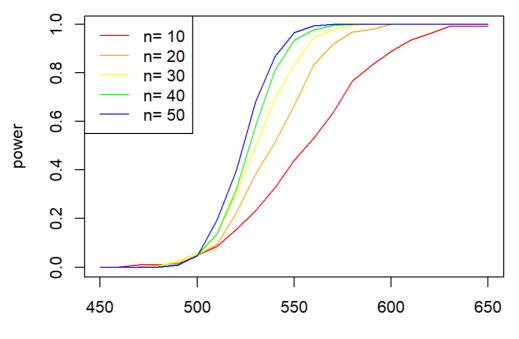
```
##### 6.3
            ###############
n \leftarrow c(seq(10, 50, 10))
m < -1000
mu0 <- 500
sigma <- 100
mu <- c(seq(450, 650, 10)) #alternatives 不同备择参数
M <- length(mu)
power <- numeric(M)</pre>
color <- c('red','orange', 'yellow','green','blue') #设置颜色
label <- c('','','','','') ##图例空值
for (j in 1:length(n))
  n1 <- n[j]
  for (i in 1:M) {
   mu1 <- mu[i]
    pvalues <- replicate(m, expr = {</pre>
     #simulate under alternative mu1
     x \leftarrow rnorm(n1, mean = mu1, sd = sigma)
      ttest <- t.test(x,</pre>
                     alternative = "greater", mu = mu0)
     ttest$p.value } )
   power[i] <- mean(pvalues <= .05)</pre>
  }
 if (j==1) {plot(mu, power, xlab = bquote(mu),col=color[j],'l')} #n=10 \mathbb{H}plot
  lines(mu, power, col=color[j],'l') #n=20..50 用lines
  label[j] <- paste('n=',n[j]) #图例赋值
}
legend("topleft", label, col=color, lty=1) #将图例置于左上角
# 对正态样本做均值t检验(eg6.7),估计功效并绘制经验功效曲线(eg6.9),
# 经验功效在接近mu0=500时较小, 当mu远离mu0时开始增大,逐渐趋近于1。
# 若改变样本量的大小,则样本量越大,功效增大越快。
```



```
##### 6.4 ############
# 使用MC方法得到经验置信水平,理论上应接近1-α=0.95
m < -1000
alpha <- .05
mean <- 0
sd <- 1
CL <- replicate(m, expr={</pre>
 # 生成对数正态分布的样本
 x <- rlnorm(n, mean, sd)</pre>
  # 构造均值µ的置信区间上下界
  \max \leftarrow \sup(\log(x))/n + qt(1-alpha/2, n-1)*sqrt(var(\log(x)))/sqrt(n)
  min \leftarrow sum(log(x))/n-qt(1-alpha/2, n-1)*sqrt(var(log(x)))/sqrt(n)
  if(0 > min \& 0 < max) \{ beta <- 1 \} else \{ beta <- 0 \}
  beta
  })
mean(CL) #计算经验置信水平(接近0.95)
```

```
##### 6.8 ############
# 双样本等方差的Count5检验函数 返回值1(拒绝H0)或0(接受H0)
count5test <- function(x, y) {</pre>
  X \leftarrow x - mean(x)
 Y \leftarrow y - mean(y)
 outx <- sum(X > max(Y)) + sum(X < min(Y))
  outy \leftarrow sum(Y > max(X)) + sum(Y < min(X))
  # return 1 (reject) or 0 (do not reject H0)
  return(as.integer(max(c(outx, outy)) > 5))
}
m < -1000
n \leftarrow c(10, 20, 100)
alpha <- .055
sigma1 <- 1
sigma2 <- 1.5
power.C5 <- numeric(3)</pre>
power.F <- numeric(3)</pre>
for (i in 1:length(n)) {
  n1 <- n2 <- n[i]
 rep <- replicate(m, expr={</pre>
  x \leftarrow rnorm(n1, 0, sigma1)
```

```
y <- rnorm(n2, 0, sigma2)
C5 <- count5test(x, y) ##C5-test
Ftest <- var.test(x, y, ratio = 1, "two.sided", conf.level = 1-alpha)
F <- Ftest$p.value < alpha ##F-test
c(C5, F)})
power.C5[i] <- mean(rep[1,])
power.F[i] <- mean(rep[2,])
}

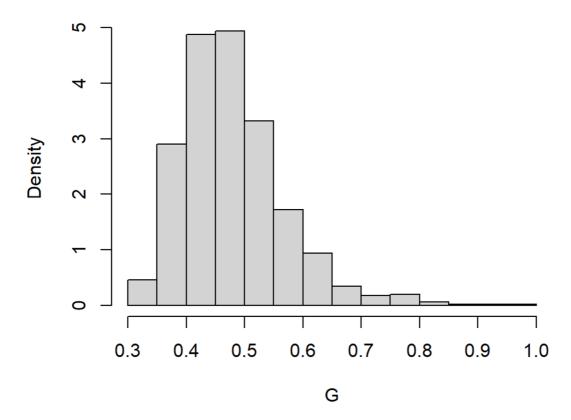
> data.frame(n=c(10,20,100), power.C5, power.F)
    n power.C5 power.F
1 10 0.116 0.231
2 20 0.323 0.423
3 100 0.854 0.977
# 在样本量分别为10,20,100时,F检验的功效都高于C5检验
```

```
> ##### 6.9 ############
> # 利用次序统计量表示基尼系数(分配公平程度)
> #基于标准对数正态分布
> n <- 20
> m <- 1000
> G <- numeric(m)</pre>
> G <- replicate(m,expr={</pre>
  g <- numeric(n)</pre>
  for (i in 1:n){
     x <- sort(rlnorm(n, mean=0, sd=1)) #排序
    mu <- mean(x)
+
     g[i] \leftarrow (2*i-n-1)*x[i]/(n^2*mu)
+
      sum(g)
> G.mean <- mean(G)</pre>
> G.mean
[1] 0.4785041
> G.median <- median(G)</pre>
> G.median
[1] 0.4638238
> G.quantile \leftarrow quantile(G,c(seq(0.1,0.9,0.1)))
> G.quantile
      10%
                20% 30%
                                                50%
                                40%
0.3858744 0.4108444 0.4311688 0.4482350 0.4638238
                         80%
               70%
                                    90%
0.4860891 0.5087767 0.5405828 0.5846074
> hist(G, prob=TRUE)
> #基于均匀分布
> n <- 20
> m <- 1000
> G <- numeric(m)</pre>
> G <- replicate(m,expr={</pre>
+ g <- numeric(n)</pre>
  x <- sort(runif(n))</pre>
+ mu \leftarrow mean(x)
  for (i in 1:n){
    g[i] \leftarrow (2*i-n-1)*x[i]/(n^2*mu)
  }
    sum(g)
> G.mean <- mean(G)</pre>
> G.mean
```

```
[1] 0.3217252
> G.median <- median(G)</pre>
> G.median
[1] 0.3212548
> G.quantile \leftarrow quantile(G,c(seq(0.1,0.9,0.1)))
> G.quantile
      10% 20% 30% 40%
                                         50%
0.2561058 0.2789920 0.2950767 0.3092741 0.3212548
      60% 70% 80% 90%
0.3333215 0.3481908 0.3641689 0.3865076
> hist(G, prob=TRUE)
> #基于伯努利分布
> n <- 20
> m <- 1000
> G <- numeric(m)</pre>
> G <- replicate(m,expr={</pre>
+ g <- numeric(n)</pre>
+ x \leftarrow sort(rbinom(n,1,0.1))
+ mu \leftarrow mean(x)
+ for (i in 1:n){
    if (mu==0){g[i]=0}
   else{g[i] <- (2*i-n-1)*x[i]/(n^2*mu)}
+
+ }
  sum(g)})
> G.mean <- mean(G)</pre>
> G.mean
[1] 0.7926
> G.median <- median(G)</pre>
> G.median
[1] 0.9
> G.quantile \leftarrow quantile(G,c(seq(0.1,0.9,0.1)))
> G.quantile
10% 20% 30% 40% 50% 60% 70% 80% 90%
0.00 0.80 0.85 0.85 0.90 0.90 0.90 0.95 0.95
> hist(G, prob=TRUE)
> ## 基于密度直方图,可以看出贫富差距悬殊。
> # 事实上由于x只取值于0或1,直观上可以推断密度直方图会趋于极端。
```

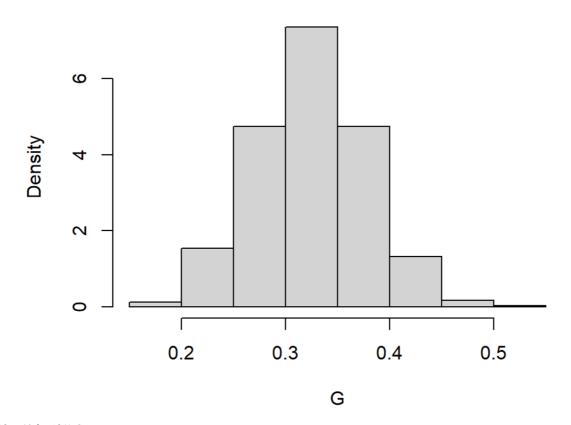
基于标准对数正态分布

Histogram of G



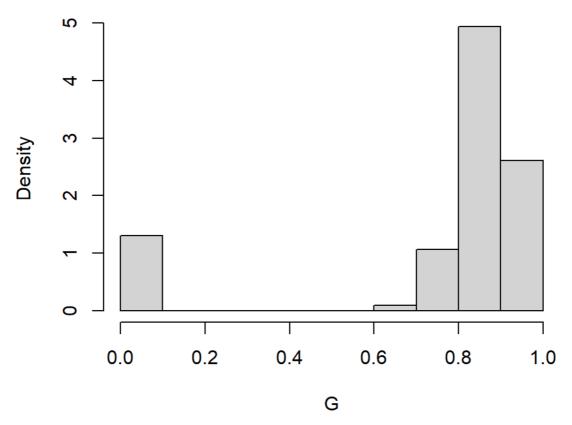
基于均匀分布

Histogram of G



基于伯努利分布

Histogram of G



```
> #####
         6A显著性水平的t-test ##############
> #用MC模拟研究 经验error I 是否约等于理论显著水平α
> #注: t-test对微小的正态性偏离是稳定的
> #基于卡方分布
> n <- 20
> alpha <- .05
> mu0 <- 1
> m <- 10000
                   #number of replicates
> p <- numeric(m)</pre>
                   #storage for p-values
> for (j in 1:m) {
+ x <- rchisq(n, df=1)
  ttest <- t.test(x, alternative = "greater", mu = mu0)</pre>
   p[j] <- ttest$p.value</pre>
> p.hat <- mean(p < alpha) #观测的第一类错误率 (理论上应该接α=0.05)
> se.hat <- sqrt(p.hat * (1 - p.hat) / m) #估计的标准误差
> print(c(p.hat, se.hat))
[1] 0.01300000 0.00113274
> # p-value < 0.02,和0.05差距较大
> #基于均匀分布
> n <- 20
> alpha <- .05
> mu0 <- 1
                  #number of replicates
> m <- 10000
> p <- numeric(m)</pre>
                   #storage for p-values
> for (j in 1:m) {
+ x <- runif(n, 0, 2)
  ttest <- t.test(x, alternative = "greater", mu = mu0)</pre>
```

```
+ p[j] <- ttest$p.value</pre>
+ }
> p.hat <- mean(p < alpha) #观测的第一类错误率 (理论上应该接α=0.05)
> se.hat <- sqrt(p.hat * (1 - p.hat) / m) #估计的标准误差
> print(c(p.hat, se.hat))
[1] 0.048700000 0.002152401
> # p-value 接近0.05
> #基于指数分布
> n <- 20
> alpha <- .05
> mu0 <- 1
                   #number of replicates
> m <- 10000
> p <- numeric(m)</pre>
                    #storage for p-values
> for (j in 1:m) {
+ x \leftarrow rexp(n, 1)
  ttest <- t.test(x, alternative = "greater", mu = mu0)</pre>
+ p[j] <- ttest$p.value</pre>
+ }
> p.hat <- mean(p < alpha) #观测的第一类错误率 (理论上应该接α=0.05)
> se.hat <- sqrt(p.hat * (1 - p.hat) / m) #估计的标准误差
> print(c(p.hat, se.hat))
[1] 0.018900000 0.001361719
> # p-value < 0.02,和0.05差距较大
```

```
> ##### 6B ##############
> library(MASS) #加载MASS包
> #二元正态下, pearson比另两种非参数检验得到的pho更接近
> mean<-c(1, 1) #指定均值向量
> sigma<-matrix(c(1, 0.5, 0.5, 1), nrow=2, ncol=2) #指定协方差矩阵
> mydata <- mvrnorm(100000, mean, sigma)</pre>
> cor.test(mydata[,1],mydata[,2],method = "pearson")
0.5018238
> cor.test(mydata[,1],mydata[,2],method = "spearman")
     rho
0.4856505
> cor.test(mydata[,1],mydata[,2],method = "kendall")
     tau
0.3355287
> #在数据非线性单调情况下,非参检验功效更高
> x <- runif(1000,0,1)
> y <- x^6
> cor.test(x,y,method = "pearson")
     cor
0.7786169
> cor.test(x,y,method = "spearman")
rho
 1
> cor.test(x,y,method = "kendall")
sample estimates:
```

tau

1