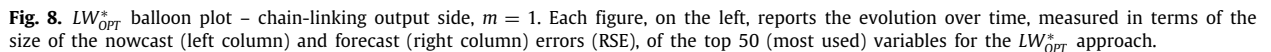
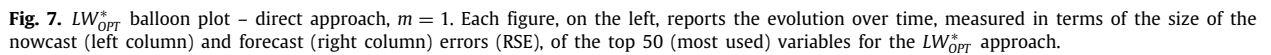


Fig. 6. Fluctuation test for the LW indirect GDP predictor by the expenditure approach. The solid line in each figure is the average difference between the squared nowcast (forecast) errors of two LW methods and the univariate benchmark, normalized by its estimated standard deviation, computed over a rolling window of eight quarters. The zero horizontal line indicates equal performance, and the dotted lines indicate the 5% critical values. LW_{OPT}^* outperforms (underperforms) the AR(BIC) locally, at the 5% significance level, when the solid line is above (below) the upper (lower) dashed line.



quarter than in nowcasting. That is, they better reflect the development of economic conditions at longer horizons.

Obviously, the nowcast and forecast errors are strongly procyclical, being larger during the Great Recession and

the sovereign debt crisis. It was particularly hard to forecast the Great Recession since its inception, and in particular the three quarters from 2008.q4–2009.q2, which marked a deep trough of economic activity and the initial recovery.

The plots also show that in $m = 3$, the hard indicators relating to the freight of goods, including truck flow and industrial production, have the largest predictive accuracy for forecasting one quarter ahead.

7. Discussion and concluding remarks

The paper considered the problem of nowcasting and forecasting production and expenditure national accounts, using a large-dimensional set of economic indicators available at the monthly frequency.

Our methodology relies on estimating all the possible mixed-frequency bivariate models of the quarterly GDP component and each monthly indicator in turn, taking into account the temporal aggregation constraints and the ragged-edge structure of the data. The different nowcasts and forecasts are then combined with weights reflecting their local accuracy in out-of-sample predictions. A recursive pseudo-real-time exercise showed that the Ledoit–Wolf shrinkage approach, see [Ledoit and Wolf \(2004a, 2004b\)](#), provides a very effective solution to the estimation of the model averaging weights.

With reference to the Italian case, we illustrated that our approach can keep up with the challenges posed by the dimensionality, since it can handle a large number of time series with a complexity that increases linearly with the cross-sectional dimension, while retaining the essential heterogeneity of the information about the macroeconomy. The combination with aggregation weights that reflect the predictive ability of the indicators is quintessential to the efficiency of the methodology.

Along with nowcasting and forecasting aggregate GDP (synthesis), our approach can measure the contribution to GDP growth by the components by output and expenditure (by growth accounting) and assess the comparative merits of the indirect approach vis-à-vis the direct prediction of GDP (analysis). The estimation of all possible bivariate models generates a wealth of information, making it is possible to assess when and how each individual indicator contributes to the result (indicator importance). This analysis has led to several interesting discoveries, among other things, concerning the relative contribution of soft and hard indicators as they become available in real time, as well as their predictive performance during recessionary episodes.

The results on the relative importance of the indicators presented in Section 6.2 should not be interpreted as an expression of the *sparsity* of economic information, i.e. that only a few indicators matter for forecasting the level of aggregate economic activity.

A substantial amount of literature has challenged the comparative efficiency of large-scale versus small-scale forecasting approaches. [Boivin and Ng \(2006\)](#) show that substantial cross-correlation among the idiosyncratic components leads to the deterioration of the contribution of factors extracted from large macroeconomic panels

to forecasting key variables such as GDP. [Bai and Ng \(2008\)](#) argue that supervising the factor extraction by a preliminary pre-selection of the indicators could lead to superior performance. [Bańbura and Modugno \(2012\)](#) find that for the euro area, using more disaggregated information does not improve the predictive accuracy. [Poncela and Ruiz \(2015\)](#) also point out that more information is not necessarily better, highlighting the loss in accuracy due to the increase in parameter uncertainty when higher-dimensional factor models are estimated. [Alvarez et al. \(2016\)](#) show that smaller-dimensional factor models using aggregate indicators, selected on the basis of their representativeness within economic categories, outperform larger-scale dynamic factor models for predicting the euro area GDP.

Against this background, we argue that the overrepresentation of categories of particular indicators (e.g., consumer and business surveys) and high levels of cross-sectional correlation in the idiosyncratic components are more likely to affect global approaches that use all the information at once, e.g., estimating a large dimensional factor model by principal components or the EM algorithm, than our approach. In our framework the individual indicators are entered one by one and thus the cross-correlation of the idiosyncratic effect does not play a role. Moreover, the supervision of the individual nowcasts and forecasts is guaranteed both by the estimation of the bivariate factor model, which is formulated such that the common component of the indicator under consideration is distilled, without having to specify the nature and number of the common dynamic factors, and by the model averaging methodology, which assigns weights to the nowcasts and forecasts according to the predictive accuracy manifested in the past.

Future research will be dedicated to introducing interventions to the model and the averaging method to account for the effects of the Covid-19 pandemic, and to evaluating the information content of the monthly indicators for nowcasting and forecasting GDP during the pandemic.

Declaration of competing interest

The authors declare that they have no known competing financial interests or personal relationships that could have appeared to influence the work reported in this paper.

Acknowledgments

This research has been supported by a research grant by the Italian Ministry of Economics and Finance (MEF), Direction I (Department of Treasury), as part of the collaboration agreement “Short run forecasting of macroeconomic indicators” between MEF and the University of Tor Vergata.

Table B.1

List of monthly indicators.

| Id | Variables name | Abbreviation | Sector | Timing | Country |
|----|--|--|--------------|--------|---------|
| 1 | Activity rate (15-64): Italy | Activity rate (15-64) | Labor | 1 | Italy |
| 2 | Freight in tons travelled by plane: Italy | Freight in tons - by plane | Services | 1 | Italy |
| 3 | Number of passengers travelled by plane : Italy | Passengers travelled by plane | Services | 1 | Italy |
| 4 | Arrivals in hotel, foreigners: Italy | Arrivals in hotel - foreigners | Services | 3 | Italy |
| 5 | Arrivals in hotel: Italy | Arrivals in hotel | Services | 3 | Italy |
| 6 | Attendance in hotel, foreigners: Italy | Attendance in hotel - foreigners | Services | 3 | Italy |
| 7 | Attendance in hotel: Italy | Attendance in hotel | Services | 3 | Italy |
| 8 | Loans to residents of Italy: bad debts | Loans bad debts - Italy | Finance | 2 | Italy |
| 9 | ESI - industry - Italy | ESI - industry - Italy | Survey | 0 | Italy |
| 10 | ESI - industry - France | ESI - industry - France | Survey | 0 | Europe |
| 11 | ESI - industry - Germany | ESI - industry - Germany | Survey | 0 | Europe |
| 12 | ESI - industry - Spain | ESI - industry - Spain | Survey | 0 | Europe |
| 13 | ESI - industry - UK | ESI - industry - UK | Survey | 0 | Europe |
| 14 | Current opportunity to purchase durable goods | Opport. purchase durable gds | Survey | 0 | Italy |
| 15 | Opinions on family budget | Opinions on family budget | Survey | 0 | Italy |
| 16 | Balance of payments - services - credits | BoP - services - credits | Trade | 3 | Italy |
| 17 | Balance of payments - services - debts | BoP - services - debts | Trade | 3 | Italy |
| 18 | Italy benchmark bond 10 years | Italy bond 10 years | Finance | 0 | Italy |
| 19 | BCI in construction | BCI retail Constr. | Survey | 0 | Italy |
| 20 | BCI Manufac. - Employ. expectations | BCI mfg - Employ. Exp. | Survey | 0 | Italy |
| 21 | BCI Manufac. - orders expectations | BCI mfg - orders Exp. | Survey | 0 | Italy |
| 22 | BCI Manufac. - selling prices expectations | BCI mfg - selling prices Exp. | Survey | 0 | Italy |
| 23 | BCI Manufac. - production expectations | BCI mfg - prod. Exp. | Survey | 0 | Italy |
| 24 | BCI Manufac. - goods - Employ. expectations | BCI mfg - Cons. gds - Employ. Exp. | Survey | 0 | Italy |
| 25 | BCI Manufac. - energy - Employ. expectations | BCI mfg - energy - Employ. Exp. | Survey | 0 | Italy |
| 26 | BCI Manufac. - intermediate goods - Employ. expectations | BCI mfg - Interm. gds - Employ. Exp. | Survey | 0 | Italy |
| 27 | BCI Manufac. - capital goods - Employ. expectations | BCI mfg - Capital gds - Employ. Exp. | Survey | 0 | Italy |
| 28 | BCI Manufac. - opinions on orders | BCI mfg -Int. orders | Survey | 0 | Italy |
| 29 | BCI Manufac. - opinions on production levels | BCI mfg -Int. prod. levels | Survey | 0 | Italy |
| 30 | BCI in the retail sales sector | BCI retail sales | Survey | 0 | Italy |
| 31 | BCI in the services sector | BCI retail services | Survey | 0 | Italy |
| 32 | BCI Manufac. - opinions on stocks of finished products | BCI mfg - finished products | Survey | 0 | Italy |
| 33 | Car Registration - Alfa Romeo | Car Reg. - Alfa Romeo | Industry | 1 | Italy |
| 34 | Car Registration - commercial vehicles - total | Car Reg. - total | Industry | 1 | Italy |
| 35 | Car Registration - Ferrari | Car Reg. - Ferrari | Industry | 1 | Italy |
| 36 | Car Registration - Fiat | Car Reg. - Fiat | Industry | 1 | Italy |
| 37 | Car Registration - heavy (over 16T) | Car Reg. - commercial - heavy | Industry | 1 | Italy |
| 38 | Car Registration - Lancia | Car Reg. - Lancia | Industry | 1 | Italy |
| 39 | Car Registration - light (up to 3.5T) | Car Reg. - commercial - light | Industry | 1 | Italy |
| 40 | Car Registration - Maserati | Car Reg. - Maserati | Industry | 1 | Italy |
| 41 | Car Registration - medium (over 3.5T) | Car Reg. - commercial - medium | Industry | 1 | Italy |
| 42 | Car Registration - Other brands | Car Reg. - Other brands | Industry | 1 | Italy |
| 43 | Car Registration | Car Reg. | Industry | 1 | Italy |
| 44 | Wage Guarantee Fund - for exceptional cases | WGF - for exceptional cases | Labor | 1 | Italy |
| 45 | Wage Guarantee Fund - ordinary | WGF - ordinary | Labor | 1 | Italy |
| 46 | Wage Guarantee Fund - extraordinary | WGF - extraordinary | Labor | 1 | Italy |
| 47 | Wage Guarantee Fund - total | WGF - total | Labor | 1 | Italy |
| 48 | BCI Manufac. - goods | BCI mfg - Cons. gds | Survey | 0 | Italy |
| 49 | BCI Manufac. - intermediate goods | BCI mfg - Interm. gds | Survey | 0 | Italy |
| 50 | BCI Manufac. - capital goods | BCI mfg - Capital gds | Survey | 0 | Italy |
| 51 | BCI Manufac. - total | BCI mfg - total | Survey | 0 | Italy |
| 52 | Compensation of employees (net of WGF) - total industry | Compensation (net of WGF) - Industry | Labor | 2 | Italy |
| 53 | Confcommercio index - clothing | ConfCom Index - clothing | Services | 1 | Italy |
| 54 | Confcommercio index - accommodation | ConfCom Index - accommodation | Services | 1 | Italy |
| 55 | Confcommercio index - food | ConfCom Index - food | Services | 1 | Italy |
| 56 | Confcommercio index - total goods | ConfCom Indexprod.gds | Services | 1 | Italy |
| 57 | Confcommercio index - housing | ConfCom Index - housing | Services | 1 | Italy |
| 58 | Confcommercio index - communication | ConfCom Index - communication | Services | 1 | Italy |
| 59 | Confcommercio index - person care | ConfCom Index - person care | Services | 1 | Italy |
| 60 | Confcommercio index - durable goods | ConfCom Index - durable gds | Services | 1 | Italy |
| 61 | Confcommercio index - mobility | ConfCom Index - mobility | Services | 1 | Italy |
| 62 | Confcommercio index - non-durable goods | ConfCom Index - non-durable | Services | 1 | Italy |
| 63 | Confcommercio index - recreational | ConfCom Index - recreational | Services | 1 | Italy |
| 64 | Confcommercio index - services | ConfCom Index - services | Services | 1 | Italy |
| 65 | Confcommercio index - total | ConfCom Index prod. | Services | 1 | Italy |
| 66 | Current savings opportunity | Current savings opportunity | Survey | 0 | Italy |
| 67 | BCI in the construction sector | BCI - Constr. | Survey | 0 | Italy |
| 68 | CCI | CCI | Survey | 0 | Italy |
| 69 | CCI - France | CCI - France | Survey | 0 | Europe |
| 70 | CCI - Germany | CCI - Germany | Survey | 0 | Europe |
| 71 | CCI - Spain | CCI - Spain | Survey | 0 | Europe |
| 72 | CCI - UK | CCI - UK | Survey | 0 | Europe |
| 73 | CCI - US | CCI - US | Survey | 0 | world |
| 74 | Gross demand of electricity | Gross demand - electricity | Industry | 0 | Italy |
| 75 | CCI - current climate | CCI - current climate | Survey | 0 | Italy |
| 76 | Cost of construction of a residential building | Cost residential building | Prices | 2 | Italy |
| 77 | BCI in the construction sector - prices expectations | BCI - Constr. - Prices Exp. | Survey | 0 | Italy |
| 78 | Consumer price index excluding tobacco (FOI) | CPI exc. tobacco (FOI) | Prices | 1 | Italy |
| 79 | Consumer price index - total - Germany | CPI - Germany | Prices | 1 | Europe |
| 80 | Consumer price index - core inflation | CPI - core inflation | Prices | 1 | Italy |
| 81 | Harmonized index of consumer prices - EA | HCPI - EA | Prices | 1 | Europe |
| 82 | Harmonized index of consumer prices | HCPI | Prices | 1 | Italy |
| 83 | Consumer price index - NIC | CPI - NIC | Prices | 1 | Italy |
| 84 | Consumer price index - NIC excluding tobacco | CPI - NIC exc. tobacco | Prices | 1 | Italy |
| 85 | Dummy for number of working days | n. Working Days | Working days | 0 | Italy |
| 86 | Dummy for number of working days (with bank holiday) | n. Working Days - bank holiday | Working days | 0 | Italy |
| 87 | Dummy for number of Mondays | Dummy for number of Mondays | Working days | 0 | Italy |
| 88 | Dow Jones stock market index | Dow Jones stock market index | Finance | 0 | world |
| 89 | CCI - economic climate | CCI - economic climate | Survey | 0 | Italy |
| 90 | BCI in the retail sector - expectations on the economy | BCI - retail - Exp. Economy | Survey | 0 | Italy |
| 91 | BCI in the retail sector - BCI current situation | BCI - retail - current situation | Survey | 0 | Italy |
| 92 | BCI in the retail sector | BCI - retail | Survey | 0 | Italy |
| 93 | BCI in the retail sector - Employ. expectations | BCI - retail - Empl. Exp. | Survey | 0 | Italy |
| 94 | BCI in the retail sector - orders | BCI - retail - orders | Survey | 0 | Italy |
| 95 | BCI in the retail sector - BCI stocks | BCI - retail - Int. the stocks | Survey | 0 | Italy |
| 96 | BCI in the services sector - BCI current situation | BCI - services -Int. the current situation | Survey | 0 | Italy |
| 97 | BCI in the services sector | BCI - services | Survey | 0 | Italy |

(continued on next page)

Table B.1 (continued).

| Id | Variables name | Abbreviation | Sector | Timing | Country |
|-----|--|---|---------------|--------|---------|
| 98 | BCI in the services sector - BCI demand | BCI - services -Int. the demand | Survey | 0 | Italy |
| 99 | BCI in the services sector - expectations on the demand | BCI - services - Exp. on the demand | Survey | 0 | Italy |
| 100 | BCI in the services sector - BCI Employ. | BCI - services -Int. the Employ. | Survey | 0 | Italy |
| 101 | BCI in the services sector - Employ. expectations | BCI - services - Employ. Exp. | Survey | 0 | Italy |
| 102 | Employees (15 years and over) - survey of the Labor force | Employees (> 15) - survey | Labor | 1 | Italy |
| 103 | BCI Manufac. - Employ. expectations | BCI mfg - Employ. Exp. | Survey | 0 | Italy |
| 104 | ESI | ESI | Survey | 0 | Italy |
| 105 | ESI - France | ESI - France | Survey | 0 | Europe |
| 106 | ESI - Germany | ESI - Germany | Survey | 0 | Europe |
| 107 | ESI - Spain | ESI - Spain | Survey | 0 | Europe |
| 108 | ESI - UK | ESI - UK | Survey | 0 | Europe |
| 109 | Euro stoxx 50 stock market index | Euro stoxx 50 | Finance | 0 | Europe |
| 110 | Export - goods | Export - Cons. gds | Trade | 2 | Italy |
| 111 | Export - energy | Export - energy | Trade | 2 | Italy |
| 112 | Current account BP - goods - credits | Current account BP - gds - credits | Trade | 2 | Italy |
| 113 | Export of goods - world | Export gds - world | Trade | 2 | Italy |
| 114 | Export of goods (fob) - extra EU | Export gds - extra EU | Trade | 2 | Italy |
| 115 | Export of goods (fob) - EU | Export gds - EU | Trade | 2 | Italy |
| 116 | Export - intermediate goods | Export - Interm. gds | Trade | 2 | Italy |
| 117 | Export - capital goods | Export - Capital gds | Trade | 2 | Italy |
| 118 | Current account BP - services - credits | BP - services - credits | Trade | 2 | Italy |
| 119 | CCI - future climate | CCI - future climate | Survey | 0 | Italy |
| 120 | Maastricht public debt: Italy | Maastricht public debt | Finance | 2 | Italy |
| 121 | Export of goods (fob) - world - volume | Export gds - world | Trade | 2 | Italy |
| 122 | Export of goods (fob) - extra EU - volume | Export gds - extra EU | Trade | 2 | Italy |
| 123 | Export of goods (fob) - EU - volume | Export gds - EU | Trade | 2 | Italy |
| 124 | Imports of goods (fob) - world - volume | Imports gds - world | Trade | 2 | Italy |
| 125 | Imports of goods (fob) - extra EU - volume | Imports gds - extra EU | Trade | 2 | Italy |
| 126 | Imports of goods (fob) - EU - volume | Imports gds - EU | Trade | 2 | Italy |
| 127 | Inactivity rate (15-64) | Inactivity rate (15-64) | Labor | 1 | Italy |
| 128 | Inactive population | Inactive population | Labor | 1 | Italy |
| 129 | IPI: Italy | IPI | Industry | 2 | Italy |
| 130 | IPI - Austria | IPI - Austria | Industry | 2 | Europe |
| 131 | IPI - Belgium | IPI - Belgium | Industry | 2 | Europe |
| 132 | IPI - Denmark | IPI - Denmark | Industry | 2 | Europe |
| 133 | IPI - energy | IPI - energy | Industry | 2 | Italy |
| 134 | IPI - Finland | IPI - Finland | Industry | 2 | Europe |
| 135 | IPI - France | IPI - France | Industry | 2 | Europe |
| 136 | IPI - Greece | IPI - Greece | Industry | 2 | Europe |
| 137 | IPI - Germany | IPI - Germany | Industry | 2 | Europe |
| 138 | IPI - Ireland | IPI - Ireland | Industry | 2 | Europe |
| 139 | IPI - Japan | IPI - Japan | Industry | 2 | world |
| 140 | IPI - Korea | IPI - Korea | Industry | 2 | world |
| 141 | IPI - Mexico | IPI - Mexico | Industry | 2 | world |
| 142 | IPI - excluding energy | IPI - exc. energy | Industry | 2 | Italy |
| 143 | IPI - The Netherlands | IPI - The Netherlands | Industry | 2 | Europe |
| 144 | IPI - Norway | IPI - Norway | Industry | 2 | Europe |
| 145 | IPI - Portugal | IPI - Portugal | Industry | 2 | Europe |
| 146 | IPI - Spain | IPI - Spain | Industry | 2 | Europe |
| 147 | IPI - Sweden | IPI - Sweden | Industry | 2 | Europe |
| 148 | IPI - UK | IPI - UK | Industry | 2 | Europe |
| 149 | IPI - US | IPI - US | Industry | 2 | world |
| 150 | IPI - manufacture of motor vehicles | IPI - manufacture of motor vehicles | Industry | 2 | Italy |
| 151 | IPI - manufacture of cement | IPI - manufacture of cement | Industry | 2 | Italy |
| 152 | IPI - goods | IPI - Cons. gds | Industry | 2 | Italy |
| 153 | IPI - durable goods | IPI - durable Cons. gds | Industry | 2 | Italy |
| 154 | IPI - non-durable goods | IPI - non-durable Cons. gds | Industry | 2 | Italy |
| 155 | IPI - construction | IPI - construction | Industry | 2 | Italy |
| 156 | IPI - energy | IPI - energy | Industry | 2 | Italy |
| 157 | IPI - intermediate goods | IPI - Interm. gds | Industry | 2 | Italy |
| 158 | IPI - capital goods | IPI - Capital gds | Industry | 2 | Italy |
| 159 | IPI - manufacturing | IPI - Mfg | Industry | 2 | Italy |
| 160 | IPI - steel | IPI - steel | Industry | 2 | Italy |
| 161 | CCI | CCI | Survey | 0 | Italy |
| 162 | VAT revenues - Imports | VAT revenues - Imports | Services | 2 | Italy |
| 163 | VAT revenues - Domestic | VAT revenues - Domestic | Services | 2 | Italy |
| 164 | World trade index: Italy | World trade index | International | 2 | Italy |
| 165 | International IPI weighted for Italy - world | Intern.IPI weighted for Italy - world | International | 2 | Italy |
| 166 | International IPI weighted for Italy - UE | Intern.IPI weighted for Italy - UE | International | 2 | Italy |
| 167 | International IPI weighted for Italy - Extra UE | Intern.IPI weighted for Italy - Extra UE | International | 2 | Italy |
| 168 | Unemployed (15 years and over) - female | Unemployed (> 15) - female | Labor | 1 | Italy |
| 169 | Unemployed (15 years and over) - male | Unemployed (> 15) - male | Labor | 1 | Italy |
| 170 | Dummy for a leap day | Dummy for a leap day | Working days | 0 | Italy |
| 171 | Motorway flow of trucks - light | Motorway flow of trucks - light | Industry | 0 | Italy |
| 172 | Labor force (15 years and over) | Labor force (> 15) | Labor | 1 | Italy |
| 173 | Loans to non-financial corporation - up to 1 year | Loans no financ. - up to 1 year | Finance | 2 | Italy |
| 174 | Loans to non-financial corporation - 1 to 5 years | Loans no financ. - 1 to 5 years | Finance | 2 | Italy |
| 175 | Loans to non-financial corporation - over 5 years | Loans no financ. - over 5 years | Finance | 2 | Italy |
| 176 | Loans to households - credit - up to 1 year | HH Loans - Cons. Credit- up to 1 year | Finance | 2 | Italy |
| 177 | Loans to households - credit - 1 to 5 years | HH Loans - Cons. Credit- 1 to 5 years | Finance | 2 | Italy |
| 178 | Loans to households - credit - over 5 years | HH Loans - Cons. Credit- over 5 years | Finance | 2 | Italy |
| 179 | Loans to households - home purchase - up to 1 year | HH Loans - home purchase - up to 1 year | Finance | 2 | Italy |
| 180 | Loans to households - home purchase - over 5 years | HH Loans - home purchase - over 5 years | Finance | 2 | Italy |
| 181 | BCI Manufac. - goods - opinions on orders | BCI mfg - Cons. gds -Int. orders | Survey | 0 | Italy |
| 182 | BCI Manufac. - goods - opinions on export orders | BCI mfg - Cons. gds -Int. export orders | Survey | 0 | Italy |
| 183 | BCI Manufac. - intermediate goods - opin. on export orders | BCI mfg - Interm. gds -Int. export orders | Survey | 0 | Italy |
| 184 | BCI Manufac. - capital goods - opinions on export orders | BCI mfg - Capital gds -Int. export orders | Survey | 0 | Italy |
| 185 | BCI Manufac. - total - opinions on export orders | BCI mfg -Int. export orders | Survey | 0 | Italy |
| 186 | BCI Manufac. - goods - opinions on national orders | BCI mfg - Cons. gds -Int. national orders | Survey | 0 | Italy |
| 187 | BCI Manufac. - intermediate goods - opinions on orders | BCI mfg - Interm. gds -Int. orders | Survey | 0 | Italy |
| 188 | BCI Manufac. - intermediate goods - opinions on nat. orders | BCI mfg - Interm. gds -Int. national orders | Survey | 0 | Italy |
| 189 | BCI Manufac. - capital goods - opinions on nat. orders | BCI mfg - Capital gds -Int. national orders | Survey | 0 | Italy |
| 190 | BCI Manufac. - total - opinions on national orders | BCI mfgprod.-Int. national orders | Survey | 0 | Italy |
| 191 | BCI Manufac. - capital goods - opinions on orders | BCI mfg - Capital gds -Int. orders | Survey | 0 | Italy |
| 192 | BCI Manufac. - total - opinions on orders | BCI mfgprod.-Int. orders | Survey | 0 | Italy |
| 193 | BCI Manufac. - goods - opinions on production levels | BCI mfg - Cons. gds -Int. prod. levels | Survey | 0 | Italy |
| 194 | BCI Manufac. - intermediate goods - opin. on prod. levels | BCI mfg - Interm. gds -Int. prod. levels | Survey | 0 | Italy |
| 195 | BCI Manufac. - capital goods - opinions on production levels | BCI mfg - Capital gds -Int. prod. levels | Survey | 0 | Italy |

(continued on next page)

Table B.1 (continued).

| Id | Variables name | Abbreviation | Sector | Timing | Country |
|-----|--|-------------------------------------|----------|--------|---------|
| 196 | BCI Manufac. - totals - opinions on production levels | BCI mfg - totals -Int. prod. levels | Survey | 0 | Italy |
| 197 | BCI Manufac. - goods - opinions on stocks | BCI mfg - Cons. gds -Int. stocks | Survey | 0 | Italy |
| 198 | BCI Manufac. - intermediate goods - opinions on stocks | BCI mfg - Intern. gds -Int. stocks | Survey | 0 | Italy |
| 199 | BCI Manufac. - capital goods - opinions on stocks | BCI mfg - Capital gds -Int. stocks | Survey | 0 | Italy |
| 200 | BCI Manufac. - total - opinions on stocks | BCI mfgprod.-Int. stocks | Survey | 0 | Italy |
| 201 | M1 monetary aggregate | M1 Mon. Agg. | Finance | 2 | Italy |
| 202 | M2 monetary aggregate | M2 Mon. Agg. | Finance | 2 | Italy |
| 203 | Liabilities of mfi & positions included in m1 : Italy | Liabilities of mfi & positions (M1) | Finance | 2 | Italy |
| 204 | M3 monetary aggregate | M3 Mon. Agg. | Finance | 1 | Italy |
| 205 | Imports - s goods | Imports - Cons.s gds | Trade | 2 | Italy |
| 206 | Imports - energy | Imports - energy | Trade | 2 | Italy |
| 207 | Current account - goods - debits | Current account - gds - debits | Trade | 2 | Italy |
| 208 | Imports - world | Imports - world | Trade | 2 | Italy |
| 209 | Imports of goods (cif) - extra EU | Imports gds (cif) - extra EU | Trade | 1 | Italy |
| 210 | Imports of goods (cif) - EU | Imports gds (cif) - EU | Trade | 2 | Italy |
| 211 | Stock market index | Stock market index | Finance | 0 | Italy |
| 212 | Imports - intermediate goods | Imports - Intern. gds | Trade | 2 | Italy |
| 213 | Imports - capital goods | Imports - Capital gds | Trade | 2 | Italy |
| 214 | Current account - services - debits | Current account - services - debits | Trade | 2 | Italy |
| 215 | Population aged 65: Italy | Population aged 65 | Labor | 0 | Italy |
| 216 | Nominal effective exchange rate | Nominal effective RX | Finance | 0 | Italy |
| 217 | Nominal effective exchange rate (\$) | Nominal effective RX (\$) | Finance | 0 | Italy |
| 218 | Nominal effective exchange rate | Nominal effective RX | Finance | 0 | Italy |
| 219 | New orders - manufacturing - US | New orders - Mfg - US | Industry | 2 | world |
| 220 | New orders - vehicles | New orders - vehicles | Industry | 2 | Italy |
| 221 | New orders - total - industry | New orders - industry | Industry | 2 | Italy |
| 222 | New orders - total - internal market | New orders - internal | Industry | 2 | Italy |
| 223 | New orders - total - external market | New orders - external | Industry | 2 | Italy |
| 224 | CCI - personal climate | CCI - personal climate | Survey | 0 | Italy |
| 225 | Motorway flow of trucks - heavy | Motorway flow of trucks - heavy | Industry | 0 | Italy |
| 226 | Export prices - total - world | Export prices prod.- world | Prices | 2 | Italy |
| 227 | Export prices - total - extra EU | Export prices prod.- extra EU | Prices | 2 | Italy |
| 228 | Export prices - total - EU | Export prices prod.- EU | Prices | 2 | Italy |
| 229 | Export - unit values | Export - unit values | Trade | 2 | Italy |
| 230 | Raw materials price index - crude oil - world | Price Index - Crude Oil- World | Prices | 0 | world |
| 231 | Import prices - total - world | Import prices prod.- world | Prices | 2 | Italy |
| 232 | Import prices - total - extra EU | Import prices prod.- extra EU | Prices | 2 | Italy |
| 233 | Import prices - total - EU | Import prices prod.- EU | Prices | 2 | Italy |
| 234 | manufacturing PMI - work backlogs | Mfg PMI - work backlogs | Survey | 1 | Italy |
| 235 | Composite PMI - work backlogs | Comp. PMI - work backlogs | Survey | 1 | Italy |
| 236 | Composite PMI - production level | Comp. PMI - prod. level | Survey | 1 | Italy |
| 237 | Composite PMI - production level - EA | Comp. PMI - prod. level - EA | Survey | 1 | Europe |
| 238 | Composite PMI - production level - France | Comp. PMI - prod. level - France | Survey | 1 | Europe |
| 239 | Composite PMI - production level - Germany | Comp. PMI - prod. level - Germany | Survey | 1 | Europe |
| 240 | Composite PMI - production level - Spain | Comp. PMI - prod. level - Spain | Survey | 1 | Europe |
| 241 | Composite PMI - production level - UK | Comp. PMI - prod. level - UK | Survey | 1 | Europe |
| 242 | manufacturing PMI - quantity of purchases | Mfg PMI - quantity of purchases | Survey | 1 | Italy |
| 243 | Construction PMI - total | Constr. PMIprod. | Survey | 1 | Italy |
| 244 | Construction PMI - commercial | Constr. PMI - commercial | Survey | 1 | Italy |
| 245 | Construction PMI - Employ. | Constr. PMI - Employ. | Survey | 1 | Italy |
| 246 | Construction PMI - civil engineering | Constr. PMI - civil engineering | Survey | 1 | Italy |
| 247 | Construction PMI - residential | Constr. PMI - residential | Survey | 1 | Italy |
| 248 | Construction PMI - new orders | Constr. PMI - new orders | Survey | 1 | Italy |
| 249 | Construction PMI - input prices | Constr. PMI - input prices | Survey | 1 | Italy |
| 250 | Construction PMI - delivery time | Constr. PMI - delivery time | Survey | 1 | Italy |
| 251 | manufacturing PMI - Employ. | Mfg PMI - Employ. | Survey | 1 | Italy |
| 252 | Composite PMI - Employ. | Comp. PMI - Employ. | Survey | 1 | Italy |
| 253 | Services PMI - Employ. | Services PMI - Employ. | Survey | 1 | Italy |
| 254 | Manufacturing PMI - new export orders | Mfg PMI - new export orders | Survey | 1 | Italy |
| 255 | Manufacturing PMI - finished products | Mfg PMI - finished products | Survey | 1 | Italy |
| 256 | Manufacturing PMI - purchasing prices | Mfg PMI - purchasing prices | Survey | 1 | Italy |
| 257 | Services PMI - purchasing prices | Services PMI - purchasing prices | Survey | 1 | Italy |
| 258 | Manufacturing PMI | Mfg PMI | Survey | 1 | Italy |
| 259 | Manufacturing PMI - China | Mfg PMI - China | Survey | 1 | world |
| 260 | Manufacturing PMI - Denmark | Mfg PMI - Denmark | Survey | 1 | Europe |
| 261 | Manufacturing PMI - France | Mfg PMI - France | Survey | 1 | Europe |
| 262 | Manufacturing PMI - Germany | Mfg PMI - Germany | Survey | 1 | Europe |
| 263 | Manufacturing PMI - Ireland | Mfg PMI - Ireland | Survey | 1 | Europe |
| 264 | Manufacturing PMI - Japan | Mfg PMI - Japan | Survey | 1 | world |
| 265 | Manufacturing PMI - The Netherlands | Mfg PMI - The Netherlands | Survey | 1 | Europe |
| 266 | Manufacturing PMI - Spain | Mfg PMI - Spain | Survey | 1 | Europe |
| 267 | Manufacturing PMI - Sweden | Mfg PMI - Sweden | Survey | 1 | Europe |
| 268 | Manufacturing PMI - UK | Mfg PMI - UK | Survey | 1 | Europe |
| 269 | Composite PMI - new orders | Comp. PMI - new orders | Survey | 1 | Italy |
| 270 | Manufacturing PMI - new orders | Mfg PMI - new orders | Survey | 1 | Italy |
| 271 | manufacturing PMI - production levels | Mfg PMI - prod. levels | Survey | 1 | Italy |
| 272 | Composite PMI - selling prices | Comp. PMI - selling prices | Survey | 1 | Italy |
| 273 | Manufacturing PMI - selling prices | Mfg PMI - selling prices | Survey | 1 | Italy |
| 274 | Composite PMI - production levels | Comp. PMI - prod. levels | Survey | 1 | Italy |
| 275 | Manufacturing PMI - purchasing prices | Mfg PMI - purchasing prices | Survey | 1 | Italy |
| 276 | Composite PMI - purchasing prices | Comp. PMI - purchasing prices | Survey | 1 | Italy |
| 277 | Services PMI - purchasing prices | Services PMI - purchasing prices | Survey | 1 | Italy |
| 278 | Services PMI | Services PMI | Survey | 1 | Italy |
| 279 | Services PMI - new orders | Services PMI - new orders | Survey | 1 | Italy |
| 280 | Services PMI - expectations | Services PMI - Exp. | Survey | 1 | Italy |
| 281 | Services PMI - outstanding business | Services PMI - outstanding business | Survey | 1 | Italy |
| 282 | Services PMI - Germany | Services PMI - Germany | Survey | 1 | Europe |
| 283 | Manufacturing PMI - stocks | Mfg PMI - stocks | Survey | 1 | Italy |
| 284 | PMI Chicago Barometer | PMI Chicago | Survey | 1 | world |
| 285 | World price index non-fuel | World price index non-fuel | Prices | 1 | world |
| 286 | Raw materials prices - world - in euro | Raw Mater. prices - world | Prices | 1 | world |
| 287 | CCI - potential future savings | CCI - potential future savings | Survey | 0 | Italy |
| 288 | PPI | PPI | Prices | 1 | Italy |
| 289 | PPI - Austria | PPI- Austria | Prices | 1 | Europe |
| 290 | PPI - Belgium | PPI- Belgium | Prices | 1 | Europe |
| 291 | PPI - manufacturing | PPI- Mfg | Prices | 1 | Italy |
| 292 | PPI - goods | PPI- Cons. gds | Prices | 1 | Italy |
| 293 | PPI - Denmark | PPI- Denmark | Prices | 1 | Europe |

(continued on next page)

Table B.1 (continued).

| Id | Variables name | Abbreviation | Sector | Timing | Country |
|-----|--|--|----------|--------|---------|
| 294 | PPI - internal market | PPI- internal | Prices | 1 | Italy |
| 295 | PPI - durable goods | PPI- durable Cons. gds | Prices | 1 | Italy |
| 296 | PPI - energy | PPI- energy | Prices | 1 | Italy |
| 297 | PPI - external market - EA | PPI- external - EA | Prices | 1 | Italy |
| 298 | PPI - Finland | PPI- Finland | Prices | 1 | Europe |
| 299 | PPI - France | PPI- France | Prices | 1 | Europe |
| 300 | PPI - Greece | PPI- Greece | Prices | 1 | Europe |
| 301 | PPI - Germany | PPI- Germany | Prices | 1 | Europe |
| 302 | PPI - Hong Kong | PPI- Hong Kong | Prices | 1 | world |
| 303 | PPI - intermediate goods | PPI- Interim. gds | Prices | 1 | Italy |
| 304 | PPI - capital goods | PPI- Capital gds | Prices | 1 | Italy |
| 305 | PPI - Korea | PPI- Korea | Prices | 1 | world |
| 306 | PPI - Mexico | PPI- Mexico | Prices | 1 | world |
| 307 | PPI - The Netherlands | PPI- The Netherlands | Prices | 1 | Europe |
| 308 | PPI - external market | PPI- external market | Prices | 1 | Italy |
| 309 | PPI - non-durable goods | PPI- non-durable Cons. gds | Prices | 1 | Italy |
| 310 | PPI - external market - extra EU | PPI- external market - extra EU | Prices | 1 | Italy |
| 311 | PPI - Norway | PPI- Norway | Prices | 1 | Europe |
| 312 | PPI - Portugal | PPI- Portugal | Prices | 1 | Europe |
| 313 | International PPI weighted for Italy | Intern. PPI weighted for Italy | Prices | 1 | Italy |
| 314 | Competitiveness index - Italy | Competitiveness index - Italy | Prices | 1 | Italy |
| 315 | International PPI weighted for Italy - EA | Intern. PPI weighted for Italy -EA | Prices | 1 | Italy |
| 316 | International PPI weighted for Italy | Intern. PPI weighted for Italy | Prices | 1 | Italy |
| 317 | Euro zone PPI weighted for Italy | Euro zone PPI weighted for Italy | Prices | 1 | Italy |
| 318 | Extra euro PPI weighted for Italy | Extra euro PPI weighted for Italy | Prices | 1 | Italy |
| 319 | International PPI weighted for Italy - Extra EA | Intern. PPI weighted for Italy - ExtraEA | Prices | 1 | Italy |
| 320 | PPI - Spain | PPI- Spain | Prices | 1 | Europe |
| 321 | PPI - Sweden | PPI- Sweden | Prices | 1 | Europe |
| 322 | PPI - Switzerland | PPI- Switzerland | Prices | 1 | Europe |
| 323 | PPI - UK | PPI- UK | Prices | 1 | Europe |
| 324 | PPI - US | PPI- US | Prices | 1 | world |
| 325 | PPI - chemical products | PPI- chemical products | Prices | 1 | Italy |
| 326 | PPI - clothing | PPI- clothing | Prices | 1 | Italy |
| 327 | PPI - goods | PPI- Cons. gds | Prices | 1 | Italy |
| 328 | PPI - energy | PPI- energy | Prices | 1 | Italy |
| 329 | PPI - food | PPI- food | Prices | 1 | Italy |
| 330 | PPI - intermediate goods | PPI- Interim. gds | Prices | 1 | Italy |
| 331 | PPI - capital goods | PPI- Capital gds | Prices | 1 | Italy |
| 332 | PPI - manufacturing | PPI- Mfg | Prices | 1 | Italy |
| 333 | PPI - mining | PPI- mining | Prices | 1 | Italy |
| 334 | PPI - transports | PPI- transports | Prices | 1 | Italy |
| 335 | PPI - textile | PPI- textile | Prices | 1 | Italy |
| 336 | CCI - price expectations | CCI - price Exp. | Prices | 1 | Italy |
| 339 | CCI - expectations on the country current situation | CCI - Exp. situation of the country | Survey | 0 | Italy |
| 340 | Germany benchmark bond 10 years | Germany bond 10 years | Finance | 0 | Europe |
| 341 | 6-month interest rate - Eurozone - offer | 6-month interest rate - EA - offer | Finance | 0 | Europe |
| 342 | Real effective exchange rate | Real effective RX | Finance | 1 | Europe |
| 343 | Real effective exchange rate - EU | Real effective RX - EU | Finance | 1 | Europe |
| 344 | Real effective exchange rate - extra EU | Real effective RX - extra EU | Finance | 1 | Europe |
| 345 | Retail sales - food - volume | Retail sales - food | Services | 2 | Italy |
| 346 | Retail sales - food - value | Retail sales - food - value | Services | 2 | Italy |
| 347 | Main Refinancing Operations Fixed Rate - EA | Refinancing Operations Fixed Rate - EA | Finance | 0 | Europe |
| 348 | 3-month interest rate - Euribor | 3-month interest rate - Euribor | Finance | 0 | Europe |
| 349 | Nonfinancial corporations loans: new business | Nonfinan. loans: new business | Finance | 0 | Italy |
| 350 | 3-month interest rate - US | 3-month interest rate - US | Finance | 0 | world |
| 351 | Retail sales - deflated | Retail sales - deflated | Services | 2 | Italy |
| 352 | Exchange rate - Australian dollar/US dollar | RX - Australian dollar/US dollar | Finance | 0 | world |
| 353 | Exchange rate - Brazilian real/US dollar | RX - Brazilian real/US dollar | Finance | 0 | world |
| 354 | Exchange rate - Chinese yuan/US dollar | RX - Chinese yuan/US dollar | Finance | 0 | world |
| 355 | Exchange rate - Canadian dollar/US dollar | RX - Canadian dollar/US dollar | Finance | 0 | world |
| 356 | Exchange rate - Danish kroner/US dollar | RX - Danish kroner/US dollar | Finance | 0 | Europe |
| 357 | Exchange rate - Hong Kong dollar/US dollar | RX - Hong Kong dollar/US dollar | Finance | 0 | world |
| 358 | Exchange rate - Irish pound/US dollar | RX - Irish pound/US dollar | Finance | 0 | Europe |
| 359 | Exchange rate - Indian rupee/US dollar | RX - Indian rupee/US dollar | Finance | 0 | world |
| 360 | Exchange rate - Japanese yen/US dollar | RX - Japanese yen/US dollar | Finance | 0 | world |
| 361 | Exchange rate - Korean won/US dollar | RX - Korean won/US dollar | Finance | 0 | world |
| 362 | Exchange rate - Mexican peso/US dollar | RX - Mexican peso/US dollar | Finance | 0 | world |
| 363 | Exchange rate - Norwegian krone/US dollar | RX - Norwegian krone/US dollar | Finance | 0 | Europe |
| 364 | Exchange rate - New Zealand dollar/US dollar | RX - New Zealand dollar/US dollar | Finance | 0 | world |
| 365 | Exchange rate - Polish zloty/US dollar | RX - Polish zloty/US dollar | Finance | 0 | Europe |
| 366 | Exchange rate - Romanian leu/US dollar | RX - Romanian leu/US dollar | Finance | 0 | Europe |
| 367 | Exchange rate - Russian rouble/US dollar | RX - Russian rouble/US dollar | Finance | 0 | world |
| 368 | Exchange rate - Singapore dollar/US dollar | RX - Singapore dollar/US dollar | Finance | 0 | world |
| 369 | Exchange rate - Swedish krone/US dollar | RX - Swedish krone/US dollar | Finance | 0 | Europe |
| 370 | Exchange rate - Swiss franc/US dollar | RX - Swiss franc/US dollar | Finance | 0 | Europe |
| 371 | Exchange rate - Turkish lira/US dollar | RX - Turkish lira/US dollar | Finance | 0 | Europe |
| 372 | Exchange rate - pound sterling/US dollar | RX - pound sterling/US dollar | Finance | 0 | Europe |
| 373 | Exchange rate - US dollar/euro | RX - US dollar/euro | Finance | 0 | world |
| 374 | BCI in the services sector - Employ. expectations | BCI - services - Employ. Exp. | Survey | 0 | Italy |
| 375 | HH BCI | HH BCI | Survey | 0 | Italy |
| 376 | BCI economic situation of the households | BCI economic situation of the country | Survey | 0 | Italy |
| 377 | S&P 500 stock market index | S&P 500 stock market index | Finance | 0 | world |
| 378 | BCI general tendency of the economy - goods | BCI tendency - Economy -Cons. gds | Survey | 0 | Italy |
| 379 | BCI general tendency of the economy - intermediate goods | BCI tendency - Economy -Interm. gds | Survey | 0 | Italy |
| 380 | BCI general tendency of the economy - capital goods | BCI tendency - Economy -Capital gds | Survey | 0 | Italy |
| 381 | BCI general tendency of the economy - total | BCI tendency - Economy -total | Survey | 0 | Italy |
| 382 | BCI tendency of the liquidity - goods | BCI tendency - demand - Cons. gds | Survey | 0 | Italy |
| 383 | BCI tendency of the liquidity - intermediate goods | BCI tendency - demand - Interm. gds | Survey | 0 | Italy |
| 384 | BCI tendency of the liquidity - capital goods | BCI tendency - demand - Capital gds | Survey | 0 | Italy |
| 385 | BCI tendency of the liquidity - total | BCI tendency - demand - total | Survey | 0 | Italy |
| 386 | BCI tendency of Employ. in 3 months - goods | BCI tendency - Employ. - Cons. gds | Survey | 0 | Italy |
| 387 | BCI tendency of Employ. in 3 months - interm. goods | BCI tendency - Employ. - Interm. gds | Survey | 0 | Italy |
| 388 | BCI tendency of Employ. in 3 months - capital goods | BCI tendency - Employ. - Capital gds | Survey | 0 | Italy |
| 389 | BCI tendency of Employ. in 3 months - total | BCI tendency - Employ. - total | Survey | 0 | Italy |
| 390 | BCI tendency orders and demand - goods | BCI tendency - liquidity - Cons. gds | Survey | 0 | Italy |
| 391 | BCI tendency orders and demand - intermediate goods | BCI tendency - liquidity - Interm. gds | Survey | 0 | Italy |
| 392 | BCI tendency orders and demand - capital goods | BCI tendency - liquidity - Capital gds | Survey | 0 | Italy |
| 393 | BCI tendency orders and demand - total | BCI tendency - liquidity - total | Survey | 0 | Italy |
| 394 | BCI tendency selling prices - goods | BCI tendency prices - Cons. gds | Survey | 0 | Italy |

(continued on next page)

Table B.1 (continued).

| Id | Variables name | Abbreviation | Sector | Timing | Country |
|-----|--|--|----------|--------|---------|
| 395 | BCI tendency selling prices - intermediate goods | BCI tendency prices - Intermed. gds | Survey | 0 | Italy |
| 396 | BCI tendency selling prices - capital goods | BCI tendency prices - Capital gds | Survey | 0 | Italy |
| 397 | BCI tendency selling prices - total | BCI tendency prices - total | Survey | 0 | Italy |
| 398 | BCI tendency production - goods | BCI tendency Prod. - Cons. gds | Survey | 0 | Italy |
| 399 | BCI tendency production - intermediate goods | BCI tendency Prod. - intermediate gds | Survey | 0 | Italy |
| 400 | BCI tendency production - capital goods | BCI tendency Prod. - Capital gds | Survey | 0 | Italy |
| 401 | BCI tendency production - total | BCI tendency Prod. - total | Survey | 0 | Italy |
| 402 | Trade balance | Trade balance | Trade | 2 | Italy |
| 403 | Turnover - total - industry excluding construction | Turnover - industry exc. constr. | Trade | 2 | Italy |
| 404 | Turnover - internal market | Turnover - internal market | Industry | 2 | Italy |
| 405 | Turnover - external market | Turnover - external market | Industry | 2 | Italy |
| 406 | Unemployed (15 years and over) - survey of labor force | Unemployed (>15) - survey | Industry | 1 | Italy |
| 407 | CCI - Employ. expectations | CCI - Employ. Exp. | Labor | 0 | Italy |
| 408 | Unemploy. rate (15 years and over) | Unemploy. rate (> 15) | Labor | 1 | Italy |
| 409 | Unemploy. rate (15 - 24) | Unemploy. rate (15 - 24) | Labor | 1 | Italy |
| 410 | Unemploy. rate - female | Unemploy. rate - female | Labor | 1 | Italy |
| 411 | Unemploy. rate - male | Unemploy. rate - male | Labor | 1 | Italy |
| 412 | Consumer price index - US | CPI - US | Prices | 1 | world |
| 413 | Retail sales - food - value | Retail sales - food - value | Services | 2 | Italy |
| 414 | Retail sales - non-food - value | Retail sales - non-food - value | Services | 2 | Italy |
| 415 | Retail sales - total - value | Retail salesprod. - value | Services | 2 | Italy |
| 416 | Imports - unit values | Imports - unit values | Trade | 2 | Italy |
| 417 | Imports - goods - average unit values | Imports - Cons. gds -avg | Trade | 2 | Italy |
| 418 | Imports - energy - average unit values | Imports - energy -avg | Trade | 2 | Italy |
| 419 | Imports - intermediate goods - average unit values | Imports - Intermed. gds -avg | Trade | 2 | Italy |
| 420 | Imports - capital goods - average unit values | Imports - Capital gds -avg | Trade | 2 | Italy |
| 421 | Liquidity level operative requirements - goods | Liquidity level - Cons. gds | Survey | 0 | Italy |
| 422 | Liquidity level operative requirements - intermediate goods | Liquidity level - Intermed. gds | Survey | 0 | Italy |
| 423 | Liquidity level operative requirements - capital goods | Liquidity level - Capital gds | Survey | 0 | Italy |
| 424 | Liquidity level operative requirements - manufacturing | Liquidity level - Mfg | Survey | 0 | Italy |
| 425 | Compensation of employees - financial and insurance activity | Comp. of Employ. - financ. activity | Labor | 2 | Italy |
| 426 | Compensation of employees - wholesale and retail sales | Comp. of employ. - retail sales | Labor | 2 | Italy |
| 427 | Compensation of employees - construction | Comp. of employ. - constr. | Labor | 2 | Italy |
| 428 | Compensation of employees - energy | Comp. of employ. - energy | Labor | 2 | Italy |
| 429 | Compensation of employees - industry excluding construction | Comp. of Employ. - industry exc. Constr. | Labor | 2 | Italy |
| 430 | Compensation of employees - services | Comp. of employ. - services | Labor | 2 | Italy |
| 431 | Industries output price - retail: Ireland | Ind. output price - retail: Ireland | Prices | 2 | Europe |
| 432 | Manufacturing price index: Japan | Mfg price index: Japan | Prices | 2 | world |
| 433 | World trade index | World trade index | Trade | 2 | world |

Appendix A. Ledoit–wolf estimator of the shrinkage intensity parameter

This appendix reviews the [Ledoit and Wolf \(2004a\)](#) optimal estimator of the shrinkage intensity parameter λ .

Suppressing reference to the generic i th GDP component and referring to the (h, k) pair of indicators, we let $q_{hk,\tau} = (v_{h\tau} - \bar{v}_h)(v_{k\tau} - \bar{v}_k) - \hat{s}_{hk}$, $\tau = 1, \dots, T$, and define

$$\hat{\pi}_{hk} = \frac{1}{T} \sum_{\tau} q_{hk,\tau}^2, \quad \hat{t}_{hh,hk} = \frac{1}{T} \sum_{\tau} q_{hh,\tau} q_{hk,\tau}$$

$$(h, k) = 1, \dots, N,$$

$$\hat{\pi} = \sum_{h=1}^N \sum_{k=1}^N \hat{\pi}_{hk},$$

$$\hat{\rho} = \sum_{h=1}^N \hat{\pi}_{hh} + \sum_{h=1}^N \sum_{k=1, k \neq h}^N \frac{\bar{r}}{2} \left(\sqrt{\frac{\hat{s}_{kk}}{\hat{s}_{hh}}} \hat{t}_{hh,hk} + \sqrt{\frac{\hat{s}_{hh}}{\hat{s}_{kk}}} \hat{t}_{kk,hk} \right);$$

$\hat{\pi}$ estimates the sum of the variances of the elements of the sample covariance matrix, whereas $\hat{\rho}$ estimates the sum of the asymptotic covariances of the elements of the shrinkage target with those of the sample covariance matrix.

Finally,

$$\hat{\gamma} = \sum_{h=1}^N \sum_{k=1}^N (\hat{s}_{hk} - \tilde{\omega}_{hk})^2$$

estimates the deviation of the shrinkage target from the sample covariance matrix.

The Ledoit–Wolf estimator of λ is

$$\lambda^* = \max \left\{ 0, \min \left\{ \frac{1}{T} \frac{\hat{\pi} - \hat{\rho}}{\hat{\gamma}}, 1 \right\} \right\}.$$

Appendix B. List of monthly indicators

Here, we provide the complete list of the monthly indicators used for the estimation of monthly GDP and its components, with their group or sector in the fourth column, the publication delay in months with respect to the reference month, and their reference area in the last column.

Appendix C. Supplementary data

Supplementary material related to this article can be found online at <https://doi.org/10.1016/j.ijforecast.2021.04.003>.

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