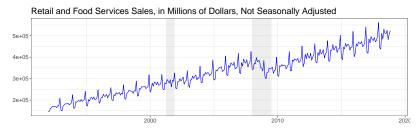
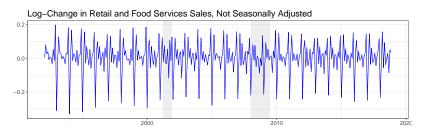
Eco 5316 Time Series Econometrics Lecture 2 Autoregressive (AR) processes

Outline

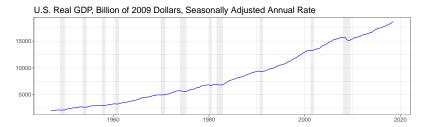
- 1. Features of Time Series
- 2. Box-Jenkins methodology
- 3. Autoregressive Model AR(p)
- 4. Autocorrelation Function (ACF)
- 5. Partial Autocorrelation Function (PACF)
- 6. Portmanteau Test Box-Pierce test and Ljung-Box test
- 7. Information Criteria Akaike (AIC) and Schwarz-Bayesian (BIC)
- 8. Example: AR model for Real GNP growth rate

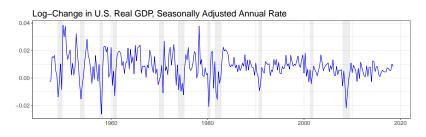
Retail and Food Services Sales https://www.quandl.com/data/FRED/RSAFSNA





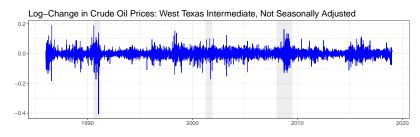
Real GDP https://www.quandl.com/data/FRED/GDPC1



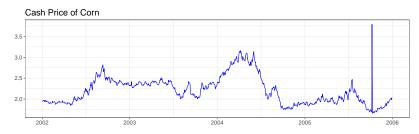


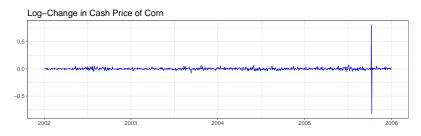
Crude Oil Prices: https://www.quandl.com/data/FRED/DCOILWTICO





Cash Price of Corn (October 8, 2005) https://www.quandl.com/data/TFGRAIN/CORN





decomposition of time series into trend, seasonal and irregular component

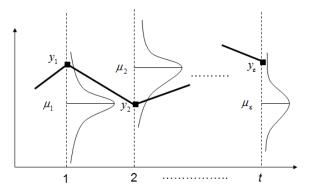
$$y_t = \mu_t + \gamma_t + \varepsilon_t$$

where

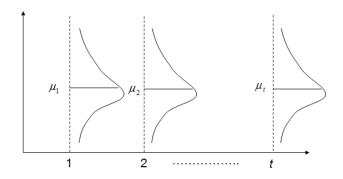
 y_t is the observed data μ_t is an slowly changing component (trend) γ_t is periodic seasonal component ε_t is irregular disturbance component

- classical approach treat trend and seasonal components as deterministic functions
- \blacktriangleright modern approach μ_t , γ_t , ε_t all contain stochastic components
- we will first look at the ways how to model the irregular component, and leave seasonal and trend components for later

Def: Stochastic process (or time series process) is a sequence of random variables $\{y_t\}$, observed time series is a particular realization of this process.

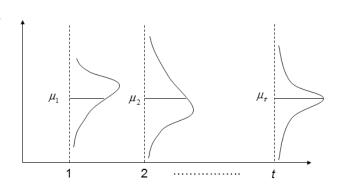


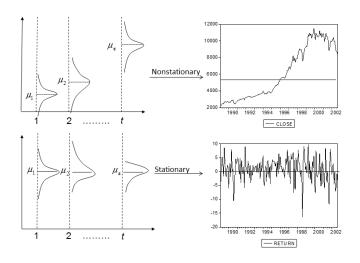
Def: Stochastic process $\{y_t\}$ is **strictly stationary** if joint distributions $F(y_{t_1},\ldots,y_{t_k})$ and $F(y_{t_1+l},\ldots,y_{t_k+l})$ are identical for all l, k and all t_1,\ldots,t_k



Def: Stochastic process $\{y_t\}$ is weakly stationary if

- (i) $E(y_t) = \mu$ for all t
- (ii) $cov(y_t, y_{t-l}) = \gamma_l$ for all t, l





- weak stationarity allows us to use sample moments to estimate population moments
- ▶ for example, given a weakly stationary time series $\{y_1,y_2,\ldots,y_t\}$ the first moment $E(y_t)$ can be estimated using $\frac{1}{t}\sum_{j=1}^t y_j$ which would make little sense if $E(y_1) \neq E(y_2) \neq \ldots \neq E(y_t)$

Def: Stochastic process $\{\varepsilon_t\}$ is called a **white noise** if ε_t are independently identically distributed with zero mean and finite variance: $E(\varepsilon_t)=0$, $Var(\varepsilon_t)=\sigma_\varepsilon^2<\infty$, $cov(\varepsilon_t,\varepsilon_s)=0$ for all $t\neq s$.

Box-Jenkins methodology to modelling weakly stationary time series

- 1. Identification
- 2. Estimation
- 3. Checking Model Adequacy

1. Indentification

- examine time series plots of the data to determine if any transformations are necessary (differencing, logarithms) to get weakly stationary time series, examine series for trend (linear/nonlinear), periods of higher volatility, seasonal patterns, structural breaks, outliers, missing data, . . .
- examine autocorrelation function (ACF) and partial autocorrelation function (PACF) of the transformed data to determine plausible models to be estimated
- use Q-statistics to test whether groups of autocorrelations are statistically significant

2. Estimation

- estimate all models considered and select the best one coefficients should be statistically significant, information criteria (AIC, SBC) should be low
- model can be estimated using either conditional likelihood method or exact likelihood method

3. Checking Model Adequacy

- perform in-sample evaluation of the estimated model
 - estimated coefficients should be consistent with the underlying assumption of stationarity
 - inspect residuals if the model was well specified residuals should be very close to white-noise
 - plot residuals, look for outliers, periods in which the model does not fit the data well, evidence of structural change
 - examine ACF and PACF of the residuals to check for significant autocorrelations
 - use Q-statistics to test whether autocorrelations of residuals are statistically significant
 - check model for parameter instability and structural change
- perform out-of-sample evaluation of the model forecast

▶ we will now look at how the Box-Jenkins methodology works in case of a simple univariate time series model - an autoregressive model

▶ simple linear regression model with cross sectional data

$$y_i = \beta_0 + \beta_1 x_i + \epsilon_i$$

suppose we are dealing with time series rather than cross sectional data, so that

$$y_t = \beta_0 + \beta_1 x_t + \epsilon_t$$

and if the explanatory variable is the lagged dependent variable $x_t = y_{t-1}$ we get

$$y_t = \beta_0 + \beta_1 y_{t-1} + \epsilon_t$$

main idea: past is prologue as it determines the present, which in turn sets the stage for future

▶ hourly time series for Akkoro Kamui's activities, before the fortress was built

$$\{y_1, y_2, \dots, y_t\} = \{drink, drink, \dots, drink\}$$

lots of time dependence here:

$$y_t = y_{t-1}$$

lacktriangle time series process $\{y_t\}$ follows autoregressive model of order 1, AR(1), if

$$y_t = \phi_0 + \phi_1 y_{t-1} + \varepsilon_t$$

or equivalently, using the lag operator

$$(1 - \phi_1 L)y_t = \phi_0 + \varepsilon_t$$

where $\{\varepsilon_t\}$ is a white noise with $E(\varepsilon_t) = 0$ and $Var(\varepsilon_t) = \sigma_{\varepsilon}^2$

▶ more generally, time series $\{y_t\}$ follows an autoregressive model of order p, AR(p), if

$$y_t = \phi_0 + \phi_1 y_{t-1} + \ldots + \phi_p y_{t-p} + \varepsilon_t$$

or equivalently, using the lag operator

$$(1-\phi_1L-\ldots-\phi_pL^p)y_t=\phi_0+\varepsilon_t$$

tools to determined the order p of the autoregressive model given $\{y_t\}$

- ► Autocorrelation Function (ACF)
- ► Partial Autocorrelation Function (PACF)
- ► Portmanteau Test Box-Pierce test and Ljung-Box test
- Information Criteria Akaike (AIC) and Schwarz-Bayesian (BIC)

Autocorrelation Function (ACF)

- lacktriangle linear dependence between y_t and y_{t-l} is given by correlation coefficient ho_l
- lacktriangle for a weakly stationary time series process $\{y_t\}$ we have

$$\rho_l = \frac{cov(y_t, y_{t-l})}{\sqrt{Var(y_t)Var(y_{t-l})}} = \frac{cov(y_t, y_{t-l})}{Var(y_t)} = \frac{\gamma_l}{\gamma_0}$$

- **b** theoretical autocorrelation function is $\{\rho_1, \rho_2, \ldots\}$
- lacktriangle given a sample $\{y_t\}_{t=1}^T$ correlation coefficients ho_l can be estimated as

$$\hat{\rho}_{l} = \frac{\sum_{t=l+1}^{T} (y_{t} - \bar{y})(y_{t-l} - \bar{y})}{\sum_{t=1}^{T} (y_{t} - \bar{y})^{2}}$$

where $\bar{y} = \frac{1}{T} \sum_{t=1}^{T} y_t$

> sample autocorrelation function is $\{\hat{\rho}_1, \hat{\rho}_2, \ldots\}$

Autocorrelation function for AR(p) model

if p=1 then $\gamma_0=Var(y_t)=\frac{\sigma_{\varepsilon}^2}{1-\phi_1^2}$ and also $\gamma_l=\phi_1\gamma_{l-1}$ for l>0, thus $\rho_l=\phi_1\rho_{l-1} \tag{1}$

and since $\rho_0=1$, we get $\rho_l=\phi_1^l$

• for weakly stationary $\{y_t\}$ it has to hold that $|\phi_1| < 1$, theoretical ACF of a stationary AR(1) thus decays exponentially, in either direct or oscillating way

Autocorrelation function for AR(p) model

ightharpoonup if p=2 theoretical ACF for AR(2) satisfies second order difference equation

$$\rho_l = \phi_1 \rho_{l-1} + \phi_2 \rho_{l-2} \tag{2}$$

or equivalently using the lag operator $(1-\phi_1L-\phi_2L^2)\rho_l=0$

solutions of the associated characteristic equation

$$1 - \phi_1 x - \phi_2 x^2 = 0$$

are
$$x_{1,2} = -\frac{\phi_1 \pm \sqrt{\phi_1^2 + 4\phi_2}}{2\phi_2}$$

- \blacktriangleright their inverses $\omega_{1,2}=1/x_{1,2}$ are called the characteristic roots of the AR(2) model
- if $D=\phi_1^2+4\phi_2>0$ then ω_1,ω_2 are real numbers, and theoretical ACF is a combination of two exponential decays
- ightharpoonup if D<0 characteristic roots are complex conjugates, and theoretical ACF will resemble a dampened sine wave
- for weak stationarity all characteristic roots need to lie inside the unit circle, that is $|\omega_i|<1$ for i=1,2
- ▶ from equation (2) we get $\rho_1 = \frac{\phi_1}{1-\phi_2}$ and $\rho_l = \rho_{l-1} + \phi_2 \rho_{l-2}$ for $l \ge 2$

Autocorrelation function for AR(p) model

 \blacktriangleright in general, theoretical ACF for AR(p) satisfies the difference equation of order p

$$(1-\phi_1L-\ldots-\phi_pL^p)\rho_l=0$$
(3)

- characteristic equation of the AR(p) model is thus $1-\phi_1x-\ldots-\phi_px^p=0$
- ► AR(p) process is weakly stationary if the characteristic roots (i.e. inverses of the solutions of the characteristic equation) lie inside of the unit circle
- plot of the theoretical ACF of a weakly stationary AR(p) process will show a mixture of exponential decays and dampened sine waves

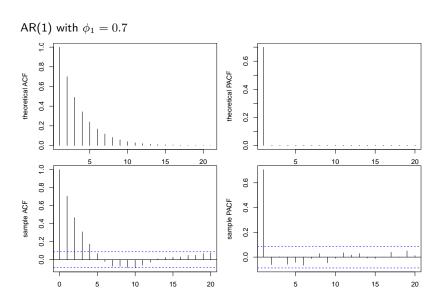
Partial autocorrelation function (PACF)

consider the following system of AR models that can be estimated by OLS

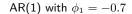
$$\begin{split} y_t &= \phi_{0,1} + \phi_{1,1} y_{t-1} + e_{1,t} \\ y_t &= \phi_{0,2} + \phi_{1,2} y_{t-1} + \phi_{2,2} y_{t-2} + e_{2,t} \\ y_t &= \phi_{0,3} + \phi_{1,3} y_{t-1} + \phi_{2,3} y_{t-2} + \phi_{3,3} y_{t-3} + e_{3,t} \\ &\vdots \end{split}$$

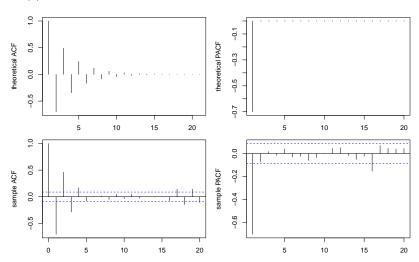
- ightharpoonup estimated coefficients $\hat{\phi}_{1,1},\hat{\phi}_{2,2},\hat{\phi}_{3,3},\ldots$ form the sample partial autocorrelation function (PACF)
- ▶ if the time series process $\{y_t\}$ comes from an AR(p) process, sample PACF should have $\hat{\phi}_{j,j}$ close to zero for j>p
- ▶ for an AR(p) with Gaussian white noise as T goes to infinity $\hat{\phi}_{p,p}$ converges to ϕ_p and $\hat{\phi}_{l,l}$ converges to 0 for l>p, in addition the asymptotic variance of $\hat{\phi}_{l,l}$ for l>p is 1/T
- \blacktriangleright this is the reason why the interval plotted by R in the plot of PACF is $0\pm 2/\sqrt{T}$
- order of the AR process can thus be determined by finding the lag after which PACF cuts off to zero

ACF and PACF for AR(1) model

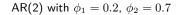


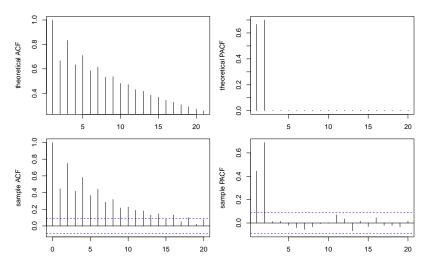
ACF and PACF for AR(1) model





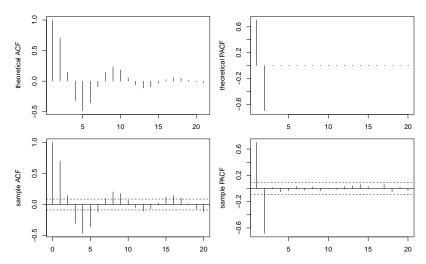
ACF and PACF for AR(2) model





ACF and PACF for AR(2) model





ACF and PACF for AR(p) model

 \blacktriangleright interactive overview of ACF and PACF for simulated AR(p) models is here

Portmanteau Test

▶ to test $H_0: \rho_1 = \ldots = \rho_m = 0$ against an alternative hypothesis $H_a: \rho_j \neq 0$ for some $j \in \{1, \ldots, m\}$ following two statistics can be used:

Box-Pierce test

$$Q^*(m) = T \sum_{l=1}^{m} \hat{\rho}_l^2$$

Ljung-Box test

$$Q(m) = T(T+2) \sum_{l=1}^{m} \frac{\hat{\rho}_{l}^{2}}{T-l}$$

- ▶ the null hypothesis is rejected at $\alpha\%$ level if the above statistics are larger than the $100(1-\alpha)$ th percentile of chi-squared distribution with m degrees of freedom
- ▶ note: Ljung-Box statistics tends to perform better in smaller samples
- ▶ the general recommendation is to use $m \approx \ln T$, but this depends on application
- ightharpoonup e.g.: for monthly data with a seasonal pattern it makes sense to set m to 12, 24 or 36, and for quarterly data with a seasonal pattern m to 4, 8, 12

Portmanteau Test

- ▶ these tests are also used for in-sample evaluation of model adequacy
- ▶ if the model was correctly specified Ljung-Box Q(m) statistics for the residuals of the estimated model follows chi-squared distribution with $m\!-\!g$ degrees of freedom where g is the number of estimated parameters
- for AR(p) that includes a constant g = p+1

Information Criteria

- in practice, there will be often several competing models that would be considered
- if these models are adequate and with very similar properties based on ACF, PACF, and Q statistics for residuals, information criteria can help decide which one is preferred
- main idea: information criteria combine the goodness of fit with a penalty for using more parameters

Information Criteria

two commonly used information criteria:

Akaike Information Criterion (AIC)

$$AIC = -\frac{2}{T}\log L + \frac{2}{T}n$$

Schwarz-Bayesian information criterion (BIC)

$$BIC = -\frac{2}{T}\log L + \frac{\log T}{T}n$$

in both expressions above T is the sample size, n is the number of parameters in the model, L is the value of the likelihood function, and \log is the natural logarithm

- ► AIC or BIC of competing models can be compared and the model that has the smallest AIC or BIC value is preferred
- ightharpoonup BIC will always select a more parsimonious model with fewer parameters than the AIC because $\log T>2$ and each additional parameter is thus penalized more heavily

Information Criteria

- fundamental difference AIC tries to select the model that most adequately approximates unknown complex data generating process with infinite number of parameters
- this true process is never in the set of candidate models that are being considered
- ▶ BIC assumes that the true model is among the set of considered candidates and tries to identify it
 - ▶ BIC performs better than AIC in large samples it is asymptotically consistent while AIC is biased toward selecting an overparameterized model
- in small samples AIC can perform better than BIC

Information Criteria

- some software packages report other information criteria in addition to AIC and BIC
- ► Hannan-Quinn information criterion (HQ)

$$HQ = -\frac{2}{T}\log L + \frac{2\log(\log T)}{T}n$$

corrected Akaike Information Criterion (AICc) which is AIC with a correction for finite sample sizes to limit overfitting; for a univariate linear model with normal residuals

$$AICc = AIC + \frac{2(n+1)(n+2)}{T - n - 2}$$

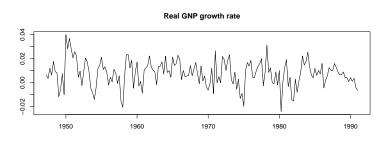
where T is the sample size and n is the number of estimated parameters

an example showing the steps of estimating and checking a model for the growth rate of GNP can be found here: lec03GNP.zip

```
str(y)
## Time-Series [1:176] from 1947 to 1991: 0.00632 0.00366 0.01202 0.00627 0.01761 ...
head(y)
## [1] 0.00632 0.00366 0.01202 0.00627 0.01761 0.00918
tail(y)
## [1] 0.00085 0.00420 0.00108 0.00358 -0.00399 -0.00650
```

▶ plot using base package

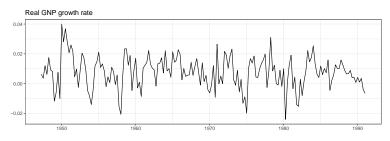
```
plot(y, xlab="", ylab="", main="Real GNP growth rate")
```



▶ plot using ggplot2 package

```
# load ggplot2
library(ggplot2)
# load ggfortify to be able to plot time series and output from acf using autoplot
library(ggfortify)

## Warning: package 'ggfortify' was built under R version 3.5.2
# define deafult theme to be BEW
theme_set(theme_bw())
# plot
autoplot(y) +
    labs(x="", y="", title="Real GNP growth rate")
```



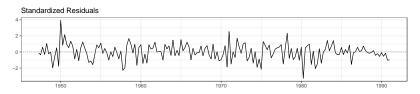
```
# plot ACF and PACF for y up to lag 24
y %>% as.data.frame() %>% acf(type="correlation",lag=24, plot=FALSE) %>% autoplot()
  1.00 -
  0.75
₽ 0.50
  0.25
  0.00
                                                10
                                                                                      20
                                                       Lag
y %>% as.data.frame() %>% acf(type="partial",lag=24, plot=FALSE) %>% autoplot()
   0.3
   0.2
   0.1
   0.0
   -0.1
                                                                                     20
                                                                 15
                                                       Lag
```

Fxample: AD model for Deal CND growth rate

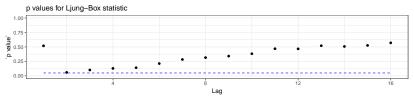
```
# estimate an AR(1) model - there is only one significant coefficient in the PACF plot for y
m1 \leftarrow arima(y, order=c(1,0,0))
# show the structure of object m1
str(m1)
## List of 14
## $ coef : Named num [1:2] 0.37865 0.00769
  ..- attr(*, "names")= chr [1:2] "ar1" "intercept"
## $ sigma2 : num 9.8e-05
## $ var.coef : num [1:2, 1:2] 4.88e-03 -1.12e-06 -1.12e-06 1.44e-06
## ..- attr(*, "dimnames")=List of 2
## ....$ : chr [1:2] "ar1" "intercept"
## ....$ : chr [1:2] "ar1" "intercept"
## $ mask : logi [1:2] TRUE TRUE
## $ loglik : num 562
## $ aic : num -1119
## $ arma : int [1:7] 1 0 0 0 4 0 0
## $ residuals: Time-Series [1:176] from 1947 to 1991: -0.00126 -0.00351 0.00586 -0.00306 0.01046 ...
## $ call : language arima(x = v, order = c(1, 0, 0))
## $ series : chr "v"
## $ code : int 0
## $ n.cond : int 0
## $ nobs : int 176
## $ model :List of 10
## ..$ phi : num 0.379
## ..$ theta: num(0)
## ..$ Delta: num(0)
  ..$ Z : num 1
##
   ..$ a : num -0.0142
##
   ..$ P : num [1, 1] 0
##
  ..$ T : num [1, 1] 0.379
##
  ..$ V : num [1, 1] 1
##
## ..$ h : num 0
  ..$ Pn : num [1, 1] 1
## - attr(*, "class")= chr "Arima"
```

```
# print out results for m1
m1
##
## Call:
## arima(x = y, order = c(1, 0, 0))
##
## Coefficients:
##
            ar1 intercept
##
         0.3787
                    0.0077
## s.e. 0.0698
                    0.0012
##
## sigma^2 estimated as 9.801e-05: log likelihood = 562.47, aic = -1118.94
```

 ${\it \# diagnostics for AR(1) model - there seems to be a problem with remaining serial correlation at lag~2 } \\ {\it ggtsdiag(m1, gof.lag=16)}$







1960

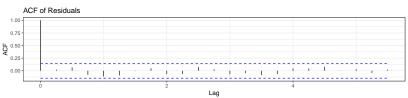
1950

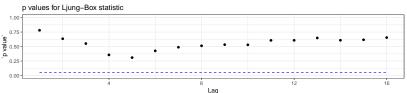
estimate an AR(2) model to deal with the problem of remaining serial correlation at lag 2 m2 <- arima(y, order=c(2,0,0))

diagnostics for AR(2) model shows that problem with remaining serial correlation at lag 2 is gone ggtsdiag(m2, gof.lag=16)



1970

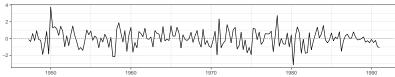




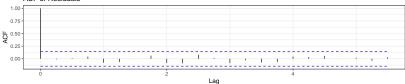
1990

```
# estimate an AR(3) model since PACF for lag 2 and 3 are comparable in size
m3 <- arima(y, order=c(3,0,0))
# diagnostics for the AR(3) model
ggtsdiag(m3, gof.lag=16)</pre>
```

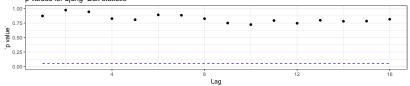




ACF of Residuals



p values for Ljung-Box statistic



```
# use AIC to choose order p of the AR model
m <- ar(v. method="mle")
str(m)
## List of 15
  $ order
             : int 3
## $ ar : num [1:3] 0.348 0.179 -0.142
## $ var.pred : num 9.43e-05
## $ x.mean : num 0.00774
## $ aic : Named num [1:13] 27.847 2.742 1.603 0 0.303 ...
## ..- attr(*, "names")= chr [1:13] "0" "1" "2" "3" ...
## $ n.used : int 176
## $ n.obs : int 176
## $ order.max : num 12
## $ partialacf : NULL
## $ resid : Time-Series [1:176] from 1947 to 1991: NA NA NA -0.00243 0.00903 ...
## $ method : chr "MLE"
## $ series : chr "y"
## $ frequency : num 4
## $ call : language ar(x = v. method = "mle")
## $ asv.var.coef: num [1:3, 1:3] 0.00555 -0.001819 -0.000724 -0.001819 0.006052 ...
## - attr(*, "class")= chr "ar"
# AIC prefers AR(3) to AR(2)
m$order
## [1] 3
m$aic
##
## 27.8466897 2.7416324 1.6032416 0.0000000 0.3027852 2.2426608 4.0520840 6.0254750 5.9046676
                    10
                              11
## 7 5718635 7 8953337 9 6788727 7 1975452
```

```
# BIC prefers AR(1) to AR(2) or AR(3)
# in general BIC puts a larger penalty on additional coefficients than AIC
BIC(m1)

## [1] -1109.431
BIC(m2)

## [1] -1107.398
BIC(m3)

## [1] -1105.832
```

0.4319577

```
# Ljung-Box test - for residuals of a model adjust the degrees of freedom m
# by subtracting the number of parameters a
# this adjustment will not make a big difference if m is large but matters if m is small
m2.LB.lag8 <- Box.test(m2$residuals, lag=8, type="Ljung")</pre>
m2.LB.lag8
##
   Box-Ljung test
##
## data: m2$residuals
## X-squared = 7.2222, df = 8, p-value = 0.5129
1-pchisq(m2.LB.lag8$statistic, df=6)
## X-squared
## 0.3007889
m2.LB.lag12 <- Box.test(m2$residuals, lag=12, type="Ljung")
m2.LB.lag12
##
   Box-Ljung test
##
## data: m2$residuals
## X-squared = 10.098, df = 12, p-value = 0.6074
1-pchisq(m2.LB.lag12$statistic, df=10)
## X-squared
```