

# Economic Forecasting

## Seasonal ARIMA models

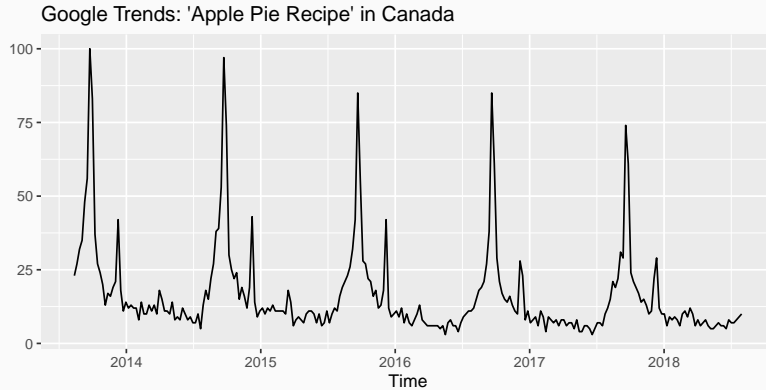
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Sebastian Fossati

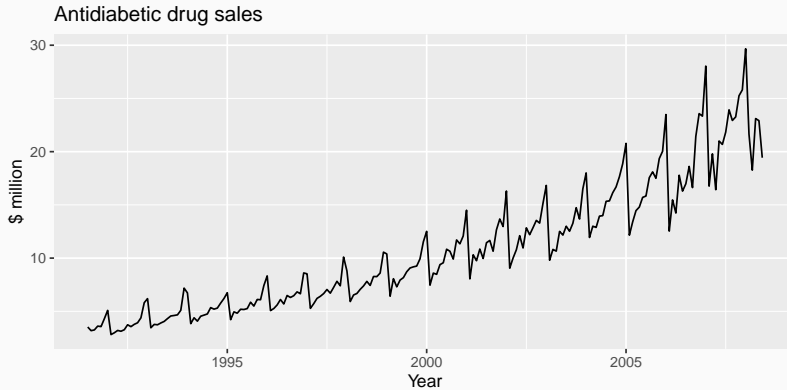
University of Alberta | E493 | 2023

- 1 Seasonal data
- 2 Time series decomposition
- 3 Seasonal ARIMA models
- 4 Automatic ARIMA modelling in R

# Seasonal data

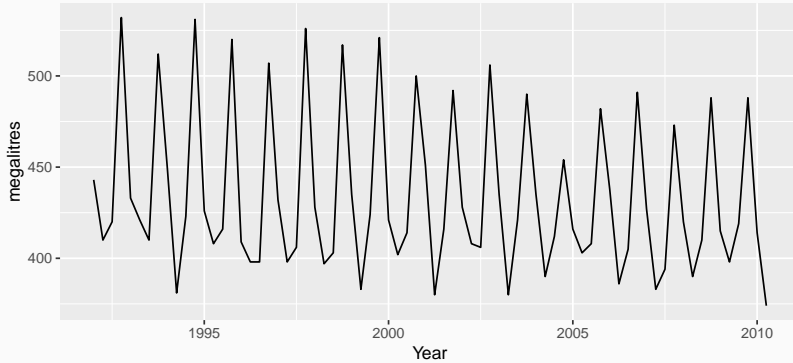


# Seasonal data



## Seasonal data

## Australian quarterly beer production

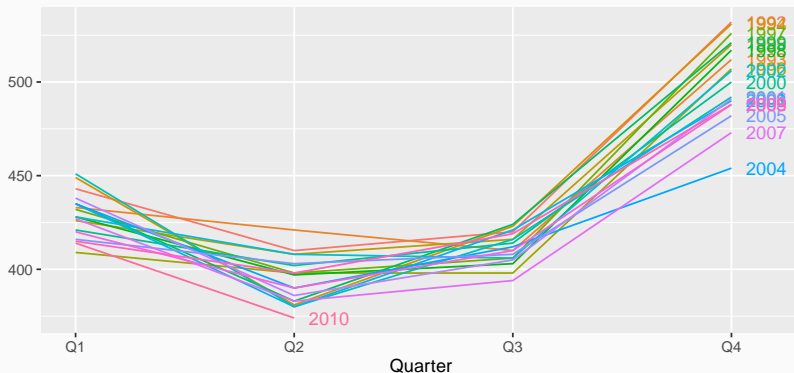


# Seasonal data

```
# season plot
```

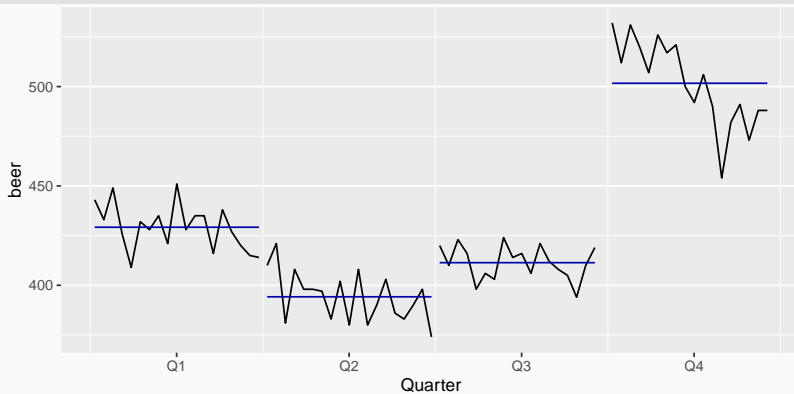
```
ggseasonplot(beer, year.labels = TRUE)
```

Seasonal plot: beer



# Seasonal data

```
# subseries plot  
ggsubseriesplot(beer)
```



**Trend:** pattern exists when there is a long-term increase or decrease in the data

**Seasonal:** pattern exists when a series is influenced by seasonal factors (eg, the quarter of the year, the month, or day of the week)

**Cyclic:** pattern exists when data exhibit rises and falls that are *not of fixed period* (duration usually of at least 2 years)



## Differences between seasonal and cyclic patterns:

- seasonal pattern constant length; cyclic pattern variable length
- average length of cycle longer than length of seasonal pattern
- magnitude of cycle more variable than magnitude of seasonal pattern

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The timing of peaks and troughs is predictable with seasonal data, but unpredictable in the long term with cyclic data.

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# Time series decomposition

$$y_t = f(S_t, T_t, R_t)$$

where  $y_t$  = data at period  $t$   
 $T_t$  = trend-cycle component at period  $t$   
 $S_t$  = seasonal component at period  $t$   
 $R_t$  = remainder component at period  $t$

# Time series decomposition

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**Additive decomposition:**  $y_t = S_t + T_t + R_t$ .

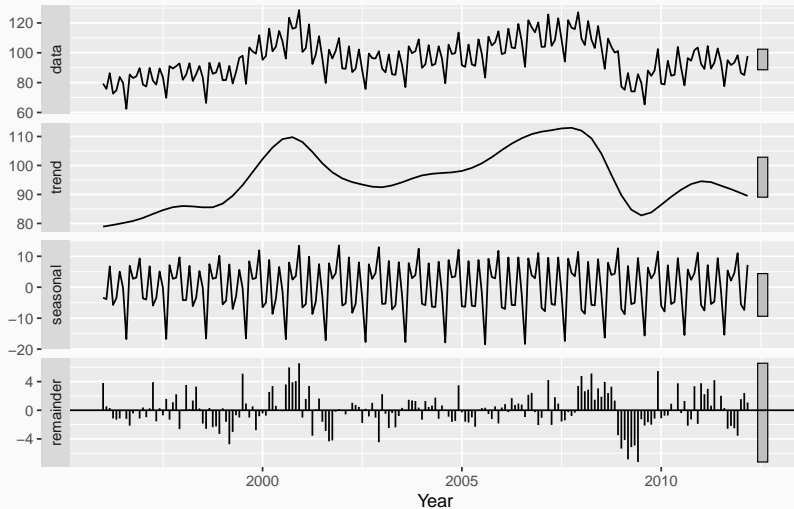
**Multiplicative decomposition:**  $y_t = S_t \times T_t \times R_t$ .

### Remarks:

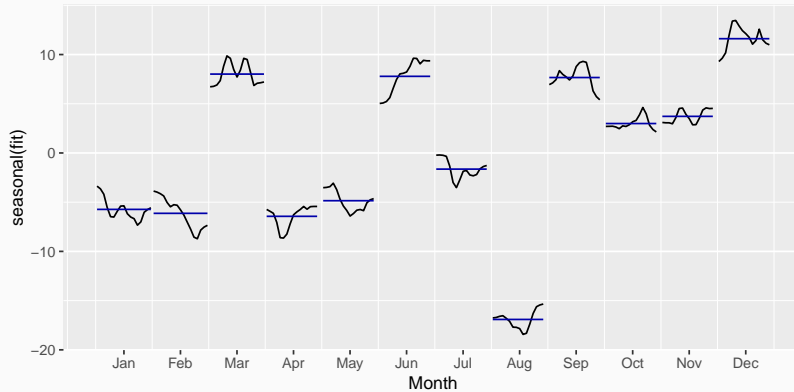
- additive model appropriate if magnitude of seasonal fluctuations does not vary with level
- if seasonal are proportional to level of series, then multiplicative model appropriate
- multiplicative decomposition more prevalent with economic series
- logs turn multiplicative relationship into an additive relationship:

$$y_t = S_t \times T_t \times E_t \quad \Rightarrow \quad \log y_t = \log S_t + \log T_t + \log R_t$$

# Euro electrical equipment



# Euro electrical equipment





# Euro electrical equipment



To decompose a time series into seasonal, trend and irregular components we can use the `stl()` function.

Helper functions:

- `seasonal()` extracts the seasonal component
- `trendcycle()` extracts the trend-cycle component
- `remainder()` extracts the remainder component
- `seasadj()` returns the seasonally adjusted series

Useful by-product of decomposition: an easy way to calculate seasonally adjusted data.

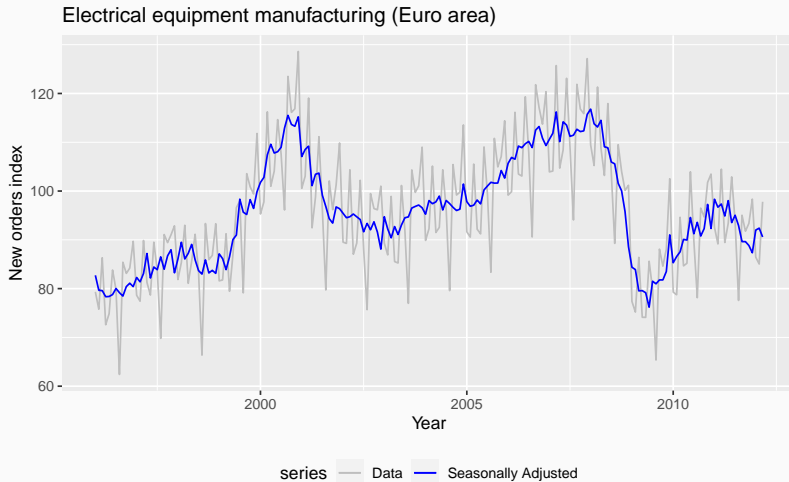
- additive decomposition: seasonally adjusted data given by

$$y_t - S_t = T_t + R_t$$

- multiplicative decomposition: seasonally adjusted data given by

$$y_t / S_t = T_t \times R_t$$

# Euro electrical equipment



## Remarks:

- we use estimates of  $S$  based on past values to seasonally adjust a current value
- seasonally adjusted series reflect **remainders** as well as **trend**

# History of time series decomposition

## History:

- Classical method originated in 1920s
- Census II method introduced in 1957 (basis for X-11 method and X-12-ARIMA, X-13-ARIMA variants)
- STL method introduced in 1983
- TRAMO/SEATS introduced in 1990s

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## National Statistics Offices:

- US Census Bureau uses X-13-ARIMA-SEATS
- Statistics Canada uses X-12-ARIMA
- EuroStat use X-13-ARIMA-SEATS

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### Autoregressive Integrated Moving Average models

#### ARIMA( $p, d, q$ ) model

AR:  $p$  = order of the autoregressive part

I:  $d$  = degree of first differencing involved

MA:  $q$  = order of the moving average part

## Non-seasonal ARIMA models

ARIMA(1,1,1) model:

$$\begin{array}{ccccc} (1 - \phi_1 B) & (1 - B)y_t & = & c + (1 + \theta_1 B)\varepsilon_t \\ \uparrow & \uparrow & & \uparrow \\ \text{AR(1)} & \text{First} & & \text{MA(1)} \\ & \text{difference} & & \end{array}$$

## Non-seasonal ARIMA models

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$$\begin{array}{ccccc}(1 - \phi_1 B) & (1 - B)y_t & = & c + (1 + \theta_1 B)\varepsilon_t \\ \uparrow & \uparrow & & \uparrow \\ \text{AR(1)} & \text{First} & & \text{MA(1)} \\ & \text{difference} & & \end{array}$$

Written out:  $y_t = c + y_{t-1} + \phi_1 y_{t-1} - \phi_1 y_{t-2} + \theta_1 \varepsilon_{t-1} + \varepsilon_t$

$$y_t - y_{t-1} = c + \phi_1(y_{t-1} - y_{t-2}) + \theta_1 \varepsilon_{t-1} + \varepsilon_t$$

$$y'_t = c + \phi_1 y'_{t-1} + \theta_1 \varepsilon_{t-1} + \varepsilon_t$$

## Seasonal ARIMA models

ARIMA	$\underbrace{(p, d, q)}$	$\underbrace{(P, D, Q)_m}$
	↑	↑
	Non-seasonal part of the model	Seasonal part of of the model

where  $m$  = number of observations per year.

## Seasonal ARIMA models

E.g.,  $\text{ARIMA}(1, 1, 1)(1, 1, 1)_4$  model (without constant)

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E.g., ARIMA(1, 1, 1)(1, 1, 1)<sub>4</sub> model (without constant)

$$(1 - \phi_1 B)(1 - \Phi_1 B^4)(1 - B)(1 - B^4)y_t = (1 + \theta_1 B)(1 + \Theta_1 B^4)\varepsilon_t.$$

## Seasonal ARIMA models

E.g., ARIMA(1, 1, 1)(1, 1, 1)<sub>4</sub> model (without constant)

$$(1 - \phi_1 B)(1 - \Phi_1 B^4)(1 - B)(1 - B^4)y_t = (1 + \theta_1 B)(1 + \Theta_1 B^4)\varepsilon_t.$$



## Seasonal ARIMA models

E.g., ARIMA(1, 1, 1)(1, 1, 1)<sub>4</sub> model (without constant)

$$(1 - \phi_1 B)(1 - \Phi_1 B^4)(1 - B)(1 - B^4)y_t = (1 + \theta_1 B)(1 + \Theta_1 B^4)\varepsilon_t.$$

All the factors can be multiplied out and the general model written as follows:

$$\begin{aligned} y_t = & (1 + \phi_1)y_{t-1} - \phi_1 y_{t-2} + (1 + \Phi_1)y_{t-4} \\ & - (1 + \phi_1 + \Phi_1 + \phi_1 \Phi_1)y_{t-5} + (\phi_1 + \phi_1 \Phi_1)y_{t-6} \\ & - \Phi_1 y_{t-8} + (\Phi_1 + \phi_1 \Phi_1)y_{t-9} - \phi_1 \Phi_1 y_{t-10} \\ & + \varepsilon_t + \theta_1 \varepsilon_{t-1} + \Theta_1 \varepsilon_{t-4} + \theta_1 \Theta_1 \varepsilon_{t-5}. \end{aligned}$$



## Common ARIMA models

The US Census Bureau uses the following models most often:

ARIMA(0,1,1)(0,1,1)<sub>m</sub> with log transformation

ARIMA(0,1,2)(0,1,1)<sub>m</sub> with log transformation

ARIMA(2,1,0)(0,1,1)<sub>m</sub> with log transformation

ARIMA(0,2,2)(0,1,1)<sub>m</sub> with log transformation

ARIMA(2,1,2)(0,1,1)<sub>m</sub> with no transformation

The seasonal part of an AR or MA model will be seen in the seasonal lags of the PACF and ACF.

**ARIMA(0,0,0)(0,0,1)<sub>12</sub> will show:**

- a spike at lag 12 in the ACF but no other significant spikes
- the PACF will show exponential decay in the seasonal lags; that is, at lags 12, 24, 36, ...

**ARIMA(0,0,0)(1,0,0)<sub>12</sub> will show:**

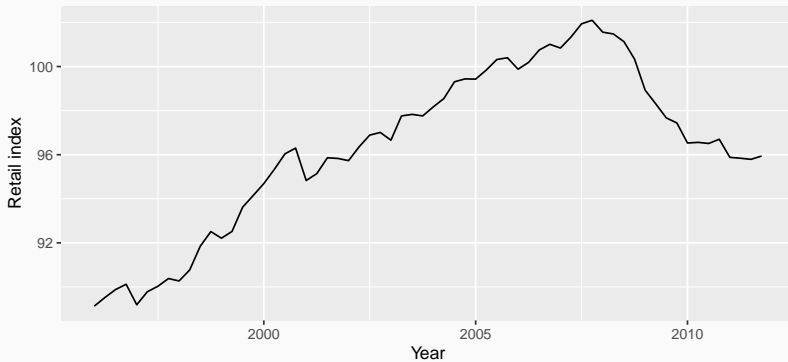
- exponential decay in the seasonal lags of the ACF
- a single significant spike at lag 12 in the PACF

# European quarterly retail trade

```
# retail data
```

```
autoplot(euretail) + xlab("Year") + ylab("Retail index") +  
  ggtitle("European quarterly retail trade")
```

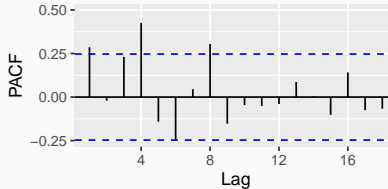
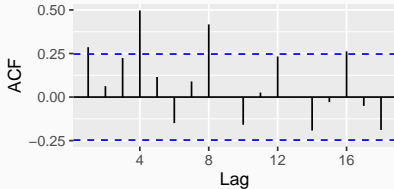
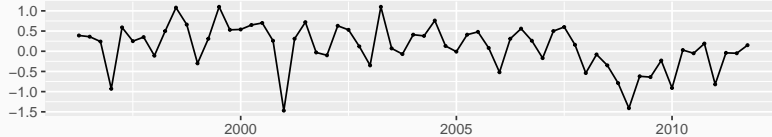
European quarterly retail trade



# European quarterly retail trade

```
# first diff
```

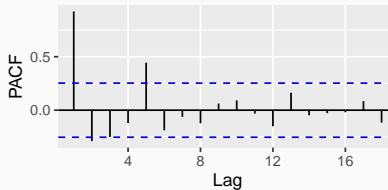
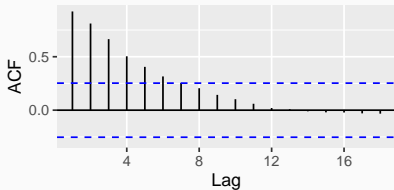
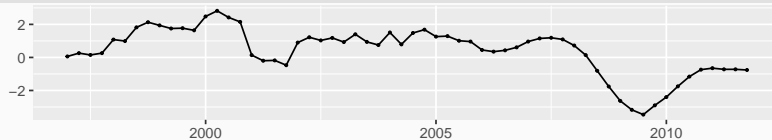
```
ggtsdisplay(diff(euretail, lag = 1))
```



# European quarterly retail trade

```
# seasonal diff
```

```
ggtsdisplay(diff(euretail, lag = 4))
```



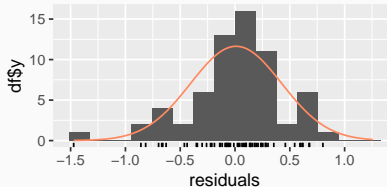
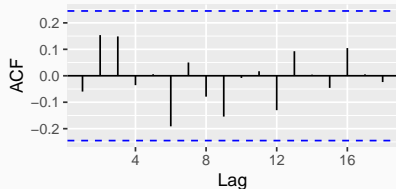
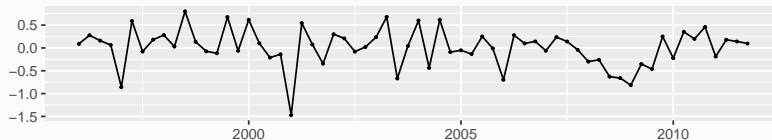
### Remarks:

- $d = 1$  seems necessary (maybe  $D = 1$  as well?)
- one significant spike at lag 1 in ACF and PACF suggests non-seasonal AR(1) or MA(1) component
- two significant spikes at lag 4 and 8 in PACF suggests seasonal AR(2) component
- initial candidate model:  $\text{ARIMA}(1,1,0)(2,0,0)_4$

# European quarterly retail trade

```
fit1 <- Arima(euretail, order=c(1,1,0), seasonal=c(2,0,0))  
checkresiduals(fit1, lag=24, test=FALSE)
```

Residuals from ARIMA(1,1,0)(2,0,0)[4]



## European quarterly retail trade

```
# residuals
checkresiduals(fit1, lag=24, plot=FALSE)

##
##  Ljung-Box test
##
## data:  Residuals from ARIMA(1,1,0)(2,0,0)[4]
## Q* = 14, df = 21, p-value = 0.9
##
## Model df: 3.    Total lags used: 24
```



Models with lowest AICc values tend to give slightly better results than the other models.

- AICc of ARIMA(1,1,0)(2,0,0)<sub>4</sub> model is 79.65
- AICc of ARIMA(0,1,1)(2,0,0)<sub>4</sub> model is 82.32
- AICc of ARIMA(1,1,1)(2,0,0)<sub>4</sub> model is 80.01
- AICc of ARIMA(2,1,0)(2,0,0)<sub>4</sub> model is 79.61

Models with lowest AICc values tend to give slightly better results than the other models.

- AICc of ARIMA(1,1,0)(2,0,0)<sub>4</sub> model is 79.65
- AICc of ARIMA(1,1,0)(3,0,0)<sub>4</sub> model is 81.98
- AICc of ARIMA(1,1,0)(1,0,0)<sub>4</sub> model is 85.94
- AICc of ARIMA(1,1,0)(1,0,1)<sub>4</sub> model is 78

## European quarterly retail trade

```
# fit model
(fit2 <- Arima(euretail, order=c(1,1,0), seasonal=c(1,0,1)))

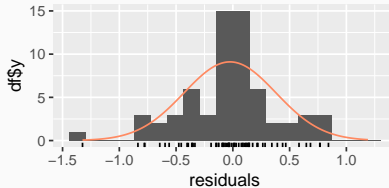
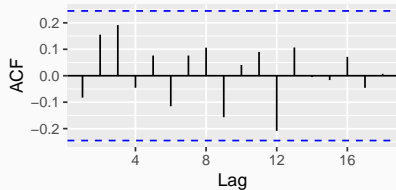
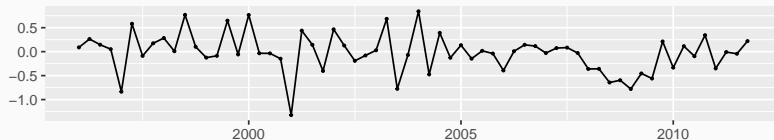
## Series: euretail
## ARIMA(1,1,0)(1,0,1)[4]
##
## Coefficients:
##          ar1    sar1    sma1
##         0.397  0.966  -0.737
## s.e.   0.127  0.044   0.161
##
## sigma^2 = 0.173: log likelihood = -34.66
## AIC=77.31   AICc=78   BIC=85.88
```

# European quarterly retail trade

```
# residuals
```

```
checkresiduals(fit2, lag=24, test=FALSE)
```

Residuals from ARIMA(1,1,0)(1,0,1)[4]



```
# residuals
checkresiduals(fit2, lag=24, plot=FALSE)

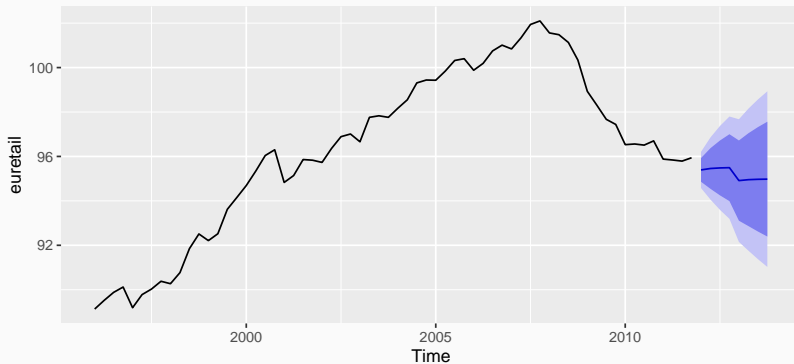
##
##  Ljung-Box test
##
## data:  Residuals from ARIMA(1,1,0)(1,0,1)[4]
## Q* = 16, df = 21, p-value = 0.8
##
## Model df: 3.    Total lags used: 24
```

# European quarterly retail trade

```
# residuals
```

```
autoplot(forecast(fit2))
```

Forecasts from ARIMA(1,1,0)(1,0,1)[4]



### Remarks:

- the textbook suggests the candidate model:  
 $\text{ARIMA}(0,1,3)(0,1,1)_4 \dots$
- AIC/AICc/BIC comparisons must have the same orders of differencing
- but **RMSE test set** comparisons can involve any models

## European quarterly retail trade

```
# data and set up
n.end <- 2003.75 # 2003Q4
h.val <- 1
# loop
pred <- matrix(rep(NA,96),32,3)
for(i in 1:32){
  tmp0 <- 1996
  tmp1 <- n.end+(i-1)*.25
  tmp <- window(euretail,tmp0,tmp1)
  pred[i,1] <- window(euretail,tmp1+h.val*.25,tmp1+h.val*.25) # actual
  # estimate models
  fit1 <- Arima(tmp, order=c(1,1,0), seasonal=c(1,0,1))
  fit2 <- Arima(tmp, order=c(0,1,3), seasonal=c(0,1,1))
  # compute forecasts (h=1)
  pred[i,2] <- forecast(fit1, h=h.val)$mean[h.val]
  pred[i,3] <- forecast(fit2, h=h.val)$mean[h.val]
}
```



## European quarterly retail trade

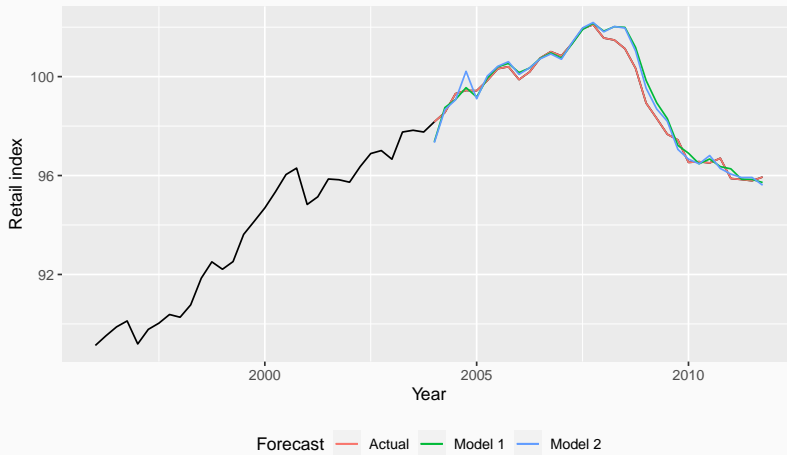
```
# compute rmse (h=1)
retailfit1 <- pred[,1] - pred[,2]
retailfit2 <- pred[,1] - pred[,3]
rmse1 <- sqrt(mean(retailfit1^2, na.rm=TRUE))
rmse2 <- sqrt(mean(retailfit2^2, na.rm=TRUE))
# display rmse
cbind(rmse1,rmse2)

##           rmse1  rmse2
## [1,] 0.3944 0.3795
```

- choose forecasting model with the smallest RMSE computed using time series cross-validation

# European quarterly retail trade

Forecasts for European quarterly retail trade ( $h=1$ )



## European quarterly retail trade

```
# data and set up
n.end <- 2003.75 # 2003Q4
h.val <- 4
# loop
pred <- matrix(rep(NA,87),29,3)
for(i in 1:29){
  tmp0 <- 1996
  tmp1 <- n.end+(i-1)*.25
  tmp <- window(euretail,tmp0,tmp1)
  pred[i,1] <- window(euretail,tmp1+h.val*.25,tmp1+h.val*.25) # actual
  # estimate models
  fit1 <- Arima(tmp, order=c(1,1,0), seasonal=c(1,0,1))
  fit2 <- Arima(tmp, order=c(0,1,3), seasonal=c(0,1,1))
  # compute forecasts (h=4)
  pred[i,2] <- forecast(fit1, h=h.val)$mean[h.val]
  pred[i,3] <- forecast(fit2, h=h.val)$mean[h.val]
}
```

## European quarterly retail trade

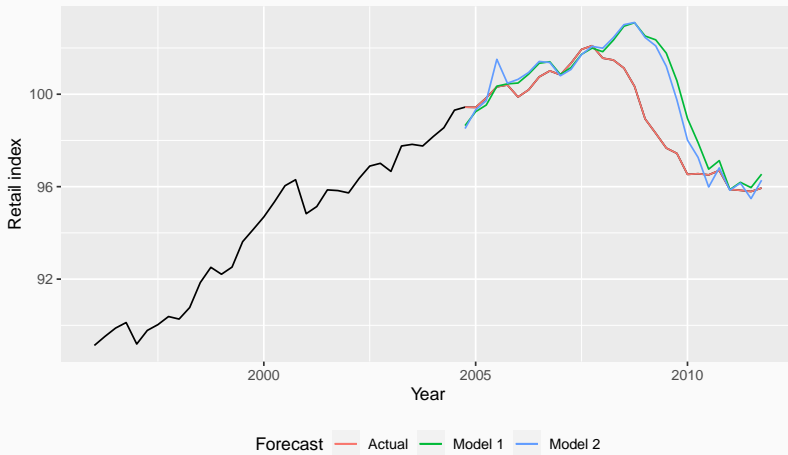
```
# compute rmse (h=4)
retailfit1 <- pred[,1] - pred[,2]
retailfit2 <- pred[,1] - pred[,3]
rmse1 <- sqrt(mean(retailfit1^2, na.rm=TRUE))
rmse2 <- sqrt(mean(retailfit2^2, na.rm=TRUE))
# display rmse
cbind(rmse1,rmse2)

##          rmse1 rmse2
## [1,] 1.643 1.489
```

- choose forecasting model with the smallest RMSE computed using time series cross-validation

# European quarterly retail trade

Forecasts for European quarterly retail trade (h=4)



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### A non-seasonal ARIMA process

$$\phi(B)(1 - B)^d y_t = c + \theta(B)\varepsilon_t$$

Need to select appropriate orders:  $p, q, d$ .

### Hyndman and Khandakar (JSS, 2008) algorithm:

- select no. differences  $d$  and  $D$  via KPSS test and seasonal strength measure
- select  $p, q$  by minimising AICc
- use stepwise search to traverse model space

## How does auto.arima() work?

**Step1:** Select model (with smallest AICc) from:

- ARIMA(2,  $d$ , 2)
- ARIMA(0,  $d$ , 0)
- ARIMA(1,  $d$ , 0)
- ARIMA(0,  $d$ , 1)

**Step 2:** Consider variations of current model:

- vary one of  $p, q$  by  $\pm 1$
- $p, q$  both vary by  $\pm 1$
- include/exclude  $c$

Model with lowest AICc becomes current model.

**Repeat Step 2 until no lower AICc can be found.**



## Choosing your own model

```
# fit model
(fit <- auto.arima(euretail))

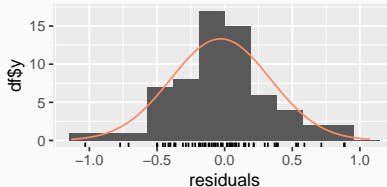
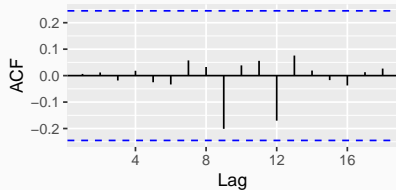
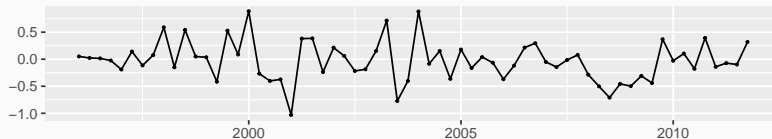
## Series: euretail
## ARIMA(0,1,3)(0,1,1)[4]
##
## Coefficients:
##          ma1      ma2      ma3      sma1
##          0.263  0.369  0.420  -0.664
## s.e.   0.124  0.126  0.129   0.155
##
## sigma^2 = 0.156:  log likelihood = -28.63
## AIC=67.26   AICc=68.39   BIC=77.65
```

# Choosing your own model

```
# check residuals
```

```
checkresiduals(fit, test=FALSE)
```

Residuals from ARIMA(0,1,3)(0,1,1)[4]



# Choosing your own model

```
# check residuals
checkresiduals(fit, plot=FALSE)

##
##  Ljung-Box test
##
## data:  Residuals from ARIMA(0,1,3)(0,1,1)[4]
## Q* = 0.51, df = 4, p-value = 1
##
## Model df: 4.    Total lags used: 8
```

## Modelling procedure with `auto.arima`

- 1 Plot the data, identify any unusual observations
- 2 If necessary, transform the data (using a Box-Cox transformation) to stabilize the variance
- 3 Use `auto.arima` to select a model
- 4 Check the residuals from your chosen model by plotting the ACF of the residuals, and doing a portmanteau test of the residuals
- 5 If they do not look like white noise, try a modified model
- 6 Once the residuals look like white noise, calculate forecasts